



LaplacesDemon Examples

Statisticat, LLC

Abstract

The **LaplacesDemon** package is a complete environment for Bayesian inference within R. Virtually any probability model may be specified. This vignette is a compendium of examples of how to specify different model forms.

Keywords: Bayesian, LaplacesDemon, LaplacesDemonCpp, R.

LaplacesDemon (Statisticat LLC. 2015), often referred to as LD, is an R package that is available at <https://web.archive.org/web/20150430054143/http://www.bayesian-inference.com/software>. **LaplacesDemonCpp** is an extension package that uses C++. A formal introduction to **LaplacesDemon** is provided in an accompanying vignette entitled “**LaplacesDemon** Tutorial”, and an introduction to Bayesian inference is provided in the “Bayesian Inference” vignette.

The purpose of this document is to provide users of the **LaplacesDemon** package with examples of a variety of Bayesian methods. It is also a testament to the diverse applicability of **LaplacesDemon** to Bayesian inference.

To conserve space, the examples are not worked out in detail, and only the minimum of necessary materials is provided for using the various methodologies. Necessary materials include the form expressed in notation, data (which is often simulated), the `Model` function, and initial values. The provided data, model specification, and initial values may be copy/pasted into an R file and updated with the `LaplacesDemon` or (usually) `LaplaceApproximation` functions. Although many of these examples update quickly, some examples are computationally intensive.

All examples are provided in R code, but the model specification function can be in another language. A goal is to provide these example model functions in C++ as well, and some are now available at <https://web.archive.org/web/20140513065103/http://www.bayesian-inference.com/cpp/LaplacesDemonExamples.txt>.

Initial values are usually hard-coded in the examples, though the Parameter-Generating Function (PGF) is also specified. It is recommended to generate initial values with the `GIV` function according to the user-specified PGF.

Notation in this vignette follows these standards: Greek letters represent parameters, lower case letters represent indices, lower case bold face letters represent scalars or vectors, probabil-

ity distributions are represented with calligraphic font, upper case letters represent index limits, and upper case bold face letters represent matrices. More information on notation is available at <https://web.archive.org/web/20150501205317/http://www.bayesian-inference.com/notation>.

This vignette may grow over time as examples of more methods become included. Contributed examples are welcome via <https://github.com/LaplacesDemonR/LaplacesDemon/issues>. All accepted contributions are, of course, credited.

Contents

- Adaptive Logistic Basis (ALB) Regression 1
- ANCOVA 2
- ANOVA, One-Way 3
- ANOVA, Two-Way 4
- Approximate Bayesian Computation (ABC) 5
- AR(p) 6
- AR(p)-ARCH(q) 7
- AR(p)-ARCH(q)-M 8
- AR(p)-GARCH(1,1) 9
- AR(p)-GARCH(1,1)-M 10
- AR(p)-TARCH(q) 11
- AR(p)-TARCH(q)-M 12
- Autoregressive Moving Average, ARMA(p,q) 13
- Beta Regression 14
- Beta-Binomial 15
- Binary Logit 16
- Binary Log-Log Link Mixture 17
- Binary Probit 18
- Binary Robit 19
- Binomial Logit 20
- Binomial Probit 21
- Binomial Robit 22

- Change Point Regression 23
- Cluster Analysis, Confirmatory (CCA) 24
- Cluster Analysis, Exploratory (ECA) 25
- Collaborative Filtering (CF) 38
- Conditional Autoregression (CAR), Poisson 26
- Conditional Predictive Ordinate (CPO) 27
- Contingency Table 28
- Dirichlet Process 25 62
- Discrete Choice, Conditional Logit 29
- Discrete Choice, Mixed Logit 30
- Discrete Choice, Multinomial Probit 31
- Distributed Lag, Koyck 32
- Dynamic Linear Model (DLM) ?? 92 93 94
- Dynamic Sparse Factor Model (DSFM) 33
- Exponential Smoothing 34
- Factor Analysis, Approximate Dynamic (ADFA) 35
- Factor Analysis, Confirmatory (CFA) 36
- Factor Analysis, Dynamic (DFA) 33
- Factor Analysis, Exploratory (EFA) 37
- Factor Analysis, Exploratory Ordinal (EOFA) 38
- Factor Regression 39
- Gamma Regression 40
- Gaussian Process Regression 44
- Geographically Weighted Regression 41
- Hidden Markov Model 42
- Hierarchical Bayes 50
- Horseshoe Regression 99
- Inverse Gaussian Regression 43
- Kriging 44

- Kriging, Predictive Process 45
- Laplace Regression 46
- LASSO 98 100
- Latent Dirichlet Allocation (LDA) 47
- Linear Regression 48
- Linear Regression, Frequentist 49
- Linear Regression, Hierarchical Bayesian 50
- Linear Regression, Multilevel 51
- Linear Regression with Full Missingness 52
- Linear Regression with Missing Response 53
- Linear Regression with Missing Response via ABB 54
- Linear Regression with Power Priors 55
- Linear Regression with Zellner's g-Prior 56
- LSTAR 57
- MANCOVA 58
- MANOVA 59
- Missing Values 52 53 54
- Mixed Logit 60
- Mixture Model, Finite 24 61
- Mixture Model, Infinite 25 62
- Mixture Model, Poisson-Gamma ??
- Model Averaging 102 101
- Multilevel Model 51
- Multinomial Logit 63
- Multinomial Logit, Nested 64
- Multinomial Probit 65
- Multiple Discrete-Continuous Choice 66
- Multivariate Binary Probit 67
- Multivariate Laplace Regression 68

- Multivariate Poisson Regression 69
- Multivariate Regression 70
- Negative Binomial Regression 71
- Normal, Multilevel 72
- Ordinal Logit 73
- Ordinal Probit 74
- Panel, Autoregressive Poisson 75
- Penalized Spline Regression 76
- Poisson Regression 77
- Poisson Regression, Overdispersed ?? 71
- Poisson-Gamma Regression ??
- Polynomial Regression 78
- Power Priors 55
- Proportional Hazards Regression, Weibull 79
- PVAR(p) 80
- Quantile Regression 81
- Revision, Normal 82
- Ridge Regression 83
- Robust Regression 84
- Seemingly Unrelated Regression (SUR) 85
- Simultaneous Equations 86
- Space-Time, Dynamic 87
- Space-Time, Nonseparable 88
- Space-Time, Separable 89
- Spatial Autoregression (SAR) 90
- STARMA(p,q) 91
- State Space Model (SSM), Dynamic Sparse Factor Model (DSFM) 33
- State Space Model (SSM), Linear Regression 92
- State Space Model (SSM), Local Level 93

- State Space Model (SSM), Local Linear Trend 94
- State Space Model (SSM), Stochastic Volatility (SV) 95
- Stochastic Volatility (SV) 95
- Survival Model 79
- T-test 3
- Threshold Autoregression (TAR) 96
- Topic Model 47
- Time Varying AR(1) with Chebyshev Series 97
- Variable Selection, BAL 98
- Variable Selection, Horseshoe 99
- Variable Selection, LASSO 100
- Variable Selection, RJ 101
- Variable Selection, SSVS 102
- VARMA(p,q) - SSVS 103
- VAR(p)-GARCH(1,1)-M 104
- VAR(p) with Minnesota Prior 105
- VAR(p) with SSVS ??
- Variety Model 66
- Weighted Regression 107
- Zellner's g-Prior 56
- Zero-Inflated Poisson (ZIP) 108

1. Adaptive Logistic Basis (ALB) Regression

Adaptive Logistic Basis (ALB) regression is an essentially automatic non-parametric approach to estimating the nonlinear relationship between each of multiple independent variables (IVs) and the dependent variable (DV). It is automatic because when using the suggested $K = 2J+1$ components (see below) given J IVs, the data determines the nonlinear relationships, whereas in other methods, such as with splines, the user must specify the number of knots and possibly consider placement of the knots. Knots do not exist in ALB. Both the DV and IVs should be centered and scaled.

1.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\
 \boldsymbol{\mu} &= \mathbf{S}\boldsymbol{\delta} \\
 \mathbf{S}_{i,m} &= \frac{\phi_{i,m}}{\sum_{m=1}^M \phi_{i,m}} \\
 \phi_{i,m} &= \exp(\alpha_m + \mathbf{X}_{i,1:J}\boldsymbol{\beta}_{1:J,m}), \quad i = 1, \dots, N, \quad m = 1, \dots, M \\
 \alpha_m &\sim \mathcal{N}(0, 10), \quad m = 1, \dots, (M-1) \\
 \alpha_M &= 0 \\
 \beta_{j,m} &\sim \mathcal{N}(0, 100), \quad j = 1, \dots, J, \quad m = 1, \dots, (M-1) \\
 \beta_{j,M} &= 0 \\
 \delta_m &\sim \mathcal{N}(\zeta, \tau^2), \quad m = 1, \dots, M \\
 \sigma &\sim \mathcal{HC}(25) \\
 \zeta &\sim \mathcal{N}(0, 10) \\
 \tau &\sim \mathcal{HC}(25)
 \end{aligned}$$

1.2. Data

```

data(demonsnacks)
N <- nrow(demonsnacks)
y <- log(demonsnacks$Calories)
X <- as.matrix(log(demonsnacks[,c(1,4,10)]+1))
J <- ncol(X)
y <- CenterScale(y)
for (j in 1:J) X[,j] <- CenterScale(X[,j])
K <- 2*J+1
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,K-1), beta=matrix(0,J,K-1),
  delta=rep(0,K), zeta=0, sigma=0, tau=0))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.delta <- grep("delta", parm.names)
pos.zeta <- grep("zeta", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.tau <- grep("tau", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(Data$K-1)
  beta <- rnorm(Data$J*(Data$K-1))
  delta <- rnorm(Data$K)
  zeta <- rnorm(1)
  sigma <- rhalfcauchy(1,5)
}

```

```

tau <- rhalfcauchy(1,5)
return(c(alpha, beta, delta, zeta, sigma, tau))
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, y=y, pos.alpha=pos.alpha, pos.beta=pos.beta,
  pos.delta=pos.delta, pos.zeta=pos.zeta, pos.sigma=pos.sigma,
  pos.tau=pos.tau)

```

1.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperparameters
  zeta <- parm[Data$pos.zeta]
  parm[Data$pos.tau] <- tau <- interval(parm[Data$pos.tau], 1e-100, Inf)
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- matrix(parm[Data$pos.beta], Data$J, Data$K-1)
  delta <- parm[Data$pos.delta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Hyperprior
  zeta.prior <- dnormv(zeta, 0, 10, log=TRUE)
  tau.prior <- dhalfcauchy(tau, 25, log=TRUE)
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 10, log=TRUE))
  beta.prior <- sum(dnormv(beta, 0, 100, log=TRUE))
  delta.prior <- sum(dnorm(delta, zeta, tau, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  phi <- cbind(exp(matrix(alpha, Data$N, Data$K-1, byrow=TRUE) +
    tcrossprod(Data$X, t(beta))),1)
  mu <- tcrossprod(phi / rowSums(phi), t(delta))
  LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + alpha.prior + beta.prior + delta.prior + zeta.prior
  sigma.prior + tau.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

1.4. Initial Values

```
Initial.Values <- c(rep(0,K), rep(0,J*(K-1)), rep(0,K-1), 0, 1, 1)
```


2. ANCOVA

This example is essentially the same as the two-way ANOVA (see section 4), except that a covariate $\mathbf{X}_{,3}$ has been added, and its parameter is δ .

2.1. Form

$$\begin{aligned} \mathbf{y}_i &\sim \mathcal{N}(\mu_i, \sigma_1^2) \\ \mu_i &= \alpha + \beta[\mathbf{X}_{i,1}] + \gamma[\mathbf{X}_{i,2}] + \delta\mathbf{X}_{i,2}, \quad i = 1, \dots, N \\ \epsilon_i &= \mathbf{y}_i - \mu_i \\ \alpha &\sim \mathcal{N}(0, 1000) \\ \beta_j &\sim \mathcal{N}(0, \sigma_2^2), \quad j = 1, \dots, J \\ \beta_J &= -\sum_{j=1}^{J-1} \beta_j \\ \gamma_k &\sim \mathcal{N}(0, \sigma_3^2), \quad k = 1, \dots, K \\ \gamma_K &= -\sum_{k=1}^{K-1} \gamma_k \\ \delta &\sim \mathcal{N}(0, 1000) \\ \sigma_m &\sim \mathcal{HC}(25), \quad m = 1, \dots, 3 \end{aligned}$$

2.2. Data

```

N <- 100
J <- 5 #Number of levels in factor (treatment) 1
K <- 3 #Number of levels in factor (treatment) 2
X <- cbind(rcat(N,rep(1/J,J)), rcat(N,rep(1/K,K)), runif(N,-2,2))
alpha <- runif(1,-1,1)
beta <- runif(J-1,-2,2)
beta <- c(beta, -sum(beta))
gamma <- runif(K-1,-2,2)
gamma <- c(gamma, -sum(gamma))
delta <- runif(1,-2,2)
y <- alpha + beta[X[,1]] + gamma[X[,2]] + delta*X[,3] + rnorm(N,0,0.1)
mon.names <- c("LP", paste("beta[",J,"]",sep=""),
  paste("gamma[",K,"]",sep=""), "s.beta", "s.gamma", "s.epsilon")
parm.names <- as.parm.names(list(alpha=0, beta=rep(0,J-1), gamma=rep(0,K-1),
  delta=0, sigma=rep(0,3)))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.delta <- grep("delta", parm.names)

```

```

pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  beta <- rnorm(Data$J-1)
  gamma <- rnorm(Data$K-1)
  delta <- rnorm(1)
  sigma <- runif(3)
  return(c(alpha, beta, gamma, delta, sigma))
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha,
  pos.beta=pos.beta, pos.gamma=pos.gamma, pos.delta=pos.delta,
  pos.sigma=pos.sigma, y=y)

```

2.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- parm[Data$pos.beta]
  beta <- c(beta, -sum(beta)) #Sum-to-zero constraint
  gamma <- parm[Data$pos.gamma]
  gamma <- c(gamma, -sum(gamma)) #Sum-to-zero constraint
  delta <- parm[Data$pos.delta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- dnorm(alpha, 0, 1000, log=TRUE)
  beta.prior <- sum(dnorm(beta, 0, sigma[2], log=TRUE))
  gamma.prior <- sum(dnorm(gamma, 0, sigma[3], log=TRUE))
  delta.prior <- dnorm(delta, 0, 1000, log=TRUE)
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  mu <- alpha + beta[Data$X[,1]] + gamma[Data$X[,2]] +
    delta*Data$X[,3]
  LL <- sum(dnorm(Data$y, mu, sigma[1], log=TRUE))
  ### Variance Components
  s.beta <- sd(beta)
  s.gamma <- sd(gamma)
  s.epsilon <- sd(Data$y - mu)
  ### Log-Posterior
  LP <- LL + alpha.prior + beta.prior + gamma.prior + delta.prior +
    sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP, beta[Data$J],
    gamma[Data$K], s.beta, s.gamma, s.epsilon),

```

```

      yhat=rnorm(length(mu), mu, sigma[1]), parm=parm)
return(Modelout)
}

```

2.4. Initial Values

```
Initial.Values <- c(0, rep(0,(J-1)), rep(0,(K-1)), 0, rep(1,3))
```

3. ANOVA, One-Way

When $J = 2$, this is a Bayesian form of a t-test.

3.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\mu, \sigma_1^2) \\
 \mu_i &= \alpha + \beta[\mathbf{x}_i], \quad i = 1, \dots, N \\
 \alpha &\sim \mathcal{N}(0, 1000) \\
 \beta_j &\sim \mathcal{N}(0, \sigma_2^2), \quad j = 1, \dots, J \\
 \beta_J &= -\sum_{j=1}^{J-1} \beta_j \\
 \sigma_{1:2} &\sim \mathcal{HC}(25)
 \end{aligned}$$

3.2. Data

```

N <- 1000
J <- 3
x <- rcat(N, rep(1/J, J))
alpha <- runif(1,-1,1)
beta <- runif(J-1,-2,2)
beta <- c(beta, -sum(beta))
y <- alpha + beta[x] + rnorm(N,0,0.2)
mon.names <- c("LP",paste("beta[",J,"]",sep=""))
parm.names <- as.parm.names(list(alpha=0, beta=rep(0,J-1), sigma=rep(0,2)))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  beta <- rnorm(Data$J-1)
  sigma <- runif(2)
  return(c(alpha, beta, sigma))
}

```

```
MyData <- list(J=J, N=N, PGF=PGF, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.beta=pos.beta,
  pos.sigma=pos.sigma, x=x, y=y)
```

3.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- parm[Data$pos.beta]
  beta <- c(beta, -sum(beta)) #Sum-to-zero constraint
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- dnorm(alpha, 0, 1000, log=TRUE)
  beta.prior <- sum(dnorm(beta, 0, sigma[2], log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  mu <- alpha + beta[Data$x]
  LL <- sum(dnorm(Data$y, mu, sigma[1], log=TRUE))
  ### Log-Posterior
  LP <- LL + alpha.prior + beta.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,beta[Data$J]),
    yhat=rnorm(length(mu), mu, sigma[1]), parm=parm)
  return(Modelout)
}
```

3.4. Initial Values

```
Initial.Values <- c(0, rep(0,(J-1)), rep(1,2))
```

4. ANOVA, Two-Way

In this representation, σ^m are the superpopulation variance components, **s.beta** and **s.gamma** are the finite-population within-variance components of the factors or treatments, and **s.epsilon** is the finite-population between-variance component.

4.1. Form

$$\mathbf{y}_i \sim \mathcal{N}(\mu_i, \sigma_1^2)$$

$$\mu_i = \alpha + \beta[\mathbf{X}_{i,1}] + \gamma[\mathbf{X}_{i,2}], \quad i = 1, \dots, N$$

$$\epsilon_i = \mathbf{y}_i - \mu_i$$

$$\begin{aligned} \alpha &\sim \mathcal{N}(0, 1000) \\ \beta_j &\sim \mathcal{N}(0, \sigma_2^2), \quad j = 1, \dots, J \\ \beta_J &= -\sum_{j=1}^{J-1} \beta_j \\ \gamma_k &\sim \mathcal{N}(0, \sigma_3^2), \quad k = 1, \dots, K \\ \gamma_K &= -\sum_{k=1}^{K-1} \gamma_k \\ \sigma_m &\sim \mathcal{HC}(25), \quad m = 1, \dots, 3 \end{aligned}$$

4.2. Data

```

N <- 1000
J <- 5 #Number of levels in factor (treatment) 1
K <- 3 #Number of levels in factor (treatment) 2
X <- cbind(rcat(N,rep(1/J,J)), rcat(N,rep(1/K,K)))
alpha <- runif(1,-1,1)
beta <- runif(J-1,-2,2)
beta <- -sum(beta)
gamma <- runif(K-1,-2,2)
gamma <- -sum(gamma)
y <- alpha + beta[X[,1]] + gamma[X[,2]] + rnorm(N,0,0.1)
mon.names <- c("LP", paste("beta[",J,"]",sep=""),
  paste("gamma[",K,"]",sep=""), "s.beta", "s.gamma", "s.epsilon")
parm.names <- as.parm.names(list(alpha=0, beta=rep(0,J-1), gamma=rep(0,K-1),
  sigma=rep(0,3)))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  beta <- rnorm(Data$J-1)
  gamma <- rnorm(Data$K-1)
  sigma <- runif(3)
  return(c(alpha, beta, gamma, sigma))
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.beta=pos.beta,
  pos.gamma=pos.gamma, pos.sigma=pos.sigma, y=y)

```

4.3. Model

```

Model <- function(parm, Data)
{

```

```

### Parameters
alpha <- parm[Data$pos.alpha]
beta <- parm[Data$pos.beta]
beta <- c(beta, -sum(beta)) #Sum-to-zero constraint
gamma <- parm[Data$pos.gamma]
gamma <- c(gamma, -sum(gamma)) #Sum-to-zero constraint
sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
parm[Data$pos.sigma] <- sigma
### Log-Prior
alpha.prior <- dnorm(alpha, 0, 1000, log=TRUE)
beta.prior <- sum(dnorm(beta, 0, sigma[2], log=TRUE))
gamma.prior <- sum(dnorm(gamma, 0, sigma[3], log=TRUE))
sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
### Log-Likelihood
mu <- alpha + beta[Data$X[,1]] + gamma[Data$X[,2]]
LL <- sum(dnorm(Data$y, mu, sigma[1], log=TRUE))
### Variance Components
s.beta <- sd(beta)
s.gamma <- sd(gamma)
s.epsilon <- sd(Data$y - mu)
### Log-Posterior
LP <- LL + alpha.prior + beta.prior + gamma.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP, beta[Data$J],
      gamma[Data$K], s.beta, s.gamma, s.epsilon),
      yhat=rnorm(length(mu), mu, sigma[1]), parm=parm)
return(Modelout)
}

```

4.4. Initial Values

```
Initial.Values <- c(0, rep(0,(J-1)), rep(0,(K-1)), rep(1,3))
```

5. Approximate Bayesian Computation (ABC)

Approximate Bayesian Computation (ABC), also called likelihood-free estimation, is not a statistical method, but a family of numerical approximation techniques in Bayesian inference. ABC is especially useful when evaluation of the likelihood, $p(\mathbf{y}|\Theta)$ is computationally prohibitive, or when suitable likelihoods are unavailable. The current example is the application of ABC in the context of linear regression. The log-likelihood is replaced with the negative sum of the distance between \mathbf{y} and \mathbf{y}^{rep} as the approximation of the log-likelihood. Distance reduces to the absolute difference. Although linear regression has an easily calculated likelihood, it is used as an example due to its generality. This example demonstrates how ABC may be estimated either with MCMC via the `LaplacesDemon` function or with Laplace Approximation via the `LaplaceApproximation` function. In this method, a tolerance (which is found often in ABC) does not need to be specified, and the logarithm of the unnormalized

joint posterior density is maximized, as usual. The negative and summed distance, above, may be replaced with the negative and summed distance between summaries of the data, rather than the data itself, but this has not been desirable in testing.

5.1. Form

$$\begin{aligned} \mathbf{y} &= \boldsymbol{\mu} + \boldsymbol{\epsilon} \\ \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \\ \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \end{aligned}$$

5.2. Data

```
data(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(log(demonsnacks[,c(1,4,10)]+1)))
J <- ncol(X)
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- c("LP","sigma")
parm.names <- as.parm.names(list(beta=rep(0,J)))
pos.beta <- grep("beta", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  return(beta)
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, y=y)
```

5.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  ### Log-Likelihood Approximation
  mu <- as.vector(tcrossprod(Data$X, t(beta)))
  epsilon <- Data$y - mu
  sigma <- sd(epsilon)
  LL <- -sum(abs(epsilon))
  ### Log-Posterior Approximation
  LP <- LL + beta.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,sigma),
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
```

```
return(Modelout)
}
```

5.4. Initial Values

```
Initial.Values <- c(rep(0,J))
```

6. AR(p)

6.1. Form

$$\mathbf{y}_t \sim \mathcal{N}(\mu_t, \sigma^2), \quad t = 1, \dots, T$$

$$\mu_t = \alpha + \sum_{p=1}^P \phi_p \mathbf{y}_{t-p}, \quad t = 1, \dots, T$$

$$\alpha \sim \mathcal{N}(0, 1000)$$

$$\phi_p \sim \mathcal{N}(0, 1000), \quad p = 1, \dots, P$$

$$\sigma \sim \mathcal{HC}(25)$$

6.2. Data

```
data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
T <- length(y)
L <- c(1,5,20) #Autoregressive lags
P <- length(L) #Autoregressive order
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, phi=rep(0,P), sigma=0))
pos.alpha <- grep("alpha", parm.names)
pos.phi <- grep("phi", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  phi <- runif(Data$P,-1,1)
  sigma <- rhalfcauchy(1,5)
  return(c(alpha, phi, sigma))
}
MyData <- list(L=L, PGF=PGF, P=P, T=T, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.phi=pos.phi,
  pos.sigma=pos.sigma, y=y)
```


6.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  phi <- parm[Data$pos.phi]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- dnormv(alpha, 0, 1000, log=TRUE)
  phi.prior <- sum(dnormv(phi, 0, 1000, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- rep(alpha, Data$T)
  for (p in 1:Data$P)
    mu[-c(1:Data$L[p])] <- mu[-c(1:Data$L[p])] +
      phi[p]*Data$y[1:(Data$T-Data$L[p])]
  LL <- sum(dnorm(Data$y[-c(1:Data$L[Data$P])], mu[-c(1:Data$L[Data$P])],
    sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + alpha.prior + phi.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

6.4. Initial Values

```
Initial.Values <- c(rep(0,P+1), 1)
```

7. AR(p)-ARCH(q)

7.1. Form

$$\begin{aligned}
 \mathbf{y}_t &\sim \mathcal{N}(\mu_t, \sigma_t^2), \quad t = 1, \dots, T \\
 \mu_t &= \alpha + \sum_{p=1}^P \phi_p \mathbf{y}_{t-p}, \quad t = 1, \dots, T \\
 \epsilon_t &= \mathbf{y}_t - \mu_t \\
 \alpha &\sim \mathcal{N}(0, 1000) \\
 \phi_p &\sim \mathcal{N}(0, 1000), \quad p = 1, \dots, P
 \end{aligned}$$

$$\sigma_t^2 = \omega + \sum_{q=1}^Q \theta_q \epsilon_{t-q}^2, \quad t = 2, \dots, T$$

$$\omega \sim \mathcal{HC}(25)$$

$$\theta_q \sim \mathcal{U}(0, 1), \quad q = 1, \dots, Q$$

7.2. Data

```

data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
T <- length(y)
L.P <- c(1,5,20) #Autoregressive lags
L.Q <- c(1,2) #Volatility lags
P <- length(L.P) #Autoregressive order
Q <- length(L.Q) #Volatility order
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, phi=rep(0,P), omega=0,
  theta=rep(0,Q)))
pos.alpha <- grep("alpha", parm.names)
pos.phi <- grep("phi", parm.names)
pos.omega <- grep("omega", parm.names)
pos.theta <- grep("theta", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  phi <- runif(Data$P,-1,1)
  omega <- rhalfcauchy(1,5)
  theta <- runif(Data$Q, 1e-10, 1-1e-5)
  return(c(alpha, phi, omega, theta))
}
MyData <- list(L.P=L.P, L.Q=L.Q, PGF=PGF, P=P, Q=Q, T=T, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.phi=pos.phi,
  pos.omega=pos.omega, pos.theta=pos.theta, y=y)

```

7.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  phi <- parm[Data$pos.phi]
  omega <- interval(parm[Data$pos.omega], 1e-100, Inf)
  parm[Data$pos.omega] <- omega
  theta <- interval(parm[Data$pos.theta], 1e-10, 1-1e-5)
  parm[Data$pos.theta] <- theta
  ### Log-Prior

```

```

alpha.prior <- dnormv(alpha, 0, 1000, log=TRUE)
phi.prior <- sum(dnormv(phi, 0, 1000, log=TRUE))
omega.prior <- dhalfcauchy(omega, 25, log=TRUE)
theta.prior <- sum(dunif(theta, 1e-10, 1-1e-5, log=TRUE))
### Log-Likelihood
mu <- rep(alpha, Data$T)
for (p in 1:Data$P)
  mu[-c(1:Data$L.P[p])] <- mu[-c(1:Data$L.P[p])] +
    phi[p]*Data$y[1:(Data$T-Data$L.P[p])]
epsilon <- Data$y - mu
sigma2 <- rep(omega, Data$T)
for (q in 1:Data$Q)
  sigma2[-c(1:Data$L.Q[q])] <- sigma2[-c(1:Data$L.Q[q])] +
    theta[q]*epsilon[1:(Data$T-Data$L.Q[q])]^2
LL <- sum(dnormv(Data$y[-c(1:Data$L.P[Data$P])],
  mu[-c(1:Data$L.P[Data$P])], sigma2[-c(1:Data$L.P[Data$P])],
  log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + phi.prior + omega.prior + theta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnormv(length(mu), mu, sigma2), parm=parm)
return(Modelout)
}

```

7.4. Initial Values

```
Initial.Values <- c(rep(0,P+1), 1, rep(0.5,Q))
```

8. AR(p)-ARCH(q)-M

8.1. Form

$$\begin{aligned}
 \mathbf{y}_t &\sim \mathcal{N}(\mu_t, \sigma_t^2), \quad t = 1, \dots, T \\
 \mu_t &= \alpha + \sum_{p=1}^P \phi_p \mathbf{y}_{t-p} + \delta \sigma_{t-1}^2, \quad t = 1, \dots, T \\
 \epsilon_t &= \mathbf{y}_t - \mu_t \\
 \alpha &\sim \mathcal{N}(0, 1000) \\
 \phi_p &\sim \mathcal{N}(0, 1000), \quad p = 1, \dots, P \\
 \delta &\sim \mathcal{N}(0, 1000) \\
 \sigma_t^2 &= \omega + \sum_{q=1}^Q \theta_q \epsilon_{t-q}^2, \quad t = 2, \dots, T
 \end{aligned}$$

$$\omega \sim \mathcal{HC}(25)$$

$$\theta_q \sim \mathcal{U}(0,1), \quad q = 1, \dots, Q$$

8.2. Data

```

data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
T <- length(y)
L.P <- c(1,5,20) #Autoregressive lags
L.Q <- c(1,2) #Volatility lags
P <- length(L.P) #Autoregressive order
Q <- length(L.Q) #Volatility order
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, phi=rep(0,P), delta=0, omega=0,
  theta=rep(0,Q)))
pos.alpha <- grep("alpha", parm.names)
pos.phi <- grep("phi", parm.names)
pos.delta <- grep("delta", parm.names)
pos.omega <- grep("omega", parm.names)
pos.theta <- grep("theta", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  phi <- runif(Data$P,-1,1)
  delta <- rnorm(1)
  omega <- rhalfcauchy(1,5)
  theta <- runif(Data$Q, 1e-10, 1-1e-5)
  return(c(alpha, phi, delta, omega, theta))
}
MyData <- list(L.P=L.P, L.Q=L.Q, PGF=PGF, P=P, Q=Q, T=T, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.phi=pos.phi,
  pos.delta=pos.delta, pos.omega=pos.omega, pos.theta=pos.theta, y=y)

```

8.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  phi <- parm[Data$pos.phi]
  delta <- parm[Data$pos.delta]
  omega <- interval(parm[Data$pos.omega], 1e-100, Inf)
  parm[Data$pos.omega] <- omega
  theta <- interval(parm[Data$pos.theta], 1e-10, 1-1e-5)
  parm[Data$pos.theta] <- theta
  ### Log-Prior

```

```

alpha.prior <- dnormv(alpha, 0, 1000, log=TRUE)
phi.prior <- sum(dnormv(phi, 0, 1000, log=TRUE))
delta.prior <- dnormv(delta, 0, 1000, log=TRUE)
omega.prior <- dhalfcauchy(omega, 25, log=TRUE)
theta.prior <- sum(dunif(theta, 1e-10, 1-1e-5, log=TRUE))
### Log-Likelihood
mu <- rep(alpha, Data$T)
for (p in 1:Data$P)
  mu[-c(1:Data$L.P[p])] <- mu[-c(1:Data$L.P[p])] +
    phi[p]*Data$y[1:(Data$T-Data$L.P[p])]
epsilon <- Data$y - mu
sigma2 <- rep(omega, Data$T)
for (q in 1:Data$Q)
  sigma2[-c(1:Data$L.Q[q])] <- sigma2[-c(1:Data$L.Q[q])] +
    theta[q]*epsilon[1:(Data$T-Data$L.Q[q])]^2
mu <- mu + delta*sigma2
LL <- sum(dnormv(Data$y[-c(1:Data$L.P[Data$P])],
  mu[-c(1:Data$L.P[Data$P])], sigma2[-c(1:Data$L.P[Data$P])],
  log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + phi.prior + delta.prior + omega.prior +
  theta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnormv(length(mu), mu, sigma2), parm=parm)
return(Modelout)
}

```

8.4. Initial Values

```
Initial.Values <- c(rep(0,P+2), 1, rep(0.5,Q))
```

9. AR(p)-GARCH(1,1)

9.1. Form

$$\begin{aligned}
 \mathbf{y}_t &\sim \mathcal{N}(\mu_t, \sigma_t^2), \quad t = 1, \dots, T \\
 \mu_t &= \alpha + \sum_{p=1}^P \phi_p \mathbf{y}_{t-p}, \quad t = 1, \dots, T \\
 \epsilon_t &= \mathbf{y}_t - \mu_t \\
 \alpha &\sim \mathcal{N}(0, 1000) \\
 \phi_p &\sim \mathcal{N}(0, 1000), \quad p = 1, \dots, P
 \end{aligned}$$

$$\sigma_t^2 = \theta_1 + \theta_2 \epsilon_{t-1}^2 + \theta_3 \sigma_{t-1}^2$$

$$\omega \sim \mathcal{HC}(25)$$

$$\theta_k = \frac{1}{1 + \exp(-\theta_k)}, \quad k = 1, \dots, 3$$

$$\theta_k \sim \mathcal{N}(0, 1000) \in [-10, 10], \quad k = 1, \dots, 3$$

9.2. Data

```

data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
T <- length(y)
L <- c(1,5,20) #Autoregressive lags
P <- length(L) #Autoregressive order
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, phi=rep(0,P), omega=0,
  theta=rep(0,2)))
pos.alpha <- grep("alpha", parm.names)
pos.phi <- grep("phi", parm.names)
pos.omega <- grep("omega", parm.names)
pos.theta <- grep("theta", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  phi <- runif(Data$P,-1,1)
  omega <- rhalfcauchy(1,5)
  theta <- runif(2, 1e-10, 1-1e-5)
  return(c(alpha, phi, omega, theta))
}
MyData <- list(L=L, P=P, PGF=PGF, T=T, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.phi=pos.phi,
  pos.omega=pos.omega, pos.theta=pos.theta, y=y)

```

9.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  phi <- parm[Data$pos.phi]
  omega <- interval(parm[Data$pos.omega], 1e-100, Inf)
  parm[Data$pos.omega] <- omega
  theta <- interval(parm[Data$pos.theta], 1e-10, 1-1e-5)
  if(sum(theta) >= 1) theta[2] <- 1 - 1e-5 - theta[1]
  parm[Data$pos.theta] <- theta
  ### Log-Prior
  alpha.prior <- dnorm(alpha, 0, 1000, log=TRUE)

```

```

phi.prior <- sum(dnormv(phi, 0, 1000, log=TRUE))
omega.prior <- dhalfcauchy(omega, 25, log=TRUE)
theta.prior <- sum(dunif(theta, 0, 1, log=TRUE))
### Log-Likelihood
mu <- rep(alpha, Data$T)
for (p in 1:Data$P)
  mu[-c(1:Data$L[p])] <- mu[-c(1:Data$L[p])] +
    phi[p]*Data$y[1:(Data$T-Data$L[p])]
epsilon <- Data$y - mu
sigma2 <- c(omega, omega + theta[1]*epsilon[-Data$T]^2)
sigma2[-1] <- sigma2[-1] + theta[2]*sigma2[-Data$T]
LL <- sum(dnormv(Data$y[-c(1:Data$L[Data$P])],
  mu[-c(1:Data$L[Data$P])], sigma2[-c(1:Data$L[Data$P])],
  log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + phi.prior + omega.prior + theta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnormv(length(mu), mu, sigma2), parm=parm)
return(Modelout)
}

```

9.4. Initial Values

```
Initial.Values <- c(0, rep(0,P), rep(0.4,3))
```

10. AR(p)-GARCH(1,1)-M

10.1. Form

$$\mathbf{y}_t \sim \mathcal{N}(\mu_t, \sigma_t^2), \quad t = 1, \dots, T$$

$$\mu_t = \alpha + \sum_{p=1}^P \phi_p \mathbf{y}_{t-p} + \delta \sigma_{t-1}^2, \quad t = 1, \dots, (T+1)$$

$$\epsilon_t = \mathbf{y}_t - \mu_t$$

$$\alpha \sim \mathcal{N}(0, 1000)$$

$$\phi_p \sim \mathcal{N}(0, 1000), \quad p = 1, \dots, P$$

$$\sigma_t^2 = \omega + \theta_1 \epsilon_{t-1}^2 + \theta_2 \sigma_{t-1}^2$$

$$\omega \sim \mathcal{HC}(25)$$

$$\theta_k \sim \mathcal{U}(0, 1), \quad k = 1, \dots, 2$$

10.2. Data

```

data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
T <- length(y)
L <- c(1,5,20) #Autoregressive lags
P <- length(L) #Autoregressive order
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, delta=0, phi=rep(0,P), omega=0,
  theta=rep(0,2)))
pos.alpha <- grep("alpha", parm.names)
pos.delta <- grep("delta", parm.names)
pos.phi <- grep("phi", parm.names)
pos.omega <- grep("omega", parm.names)
pos.theta <- grep("theta", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  delta <- rnorm(1)
  phi <- runif(Data$P,-1,1)
  omega <- rhalfcauchy(1,5)
  theta <- runif(2, 1e-10, 1-1e-5)
  return(c(alpha, delta, phi, omega, theta))
}
MyData <- list(L=L, P=P, PGF=PGF, T=T, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.delta=pos.delta,
  pos.phi=pos.phi, pos.omega=pos.omega, pos.theta=pos.theta, y=y)

```

10.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  delta <- parm[Data$pos.delta]
  phi <- parm[Data$pos.phi]
  omega <- interval(parm[Data$pos.omega], 1e-100, Inf)
  parm[Data$pos.omega] <- omega
  theta <- interval(parm[Data$pos.theta], 1e-10, 1-1e-5)
  if(sum(theta) >= 1) theta[2] <- 1 - 1e-5 - theta[1]
  parm[Data$pos.theta] <- theta
  ### Log-Prior
  alpha.prior <- dnormv(alpha, 0, 1000, log=TRUE)
  delta.prior <- dnormv(delta, 0, 1000, log=TRUE)
  phi.prior <- sum(dnormv(phi, 0, 1000, log=TRUE))
  omega.prior <- dhalfcauchy(omega, 25, log=TRUE)
  theta.prior <- sum(dunif(theta, 0, 1, log=TRUE))
}

```



```

### Log-Likelihood
mu <- rep(alpha, Data$T)
for (p in 1:Data$P)
  mu[-c(1:Data$L[p])] <- mu[-c(1:Data$L[p])] +
    phi[p]*Data$y[1:(Data$T-Data$L[p])]
epsilon <- Data$y - mu
sigma2 <- c(omega, omega + theta[1]*epsilon[-Data$T]^2)
sigma2[-1] <- sigma2[-1] + theta[2]*sigma2[-Data$T]
mu <- mu + delta*sigma2
LL <- sum(dnormv(Data$y[-c(1:Data$L[Data$P])],
  mu[-c(1:Data$L[Data$P])], sigma2[-c(1:Data$L[Data$P])],
  log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + delta.prior + phi.prior + omega.prior +
  theta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnormv(length(mu), mu, sigma2), parm=parm)
return(Modelout)
}

```

10.4. Initial Values

```
Initial.Values <- c(rep(0,2), rep(0,P), rep(0.4,3))
```

11. AR(p)-TARCH(q)

11.1. Form

$$\begin{aligned}
 \mathbf{y}_t &\sim \mathcal{N}(\mu_t, \sigma_t^2), \quad t = 2, \dots, T \\
 \mu_t &= \alpha + \phi_{p=1}^P \mathbf{y}_{t-p}, \quad t = (p+1), \dots, T \\
 \epsilon &= \mathbf{y} - \mu \\
 \delta &= (\epsilon > 0) \times 1 \\
 \sigma_t^2 &= \omega + \sum_{q=1}^Q \theta_{q,1} \delta_{t-1} \epsilon_{t-1}^2 + \theta_{q,2} (1 - \delta_{t-1}) \epsilon_{t-1}^2 \\
 \alpha &\sim \mathcal{N}(0, 1000) \\
 \phi_p &\sim \mathcal{N}(0, 1000), \quad p = 1, \dots, P \\
 \omega &\sim \mathcal{HC}(25) \\
 \theta_{q,j} &\sim \mathcal{U}(0, 1), \quad q = 1, \dots, Q, \quad j = 1, \dots, 2
 \end{aligned}$$

11.2. Data

```

data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
T <- length(y)
L.P <- c(1,5,20) #Autoregressive lags
L.Q <- c(1,2) #Volatility lags
P <- length(L.P) #Autoregressive order
Q <- length(L.Q) #Volatility order
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, phi=rep(0,P), omega=0,
  theta=matrix(0,Q,2)))
pos.alpha <- grep("alpha", parm.names)
pos.phi <- grep("phi", parm.names)
pos.omega <- grep("omega", parm.names)
pos.theta <- grep("theta", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  phi <- runif(Data$P,-1,1)
  omega <- rhalfcauchy(1,5)
  theta <- runif(Data$Q*2, 1e-10, 1-1e-5)
  return(c(alpha, phi, omega, theta))
}
MyData <- list(L.P=L.P, L.Q=L.Q, PGF=PGF, P=P, Q=Q, T=T, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.phi=pos.phi,
  pos.omega=pos.omega, pos.theta=pos.theta, y=y)

```

11.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  phi <- parm[Data$pos.phi]
  omega <- interval(parm[Data$pos.omega], 1e-100, Inf)
  parm[Data$pos.omega] <- omega
  theta <- matrix(interval(parm[Data$pos.theta], 1e-10, 1-1e-5), Data$Q,
    2)
  parm[Data$pos.theta] <- as.vector(theta)
  ### Log-Prior
  alpha.prior <- dnormv(alpha, 0, 1000, log=TRUE)
  phi.prior <- sum(dnormv(phi, 0, 1000, log=TRUE))
  omega.prior <- dhalfcauchy(omega, 25, log=TRUE)
  theta.prior <- sum(dunif(theta, 1e-10, 1-1e-5, log=TRUE))
  ### Log-Likelihood
  mu <- rep(alpha, Data$T)

```

```

for (p in 1:Data$P)
  mu[-c(1:Data$L.P[p])] <- mu[-c(1:Data$L.P[p])] +
    phi[p]*Data$y[1:(Data$T-Data$L.P[p])]
epsilon <- Data$y - mu
delta <- (epsilon > 0) * 1
sigma2 <- rep(omega, Data$T)
for (q in 1:Data$Q)
  sigma2[-c(1:Data$L.Q[q])] <- sigma2[-c(1:Data$L.Q[q])] +
    delta[1:(Data$T-Data$L.Q[q])] * theta[q,1] *
    epsilon[1:(Data$T-Data$L.Q[q])]^2 +
    (1 - delta[1:(Data$T-Data$L.Q[q])]) * theta[q,2] *
    epsilon[1:(Data$T-Data$L.Q[q])]^2
LL <- sum(dnormv(Data$y[-c(1:Data$L.P[Data$P])],
  mu[-c(1:Data$L.P[Data$P])], sigma2[-c(1:Data$L.P[Data$P])],
  log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + phi.prior + omega.prior + theta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnormv(length(mu), mu, sigma2), parm=parm)
return(Modelout)
}

```

11.4. Initial Values

```
Initial.Values <- c(0, rep(0,P), 1, rep(0.5,Q*2))
```

12. AR(p)-TARCH(q)-M

12.1. Form

$$\begin{aligned}
 \mathbf{y}_t &\sim \mathcal{N}(\mu_t, \sigma_t^2), \quad t = 2, \dots, T \\
 \mu_t &= \alpha + \phi_{p=1}^P \mathbf{y}_{t-p} + \delta_{t-1} \gamma_1 \sigma_{t-1}^2 + (1 - \delta_{t-1}) \gamma_2 \sigma_{t-1}^2, \quad t = (p+1), \dots, T \\
 \epsilon &= \mathbf{y} - \mu \\
 \delta &= (\epsilon > 0) \times 1 \\
 \sigma_t^2 &= \omega + \sum_{q=1}^Q \theta_{q,1} \delta_{t-1} \epsilon_{t-1}^2 + \theta_{q,2} (1 - \delta_{t-1}) \epsilon_{t-1}^2 \\
 \alpha &\sim \mathcal{N}(0, 1000) \\
 \gamma_k &\sim \mathcal{N}(0, 1000), \quad k = 1, \dots, 2 \\
 \phi_p &\sim \mathcal{N}(0, 1000), \quad p = 1, \dots, P \\
 \omega &\sim \mathcal{HC}(25)
 \end{aligned}$$

$$\theta_{q,j} \sim \mathcal{U}(0,1), \quad q = 1 \dots, Q, \quad j = 1, \dots, 2$$

12.2. Data

```

data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
T <- length(y)
L.P <- c(1,5,20) #Autoregressive lags
L.Q <- c(1,2) #Volatility lags
P <- length(L.P) #Autoregressive order
Q <- length(L.Q) #Volatility order
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, gamma=rep(0,2), phi=rep(0,P),
  omega=0, theta=matrix(0,Q,2)))
pos.alpha <- grep("alpha", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.phi <- grep("phi", parm.names)
pos.omega <- grep("omega", parm.names)
pos.theta <- grep("theta", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  gamma <- rnorm(2)
  phi <- runif(Data$P,-1,1)
  omega <- rhalfcauchy(1,5)
  theta <- runif(Data$Q*2, 1e-10, 1-1e-5)
  return(c(alpha, gamma, phi, omega, theta))
}
MyData <- list(L.P=L.P, L.Q=L.Q, PGF=PGF, P=P, Q=Q, T=T, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.gamma=pos.gamma,
  pos.phi=pos.phi, pos.omega=pos.omega, pos.theta=pos.theta, y=y)

```

12.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  gamma <- parm[Data$pos.gamma]
  phi <- parm[Data$pos.phi]
  omega <- interval(parm[Data$pos.omega], 1e-100, Inf)
  parm[Data$pos.omega] <- omega
  theta <- matrix(interval(parm[Data$pos.theta], 1e-10, 1-1e-5), Data$Q,
    2)
  parm[Data$pos.theta] <- as.vector(theta)
  ### Log-Prior
  alpha.prior <- dnormv(alpha, 0, 1000, log=TRUE)

```

```

gamma.prior <- sum(dnormv(gamma, 0, 1000, log=TRUE))
phi.prior <- sum(dnormv(phi, 0, 1000, log=TRUE))
omega.prior <- dhalfcauchy(omega, 25, log=TRUE)
theta.prior <- sum(dunif(theta, 1e-10, 1-1e-5, log=TRUE))
### Log-Likelihood
mu <- rep(alpha, Data$T)
for (p in 1:Data$P)
  mu[-c(1:Data$L.P[p])] <- mu[-c(1:Data$L.P[p])] +
    phi[p]*Data$y[1:(Data$T-Data$L.P[p])]
epsilon <- Data$y - mu
delta <- (epsilon > 0) * 1
sigma2 <- rep(omega, Data$T)
for (q in 1:Data$Q)
  sigma2[-c(1:Data$L.Q[q])] <- sigma2[-c(1:Data$L.Q[q])] +
    delta[1:(Data$T-Data$L.Q[q])] * theta[q,1] *
    epsilon[1:(Data$T-Data$L.Q[q])]^2 +
    (1 - delta[1:(Data$T-Data$L.Q[q])]) * theta[q,2] *
    epsilon[1:(Data$T-Data$L.Q[q])]^2
mu <- mu + delta*gamma[1]*sigma2 + (1 - delta)*gamma[2]*sigma2
LL <- sum(dnormv(Data$y[-c(1:Data$L.P[Data$P])],
  mu[-c(1:Data$L.P[Data$P])], sigma2[-c(1:Data$L.P[Data$P])],
  log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + gamma.prior + phi.prior + omega.prior +
  theta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnormv(length(mu), mu, sigma2), parm=parm)
return(Modelout)
}

```

12.4. Initial Values

```
Initial.Values <- c(rep(0,3), rep(0,P), 1, rep(0.5,Q*2))
```

13. Autoregressive Moving Average, ARMA(p,q)

13.1. Form

$$\mathbf{y}_t \sim \mathcal{N}(\mu_t, \sigma^2), \quad t = 1, \dots, T$$

$$\mu_t = \alpha + \sum_{p=1}^P \phi_p \mathbf{y}_{t-p} + \sum_{q=1}^Q \theta_q \epsilon_{t-q}$$

$$\epsilon_t = \mathbf{y}_t - \mu_t$$

$$\alpha \sim \mathcal{N}(0, 1000)$$

$$\phi_p \sim \mathcal{N}(0, 1000), \quad p = 1, \dots, P$$

$$\sigma \sim \mathcal{HC}(25)$$

$$\theta_q \sim \mathcal{N}(0, 1000), \quad q = 1, \dots, Q$$

13.2. Data

```

data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
T <- length(y)
L.P <- c(1,5,20) #Autoregressive lags
L.Q <- c(1,2) #Moving average lags
P <- length(L.P) #Autoregressive order
Q <- length(L.Q) #Moving average order
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, phi=rep(0,P), sigma=0,
  theta=rep(0,Q)))
pos.alpha <- grep("alpha", parm.names)
pos.phi <- grep("phi", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.theta <- grep("theta", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  phi <- runif(Data$P,-1,1)
  sigma <- rhalfcauchy(1,5)
  theta <- rnorm(Data$Q)
  return(c(alpha, phi, sigma, theta))
}
MyData <- list(L.P=L.P, L.Q=L.Q, PGF=PGF, P=P, Q=Q, T=T, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.phi=pos.phi,
  pos.sigma=pos.sigma, pos.theta=pos.theta, y=y)

```

13.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  phi <- parm[Data$pos.phi]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  theta <- parm[Data$pos.theta]
  ### Log-Prior
  alpha.prior <- dnormv(alpha, 0, 1000, log=TRUE)

```

```

phi.prior <- sum(dnormv(phi, 0, 1000, log=TRUE))
sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
theta.prior <- sum(dnormv(theta, 0, 1000, log=TRUE))
### Log-Likelihood
mu <- rep(alpha, Data$T)
for (p in 1:Data$P)
  mu[-c(1:Data$L.P[p])] <- mu[-c(1:Data$L.P[p])] +
    phi[p]*Data$y[1:(Data$T-Data$L.P[p])]
epsilon <- Data$y - mu
for (q in 1:Data$Q)
  mu[-c(1:Data$L.Q[q])] <- mu[-c(1:Data$L.Q[q])] +
    theta[q]*epsilon[1:(Data$T-Data$L.Q[q])]
LL <- sum(dnorm(Data$y[-c(1:Data$L.P[Data$P])],
  mu[-c(1:Data$L.P[Data$P])], sigma, log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + phi.prior + sigma.prior + theta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

13.4. Initial Values

```
Initial.Values <- c(0, rep(0,P), 1, rep(0,Q))
```

14. Beta Regression

14.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{BETA}(a, b) \\
 a &= \mu\phi \\
 b &= (1 - \mu)\phi \\
 \mu &= \Phi(\beta_1 + \beta_2\mathbf{x}), \quad \mu \in (0, 1) \\
 \beta_j &\sim \mathcal{N}(0, 10), \quad j = 1, \dots, J \\
 \phi &\sim \mathcal{HC}(25)
 \end{aligned}$$

where Φ is the normal CDF.

14.2. Data

```

N <- 100
x <- runif(N)

```

```

y <- rbeta(N, (0.5-0.2*x)*3, (1-(0.5-0.2*x))*3) mon.names <- "LP"
parm.names <- c("beta[1]", "beta[2]", "phi")
pos.beta <- grep("beta", parm.names)
pos.phi <- grep("phi", parm.names)
PGF <- function(Data) return(c(rnormv(2,0,10), rhalfcauchy(1,5)))
MyData <- list(PGF=PGF, mon.names=mon.names, parm.names=parm.names,
              pos.beta=pos.beta, pos.phi=pos.phi, x=x, y=y)

```

14.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  parm[Data$pos.phi] <- phi <- interval(parm[Data$pos.phi], 1e-100, Inf)
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 10, log=TRUE))
  phi.prior <- dhalfcauchy(phi, 25, log=TRUE)
  ### Log-Likelihood
  mu <- interval(pnorm(beta[1] + beta[2]*Data$x), 0.001, 0.999,
                reflect=FALSE)
  a <- mu * phi
  b <- (1 - mu) * phi
  LL <- sum(dbeta(Data$y, a, b, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + phi.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
                  yhat=rbeta(length(mu), a, b), parm=parm)
  return(Modelout)
}

```

14.4. Initial Values

```
Initial.Values <- c(rep(0,2), 0.01)
```

15. Beta-Binomial

15.1. Form

$$y_i \sim \text{BIN}(\mathbf{n}_i, \pi_i), \quad i = 1, \dots, N$$

$$\pi_i \sim \text{BETA}(\alpha, \beta) \in [0.001, 0.999]$$

15.2. Data

```

N <- 20
n <- round(runif(N, 50, 100))
y <- round(runif(N, 1, 10))
mon.names <- "LP"
parm.names <- as.parm.names(list(pi=rep(0,N)))
PGF <- function(Data) {
  pi <- rbeta(Data$N,1,1)
  return(pi)
}
MyData <- list(N=N, PGF=PGF, mon.names=mon.names, n=n,
  parm.names=parm.names, y=y)

```

15.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  parm[1:Data$N] <- pi <- interval(parm[1:Data$N], 0.001, 0.999)
  ### Log-Prior
  pi.prior <- sum(dbeta(pi, 1, 1, log=TRUE))
  ### Log-Likelihood
  LL <- sum(dbinom(Data$y, Data$n, pi, log=TRUE))
  ### Log-Posterior
  LP <- LL + pi.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rbinom(Data$N, Data$n, pi), parm=parm)
  return(Modelout)
}

```

15.4. Initial Values

```
Initial.Values <- c(rep(0.5,N))
```

16. Binary Logit

16.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{BERN}(\eta) \\
 \eta &= \frac{1}{1 + \exp(-\mu)} \\
 \mu &= \mathbf{X}\beta
 \end{aligned}$$

$$\beta_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

16.2. Data

```

data(demonsnacks)
J <- 3
y <- ifelse(demonsnacks$Calories <= 137, 0, 1)
X <- cbind(1, as.matrix(demonsnacks[,c(7,8)]))
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J)))
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  return(beta)
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, y=y)

```

16.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[1:Data$J]
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  eta <- invlogit(mu)
  LL <- sum(dbern(Data$y, eta, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rbern(length(eta), eta), parm=parm)
  return(Modelout)
}

```

16.4. Initial Values

```
Initial.Values <- rep(0,J)
```

17. Binary Log-Log Link Mixture

A weighted mixture of the log-log and complementary log-log link functions is used, where α is the weight. Since the log-log and complementary log-log link functions are asymmetric

(as opposed to the symmetric logit and probit link functions), it may be unknown *a priori* whether the log-log or complementary log-log will perform better.

17.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{BERN}(\eta) \\ \eta &= \alpha \exp(-\exp(\mu)) + (1 - \alpha)(1 - \exp(-\exp(\mu))) \\ \mu &= \mathbf{X}\beta \\ \alpha &\sim \mathcal{U}(0, 1) \\ \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \end{aligned}$$

17.2. Data

```
N <- 100
J <- 3
X <- cbind(1, matrix(rnorm(N*(J-1),N,J-1)))
alpha <- runif(1)
beta <- rnorm(J)
mu <- tcrossprod(X, t(beta))
eta <- alpha*invloglog(mu) + (1-alpha)*invcloglog(mu)
y <- rbern(N, eta)
mon.names <- c("LP","alpha")
parm.names <- as.parm.names(list(beta=rep(0,J), logit.alpha=0))
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  logit.alpha <- rnorm(1)
  return(c(beta, logit.alpha))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, y=y)
```

17.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  parm[Data$J+1] <- alpha <- interval(parm[Data$J+1], -700, 700)
  beta <- parm[1:Data$J]
  ### Log-Prior
  alpha.prior <- dunif(alpha, 0, 1, log=TRUE)
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
```

```

eta <- alpha*invloglog(mu) + (1-alpha)*invcloglog(mu)
LL <- sum(dbern(Data$y, eta, log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,alpha),
  yhat=rbern(length(eta), eta), parm=parm)
return(Modelout)
}

```

17.4. Initial Values

```
Initial.Values <- c(rep(0,J), 0)
```

18. Binary Probit

18.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{BERN}(\mathbf{p}) \\
 \mathbf{p} &= \phi(\boldsymbol{\mu}) \\
 \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \in [-10, 10] \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J
 \end{aligned}$$

where ϕ is the CDF of the standard normal distribution, and $J=3$.

18.2. Data

```

data(demonsnacks)
J <- 3
y <- ifelse(demonsnacks$Calories <= 137, 0, 1)
X <- cbind(1, as.matrix(demonsnacks[,c(7,8)]))
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J)))
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  return(beta)
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, y=y)

```

18.3. Model

```
Model <- function(parm, Data)
```

```

{
### Parameters
beta <- parm[1:Data$J]
### Log-Prior
beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
### Log-Likelihood
mu <- tcrossprod(Data$X, t(beta))
mu <- interval(mu, -10, 10, reflect=FALSE)
p <- pnorm(mu)
LL <- sum(dbern(Data$y, p, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rbern(length(p), p), parm=parm)
return(Modelout)
}

```

18.4. Initial Values

```
Initial.Values <- rep(0,J)
```

19. Binary Robit

19.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{BERN}(\mathbf{p}) \\
 \mathbf{p} &= \mathbf{T}_\nu(\boldsymbol{\mu}) \\
 \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \in [-10, 10] \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \nu &\sim \mathcal{U}(5, 10)
 \end{aligned}$$

where \mathbf{T}_ν is the CDF of the standard t-distribution with ν degrees of freedom.

19.2. Data

```

data(demonsnacks)
y <- ifelse(demonsnacks$Calories <= 137, 0, 1)
X <- cbind(1, as.matrix(demonsnacks[,c(7,8)]))
J <- ncol(X)
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), nu=0))
pos.beta <- grep("beta", parm.names)

```

```

pos.nu <- grep("nu", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  nu <- runif(1,5,10)
  return(c(beta, nu))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.nu=pos.nu, y=y)

```

19.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  parm[Data$pos.nu] <- nu <- interval(parm[Data$pos.nu], 1e-100, 1000)
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  nu.prior <- dunif(nu, 1e-100, 1000, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  mu <- interval(mu, -10, 10, reflect=FALSE)
  p <- pst(mu, nu=nu)
  LL <- sum(dbern(Data$y, p, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + nu.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rbern(length(p), p), parm=parm)
  return(Modelout)
}

```

19.4. Initial Values

```
Initial.Values <- c(rep(0,J), 5)
```

20. Binomial Logit

20.1. Form

$$\mathbf{y} \sim \mathcal{BIN}(\mathbf{p}, \mathbf{n})$$

$$\mathbf{p} = \frac{1}{1 + \exp(-\mu)}$$

$$\mu = \beta_1 + \beta_2 \mathbf{x}$$

$$\beta_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

20.2. Data

```
#10 Trials
exposed <- c(100,100,100,100,100,100,100,100,100,100)
deaths <- c(10,20,30,40,50,60,70,80,90,100)
dose <- c(1,2,3,4,5,6,7,8,9,10)
J <- 2 #Number of parameters
mon.names <- "LP"
parm.names <- c("beta[1]","beta[2]")
PGF <- function(Data) return(rnormv(Data$J,0,1000))
MyData <- list(J=J, PGF=PGF, n=exposed, mon.names=mon.names,
  parm.names=parm.names, x=dose, y=deaths)
```

20.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[1:Data$J]
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  ### Log-Likelihood
  mu <- beta[1] + beta[2]*Data$x
  p <- invlogit(mu)
  LL <- sum(dbinom(Data$y, Data$n, p, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rbinom(length(p), Data$n, p), parm=parm)
  return(Modelout)
}
```

20.4. Initial Values

```
Initial.Values <- rep(0,J)
```

21. Binomial Probit

21.1. Form

$$\mathbf{y} \sim \mathcal{BIN}(\mathbf{p}, \mathbf{n})$$

$$\mathbf{p} = \phi(\boldsymbol{\mu})$$

$$\mu = \beta_1 + \beta_2 \mathbf{x} \in [-10, 10]$$

$$\beta_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

where ϕ is the CDF of the standard normal distribution, and $J=2$.

21.2. Data

```
#10 Trials
exposed <- c(100,100,100,100,100,100,100,100,100,100)
deaths <- c(10,20,30,40,50,60,70,80,90,100)
dose <- c(1,2,3,4,5,6,7,8,9,10)
J <- 2 #Number of parameters
mon.names <- "LP"
parm.names <- c("beta[1]","beta[2]")
PGF <- function(Data) return(rnormv(Data$J,0,1000))
MyData <- list(J=J, PGF=PGF, n=exposed, mon.names=mon.names,
  parm.names=parm.names, x=dose, y=deaths)
```

21.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[1:Data$J]
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  ### Log-Likelihood
  mu <- beta[1] + beta[2]*Data$x
  mu <- interval(mu, -10, 10, reflect=FALSE)
  p <- pnorm(mu)
  LL <- sum(dbinom(Data$y, Data$n, p, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rbinom(length(p), Data$n, p), parm=parm)
  return(Modelout)
}
```

21.4. Initial Values

```
Initial.Values <- rep(0,J)
```

22. Binomial Robit

22.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{BLN}(\mathbf{p}, \mathbf{n}) \\ \mathbf{p} &= \mathbf{T}_\nu(\mu) \\ \mu &= \beta_1 + \beta_2 \mathbf{x} \in [-10, 10] \\ \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\ \nu &\sim \mathcal{U}(5, 10) \end{aligned}$$

where \mathbf{T}_ν is the CDF of the standard t-distribution with ν degrees of freedom.

22.2. Data

```
#10 Trials
exposed <- c(100,100,100,100,100,100,100,100,100,100)
deaths <- c(10,20,30,40,50,60,70,80,90,100)
dose <- c(1,2,3,4,5,6,7,8,9,10)
J <- 2 #Number of parameters
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,2), nu=0))
PGF <- function(Data) return(c(rnormv(Data$J,0,1000), runif(1,5,10)))
MyData <- list(J=J, PGF=PGF, n=exposed, mon.names=mon.names,
  parm.names=parm.names, x=dose, y=deaths)
```

22.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[1:Data$J]
  parm[Data$J+1] <- nu <- interval(parm[Data$J+1], 5, 10)
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  nu.prior <- dunif(nu, 5, 10, log=TRUE)
  ### Log-Likelihood
  mu <- beta[1] + beta[2]*Data$x
  mu <- interval(mu, -10, 10, reflect=FALSE)
  p <- pst(mu, nu=nu)
  LL <- sum(dbinom(Data$y, Data$n, p, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + nu.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rbinom(length(p), Data$n, p), parm=parm)
  return(Modelout)
}
```

22.4. Initial Values

```
Initial.Values <- c(rep(0,J), 5)
```

23. Change Point Regression

This example uses a popular variant of the stagnant water data set.

23.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{N}(\mu, \sigma^2) \\ \mu &= \alpha + \beta_1 \mathbf{x} + \beta_2 (\mathbf{x} - \theta) [(\mathbf{x} - \theta) > 0] \\ \alpha &\sim \mathcal{N}(0, 1000) \\ \beta &\sim \mathcal{N}(0, 1000) \\ \sigma &\sim \mathcal{HC}(25) \\ \theta &\sim \mathcal{U}(-1.3, 1.1) \end{aligned}$$

23.2. Data

```
N <- 29
y <- c(1.12, 1.12, 0.99, 1.03, 0.92, 0.90, 0.81, 0.83, 0.65, 0.67, 0.60,
      0.59, 0.51, 0.44, 0.43, 0.43, 0.33, 0.30, 0.25, 0.24, 0.13, -0.01,
      -0.13, -0.14, -0.30, -0.33, -0.46, -0.43, -0.65)
x <- c(-1.39, -1.39, -1.08, -1.08, -0.94, -0.80, -0.63, -0.63, -0.25, -0.25,
      -0.12, -0.12, 0.01, 0.11, 0.11, 0.11, 0.25, 0.25, 0.34, 0.34, 0.44,
      0.59, 0.70, 0.70, 0.85, 0.85, 0.99, 0.99, 1.19)
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, beta=rep(0,2), sigma=0, theta=0))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.theta <- grep("theta", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  beta <- rnorm(2)
  sigma <- runif(1)
  theta <- runif(1)
  return(c(alpha, beta, sigma, theta))
}
MyData <- list(N=N, PGF=PGF, mon.names=mon.names, parm.names=parm.names,
  pos.alpha=pos.alpha, pos.beta=pos.beta, pos.sigma=pos.sigma,
  pos.theta=pos.theta, x=x, y=y)
```

23.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  theta <- interval(parm[Data$pos.theta], -1.3, 1.1)
  parm[Data$pos.theta] <- theta
  ### Log-Prior
  alpha.prior <- dnormv(alpha, 0, 1000, log=TRUE)
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  theta.prior <- dunif(theta, -1.3, 1.1, log=TRUE)
  ### Log-Likelihood
  mu <- alpha + beta[1]*x + beta[2]*(x - theta)*(x - theta) > 0
  LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + alpha.prior + beta.prior + sigma.prior + theta.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

23.4. Initial Values

```
Initial.Values <- c(0.2, -0.45, 0, 0.2, 0)
```

24. Cluster Analysis, Confirmatory (CCA)

This is a parametric, model-based, cluster analysis, also called a finite mixture model or latent class cluster analysis, where the number of clusters C is fixed. When the number of clusters is unknown, exploratory cluster analysis should be used. The record-level cluster membership parameter vector, θ , is a vector of discrete parameters. Discrete parameters are not supported in all algorithms. The example below is updated with the Slice sampler.

24.1. Form

$$\begin{aligned}
 \mathbf{Y}_{i,j} &\sim \mathcal{N}(\mu_{\theta[i],j}, \sigma_{\theta[i]}^2), \quad i = 1, \dots, N, \quad j = 1, \dots, J \\
 \theta_i &\sim \mathcal{CAT}(\pi_{1:C}), \quad i = 1, \dots, N \\
 \pi_{1:C} &\sim \mathcal{D}(\alpha_{1:C}) \\
 \alpha_c &= 1
 \end{aligned}$$

$$\begin{aligned}\mu_{c,j} &\sim \mathcal{N}(0, \nu_c^2), & c = 1, \dots, C, & \quad j = 1, \dots, J \\ \sigma_c &\sim \mathcal{HC}(25), & c = 1, \dots, C \\ \nu_c &\sim \mathcal{HC}(25), & c = 1, \dots, C\end{aligned}$$

24.2. Data

```
data(demonsnacks)
Y <- as.matrix(log(demonsnacks + 1))
N <- nrow(Y)
J <- ncol(Y)
for (j in 1:J) Y[,j] <- CenterScale(Y[,j])
C <- 3 #Number of clusters
alpha <- rep(1,C) #Prior probability of cluster proportion
mon.names <- c("LP", paste("pi[", 1:C, "]", sep=""))
parm.names <- as.parm.names(list(theta=rep(0,N), mu=matrix(0,C,J),
  nu=rep(0,C), sigma=rep(0,C)))
pos.theta <- grep("theta", parm.names)
pos.mu <- grep("mu", parm.names)
pos.nu <- grep("nu", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  theta <- rcat(Data$N, p=rep(1/Data$C, Data$C))
  mu <- rnorm(Data$J*Data$J)
  nu <- runif(Data$C)
  sigma <- runif(Data$C)
  return(c(theta, mu, nu, sigma))
}
MyData <- list(C=C, J=J, N=N, PGF=PGF, Y=Y, alpha=alpha,
  mon.names=mon.names, parm.names=parm.names, pos.theta=pos.theta,
  pos.mu=pos.mu, pos.nu=pos.nu, pos.sigma=pos.sigma)
```

24.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  theta <- parm[Data$pos.theta]
  mu <- matrix(parm[Data$pos.mu], Data$C, Data$J)
  parm[Data$pos.nu] <- nu <- interval(parm[Data$pos.nu], 1e-100, Inf)
  pi <- rep(0, Data$C)
  tab <- table(theta)
  pi[as.numeric(names(tab))] <- as.vector(tab)
  pi <- pi / sum(pi)
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
}
```

```

### Log-Prior
theta.prior <- sum(dcat(theta, pi, log=TRUE))
mu.prior <- sum(dnorm(mu, 0, matrix(nu, Data$C, Data$J), log=TRUE))
nu.prior <- sum(dhalfcauchy(nu, 25, log=TRUE))
pi.prior <- ddirichlet(pi, Data$alpha, log=TRUE)
sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
### Log-Likelihood
LL <- sum(dnorm(Data$Y, mu[theta,], sigma[theta], log=TRUE))
### Log-Posterior
LP <- LL + theta.prior + mu.prior + nu.prior + pi.prior +
      sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,pi),
                 yhat=rnorm(prod(dim(mu[theta,])), mu[theta,], sigma[theta]),
                 parm=parm)
return(Modelout)
}

```

24.4. Initial Values

```
Initial.Values <- c(rcat(N,rep(1/C,C)), rep(0,C*J), rep(1,C), rep(1,C))
```

25. Cluster Analysis, Exploratory (ECA)

This is a nonparametric, model-based, cluster analysis, also called an infinite mixture model or latent class cluster analysis, where the number of clusters C is unknown, and a Dirichlet process via truncated stick-breaking is used to estimate the number of clusters. The record-level cluster membership parameter vector, θ , is a vector of discrete parameters. Discrete parameters are not supported in all algorithms. The example below is updated with the Slice sampler.

25.1. Form

$$\begin{aligned}
 \mathbf{Y}_{i,j} &\sim \mathcal{N}(\mu_{\theta[i],j}, \sigma_{\theta[i]}^2), \quad i = 1, \dots, N, \quad j = 1, \dots, J \\
 \theta_i &\sim \mathcal{CAT}(\pi_{1:C}), \quad i = 1, \dots, N \\
 \mu_{c,j} &\sim \mathcal{N}(0, \nu_c^2), \quad c = 1, \dots, C, \quad j = 1, \dots, J \\
 \sigma_c &\sim \mathcal{HC}(25), \quad c = 1, \dots, C \\
 \pi &= \text{Stick}(\delta) \\
 \delta_c &\sim \mathcal{BETA}(1, \gamma), \quad c = 1, \dots, (C - 1) \\
 \gamma &\sim \mathcal{G}(\alpha, \beta) \\
 \alpha &\sim \mathcal{HC}(25) \\
 \beta &\sim \mathcal{HC}(25)
 \end{aligned}$$

$$\nu_c \sim \mathcal{HC}(25), \quad c = 1, \dots, C$$

25.2. Data

```

data(demonsnacks)
Y <- as.matrix(log(demonsnacks + 1))
N <- nrow(Y)
J <- ncol(Y)
for (j in 1:J) Y[,j] <- CenterScale(Y[,j])
C <- 5 #Maximum number of clusters to explore
mon.names <- c("LP", paste("pi[", 1:C, "]", sep=""))
parm.names <- as.parm.names(list(theta=rep(0,N), delta=rep(0,C-1),
  mu=matrix(0,C,J), nu=rep(0,C), sigma=rep(0,C), alpha=0, beta=0,
  gamma=0))
pos.theta <- grep("theta", parm.names)
pos.delta <- grep("delta", parm.names)
pos.mu <- grep("mu", parm.names)
pos.nu <- grep("nu", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
PGF <- function(Data) {
  mu <- rnorm(Data$C*Data$J)
  nu <- runif(Data$C)
  sigma <- runif(Data$C)
  alpha <- runif(1)
  beta <- runif(1)
  gamma <- rgamma(1, alpha, beta)
  delta <- rev(sort(rbeta(Data$C-1, 1, gamma)))
  theta <- rcat(Data$N, Stick(delta))
  return(c(theta, delta, mu, nu, sigma, alpha, beta, gamma))
}
MyData <- list(C=C, J=J, N=N, PGF=PGF, Y=Y, mon.names=mon.names,
  parm.names=parm.names, pos.theta=pos.theta, pos.delta=pos.delta,
  pos.mu=pos.mu, pos.nu=pos.nu, pos.sigma=pos.sigma,
  pos.alpha=pos.alpha, pos.beta=pos.beta, pos.gamma=pos.gamma)

```

25.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperhyperparameters
  alpha <- interval(parm[Data$pos.alpha], 1e-100, Inf)
  parm[Data$pos.alpha] <- alpha
  beta <- interval(parm[Data$pos.beta], 1e-100, Inf)

```

```

parm[Data$pos.beta] <- beta
### Hyperparameters
delta <- interval(parm[Data$pos.delta], 1e-10, 1-1e-10)
parm[Data$pos.delta] <- delta
gamma <- interval(parm[Data$pos.gamma], 1e-100, Inf)
parm[Data$pos.nu] <- nu <- interval(parm[Data$pos.nu], 1e-100, Inf)
### Parameters
theta <- parm[Data$pos.theta]
mu <- matrix(parm[Data$pos.mu], Data$C, Data$J)
pi <- Stick(delta)
sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
parm[Data$pos.sigma] <- sigma
### Log-Hyperhyperprior
alpha.prior <- dhalfcauchy(alpha, 25, log=TRUE)
beta.prior <- dhalfcauchy(beta, 25, log=TRUE)
### Log-Hyperprior
delta.prior <- dStick(delta, gamma, log=TRUE)
gamma.prior <- dgamma(gamma, alpha, beta, log=TRUE)
nu.prior <- sum(dhalfcauchy(nu, 25, log=TRUE))
### Log-Prior
theta.prior <- sum(dcat(theta, pi, log=TRUE))
mu.prior <- sum(dnorm(mu, 0, matrix(nu, Data$C, Data$J), log=TRUE))
sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
### Log-Likelihood
LL <- sum(dnorm(Data$Y, mu[theta,], sigma[theta], log=TRUE))
### Log-Posterior
LP <- LL + theta.prior + delta.prior + mu.prior + nu.prior +
      alpha.prior + beta.prior + gamma.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,pi),
                 yhat=rnorm(prod(dim(mu[theta,])), mu[theta,], sigma[theta]),
                 parm=parm)
return(Modelout)
}

```

25.4. Initial Values

```

Initial.Values <- c(rcat(N, rev(sort(rStick(C-1,1)))), rep(0.5,C-1),
                  rep(0,C*J), rep(1,C), rep(1,C), rep(1,3))

```

26. Conditional Autoregression (CAR), Poisson

This CAR example is a slightly modified form of example 7.3 (Model A) in Congdon (2003). The Scottish lip cancer data also appears in the WinBUGS (Spiegelhalter, Thomas, Best, and Lunn 2003) examples and is a widely analyzed example. The data y consists of counts for $i = 1, \dots, 56$ counties in Scotland. A single predictor x is provided. The errors, ϵ , are

allowed to include spatial effects as smoothing by spatial effects from areal neighbors. The vector ϵ_μ is the mean of each area's error, and is a weighted average of errors in contiguous areas. Areal neighbors are indicated in adjacency matrix \mathbf{A} .

26.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{P}(\lambda) \\ \lambda &= \exp(\log(\mathbf{E}) + \beta_1 + \beta_2 \mathbf{x} + \epsilon) \\ \epsilon &\sim \mathcal{N}(\epsilon_\mu, \sigma^2) \\ \epsilon_{\mu[i]} &= \rho \sum_{j=1}^J \mathbf{A}_{i,j} \epsilon_j, \quad i = 1, \dots, N \\ \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\ \rho &\sim \mathcal{U}(-1, 1) \\ \sigma &\sim \mathcal{HC}(25) \end{aligned}$$

26.2. Data

```
N <- 56 #Number of areas
NN <- 264 #Number of adjacent areas
y <- c(9,39,11,9,15,8,26,7,6,20,13,5,3,8,17,9,2,7,9,7,16,31,11,7,19,15,7,
      10,16,11,5,3,7,8,11,9,11,8,6,4,10,8,2,6,19,3,2,3,28,6,1,1,1,1,0,0)
E <- c( 1.4,8.7,3.0,2.5,4.3,2.4,8.1,2.3,2.0,6.6,4.4,1.8,1.1,3.3,7.8,4.6,
      1.1,4.2,5.5,4.4,10.5,22.7,8.8,5.6,15.5,12.5,6.0,9.0,14.4,10.2,4.8,
      2.9,7.0,8.5,12.3,10.1,12.7,9.4,7.2,5.3,18.8,15.8,4.3,14.6,50.7,8.2,
      5.6,9.3,88.7,19.6,3.4,3.6,5.7,7.0,4.2,1.8) #Expected
x <- c(16,16,10,24,10,24,10,7,7,16,7,16,10,24,7,16,10,7,7,10,7,16,10,7,1,1,
      7,7,10,10,7,24,10,7,7,0,10,1,16,0,1,16,16,0,1,7,1,1,0,1,1,0,1,1,16,10)
A <- matrix(0, N, N)
A[1,c(5,9,11,19)] <- 1 #Area 1 is adjacent to areas 5, 9, 11, and 19
A[2,c(7,10)] <- 1 #Area 2 is adjacent to areas 7 and 10
A[3,c(6,12)] <- 1; A[4,c(18,20,28)] <- 1; A[5,c(1,11,12,13,19)] <- 1
A[6,c(3,8)] <- 1; A[7,c(2,10,13,16,17)] <- 1; A[8,6] <- 1
A[9,c(1,11,17,19,23,29)] <- 1; A[10,c(2,7,16,22)] <- 1
A[11,c(1,5,9,12)] <- 1; A[12,c(3,5,11)] <- 1; A[13,c(5,7,17,19)] <- 1
A[14,c(31,32,35)] <- 1; A[15,c(25,29,50)] <- 1
A[16,c(7,10,17,21,22,29)] <- 1; A[17,c(7,9,13,16,19,29)] <- 1
A[18,c(4,20,28,33,55,56)] <- 1; A[19,c(1,5,9,13,17)] <- 1
A[20,c(4,18,55)] <- 1; A[21,c(16,29,50)] <- 1; A[22,c(10,16)] <- 1
A[23,c(9,29,34,36,37,39)] <- 1; A[24,c(27,30,31,44,47,48,55,56)] <- 1
A[25,c(15,26,29)] <- 1; A[26,c(25,29,42,43)] <- 1
A[27,c(24,31,32,55)] <- 1; A[28,c(4,18,33,45)] <- 1
A[29,c(9,15,16,17,21,23,25,26,34,43,50)] <- 1
```



```

A[30,c(24,38,42,44,45,56)] <- 1; A[31,c(14,24,27,32,35,46,47)] <- 1
A[32,c(14,27,31,35)] <- 1; A[33,c(18,28,45,56)] <- 1
A[34,c(23,29,39,40,42,43,51,52,54)] <- 1; A[35,c(14,31,32,37,46)] <- 1
A[36,c(23,37,39,41)] <- 1; A[37,c(23,35,36,41,46)] <- 1
A[38,c(30,42,44,49,51,54)] <- 1; A[39,c(23,34,36,40,41)] <- 1
A[40,c(34,39,41,49,52)] <- 1; A[41,c(36,37,39,40,46,49,53)] <- 1
A[42,c(26,30,34,38,43,51)] <- 1; A[43,c(26,29,34,42)] <- 1
A[44,c(24,30,38,48,49)] <- 1; A[45,c(28,30,33,56)] <- 1
A[46,c(31,35,37,41,47,53)] <- 1; A[47,c(24,31,46,48,49,53)] <- 1
A[48,c(24,44,47,49)] <- 1; A[49,c(38,40,41,44,47,48,52,53,54)] <- 1
A[50,c(15,21,29)] <- 1; A[51,c(34,38,42,54)] <- 1
A[52,c(34,40,49,54)] <- 1; A[53,c(41,46,47,49)] <- 1
A[54,c(34,38,49,51,52)] <- 1; A[55,c(18,20,24,27,56)] <- 1
A[56,c(18,24,30,33,45,55)] <- 1
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,2), epsilon=rep(0,N), rho=0,
  sigma=0))
pos.beta <- grep("beta", parm.names)
pos.epsilon <- grep("epsilon", parm.names)
pos.rho <- grep("rho", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(2)
  epsilon <- rnorm(Data$N)
  rho <- runif(1,-1,1)
  sigma <- runif(1)
  return(c(beta, epsilon, rho, sigma))
}
MyData <- list(A=A, E=E, N=N, PGF=PGF, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.epsilon=pos.epsilon,
  pos.rho=pos.rho, pos.sigma=pos.sigma, x=x, y=y)

```

26.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  epsilon <- parm[Data$pos.epsilon]
  parm[Data$pos.rho] <- rho <- interval(parm[Data$pos.rho], -1, 1)
  epsilon.mu <- rho * rowSums(epsilon * Data$A)
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  epsilon.prior <- sum(dnorm(epsilon, epsilon.mu, sigma, log=TRUE))
  rho.prior <- dunif(rho, -1, 1, log=TRUE)
}

```

```

sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
### Log-Likelihood
lambda <- exp(log(Data$E) + beta[1] + beta[2]*Data$x/10 + epsilon)
LL <- sum(dpois(Data$y, lambda, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + epsilon.prior + rho.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rpois(length(lambda), lambda), parm=parm)
return(Modelout)
}

```

26.4. Initial Values

```
Initial.Values <- c(rep(0,2), rep(0,N), 0, 1)
```

27. Conditional Predictive Ordinate

For a more complete introduction to the conditional predictive ordinate (CPO), see the vignette entitled “Bayesian Inference”. Following is a brief guide to the applied use of CPO.

To include CPO in any model that is to be updated with MCMC, calculate and monitor the record-level inverse of the likelihood, InvL_i for records $i = 1, \dots, N$. CPO_i is the inverse of the posterior mean of InvL_i . The inverse CPO_i , or ICPO_i , is the posterior mean of InvL_i . ICPOs larger than 40 can be considered as possible outliers, and higher than 70 as extreme values.

Here, CPO is added to the linear regression example in section 48. In this data, record 6 is a possible outlier, and record 8 is an extreme value.

27.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\mu, \sigma^2) \\
 \mu &= \mathbf{X}\beta \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \sigma &\sim \mathcal{HC}(25)
 \end{aligned}$$

27.2. Data

```

data(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(log(demonsnacks[,c(1,4,10)]+1)))
J <- ncol(X)
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- c("LP", as.parm.names(list(InvL=rep(0,N))))
parm.names <- as.parm.names(list(beta=rep(0,J), sigma=0))

```

```

pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  sigma <- runif(1)
  return(c(beta, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma, y=y)

```

27.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  LL <- dnorm(Data$y, mu, sigma, log=TRUE)
  InvL <- 1 / exp(LL)
  LL <- sum(LL)
  ### Log-Posterior
  LP <- LL + beta.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,InvL),
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

27.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1)
```

28. Contingency Table

The two-way contingency table, matrix \mathbf{Y} , can easily be extended to more dimensions. Contingency table \mathbf{Y} has J rows and K columns. The cell counts are fit with Poisson regression, according to intercept α , main effects β_j for each row, main effects γ_k for each column, and interaction effects $\delta_{j,k}$ for dependence effects. An omnibus (all cells) test of independence is done by estimating two models (one with δ , and one without), and a large enough Bayes Factor indicates a violation of independence when the model with δ fits better than the model

without δ . In an ANOVA-like style, main effects contrasts can be used to distinguish rows or groups of rows from each other, as well as with columns. Likewise, interaction effects contrasts can be used to test independence in groups of $\delta_{j,k}$ elements. Finally, single-cell interactions can be used to indicate violations of independence for a given cell, such as when zero is not within its 95% probability interval.

28.1. Form

$$\begin{aligned} \mathbf{Y}_{j,k} &\sim \mathcal{P}(\lambda_{j,k}), \quad j = 1, \dots, J, \quad k = 1, \dots, K \\ \lambda_{j,k} &= \exp(\alpha + \beta_j + \gamma_k + \delta_{j,k}), \quad j = 1, \dots, J, \quad k = 1, \dots, K \\ \alpha &\sim \mathcal{N}(0, 1000) \\ \beta_j &\sim \mathcal{N}(0, \beta_\sigma^2), \quad j = 1, \dots, J \\ \beta_J &= -\sum_{j=1}^{J-1} \beta_j \\ \beta_\sigma &\sim \mathcal{HC}(25) \\ \gamma_k &\sim \mathcal{N}(0, \gamma_\sigma^2), \quad k = 1, \dots, K \\ \gamma_K &= -\sum_{k=1}^{K-1} \gamma_k \\ \gamma_\sigma &\sim \mathcal{HC}(25) \\ \delta_{j,k} &\sim \mathcal{N}(0, \delta_\sigma^2) \\ \delta_{J,K} &= -\sum \delta_{-J,-K} \\ \delta_\sigma &\sim \mathcal{HC}(25) \end{aligned}$$

28.2. Data

```
J <- 4 #Rows
K <- 4 #Columns
Y <- matrix(c(20,94,84,17,68,7,119,26,5,16,29,14,15,10,54,14), 4, 4)
rownames(Y) <- c("Black", "Blond", "Brunette", "Red")
colnames(Y) <- c("Blue", "Brown", "Green", "Hazel")
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, beta=rep(0,J-1),
  gamma=rep(0,K-1), delta=rep(0,J*K-1), b.sigma=0, g.sigma=0,
  d.sigma=0))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.delta <- grep("delta", parm.names)
pos.b.sigma <- grep("b.sigma", parm.names)
pos.g.sigma <- grep("g.sigma", parm.names)
```

```

pos.d.sigma <- grep("d.sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1,log(mean(Y)),1)
  beta <- rnorm(Data$J-1)
  gamma <- rnorm(Data$K-1)
  delta <- rnorm(Data$J*Data$K-1)
  sigma <- runif(3)
  return(c(alpha, beta, gamma, delta, sigma))
}
MyData <- list(J=J, K=K, PGF=PGF, Y=Y, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.beta=pos.beta,
  pos.gamma=pos.gamma, pos.delta=pos.delta, pos.b.sigma=pos.b.sigma,
  pos.g.sigma=pos.g.sigma, pos.d.sigma=pos.d.sigma)

```

28.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperparameters
  beta.sigma <- interval(parm[Data$pos.b.sigma], 1e-100, Inf)
  parm[Data$pos.b.sigma] <- beta.sigma
  gamma.sigma <- interval(parm[Data$pos.g.sigma], 1e-100, Inf)
  parm[Data$pos.g.sigma] <- gamma.sigma
  delta.sigma <- interval(parm[Data$pos.d.sigma], 1e-100, Inf)
  parm[Data$pos.d.sigma] <- delta.sigma
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- parm[Data$pos.beta]
  beta <- c(beta, -sum(beta))
  gamma <- parm[Data$pos.gamma]
  gamma <- c(gamma, -sum(gamma))
  delta <- parm[Data$pos.delta]
  delta <- c(delta, -sum(delta))
  delta <- matrix(delta, Data$J, Data$K)
  ### Log-Hyperprior
  beta.sigma.prior <- dhalfcauchy(beta.sigma, 25, log=TRUE)
  gamma.sigma.prior <- dhalfcauchy(gamma.sigma, 25, log=TRUE)
  delta.sigma.prior <- dhalfcauchy(delta.sigma, 25, log=TRUE)
  ### Log-Prior
  alpha.prior <- dnorm(alpha, 0, 1000, log=TRUE)
  beta.prior <- sum(dnorm(beta, 0, beta.sigma, log=TRUE))
  gamma.prior <- sum(dnorm(gamma, 0, gamma.sigma, log=TRUE))
  delta.prior <- sum(dnorm(delta, 0, delta.sigma, log=TRUE))
  ### Log-Likelihood
  beta <- matrix(beta, Data$J, Data$K)
  gamma <- matrix(gamma, Data$J, Data$K, byrow=TRUE)

```

```

lambda <- exp(alpha + beta + gamma + delta)
LL <- sum(dpois(Data$Y, lambda, log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior + beta.sigma.prior +
      gamma.prior + gamma.sigma.prior + delta.prior +
      delta.sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
                yhat=rpois(length(lambda), lambda),
                parm=parm)
return(Modelout)
}

```

28.4. Initial Values

```

Initial.Values <- c(log(mean(Y)), rep(0,J-1), rep(0,K-1), rep(0,J*K-1),
                  rep(1,3))

```

29. Discrete Choice, Conditional Logit

29.1. Form

$$\mathbf{y}_i \sim \mathcal{CAT}(\mathbf{p}_{i,1:J}), \quad i = 1, \dots, N, \quad j = 1, \dots, J$$

$$\mathbf{p}_{i,j} = \frac{\phi_{i,j}}{\sum_{j=1}^J \phi_{i,j}}$$

$$\phi = \exp(\mu)$$

$$\mu_{i,j} = \beta_{j,1:K} \mathbf{X}_{i,1:K} + \gamma \mathbf{Z}_{i,1:C} \in [-700, 700], \quad j = 1, \dots, (J-1)$$

$$\mu_{i,J} = \gamma \mathbf{Z}_{i,1:C}$$

$$\beta_{j,k} \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, (J-1)$$

$$\gamma_c \sim \mathcal{N}(0, 1000)$$

29.2. Data

```

data(demonchoice)
y <- as.numeric(demonchoice[,1])
X <- cbind(1, as.matrix(demonchoice[,2:3]))
Z <- as.matrix(demonchoice[,4:9])
for (j in 2:ncol(X)) X[,j] <- CenterScale(X[,j])
for (j in 1:ncol(Z)) Z[,j] <- CenterScale(Z[,j])
N <- length(y) #Number of records
J <- length(unique(y)) #Number of categories in y

```

```

K <- ncol(X) #Number of individual attributes (including the intercept)
C <- ncol(Z) #Number of choice-based attributes (intercept is not included)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=matrix(0,J-1,K), gamma=rep(0,C)))
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
PGF <- function(Data) {
  beta <- rnorm((Data$J-1)*Data$K)
  gamma <- rnorm(Data$C)
  return(c(beta, gamma))
}
MyData <- list(C=C, J=J, K=K, N=N, PGF=PGF, X=X, Z=Z, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.gamma=pos.gamma, y=y)

```

29.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- matrix(parm[Data$pos.beta], Data$J-1, Data$K)
  gamma <- parm[Data$pos.gamma]
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  gamma.prior <- sum(dnormv(gamma, 0, 1000, log=TRUE))
  ### Log-Likelihood
  mu <- matrix(tcrossprod(gamma, Data$Z), Data$N, Data$J)
  mu[, -Data$J] <- mu[, -Data$J] + tcrossprod(Data$X, beta)
  mu <- interval(mu, -700, 700, reflect=FALSE)
  phi <- exp(mu)
  p <- phi / rowSums(phi)
  LL <- sum(dcat(Data$y, p, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + gamma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP, yhat=rcat(nrow(p), p),
    parm=parm)
  return(Modelout)
}

```

29.4. Initial Values

```

Initial.Values <- c(rep(0,(J-1)*K), rep(0,C))

```

30. Discrete Choice, Mixed Logit

30.1. Form

$$\mathbf{y}_i \sim \mathcal{CAT}(\mathbf{p}_{i,1:J}), \quad i = 1, \dots, N$$

$$\mathbf{p}_{i,j} = \frac{\phi_{i,j}}{\sum_{j=1}^J \phi_{i,j}}$$

$$\phi = \exp(\boldsymbol{\mu})$$

$$\mu_{i,j} = \beta_{j,1:K,i} \mathbf{X}_{i,1:K} + \gamma \mathbf{Z}_{i,1:C} \in [-700, 700], \quad i = 1, \dots, N, \quad j = 1, \dots, (J-1)$$

$$\mu_{i,J} = \gamma \mathbf{Z}_{i,1:C}$$

$$\beta_{j,k,i} \sim \mathcal{N}(\zeta_{j,k}^\mu, \zeta_{j,k}^\sigma 2_{j,k}), \quad i = 1, \dots, N, \quad j = 1, \dots, (J-1), \quad k = 1, \dots, K$$

$$\gamma_c \sim \mathcal{N}(0, 1000), \quad c = 1, \dots, C$$

$$\zeta_{j,k}^\mu \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, (J-1), \quad k = 1, \dots, K$$

$$\zeta_{j,k}^\sigma \sim \mathcal{HC}25), \quad j = 1, \dots, (J-1), \quad k = 1, \dots, K$$

30.2. Data

```

data(demonchoice)
y <- as.numeric(demonchoice[,1])
X <- cbind(1, as.matrix(demonchoice[,2:3]))
Z <- as.matrix(demonchoice[,4:9])
for (j in 2:ncol(X)) X[,j] <- CenterScale(X[,j])
for (j in 1:ncol(Z)) Z[,j] <- CenterScale(Z[,j])
N <- length(y)
J <- length(unique(y)) #Number of categories in y
K <- ncol(X) #Number of predictors (including the intercept)
C <- ncol(Z) #Number of choice-based attributes (intercept is not included)
S <- diag(J-1)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=array(0, dim=c(J-1,K,N)),
  gamma=rep(0,C), zeta.mu=matrix(0,J-1,K), zeta.sigma=matrix(0,J-1,K)))
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.zeta.mu <- grep("zeta.mu", parm.names)
pos.zeta.sigma <- grep("zeta.sigma", parm.names)
PGF <- function(Data) {
  zeta.mu <- matrix(rnorm((Data$J-1)*Data$K), Data$J-1, Data$K)
  zeta.sigma <- matrix(runif((Data$J-1)*Data$K), Data$J-1, Data$K)
  beta <- array(rnorm((Data$J-1)*Data$K*Data$N),
    dim=c(Data$J-1, Data$K, Data$N))
  gamma <- rnorm(Data$C)
  return(c(beta, gamma, as.vector(zeta.mu), as.vector(zeta.sigma)))
}
MyData <- list(C=C, J=J, K=K, N=N, PGF=PGF, S=S, X=X, Z=Z,

```



```

mon.names=mon.names, parm.names=parm.names, pos.beta=pos.beta,
pos.gamma=pos.gamma, pos.zeta.mu=pos.zeta.mu,
pos.zeta.sigma=pos.zeta.sigma, y=y)

```

30.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- array(parm[Data$pos.beta], dim=c(Data$J-1, Data$K, Data$N))
  gamma <- parm[Data$pos.gamma]
  zeta.mu <- matrix(parm[Data$pos.zeta.mu], Data$J-1, Data$K)
  zeta.sigma <- matrix(interval(parm[Data$pos.zeta.sigma], 1e-100, Inf),
    Data$J-1, Data$K)
  parm[Data$pos.zeta.sigma] <- as.vector(zeta.sigma)
  ### Log-Hyperprior
  zeta.mu.prior <- sum(dnormv(zeta.mu, 0, 1000, log=TRUE))
  zeta.sigma.prior <- sum(dhalfcauchy(zeta.sigma, 25, log=TRUE))
  ### Log-Prior
  beta.prior <- sum(dnorm(beta, zeta.mu, zeta.sigma, log=TRUE))
  gamma.prior <- sum(dnormv(gamma, 0, 1000, log=TRUE))
  ### Log-Likelihood
  mu <- matrix(tcrossprod(Data$Z, t(gamma)), Data$N, Data$J)
  for (j in 1:(Data$J-1)) mu[,j] <- rowSums(Data$X * t(beta[j, , ]))
  mu <- interval(mu, -700, 700, reflect=FALSE)
  phi <- exp(mu)
  p <- phi / rowSums(phi)
  LL <- sum(dcat(Data$y, p, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + gamma.prior + zeta.mu.prior + zeta.sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP, yhat=rcat(nrow(p), p),
    parm=parm)
  return(Modelout)
}

```

30.4. Initial Values

```

Initial.Values <- c(rep(0, (J-1)*K*N), rep(0, C), rep(0, (J-1)*K),
  rep(1, (J-1)*K))

```

31. Discrete Choice, Multinomial Probit

31.1. Form

$$\begin{aligned} \mathbf{W}_{i,1:(J-1)} &\sim \mathcal{N}_{J-1}(\mu_{i,1:(J-1)}, \Sigma), \quad i = 1, \dots, N \\ \mathbf{W}_{i,j} &\in \begin{cases} [0,10] & \text{if } \mathbf{y}_i = j \\ [-10,0] & \end{cases} \\ \mu_{1:N,j} &= \mathbf{X}\beta_{j,1:K} + \mathbf{Z}\gamma \\ \Sigma &= \mathbf{U}^T\mathbf{U} \\ \beta_{j,k} &\sim \mathcal{N}(0,10), \quad j = 1, \dots, (J-1), \quad k = 1, \dots, K \\ \gamma_c &\sim \mathcal{N}(0,10), \quad c = 1, \dots, C \\ \mathbf{U}_{j,k} &\sim \mathcal{N}(0,1), \quad j = 1, \dots, (J-1), \quad k = 1, \dots, (J-1), \quad j \geq k, \quad j \neq k = 1 \end{aligned}$$

31.2. Data

```

data(demonchoice)
y <- as.numeric(demonchoice[,1])
X <- cbind(1, as.matrix(demonchoice[,2:3]))
Z <- as.matrix(demonchoice[,4:9])
for (j in 2:ncol(X)) X[,j] <- CenterScale(X[,j])
for (j in 1:ncol(Z)) Z[,j] <- CenterScale(Z[,j])
J <- length(unique(y)) #Number of categories in y
K <- ncol(X) #Number of predictors (including the intercept)
C <- ncol(Z) #Number of choice-based attributes (intercept is not included)
S <- diag(J-1)
U <- matrix(NA, J-1, J-1)
U[upper.tri(U, diag=TRUE)] <- 0
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=matrix(0, (J-1), K), gamma=rep(0, C),
  U=U, W=matrix(0, N, J-1)))
parm.names <- parm.names[-which(parm.names == "U[1,1]")]
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.U <- grep("U", parm.names)
pos.W <- grep("W", parm.names)
PGF <- function(Data) {
  beta <- rnorm((Data$J-1)*Data$K)
  gamma <- rnorm(Data$C)
  U <- rnorm((Data$J-2) + (factorial(Data$J-1) /
    (factorial(Data$J-1-2)*factorial(2))), 0, 1)
  W <- matrix(runif(Data$N*(Data$J-1), -10, 0), Data$N, Data$J-1)
  Y <- as.indicator.matrix(Data$y)
  W <- ifelse(Y[, -Data$J] == 1, abs(W), W)
  return(c(beta, gamma, U, as.vector(W)))}
MyData <- list(C=C, J=J, K=K, N=N, PGF=PGF, S=S, X=X, Z=Z,

```

```
mon.names=mon.names, parm.names=parm.names, pos.beta=pos.beta,
pos.gamma=pos.gamma, pos.U=pos.U, pos.W=pos.W, y=y)
```

31.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  beta <- matrix(parm[Data$pos.beta], Data$J-1, Data$K)
  gamma <- parm[Data$pos.gamma]
  u <- c(0, parm[Data$pos.U])
  U <- diag(Data$J-1)
  U[upper.tri(U, diag=TRUE)] <- u
  diag(U) <- exp(diag(U))
  Sigma <- t(U) %*% U
  Sigma[1,] <- Sigma[,1] <- U[1,]
  W <- matrix(parm[Data$pos.W], Data$N, Data$J-1)
  Y <- as.indicator.matrix(Data$y)
  temp <- which(Y[, -c(Data$J)] == 1)
  W[temp] <- interval(W[temp], 0, 10)
  temp <- which(Y[, -c(Data$J)] == 0)
  W[temp] <- interval(W[temp], -10, 0)
  parm[Data$pos.W] <- as.vector(W)
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 10, log=TRUE))
  gamma.prior <- sum(dnormv(gamma, 0, 10, log=TRUE))
  U.prior <- sum(dnorm(u[-1], 0, 1, log=TRUE))
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, beta) +
    as.vector(tcrossprod(Data$Z, t(gamma)))
  #eta <- exp(cbind(mu,0))
  #p <- eta / rowSums(eta)
  LL <- sum(dmvn(W, mu, Sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + gamma.prior + U.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=max.col(cbind(rmvn(nrow(mu), mu, Sigma),0)), parm=parm)
  return(Modelout)
}
```

31.4. Initial Values

```
Initial.Values <- GIV(Model, MyData, PGF=TRUE)
```

32. Distributed Lag, Koyck

This example applies Koyck or geometric distributed lags to $k = 1, \dots, K$ discrete events in covariate \mathbf{x} , transforming the covariate into a $N \times K$ matrix \mathbf{X} and creates a $N \times K$ lag matrix \mathbf{L} .

32.1. Form

$$\mathbf{y} \sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2)$$

$$\mu_t = \alpha + \phi \mathbf{y}_{t-1} + \sum_{k=1}^K \mathbf{X}_{t,k} \beta \lambda^{L[t,k]}, \quad k = 1, \dots, K, \quad t = 2, \dots, T$$

$$\mu_1 = \alpha + \sum_{k=1}^K \mathbf{X}_{1,k} \beta \lambda^{L[1,k]}, \quad k = 1, \dots, K$$

$$\alpha \sim \mathcal{N}(0, 1000)$$

$$\beta \sim \mathcal{N}(0, 1000)$$

$$\lambda \sim \mathcal{U}(0, 1)$$

$$\phi \sim \mathcal{N}(0, 1000)$$

$$\sigma \sim \mathcal{HC}(25)$$

32.2. Data

```
data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
x <- (y > 0.01)*1 #Made-up distributed lag IV
T <- length(y)
K <- length(which(x != 0))
L <- X <- matrix(0, T, K)
for (i in 1:K) {
  X[which(x != 0)[i]:T,i] <- x[which(x != 0)[i]]
  L[(which(x != 0)[i]):T,i] <- 0:(T - which(x != 0)[i])}
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, beta=0, lambda=0, phi=0, sigma=0))
PGF <- function(Data) {
  alpha <- rnorm(1)
  beta <- rnorm(1)
  lambda <- runif(1)
  phi <- rnorm(1)
  sigma <- runif(1)
  return(c(alpha, beta, lambda, phi, sigma))
}
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
```

```

pos.lambda <- grep("lambda", parm.names)
pos.phi <- grep("phi", parm.names)
pos.sigma <- grep("sigma", parm.names)
MyData <- list(L=L, PGF=PGF, T=T, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.beta=pos.beta,
  pos.lambda=pos.lambda, pos.phi=pos.phi, pos.sigma=pos.sigma, y=y)

```

32.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- parm[Data$pos.beta]
  lambda <- interval(parm[Data$pos.lambda], 0, 1)
  parm[Data$pos.lambda] <- lambda
  phi <- parm[Data$pos.phi]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- dnormv(alpha, 0, 1000, log=TRUE)
  beta.prior <- dnormv(beta, 0, 1000, log=TRUE)
  lambda.prior <- dunif(lambda, 0, 1, log=TRUE)
  phi.prior <- dnormv(phi, 0, 1000, log=TRUE)
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- c(alpha, alpha + phi*Data$y[-Data$T]) +
    rowSums(Data$X * beta * lambda^Data$L)
  LL <- sum(dnorm(Data$y[-1], mu[-1], sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + alpha.prior + beta.prior + lambda.prior + phi.prior +
    sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

32.4. Initial Values

```

Initial.Values <- c(rep(0,2), 0.5, 0, 1)

```

33. Dynamic Sparse Factor Model (DSFM)

33.1. Form

$$\mathbf{Y}_{t,j} \sim \mathcal{N}(\alpha t, j + \mathbf{F}_{t,1:P} \Lambda_{1:P,1:j,t}, \Sigma_{t,j}^2), \quad t = 1, \dots, T, \quad j = 1, \dots, J$$

$$\alpha_{t,j} \sim \mathcal{N}(\alpha_j^\mu + \alpha_j^\phi (\alpha_{t-1,j} - \alpha^m u_j), \alpha^\sigma 2_j)$$

$$\mathbf{F}_{t,1:P} \sim \mathcal{N}_P(\mathbf{F}^\phi \mathbf{F}_{t-1,1:P}, \mathbf{f}_{t,1:P}^\Sigma)$$

$$\mathbf{f}_{t,1:P}^\Sigma = t(\mathbf{f}_{1:P,1:P,t}^{\mathbf{U}}) \mathbf{f}_{1:P,1:P,t}^{\mathbf{U}}$$

$$\mathbf{f}_{p,q,t}^{\mathbf{U}} \sim \mathcal{N}(\mathbf{f}_{p,q}^{\mathbf{u}\mu} + \mathbf{f}_{p,q}^{\mathbf{u}\phi} (\mathbf{f}_{p,q,t-1}^{\mathbf{U}} - \mathbf{f}_{p,q}^{\mathbf{u}\mu}), \mathbf{f}_{p,q}^{\mathbf{u}\sigma^2})$$

$$\Lambda_{p,j,t} \sim \mathcal{N}(\lambda_{p,j}^\mu + \lambda_{p,j}^\phi (\Lambda_{p,j,t-1} - \lambda^m u_{p,j}), \lambda^\sigma 2_{p,j})$$

$$\Sigma_{t,j} = \exp(\log(\Sigma_{t,j}))$$

$$\log(\Sigma_{t,j}) \sim \mathcal{N}(\sigma_j^\mu + \sigma_j^\phi (\log(\Sigma_{t-1,j}) - \sigma^m u_j), \sigma^\sigma 2_j)$$

$$\alpha_j^0 \sim \mathcal{N}(0, 1), \quad j = 1, \dots, J$$

$$\alpha_j^\mu \sim \mathcal{N}(0, 1), \quad j = 1, \dots, J$$

$$\frac{\alpha_j^\phi + 1}{2} \sim \mathcal{BETA}(20, 1.5), \quad j = 1, \dots, J$$

$$\alpha_j^\sigma \sim \mathcal{HC}(5), \quad j = 1, \dots, J$$

$$\mathbf{f}_j^0 \sim \mathcal{N}(0, 1), \quad j = 1, \dots, J$$

$$\frac{\mathbf{f}_j^\phi + 1}{2} \sim \mathcal{BETA}(1, 1), \quad j = 1, \dots, J$$

$$\mathbf{f}_j^{\mathbf{u}0} \sim \mathcal{N}(0, 1), \quad j = 1, \dots, J$$

$$\mathbf{f}_j^{\mathbf{u}\mu} \sim \mathcal{N}(0, 1), \quad j = 1, \dots, J$$

$$\frac{\mathbf{f}_j^{\mathbf{u}\phi} + 1}{2} \sim \mathcal{BETA}(20, 1.5), \quad j = 1, \dots, J$$

$$\mathbf{f}_j^{\mathbf{u}\sigma} \sim \mathcal{HC}(1), \quad j = 1, \dots, J$$

$$\lambda_j^0 \sim \mathcal{N}(0, 1), \quad j = 1, \dots, J$$

$$\lambda_j^d \sim \mathcal{U}(0, |\lambda_j^\mu| + 3 \sqrt{\frac{\lambda_j^\sigma}{1 - \lambda_j^\phi \lambda_j^\phi}}), \quad j = 1, \dots, J$$

$$\lambda_j^\mu \sim \mathcal{N}(0, 1), \quad j = 1, \dots, J$$

$$\frac{\lambda_j^\phi + 1}{2} \sim \mathcal{BETA}(20, 1.5), \quad j = 1, \dots, J$$

$$\lambda_j^\sigma \sim \mathcal{HC}(1), \quad j = 1, \dots, J$$

$$\log(\sigma_j^0) \sim \mathcal{N}(0, 1), \quad j = 1, \dots, J$$

$$\log(\sigma_j^\mu) \sim \mathcal{N}(0, 1), \quad j = 1, \dots, J$$

$$\frac{\log(\sigma_j^\phi) + 1}{2} \sim \mathcal{BETA}(20, 1.5), \quad j = 1, \dots, J$$

$$\log(\sigma_j^\sigma) \sim \mathcal{HC}(1), \quad j = 1, \dots, J$$

33.2. Data

```

data(demonfx)
Y.orig <- demonfx
Y <- log(as.matrix(Y.orig[1:20,1:3]))
Y.means <- colMeans(Y)
Y <- Y - matrix(Y.means, nrow(Y), ncol(Y), byrow=TRUE) #Center
Y.scales <- sqrt(.colVars(Y))
Y <- Y / matrix(Y.scales, nrow(Y), ncol(Y), byrow=TRUE) #Scale
T <- nrow(Y) #Number of time-periods
J <- ncol(Y) #Number of time-series
P <- 2 #Number of dynamic factors
mon.names <- "LP"
U1 <- matrix(NA,P,P); U2 <- matrix(NA,P,J)
U1[upper.tri(U1, diag=TRUE)] <- 0; U2[upper.tri(U2)] <- 0
Lambda <- array(NA, dim=c(P,J,T))
U <- array(NA, dim=c(P,P,T))
for (t in 1:T) {
  U[ , , t] <- U1
  Lambda[ , , t] <- U2}
parm.names <- as.parm.names(list(alpha0=rep(0,J), Alpha=matrix(0,T,J),
  alpha.mu=rep(0,J), alpha.phi=rep(0,J), alpha.sigma=rep(0,J),
  f0=rep(0,P), F=matrix(0,T,P), f.phi=rep(0,P), f.u0=U1, f.U=U,
  f.u.mu=U1, f.u.phi=U1, f.u.sigma=U1, lambda0=U2, Lambda=Lambda,
  lambda.d=U2, lambda.mu=U2, lambda.phi=U2, lambda.sigma=U2,
  lsigma0=rep(0,J), lSigma=matrix(0,T,J),
  lsigma.mu=rep(0,J), lsigma.phi=rep(0,J), lsigma.sigma=rep(0,J)))
pos.alpha0 <- grep("alpha0", parm.names)
pos.Alpha <- grep("Alpha", parm.names)
pos.alpha.mu <- grep("alpha.mu", parm.names)
pos.alpha.phi <- grep("alpha.phi", parm.names)
pos.alpha.sigma <- grep("alpha.sigma", parm.names)
pos.f0 <- grep("f0", parm.names)
pos.F <- grep("F", parm.names)
pos.f.phi <- grep("f.phi", parm.names)
pos.f.u0 <- grep("f.u0", parm.names)
pos.f.U <- grep("f.U", parm.names)
pos.f.u.mu <- grep("f.u.mu", parm.names)
pos.f.u.phi <- grep("f.u.phi", parm.names)
pos.f.u.sigma <- grep("f.u.sigma", parm.names)
pos.lambda0 <- grep("lambda0", parm.names)
pos.Lambda <- grep("Lambda", parm.names)

```

```

pos.lambda.d <- grep("lambda.d", parm.names)
pos.lambda.mu <- grep("lambda.mu", parm.names)
pos.lambda.phi <- grep("lambda.phi", parm.names)
pos.lambda.sigma <- grep("lambda.sigma", parm.names)
pos.lsigma0 <- grep("lsigma0", parm.names)
pos.lSigma <- grep("lSigma", parm.names)
pos.lsigma.mu <- grep("lsigma.mu", parm.names)
pos.lsigma.phi <- grep("lsigma.phi", parm.names)
pos.lsigma.sigma <- grep("lsigma.sigma", parm.names)
PGF <- function(Data) {
  alpha0 <- rnorm(Data$J)
  Alpha <- rnorm(Data$T*Data$J)
  alpha.mu <- rnorm(Data$J)
  alpha.phi <- rbeta(Data$J, 20, 1.5) * 2 - 1
  alpha.sigma <- runif(Data$J)
  f0 <- rnorm(Data$P)
  F <- rnorm(Data$T*Data$P)
  f.phi <- rbeta(Data$P, 1, 1) * 2 - 1
  f.u0 <- rnorm(length(Data$pos.f.u0))
  f.U <- rnorm(length(Data$pos.f.U))
  f.u.mu <- rnorm(length(Data$pos.f.u.mu))
  f.u.phi <- runif(length(Data$pos.f.u.phi))
  f.u.sigma <- runif(length(Data$pos.f.u.sigma))
  lambda0 <- rnorm(length(Data$pos.lambda0))
  Lambda <- rnorm(length(Data$pos.Lambda))
  lambda.mu <- rnorm(length(Data$pos.lambda.mu))
  lambda.phi <- rbeta(length(Data$pos.lambda.phi), 20, 1.5)
  lambda.sigma <- runif(length(Data$pos.lambda.sigma))
  lambda.d <- runif(length(Data$pos.lambda.d), 0, abs(lambda.mu) +
    3*sqrt(lambda.sigma/(1-lambda.phi^2)))
  lsigma0 <- rnorm(Data$J)
  lSigma <- rnorm(Data$T*Data$J)
  lsigma.mu <- rnorm(Data$J)
  lsigma.phi <- rbeta(Data$J, 20, 1.5) * 2 - 1
  lsigma.sigma <- runif(Data$J)
  return(c(alpha0, Alpha, alpha.mu, alpha.phi, alpha.sigma, f0, F,
    f.phi, f.u0, f.U, f.u.mu, f.u.phi, f.u.sigma, lambda0, Lambda,
    lambda.d, lambda.mu, lambda.phi, lambda.sigma, lsigma0, lSigma,
    lsigma.mu, lsigma.phi, lsigma.sigma))
}
MyData <- list(J=J, P=P, PGF=PGF, T=T, Y=Y, mon.names=mon.names,
  parm.names=parm.names, pos.alpha0=pos.alpha0, pos.Alpha=pos.Alpha,
  pos.alpha.mu=pos.alpha.mu, pos.alpha.phi=pos.alpha.phi,
  pos.alpha.sigma=pos.alpha.sigma, pos.f0=pos.f0, pos.F=pos.F,
  pos.f.phi=pos.f.phi, pos.f.u0=pos.f.u0, pos.f.U=pos.f.U,
  pos.f.u.mu=pos.f.u.mu, pos.f.u.phi=pos.f.u.phi,
  pos.f.u.sigma=pos.f.u.sigma, pos.lambda0=pos.lambda0,

```



```

pos.Lambda=pos.Lambda, pos.lambda.d=pos.lambda.d,
pos.lambda.mu=pos.lambda.mu, pos.lambda.phi=pos.lambda.phi,
pos.lambda.sigma=pos.lambda.sigma, pos.lsigma0=pos.lsigma0,
pos.lSigma=pos.lSigma, pos.lsigma.mu=pos.lsigma.mu,
pos.lsigma.phi=pos.lsigma.phi, pos.lsigma.sigma=pos.lsigma.sigma)

```

33.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha0 <- parm[Data$pos.alpha0]
  Alpha <- matrix(parm[Data$pos.Alpha], Data$T, Data$J)
  alpha.mu <- parm[Data$pos.alpha.mu]
  alpha.phi <- interval(parm[Data$pos.alpha.phi], -1, 1)
  parm[Data$pos.alpha.phi] <- alpha.phi
  alpha.sigma <- interval(parm[Data$pos.alpha.sigma], 1e-100, Inf)
  parm[Data$pos.alpha.sigma] <- alpha.sigma
  f0 <- parm[Data$pos.f0]
  F <- matrix(parm[Data$pos.F], Data$T, Data$P)
  f.phi <- interval(parm[Data$pos.f.phi], -1, 1)
  parm[Data$pos.f.phi] <- f.phi
  f.u0 <- parm[Data$pos.f.u0]
  f.U <- parm[Data$pos.f.U]
  f.u.mu <- parm[Data$pos.f.u.mu]
  f.u.phi <- interval(parm[Data$pos.f.u.phi], -1, 1)
  parm[Data$pos.f.u.phi] <- f.u.phi
  f.u.sigma <- interval(parm[Data$pos.f.u.sigma], 1e-100, Inf)
  parm[Data$pos.f.u.sigma] <- f.u.sigma
  lambda0 <- parm[Data$pos.lambda0]
  Lambda <- parm[Data$pos.Lambda]
  lambda.mu <- parm[Data$pos.lambda.mu]
  lambda.phi <- interval(parm[Data$pos.lambda.phi], -1, 1)
  parm[Data$pos.lambda.phi] <- lambda.phi
  lambda.sigma <- interval(parm[Data$pos.lambda.sigma], 1e-100, Inf)
  parm[Data$pos.lambda.sigma] <- lambda.sigma
  lambda.d <- parm[Data$pos.lambda.d]
  for (i in 1:length(lambda.d))
    lambda.d[i] <- interval(lambda.d[i], 0, abs(lambda.mu[i]) +
      3*sqrt(lambda.sigma[i]/(1-lambda.phi[i]^2)))
  parm[Data$pos.lambda.d] <- lambda.d
  lsigma0 <- parm[Data$pos.lsigma0]
  lSigma <- matrix(parm[Data$pos.lSigma], Data$T, Data$J)
  lsigma.mu <- parm[Data$pos.lsigma.mu]
  lsigma.phi <- interval(parm[Data$pos.lsigma.phi], -1, 1)
  parm[Data$pos.lsigma.phi] <- lsigma.phi

```

```

lsigma.sigma <- interval(parm[Data$pos.lsigma.sigma], 1e-100, Inf)
parm[Data$pos.lsigma.sigma] <- lsigma.sigma
### Log-Prior
alpha0.prior <- sum(dnorm(alpha0, 0, 1, log=TRUE))
Alpha.prior <- sum(dnorm(Alpha,
  matrix(alpha.mu, Data$T, Data$J, byrow=TRUE) +
  matrix(alpha.phi, Data$T, Data$J, byrow=TRUE) *
  (rbind(alpha0, Alpha[-Data$T,]) -
  matrix(alpha.mu, Data$T, Data$J, byrow=TRUE)),
  matrix(alpha.sigma, Data$T, Data$J, byrow=TRUE), log=TRUE))
alpha.mu.prior <- sum(dnorm(alpha.mu, 0, 1, log=TRUE))
alpha.phi.prior <- sum(dbeta((alpha.phi + 1) / 2, 20, 1.5, log=TRUE))
alpha.sigma.prior <- sum(dhalfcauchy(alpha.sigma, 5, log=TRUE))
f0.prior <- sum(dnorm(f0, 0, 1, log=TRUE))
f.phi.prior <- sum(dbeta((f.phi + 1) / 2, 1, 1, log=TRUE))
f.u0.prior <- sum(dnorm(f.u0, 0, 1, log=TRUE))
f.U.prior <- sum(dnorm(matrix(f.U, nrow=Data$T, byrow=TRUE),
  matrix(f.u.mu, Data$T, Data$P*(Data$P-1)/2+Data$P, byrow=TRUE) +
  matrix(f.u.phi, Data$T, Data$P*(Data$P-1)/2+Data$P, byrow=TRUE) *
  (rbind(f.u0, matrix(f.U, nrow=Data$T, byrow=TRUE)[-Data$T,]) -
  matrix(f.u.mu, Data$T, Data$P*(Data$P-1)/2+Data$P, byrow=TRUE)),
  matrix(f.u.sigma, Data$T, Data$P*(Data$P-1)/2+Data$P, byrow=TRUE),
  log=TRUE))
f.u.mu.prior <- sum(dnorm(f.u.mu, 0, 1, log=TRUE))
f.u.phi.prior <- sum(dbeta((f.u.phi + 1) / 2, 20, 1.5, log=TRUE))
f.u.sigma.prior <- sum(dhalfcauchy(f.u.sigma, 1, log=TRUE))
lambda0.prior <- sum(dnorm(lambda0, 0, 1, log=TRUE))
Lambda.prior <- sum(dnorm(matrix(Lambda, nrow=Data$T, byrow=TRUE),
  matrix(lambda.mu, Data$T, length(lambda.mu), byrow=TRUE) +
  (rbind(lambda0, matrix(Lambda, nrow=Data$T, byrow=TRUE))[-(Data$T+1),]
  -
  matrix(lambda.mu, Data$T, length(lambda.mu), byrow=TRUE)),
  matrix(lambda.sigma, Data$T, length(lambda.sigma), byrow=TRUE),
  log=TRUE))
lambda.d.prior <- sum(dunif(lambda.d, 0, abs(lambda.mu) +
  3*sqrt(lambda.sigma/(1-lambda.phi^2)), log=TRUE))
lambda.mu.prior <- sum(dnorm(lambda.mu, 0, 1, log=TRUE))
lambda.phi.prior <- sum(dbeta((lambda.phi + 1) / 2, 20, 1.5, log=TRUE))
lambda.sigma.prior <- sum(dhalfcauchy(lambda.sigma, 1, log=TRUE))
lsigma0.prior <- sum(dnorm(lsigma0, 0, 1, log=TRUE))
lSigma.prior <- sum(dnorm(lSigma,
  matrix(lsigma.mu, Data$T, Data$J, byrow=TRUE) +
  matrix(lsigma.phi, Data$T, Data$J, byrow=TRUE) *
  (rbind(lsigma0, lSigma[-Data$T,]) -
  matrix(lsigma.mu, Data$T, Data$J, byrow=TRUE)),
  matrix(lsigma.sigma, Data$T, Data$J, byrow=TRUE), log=TRUE))
lsigma.mu.prior <- sum(dnorm(lsigma.mu, 0, 1, log=TRUE))

```

```

lsigma.phi.prior <- sum(dbeta((lsigma.phi + 1) / 2, 20, 1.5, log=TRUE))
lsigma.sigma.prior <- sum(dhalfcauchy(lsigma.sigma, 1, log=TRUE))
### Log-Likelihood
LL <- 0; Yhat <- Data$Y; F.prior <- 0
for (t in 1:Data$T) {
  f.U.temp <- matrix(0, Data$P, Data$P)
  f.U.temp[upper.tri(f.U.temp, diag=TRUE)] <- matrix(f.U, nrow=Data$T,
    byrow=TRUE)[t,]
  diag(f.U.temp) <- exp(diag(f.U.temp))
  f.Sigma <- as.symmetric.matrix(t(f.U.temp) %**% f.U.temp)
  F.prior <- F.prior + dmvn(F[t,], rbind(f0, F)[t,] %**% diag(f.phi),
    f.Sigma, log=TRUE)
  Lambda.temp <- matrix(1, Data$P, Data$J)
  Lambda.temp[lower.tri(Lambda.temp)] <- 0
  Lambda.temp[upper.tri(Lambda.temp)] <- matrix(Lambda,
    nrow=Data$T, byrow=TRUE)[t,]*(abs(matrix(Lambda,
    nrow=Data$T, byrow=TRUE)[t,]) > lambda.d)
  mu <- Alpha[t,] + F[t,] %**% Lambda.temp
  LL <- LL + sum(dnorm(Data$Y[t,], mu, exp(lSigma[t,]), log=TRUE))
  Yhat[t,] <- rnorm(Data$J, mu, exp(lSigma[t,])) #Fitted
}
### Log-Posterior
LP <- LL + alpha0.prior + Alpha.prior + alpha.mu.prior +
  alpha.phi.prior + alpha.sigma.prior + f0.prior + F.prior +
  f.phi.prior + f.u0.prior + f.U.prior + f.u.mu.prior +
  f.u.phi.prior + f.u.sigma.prior + lambda0.prior +
  Lambda.prior + lambda.d.prior + lambda.mu.prior +
  lambda.phi.prior + lambda.sigma.prior + lsigma0.prior +
  lSigma.prior + lsigma.mu.prior + lsigma.phi.prior +
  lsigma.sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP, yhat=Yhat, parm=parm)
return(Modelout)
}

```

33.4. Initial Values

```

Initial.Values <- c(rnorm(J), rnorm(T*J), rnorm(J), runif(J), runif(J),
  rnorm(P), rnorm(T*P), rbeta(P,1,1)*2-1, rnorm(P*(P-1)/2+P),
  rnorm((P*(P-1)/2+P)*T), rnorm(P*(P-1)/2+P),
  rbeta(P*(P-1)/2+P,1,1)*2-1, runif(P*(P-1)/2+P),
  rnorm(P*J-P-P*(P-1)/2), rnorm((P*J-P-P*(P-1)/2)*T),
  runif(P*J-P-P*(P-1)/2,0,1e-3), rnorm(P*J-P-P*(P-1)/2),
  rbeta(P*J-P-P*(P-1)/2,20,1.5)*2-1, runif(P*J-P-P*(P-1)/2),
  rnorm(J), rnorm(T*J), rnorm(J), rbeta(J,20,1.5)*2-1, runif(J))

```

34. Exponential Smoothing

34.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\ \mu_t &= \alpha \mathbf{y}_{t-1} + (1 - \alpha) \mu_{t-1}, \quad t = 2, \dots, T \\ \alpha &\sim \mathcal{U}(0, 1) \\ \sigma &\sim \mathcal{HC} \end{aligned}$$

34.2. Data

```
data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
mon.names <- "LP"
parm.names <- c("alpha", "sigma")
PGF <- function(Data) {
  alpha <- runif(1)
  sigma <- runif(1)
  return(c(alpha, sigma))
}
MyData <- list(PGF=PGF, T=T, mon.names=mon.names, parm.names=parm.names,
              y=y)
```

34.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  parm[1] <- alpha <- interval(parm[1], 0, 1)
  parm[2] <- sigma <- interval(parm[2], 1e-100, Inf)
  ### Log-Prior
  alpha.prior <- dunif(alpha, 0, 1, log=TRUE)
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- y
  mu[-1] <- alpha*Data$y[-1]
  mu[-1] <- mu[-1] + (1 - alpha) * mu[-Data$T]
  LL <- sum(dnorm(Data$y[-1], mu[-Data$T], sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + alpha.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
                  yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
```

```
}
```

34.4. Initial Values

```
Initial.Values <- c(0.5, 1)
```

35. Factor Analysis, Approximate Dynamic

The Approximate Dynamic Factor Analysis (ADFA) model has many names, including the approximate factor model and approximate dynamic factor model. An ADFA is a Dynamic Factor Analysis (DFA) in which the factor scores of the dynamic factors are approximated with principal components. This is a combination of principal components and common factor analysis, in which the factor loadings of common factors are estimated from the data and factor scores are estimated from principal components. This is a two-stage model: principal components are estimated in the first stage and a decision is made regarding how many principal components to retain, and ADFA is estimated in the second stage. For more information on DFA, see section 33.

35.1. Form

$$\begin{aligned} \mathbf{Y}_{t,j} &\sim \mathcal{N}(\mu_{t,j}, \sigma_j^2), \quad t = 2, \dots, T, \quad j = 1, \dots, J \\ \mu_t &= \mathbf{F}_{t-1} \Lambda \\ \Lambda_{p,j} &\sim \mathcal{N}(0, 1), \quad p = 1, \dots, P, \quad j = 1, \dots, J \\ \sigma_j &\sim \mathcal{HC}(25), \quad j = 1, \dots, J \end{aligned}$$

35.2. Data

```
data(demonfx)
Y.orig <- as.matrix(demonfx)
Y <- diff(log(Y.orig[1:100,]))
Y.scales <- sqrt(.colVars(Y))
Y <- Y / matrix(Y.scales, nrow(Y), ncol(Y), byrow=TRUE)
T <- nrow(Y) #Number of time-periods
J <- ncol(Y) #Number of time-series
P <- 7 #Number of approximate factors
PCA <- prcomp(Y, scale=TRUE)
F <- PCA$x[,1:P]
mon.names <- "LP"
parm.names <- as.parm.names(list(Lambda=matrix(0,P,J), sigma=rep(0,J)))
pos.Lambda <- grep("Lambda", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  Lambda <- rnorm(Data$P*Data$J)
```

```

sigma <- runif(Data$J)
return(c(Lambda, sigma))
}
MyData <- list(F=F, J=J, P=P, PGF=PGF, T=T, Y=Y, mon.names=mon.names,
  parm.names=parm.names, pos.Lambda=pos.Lambda, pos.sigma=pos.sigma)

```

35.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  Lambda <- matrix(parm[Data$pos.Lambda], Data$P, Data$J)
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  Lambda.prior <- sum(dnorm(Lambda, 0, 1, log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  mu <- tcrossprod(rbind(rep(0,Data$P), F[-Data$T,]), t(Lambda))
  Sigma <- matrix(sigma, Data$T, Data$J, byrow=TRUE)
  LL <- sum(dnorm(Data$Y[-1,], mu[-1,], Sigma[-1,], log=TRUE))
  ### Log-Posterior
  LP <- LL + Lambda.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(prod(dim(mu)), mu, Sigma), parm=parm)
  return(Modelout)
}

```

35.4. Initial Values

```
Initial.Values <- c(rep(0,P*J), rep(1,J))
```

36. Factor Analysis, Confirmatory

Factor scores are in matrix \mathbf{F} , factor loadings for each variable are in vector λ , and \mathbf{f} is a vector that indicates which variable loads on which factor.

36.1. Form

$$\begin{aligned}
 \mathbf{Y}_{i,m} &\sim \mathcal{N}(\mu_{i,m}, \sigma_m^2), \quad i = 1, \dots, N, \quad m = 1, \dots, M \\
 \mu &= \mathbf{F}_{1:N,\mathbf{f}} \lambda^T \\
 \mathbf{F}_{i,1:P} &\sim \mathcal{N}_P(0, \Omega^{-1}), \quad i = 1, \dots, N \\
 \lambda_m &\sim \mathcal{N}(0, 1), \quad m = 1, \dots, M
 \end{aligned}$$

$$\sigma_m \sim \mathcal{HC}(25), \quad m = 1, \dots, M$$

$$\Omega \sim \mathcal{W}_N(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_P$$

36.2. Data

```

data(swiss)
Y <- cbind(swiss$Agriculture, swiss$Examination, swiss$Education,
           swiss$Catholic, swiss$Infant.Mortality)
M <- ncol(Y) #Number of variables
N <- nrow(Y) #Number of records
P <- 3 #Number of factors
f <- c(1,3,2,2,1) #Indicator f for the factor for each variable m
for (m in 1:M) Y[,m] <- CenterScale(Y[,m])
S <- diag(P)
mon.names <- "LP"
parm.names <- as.parm.names(list(F=matrix(0,N,P), lambda=rep(0,M),
                                U=diag(P), sigma=rep(0,M)), uppertri=c(0,0,1,0))
pos.F <- grep("F", parm.names)
pos.lambda <- grep("lambda", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  U <- rwishartc(Data$N, Data$S)
  F <- as.vector(rmvnpc(Data$N, rep(0,Data$P), U))
  U <- U[upper.tri(U, diag=TRUE)]
  lambda <- rnorm(Data$M)
  sigma <- runif(Data$M)
  return(c(F, lambda, U, sigma))
}
MyData <- list(M=M, N=N, P=P, PGF=PGF, S=S, Y=Y, f=f, mon.names=mon.names,
              parm.names=parm.names, pos.F=pos.F, pos.lambda=pos.lambda,
              pos.sigma=pos.sigma)

```

36.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  lambda <- parm[Data$pos.lambda]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  F <- matrix(parm[Data$pos.F], Data$N, Data$P)
  U <- as.parm.matrix(U, Data$P, parm, Data, chol=TRUE)
  diag(U) <- exp(diag(U))
  ### Log-Prior
  lambda.prior <- sum(dnorm(lambda, 0, 1, log=TRUE))
}

```

```

sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
U.prior <- dwishartc(U, Data$N, Data$S, log=TRUE)
F.prior <- sum(dmvnnc(F, rep(0,Data$P), U, log=TRUE))
### Log-Likelihood
mu <- F[,Data$f] * matrix(lambda, Data$N, Data$M, byrow=TRUE)
Sigma <- matrix(sigma, Data$N, Data$M, byrow=TRUE)
LL <- sum(dnorm(Data$Y, mu, Sigma, log=TRUE))
### Log-Posterior
LP <- LL + lambda.prior + sigma.prior + F.prior + U.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(prod(dim(mu)), mu, Sigma), parm=parm)
return(Modelout)
}

```

36.4. Initial Values

```

Initial.Values <- c(rep(0,N*P), rep(0,M), upper.triangle(S, diag=TRUE),
  rep(1,M))

```

37. Factor Analysis, Exploratory

Factor scores are in matrix \mathbf{F} and factor loadings are in matrix Λ .

37.1. Form

$$\mathbf{Y}_{i,m} \sim \mathcal{N}(\mu_{i,m}, \sigma_m^2), \quad i = 1, \dots, N, \quad m = 1, \dots, M$$

$$\mu = \mathbf{F}\Lambda$$

$$\mathbf{F}_{i,1:P} \sim \mathcal{N}_P(0, \Omega^{-1}), \quad i = 1, \dots, N$$

$$\Lambda_{p,m} \sim \mathcal{N}(0, 1), \quad p = 1, \dots, P, \quad m = (p+1), \dots, M$$

$$\Omega \sim \mathcal{W}_N(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_P$$

$$\sigma_m \sim \mathcal{HC}(25), \quad m = 1, \dots, M$$

37.2. Data

```

data(USJudgeRatings)
Y <- as.matrix(USJudgeRatings)
for (m in 1:M) Y[,m] <- CenterScale(Y[,m])
M <- ncol(Y) #Number of variables
N <- nrow(Y) #Number of records
P <- 3 #Number of factors
Lambda <- matrix(NA, P, M)
Lambda[upper.tri(Lambda)] <- 0

```



```

S <- diag(P)
mon.names <- "LP"
parm.names <- as.parm.names(list(F=matrix(0,N,P), Lambda=Lambda, U=S,
  sigma=rep(0,M)), uppertri=c(0,0,1,0))
pos.F <- grep("F", parm.names)
pos.Lambda <- grep("Lambda", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  U <- rwishartc(Data$N, Data$S)
  F <- as.vector(rmvnpc(Data$N, rep(0,Data$P), U))
  Lambda <- rnorm(Data$P*Data$M-Data$P-Data$P*(Data$P-1)/2,0,1)
  sigma <- runif(Data$M)
  return(c(F, Lambda, U[upper.tri(U, diag=TRUE)], sigma))
}
MyData <- list(M=M, N=N, P=P, PGF=PGF, S=S, Y=Y, mon.names=mon.names,
  parm.names=parm.names, pos.F=pos.F, pos.Lambda=pos.Lambda,
  pos.sigma=pos.sigma)

```

37.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  F <- matrix(parm[Data$pos.F], Data$N, Data$P)
  lambda <- parm[Data$pos.Lambda]
  U <- as.parm.matrix(U, Data$P, parm, Data, chol=TRUE)
  diag(U) <- exp(diag(U))
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  F.prior <- sum(dmvnpc(F, rep(0,Data$P), U, log=TRUE))
  Lambda.prior <- sum(dnorm(lambda, 0, 1, log=TRUE))
  U.prior <- dwishartc(U, Data$N, Data$S, log=TRUE)
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  Lambda <- matrix(1, Data$P, Data$M)
  Lambda[lower.tri(Lambda)] <- 0
  Lambda[upper.tri(Lambda)] <- lambda
  mu <- tcrossprod(F, t(Lambda))
  Sigma <- matrix(sigma, Data$N, Data$M, byrow=TRUE)
  LL <- sum(dnorm(Data$Y, mu, Sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + F.prior + Lambda.prior + U.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(prod(dim(mu)), mu, Sigma), parm=parm)
  return(Modelout)
}

```

```
}
```

37.4. Initial Values

```
Initial.Values <- c(rep(0,N*P), rep(0,P*M-P-P*(P-1)/2), rep(0,P*(P-1)/2+P),
  rep(1,M))
```

38. Factor Analysis, Exploratory Ordinal

This exploratory ordinal factor analysis (EOFA) model form is also suitable for collaborative filtering.

38.1. Form

$$\begin{aligned}
 \mathbf{Y}_{i,m} &\sim \mathcal{CAT}(\mathbf{P}_{i,m,1:K}), \quad i = 1, \dots, N, \quad m = 1, \dots, M \\
 \mathbf{P}_{:,K} &= 1 - Q_{:,(K-1)} \\
 \mathbf{P}_{:,k} &= |Q_{:,k} - Q_{:,(k-1)}|, \quad k = 2, \dots, (K-1) \\
 \mathbf{P}_{:,1} &= Q_{:,1} \\
 Q &= \phi(\mu) \\
 \mu_{:,k} &= \alpha_k - \mathbf{F}\Lambda, \quad k = 1, \dots, (K-1) \\
 \mathbf{F}_{i,1:P} &\sim \mathcal{N}_P(0, \Omega^{-1}), \quad i = 1, \dots, N \\
 \gamma_p &= 0, \quad p = 1, \dots, P \\
 \Lambda_{p,m} &\sim \mathcal{N}(0, 1), \quad p = 1, \dots, P, \quad m = (p+1), \dots, M \\
 \Omega &\sim \mathcal{W}_N(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_P \\
 \alpha_k &\sim \mathcal{N}(0, 1) \in [(k-1), k] \in [-5, 5], \quad k = 2, \dots, (K-1)
 \end{aligned}$$

38.2. Data

```
M <- 10 #Number of variables
N <- 20 #Number of records
K <- 3 #Number of discrete values
P <- 3 #Number of factors
alpha <- sort(rnorm(K-1))
Lambda <- matrix(1, P, M)
Lambda[lower.tri(Lambda)] <- 0
Lambda[upper.tri(Lambda)] <- rnorm(P*M-P-P*(P-1)/2)
Omega <- runif(P)
F <- rmvnp(N, rep(0,P), Omega)
mu <- aperm(array(alpha, dim=c(K-1, M, N)), perm=c(3,2,1))
mu <- mu - array(tcrossprod(F, t(Lambda)), dim=c(N, M, K-1))
```

```

Pr <- Q <- pnorm(mu)
Pr[, , -1] <- abs(Q[, , -1] - Q[, , -(K-1)])
Pr <- array(Pr, dim=c(N, M, K))
Pr[, , K] <- 1 - Q[, , (K-1)]
dim(Pr) <- c(N*M, K)
Y <- matrix(rcat(nrow(Pr), Pr), N, M) #Make sure Y has all values
S <- diag(P)
Lambda <- matrix(0, P, M)
Lambda[lower.tri(Lambda, diag=TRUE)] <- NA
mon.names <- "LP"
parm.names <- as.parm.names(list(F=matrix(0,N,P), Omega=rep(0,P),
  Lambda=Lambda, alpha=rep(0,K-1)))
pos.F <- grep("F", parm.names)
pos.Omega <- grep("Omega", parm.names)
pos.Lambda <- grep("Lambda", parm.names)
pos.alpha <- grep("alpha", parm.names)
PGF <- function(Data) {
  Omega <- runif(Data$P)
  F <- as.vector(rmvnp(Data$N, rep(0,Data$P), diag(Omega)))
  Lambda <- rnorm(Data$P*Data$M-Data$P-Data$P*(Data$P-1)/2)
  alpha <- sort(rnorm(Data$K-1))
  return(c(F, Omega, Lambda, alpha))
}
MyData <- list(K=K, M=M, N=N, P=P, PGF=PGF, S=S, Y=Y,
  mon.names=mon.names, parm.names=parm.names, pos.F=pos.F,
  pos.Omega=pos.Omega, pos.Lambda=pos.Lambda, pos.alpha=pos.alpha)

```

38.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  F <- matrix(parm[Data$pos.F], Data$N, Data$P)
  Omega <- interval(parm[Data$pos.Omega], 1e-100, Inf)
  parm[Data$pos.Omega] <- Omega
  lambda <- parm[Data$pos.Lambda]
  alpha <- sort(interval(parm[Data$pos.alpha], -5, 5))
  parm[Data$pos.alpha] <- alpha
  ### Log-Prior
  F.prior <- sum(dmvnp(F, rep(0,Data$P), diag(Omega), log=TRUE))
  Omega.prior <- dwishart(diag(Omega), Data$N, Data$S, log=TRUE)
  Lambda.prior <- sum(dnorm(lambda, 0, 1, log=TRUE))
  alpha.prior <- sum(dnormv(alpha, 0, 10, log=TRUE))
  ### Log-Likelihood
  Lambda <- matrix(1, Data$P, Data$M)
  Lambda[lower.tri(Lambda)] <- 0

```

```

Lambda[upper.tri(Lambda)] <- lambda
mu <- aperm(array(alpha, dim=c(Data$K-1, Data$M, Data$N)),
            perm=c(3,2,1))
mu <- mu - array(tcrossprod(F, t(Lambda)),
                dim=c(Data$N, Data$M, Data$K-1))
P <- Q <- pnorm(mu)
P[ , , -1] <- abs(Q[ , , -1] - Q[ , , -(Data$K-1)])
P <- array(P, dim=c(Data$N, Data$M, Data$K))
P[ , , Data$K] <- 1 - Q[ , , (Data$K-1)]
y <- as.vector(Data$Y)
dim(P) <- c(Data$N*Data$M, Data$K)
LL <- sum(dcat(y, P, log=TRUE))
### Log-Posterior
LP <- LL + F.prior + Omega.prior + Lambda.prior + alpha.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
                yhat=matrix(rcat(nrow(P), P), Data$N, Data$M), parm=parm)
return(Modelout)
}

```

38.4. Initial Values

```

Initial.Values <- c(rep(0,N*P), rep(0,P), rep(0,P*M-P-P*(P-1)/2),
                  seq(from=-1, to=1, len=K-1))

```

39. Factor Regression

This example of factor regression is constrained to the case where the number of factors is equal to the number of independent variables (IVs) less the intercept. The purpose of this form of factor regression is to orthogonalize the IVs with respect to \mathbf{y} , rather than variable reduction. This method is the combination of confirmatory factor analysis in section 36 and linear regression in section 48.

39.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\nu, \sigma_{J+1}^2) \\
 \nu &= \mathbf{F}\beta \\
 \mu_{i,1} &= 1 \\
 \mu_{i,j+1} &= \mu_{i,j}, \quad j = 1, \dots, J \\
 \mathbf{X}_{i,j} &\sim \mathcal{N}(\mu_{i,j}, \sigma_j^2), \quad i = 1, \dots, N, \quad j = 2, \dots, J \\
 \mu_{i,j} &= \lambda_j \mathbf{F}_{i,j}, \quad i = 1, \dots, N, \quad j = 2, \dots, J \\
 \mathbf{F}_{i,1:J} &\sim \mathcal{N}_{J-1}(0, \Omega^{-1}), \quad i = 1, \dots, N \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J
 \end{aligned}$$

$$\lambda_j \sim \mathcal{N}(0, 1), \quad j = 1, \dots, (J - 1)$$

$$\sigma_j \sim \mathcal{HC}(25), \quad j = 1, \dots, (J + 1)$$

$$\Omega \sim \mathcal{W}_N(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_J$$

39.2. Data

```

data(demonsnacks)
N <- nrow(demonsnacks)
y <- log(demonsnacks$Calories)
X <- as.matrix(log(demonsnacks[,c(1,4,10)]+1))
J <- ncol(X)
for (j in 1:J) X[,j] <- CenterScale(X[,j])
S <- diag(J)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J+1), lambda=rep(0,J),
  sigma=rep(0,J+1), F=matrix(0,N,J), Omega=rep(0,J)))
pos.beta <- grep("beta", parm.names)
pos.lambda <- grep("lambda", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.F <- grep("F", parm.names)
pos.Omega <- grep("Omega", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J+1)
  lambda <- rnorm(Data$J)
  sigma <- runif(Data$J+1)
  Omega <- runif(Data$J)
  F <- as.vector(rmvnp(Data$N, rep(0,Data$J), diag(Omega)))
  return(c(beta, lambda, sigma, F, Omega))
}
MyData <- list(J=J, N=N, PGF=PGF, S=S, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.lambda=pos.lambda,
  pos.sigma=pos.sigma, pos.F=pos.F, pos.Omega=pos.Omega, y=y)

```

39.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  lambda <- parm[Data$pos.lambda]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  F <- matrix(Data$pos.F, Data$N, Data$J)
  Omega <- interval(parm[Data$pos.Omega], 1e-100, Inf)
  parm[Data$pos.Omega] <- Omega

```

```

### Log-Prior
beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
lambda.prior <- sum(dnorm(lambda, 0, 1, log=TRUE))
sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
F.prior <- sum(dmvnp(F, rep(0,Data$J), diag(Omega), log=TRUE))
Omega.prior <- dwishart(diag(Omega), Data$N, Data$S, log=TRUE)
### Log-Likelihood
mu <- F * matrix(lambda, Data$N, Data$J, byrow=TRUE)
nu <- tcrossprod(cbind(1,F), t(beta))
LL <- sum(dnorm(Data$X, mu, matrix(sigma[1:Data$J], Data$N, Data$J,
  byrow=TRUE), log=TRUE))
LL <- LL + dnorm(Data$y, nu, sigma[Data$J+1], log=TRUE)
### Log-Posterior
LP <- LL + beta.prior + lambda.prior + sigma.prior + F.prior
  Omega.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(Data$N, nu, sigma[Data$J+1]), parm=parm)
return(Modelout)
}

```

39.4. Initial Values

```
Initial.Values <- c(rep(0,J+1), rep(0,J), rep(0,J+1), rep(0,N*J), rep(1,J))
```

40. Gamma Regression

40.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{G}(\lambda\tau, \tau) \\
 \lambda &= \exp(\mathbf{X}\beta) \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \tau &\sim \mathcal{HC}(25)
 \end{aligned}$$

40.2. Data

```

N <- 20
J <- 3
X <- matrix(runif(N*J,-2,2),N,J); X[,1] <- 1
beta <- runif(J,-2,2)
y <- round(exp(tcrossprod(X, t(beta)))) + 0.1 #Must be > 0
mon.names <- c("LP","sigma2")
parm.names <- as.parm.names(list(beta=rep(0,J), tau=0))

```

```

pos.beta <- grep("beta", parm.names)
pos.tau <- grep("tau", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  tau <- runif(1)
  return(c(beta, tau))
}
MyData <- list(J=J, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.tau=pos.tau, y=y)

```

40.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  tau <- interval(parm[Data$pos.tau], 1e-100, Inf)
  parm[Data$pos.tau] <- tau
  sigma2 <- 1/tau
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  tau.prior <- dhalfcauchy(tau, 25, log=TRUE)
  ### Log-Likelihood
  lambda <- exp(tcrossprod(Data$X, t(beta)))
  LL <- sum(dgamma(Data$y, tau*lambda, tau, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + tau.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,sigma2),
    yhat=rgamma(nrow(lambda), tau*lambda, tau), parm=parm)
  return(Modelout)
}

```

40.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1)
```

41. Geographically Weighted Regression

41.1. Form

$$y_{i,k} \sim \mathcal{N}(\mu_{i,k}, \tau_{i,k}^{-1}), \quad i = 1, \dots, N, \quad k = 1, \dots, N$$

$$\mu_{i,1:N} = \mathbf{X}_i \beta_{i,1:J}$$

$$\begin{aligned}\tau &= \frac{1}{\sigma^2} \mathbf{w} \nu \\ \mathbf{w} &= \frac{\exp(-0.5 \mathbf{Z}^2)}{\mathbf{h}} \\ \alpha &\sim \mathcal{U}(1.5, 100) \\ \beta_{i,j} &\sim \mathcal{N}(0, 1000), \quad i = 1, \dots, N, \quad j = 1, \dots, J \\ \mathbf{h} &\sim \mathcal{N}(0.1, 1000) \in [0.1, \infty] \\ \nu_{i,k} &\sim \mathcal{G}(\alpha, 2), \quad i = 1, \dots, N, \quad k = 1, \dots, N \\ \sigma_i &\sim \mathcal{HC}(25), \quad i = 1, \dots, N\end{aligned}$$

41.2. Data

```
crime <- c(18.802, 32.388, 38.426, 0.178, 15.726, 30.627, 50.732,
  26.067, 48.585, 34.001, 36.869, 20.049, 19.146, 18.905, 27.823,
  16.241, 0.224, 30.516, 33.705, 40.970, 52.794, 41.968, 39.175,
  53.711, 25.962, 22.541, 26.645, 29.028, 36.664, 42.445, 56.920,
  61.299, 60.750, 68.892, 38.298, 54.839, 56.706, 62.275, 46.716,
  57.066, 54.522, 43.962, 40.074, 23.974, 17.677, 14.306, 19.101,
  16.531, 16.492)
income <- c(21.232, 4.477, 11.337, 8.438, 19.531, 15.956, 11.252,
  16.029, 9.873, 13.598, 9.798, 21.155, 18.942, 22.207, 18.950,
  29.833, 31.070, 17.586, 11.709, 8.085, 10.822, 9.918, 12.814,
  11.107, 16.961, 18.796, 11.813, 14.135, 13.380, 17.017, 7.856,
  8.461, 8.681, 13.906, 14.236, 7.625, 10.048, 7.467, 9.549,
  9.963, 11.618, 13.185, 10.655, 14.948, 16.940, 18.739, 18.477,
  18.324, 25.873)
housing <- c(44.567, 33.200, 37.125, 75.000, 80.467, 26.350, 23.225,
  28.750, 18.000, 96.400, 41.750, 47.733, 40.300, 42.100, 42.500,
  61.950, 81.267, 52.600, 30.450, 20.300, 34.100, 23.600, 27.000,
  22.700, 33.500, 35.800, 26.800, 27.733, 25.700, 43.300, 22.850,
  17.900, 32.500, 22.500, 53.200, 18.800, 19.900, 19.700, 41.700,
  42.900, 30.600, 60.000, 19.975, 28.450, 31.800, 36.300, 39.600,
  76.100, 44.333)
easting <- c(35.62, 36.50, 36.71, 33.36, 38.80, 39.82, 40.01, 43.75,
  39.61, 47.61, 48.58, 49.61, 50.11, 51.24, 50.89, 48.44, 46.73,
  43.44, 43.37, 41.13, 43.95, 44.10, 43.70, 41.04, 43.23, 42.67,
  41.21, 39.32, 41.09, 38.3, 41.31, 39.36, 39.72, 38.29, 36.60,
  37.60, 37.13, 37.85, 35.95, 35.72, 35.76, 36.15, 34.08, 30.32,
  27.94, 27.27, 24.25, 25.47, 29.02)
northing <- c(42.38, 40.52, 38.71, 38.41, 44.07, 41.18, 38.00, 39.28,
  34.91, 36.42, 34.46, 32.65, 29.91, 27.80, 25.24, 27.93, 31.91,
  35.92, 33.46, 33.14, 31.61, 30.40, 29.18, 28.78, 27.31, 24.96,
  25.90, 25.85, 27.49, 28.82, 30.90, 32.88, 30.64, 30.35, 32.09,
  34.08, 36.12, 36.30, 36.40, 35.60, 34.66, 33.92, 30.42, 28.26,
```



```

29.85, 28.21, 26.69, 25.71, 26.58)
N <- length(crime)
J <- 3 #Number of predictors, including the intercept
X <- matrix(c(rep(1,N), income, housing),N,J)
D <- as.matrix(dist(cbind(northing,easting), diag=TRUE, upper=TRUE))
Z <- D / sd(as.vector(D))
y <- matrix(0,N,N); for (i in 1:N) {for (k in 1:N) {y[i,k] <- crime[k]}}
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=0, beta=matrix(0,N,J), H=0,
  nu=matrix(0,N,N), sigma=rep(0,N)))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.H <- grep("H", parm.names)
pos.nu <- grep("nu", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- runif(1,1.5,100)
  beta <- rnorm(Data$N*Data$J)
  H <- runif(1,0.1,1000)
  nu <- rgamma(Data$N*Data$N,alpha,2)
  sigma <- runif(Data$N)
  return(c(alpha, beta, H, nu, sigma))
}
MyData <- list(J=J, N=N, PGF=PGF, X=X, Z=Z, latitude=northing,
  longitude=easting, mon.names=mon.names, parm.names=parm.names,
  pos.alpha=pos.alpha, pos.beta=pos.beta, pos.H=pos.H, pos.nu=pos.nu,
  pos.sigma=pos.sigma, y=y)

```

41.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- interval(parm[Data$pos.alpha], 1.5, 100)
  parm[Data$pos.alpha] <- alpha
  beta <- matrix(parm[Data$pos.beta], Data$N, Data$J)
  parm[Data$pos.H] <- H <- interval(parm[Data$pos.H], 0.1, Inf)
  parm[Data$pos.nu] <- nu <- interval(parm[Data$pos.nu], 1e-100, Inf)
  nu <- matrix(nu, Data$N, Data$N)      sigma <- interval(parm[Data$pos.sigma],
1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- dunif(alpha, 1.5, 100, log=TRUE)
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  h.prior <- dhalfnorm(H-0.1, 1000, log=TRUE)
  nu.prior <- sum(dgamma(nu, alpha, 2, log=TRUE))

```

```

sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
### Log-Likelihood
w <- exp(-0.5 * Data$Z^2) / H
tau <- (1/sigma^2) * w * nu
mu <- tcrossprod(Data$X, beta)
LL <- sum(dnormp(Data$y, mu, tau, log=TRUE))
#WSE <- w * nu * (Data$y - mu)^2; w.y <- w * nu * Data$y
#WMSE <- rowMeans(WSE); y.w <- rowSums(w.y) / rowSums(w)
#LAR2 <- 1 - WMSE / sd(y.w)^2
### Log-Posterior
LP <- LL + alpha.prior + beta.prior + h.prior + nu.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnormp(prod(dim(mu)), mu, tau), parm=parm)
return(Modelout)
}

```

41.4. Initial Values

```
Initial.Values <- c(runif(1,1.5,100), rep(0,N*J), 1, rep(1,N*N), rep(1,N))
```

42. Hidden Markov Model

42.1. Form

This introductory hidden Markov model (HMM) includes N discrete states.

$$\begin{aligned}
 \mathbf{y}_t &\sim \mathcal{N}(\mu_\theta, \sigma_\theta^2), \quad t = 1, \dots, T \\
 \mu &\sim \mathcal{N}(\mu_0, \sigma^2) \\
 \sigma^2 &\sim \mathcal{HC}(25) \\
 \theta_t &\sim \mathcal{CAT}(\phi_{\theta_{t-1}, 1:N}), \quad t = 1, \dots, T \\
 \phi_{i, 1:N} &\sim \mathcal{D}(\alpha_{1:N}), \quad i = 1, \dots, N \\
 \mu_0 &\sim \mathcal{N}(0, 1000) \\
 \sigma_0^2 &\sim \mathcal{HC}(25)
 \end{aligned}$$

42.2. Data

```

data(demonfx)
y <- as.vector(log(as.matrix(demonfx[1:50,1])))
T <- length(y) #Number of time-periods
N <- 2 #Number of discrete (hidden) states
alpha <- matrix(1,N,N) #Concentration hyperparameter
mon.names <- "LP"

```

```

parm.names <- as.parm.names(list(mu0=rep(0,N), mu1=rep(0,N),
  phi=matrix(0,N,N), sigma2=rep(0,N), theta=rep(0,T)))
pos.mu0 <- grep("mu0", parm.names)
pos.mu1 <- grep("mu1", parm.names)
pos.phi <- grep("phi", parm.names)
pos.sigma2 <- grep("sigma2", parm.names)
pos.theta <- grep("theta", parm.names)
PGF <- function(Data) {
  mu0 <- sort(runif(Data$N, min(Data$y), max(Data$y)))
  mu1 <- sort(runif(Data$N, min(Data$y), max(Data$y)))
  phi <- matrix(runif(Data$N*Data$N), Data$N, Data$N)
  phi <- as.vector(phi / rowSums(phi))
  sigma2 <- runif(Data$N)
  theta <- rcat(Data$T, rep(1/Data$N,Data$N))
  return(c(mu0, mu1, phi, sigma2, theta))
}
MyData <- list(N=N, PGF=PGF, T=T, alpha=alpha, mon.names=mon.names,
  parm.names=parm.names, pos.mu0=pos.mu0, pos.mu1=pos.mu1,
  pos.phi=pos.phi, pos.sigma2=pos.sigma2, pos.theta=pos.theta, y=y)

```

42.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  mu0 <- interval(parm[Data$pos.mu0], min(Data$y), max(Data$y))
  parm[Data$pos.mu0] <- mu0
  mu <- interval(parm[Data$pos.mu1], min(Data$y), max(Data$y))
  parm[Data$pos.mu1] <- mu <- sort(mu)
  phi <- matrix(abs(parm[Data$pos.phi]), Data$N, Data$N)
  parm[Data$pos.phi] <- phi <- phi / rowSums(phi)
  sigma2 <- interval(parm[Data$pos.sigma2], 1e-100, Inf)
  parm[Data$pos.sigma2] <- sigma2
  theta <- parm[Data$pos.theta]
  ### Log-Hyperprior
  mu0.prior <- sum(dnormv(mu0, 0, 1000, log=TRUE))
  ### Log-Prior
  mu.prior <- sum(dnormv(mu, mu0, sigma2, log=TRUE))
  phi.prior <- 0
  for (i in 1:Data$N)
    phi.prior <- phi.prior + sum(ddirichlet(phi[i,], Data$alpha[i,],
      log=TRUE))
  sigma2.prior <- sum(dhalfcauchy(sigma2, 25, log=TRUE))
  theta.prior <- sum(dcat(theta, rbind(rep(1/Data$N,Data$N),
    phi[theta[-Data$T],]), log=TRUE))
  ### Log-Likelihood

```

```

LL <- sum(dnormv(Data$y, mu[theta], sigma2[theta], log=TRUE))
### Log-Posterior
LP <- LL + mu0.prior + mu.prior + phi.prior + sigma2.prior + theta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnormv(length(theta), mu[theta], sigma2[theta]), parm=parm)
return(Modelout)
}

```

42.4. Initial Values

```

Initial.Values <- c(sort(runif(N, min(y), max(y))),
  sort(runif(N, min(y), max(y))), runif(N*N), runif(N),
  rcat(T, rep(1/N,N)))

```

43. Inverse Gaussian Regression

43.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}^{-1}(\boldsymbol{\mu}, \lambda) \\
 \boldsymbol{\mu} &= \exp(\mathbf{X}\boldsymbol{\beta}) + C \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \lambda &\sim \mathcal{HC}(25)
 \end{aligned}$$

where C is a small constant, such as 1.0E-10.

43.2. Data

```

N <- 100
J <- 3 #Number of predictors, including the intercept
X <- matrix(1,N,J)
for (j in 2:J) {X[,j] <- rnorm(N,runif(1,-3,3),runif(1,0.1,1))}
beta.orig <- runif(J,-3,3)
e <- rnorm(N,0,0.1)
y <- exp(tcrossprod(X, t(beta.orig)) + e)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), lambda=0))
pos.beta <- grep("beta", parm.names)
pos.lambda <- grep("lambda", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  lambda <- runif(1)
  return(c(beta, lambda))
}

```

```
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.lambda=pos.lambda, y=y)
```

43.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  lambda <- interval(parm[Data$pos.lambda], 1e-100, Inf)
  parm[Data$pos.lambda] <- lambda
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  lambda.prior <- dhalfcauchy(lambda, 25, log=TRUE)
  ### Log-Likelihood
  mu <- exp(tcrossprod(Data$X, t(beta))) + 1.0E-10
  LL <- sum(dinvgaussian(Data$y, mu, lambda, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + lambda.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rinvgaussian(length(mu), mu, lambda), parm=parm)
  return(Modelout)
}
```

43.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1)
```

44. Kriging

This is an example of universal kriging of \mathbf{y} given \mathbf{X} , regression effects β , and spatial effects ζ . Euclidean distance between spatial coordinates (longitude and latitude) is used for each of $i = 1, \dots, N$ records of \mathbf{y} . An additional record is created from the same data-generating process to compare the accuracy of interpolation. For the spatial component, ϕ is the rate of spatial decay and κ is the scale. κ is often difficult to identify, so it is set to 1 (Gaussian), but may be allowed to vary up to 2 (Exponential). In practice, ϕ is also often difficult to identify. While Σ is spatial covariance, spatial correlation is $\rho = \exp(-\phi\mathbf{D})$. To extend this to a large data set, consider the predictive process kriging example in section 45.

44.1. Form

$$\mathbf{y} \sim \mathcal{N}(\mu, \sigma_1^2)$$

$$\mu = \mathbf{X}\beta + \zeta$$

$$\begin{aligned}
\mathbf{y}^{new} &= \mathbf{X}\beta + \sum_{i=1}^N \left(\frac{\rho_i}{\sum \rho} \zeta_i \right) \\
\rho &= \exp(-\phi \mathbf{D}^{new})^\kappa \\
\zeta &\sim \mathcal{N}_N(\zeta_\mu, \Sigma) \\
\Sigma &= \sigma_2^2 \exp(-\phi \mathbf{D})^\kappa \\
\beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, 2 \\
\sigma_j &\sim \mathcal{HC}(25) \in [0.1, 10], \quad j = 1, \dots, 2 \\
\phi &\sim \mathcal{U}(1, 5) \\
\zeta_\mu &= 0 \\
\kappa &= 1
\end{aligned}$$

44.2. Data

```

N <- 20
longitude <- runif(N+1,0,100)
latitude <- runif(N+1,0,100)
D <- as.matrix(dist(cbind(longitude,latitude), diag=TRUE, upper=TRUE))
Sigma <- 10000 * exp(-1.5 * D)
zeta <- colMeans(rmvn(1000, rep(0,N+1), Sigma))
beta <- c(50,2)
X <- matrix(runif((N+1)*2,-2,2),(N+1),2); X[,1] <- 1
mu <- as.vector(tcrossprod(X, t(beta)))
y <- mu + zeta
longitude.new <- longitude[N+1]; latitude.new <- latitude[N+1]
Xnew <- X[N+1,]; ynew <- y[N+1]
longitude <- longitude[1:N]; latitude <- latitude[1:N]
X <- X[1:N,]; y <- y[1:N]
D <- as.matrix(dist(cbind(longitude,latitude), diag=TRUE, upper=TRUE))
D.new <- sqrt((longitude - longitude.new)^2 + (latitude - latitude.new)^2)
mon.names <- c("LP","ynew")
parm.names <- as.parm.names(list(zeta=rep(0,N), beta=rep(0,2),
  sigma=rep(0,2), phi=0))
pos.zeta <- grep("zeta", parm.names)
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.phi <- grep("phi", parm.names)
PGF <- function(Data) {
  beta <- rnorm(2)
  sigma <- runif(2,0.1,10)
  phi <- runif(1,1,5)
  kappa <- 1
  zeta <- rmvn(1, rep(0,Data$N),
    sigma[2]*sigma[2]*exp(-phi*Data$D)^kappa)

```

```

    return(c(zeta, beta, sigma, phi))
  }
MyData <- list(D=D, D.new=D.new, latitude=latitude, longitude=longitude,
  N=N, PGF=PGF, X=X, Xnew=Xnew, mon.names=mon.names,
  parm.names=parm.names, pos.zeta=pos.zeta, pos.beta=pos.beta,
  pos.sigma=pos.sigma, pos.phi=pos.phi, y=y)

```

44.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  zeta <- parm[Data$pos.zeta]
  kappa <- 1
  sigma <- interval(parm[Data$pos.sigma], 0.1, 10)
  parm[Data$pos.sigma] <- sigma
  parm[Data$pos.phi] <- phi <- interval(parm[Data$pos.phi], 1, 5)
  Sigma <- sigma[2]*sigma[2] * exp(-phi * Data$D)^kappa
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  zeta.prior <- dmvn(zeta, rep(0, Data$N), Sigma, log=TRUE)
  sigma.prior <- sum(dhalfcauchy(sigma - 1, 25, log=TRUE))
  phi.prior <- dunif(phi, 1, 5, log=TRUE)
  ### Interpolation
  rho <- exp(-phi * Data$D.new)^kappa
  ynew <- rnorm(1, sum(beta * Data$Xnew) + sum(rho / sum(rho) * zeta),
    sigma[1])
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta)) + zeta
  LL <- sum(dnorm(Data$y, mu, sigma[1], log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + zeta.prior + sigma.prior + phi.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,ynew),
    yhat=rnorm(length(mu), mu, sigma[1]), parm=parm)
  return(Modelout)
}

```

44.4. Initial Values

```
Initial.Values <- c(rep(0,N), rep(0,2), rep(1,2), 1)
```

45. Kriging, Predictive Process

The first K of N records in \mathbf{y} are used as knots for the parent process, and the predictive

process involves records $(K + 1), \dots, N$. For more information on kriging, see section 44.

45.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\mu, \sigma_1^2) \\
 \mu_{1:K} &= \mathbf{X}_{1:K,1:J}\beta + \zeta \\
 \mu_{(K+1):N} &= \mathbf{X}_{(K+1):N,1:J}\beta + \sum_{p=1}^{N-K} \frac{\lambda_{p,1:K}}{\sum_{q=1}^{N-K} \lambda_{q,1:K}} \zeta^T \\
 \lambda &= \exp(-\phi \mathbf{D}_P)^\kappa \\
 \mathbf{y}^{new} &= \mathbf{X}\beta + \sum_{k=1}^K \left(\frac{\rho_k}{\sum \rho} \zeta_k \right) \\
 \rho &= \exp(-\phi \mathbf{D}^{new})^\kappa \\
 \zeta &\sim \mathcal{N}_K(0, \Sigma) \\
 \Sigma &= \sigma_2^2 \exp(-\phi \mathbf{D})^\kappa \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, 2 \\
 \sigma_j &\sim \mathcal{HC}(25), \quad j = 1, \dots, 2 \\
 \phi &\sim \mathcal{N}(0, 1000) \in [1, 5] \\
 \kappa &= 1
 \end{aligned}$$

45.2. Data

```

N <- 100
K <- 30 #Number of knots
longitude <- runif(N+1,0,100)
latitude <- runif(N+1,0,100)
D <- as.matrix(dist(cbind(longitude,latitude), diag=TRUE, upper=TRUE))
Sigma <- 10000 * exp(-1.5 * D)
zeta <- colMeans(rmvn(1000, rep(0,N+1), Sigma))
beta <- c(50,2)
X <- matrix(runif((N+1)*2,-2,2),(N+1),2); X[,1] <- 1
mu <- as.vector(tcrossprod(X, t(beta)))
y <- mu + zeta
longitude.new <- longitude[N+1]; latitude.new <- latitude[N+1]
Xnew <- X[N+1,]; ynew <- y[N+1]
longitude <- longitude[1:N]; latitude <- latitude[1:N]
X <- X[1:N,]; y <- y[1:N]
D <- as.matrix(dist(cbind(longitude[1:K],latitude[1:K]), diag=TRUE,
  upper=TRUE))
D.P <- matrix(0, N-K, K)
for (i in (K+1):N) {

```



```

D.P[K+1-i,] <- sqrt((longitude[1:K] - longitude[i])^2 +
  (latitude[1:K] - latitude[i])^2)}
D.new <- sqrt((longitude[1:K] - longitude.new)^2 +
  (latitude[1:K] - latitude.new)^2)
mon.names <- c("LP","ynew")
parm.names <- as.parm.names(list(zeta=rep(0,K), beta=rep(0,2),
  sigma=rep(0,2), phi=0))
pos.zeta <- grep("zeta", parm.names)
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.phi <- grep("phi", parm.names)
PGF <- function(Data) {
  beta <- rnorm(2)
  sigma <- runif(2,0.1,10)
  phi <- runif(1,1,5)
  kappa <- 1
  zeta <- rmvn(1, rep(0,Data$K),
    sigma[2]*sigma[2]*exp(-phi*Data$D)^kappa)
  return(c(zeta, beta, sigma, phi))
}
MyData <- list(D=D, D.new=D.new, D.P=D.P, K=K, N=N, PGF=PGF, X=X,
  Xnew=Xnew, latitude=latitude, longitude=longitude,
  mon.names=mon.names, parm.names=parm.names, pos.zeta=pos.zeta,
  pos.beta=pos.beta, pos.sigma=pos.sigma, pos.phi=pos.phi, y=y)

```

45.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  zeta <- parm[Data$pos.zeta]
  kappa <- 1
  sigma <- interval(parm[Data$pos.sigma], 1, Inf)
  parm[Data$pos.sigma] <- sigma
  parm[Data$pos.phi] <- phi <- interval(parm[Data$pos.phi], 1, 5)
  Sigma <- sigma[2]*sigma[2] * exp(-phi * Data$D)^kappa
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  zeta.prior <- dmvn(zeta, rep(0, Data$K), Sigma, log=TRUE)
  sigma.prior <- sum(dhalfcauchy(sigma - 1, 25, log=TRUE))
  phi.prior <- dunif(phi, 1, 5, log=TRUE)
  ### Interpolation
  rho <- exp(-phi * Data$D.new)^kappa
  ynew <- rnorm(1, sum(beta * Data$Xnew) + sum(rho / sum(rho) * zeta),
    sigma)
}

```

```

### Log-Likelihood
mu <- tcrossprod(Data$X, t(beta))
mu[1:Data$K] <- mu[1:Data$K] + zeta
lambda <- exp(-phi * Data$D.P)^kappa
mu[(Data$K+1):Data$N] <- mu[(Data$K+1):Data$N] +
  rowSums(lambda / rowSums(lambda) *
    matrix(zeta, Data$N - Data$K, Data$K, byrow=TRUE))
LL <- sum(dnorm(Data$y, mu, sigma[1], log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + zeta.prior + sigma.prior + phi.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,ynew),
  yhat=rnorm(length(mu), mu, sigma[1]), parm=parm)
return(Modelout)
}

```

45.4. Initial Values

```
Initial.Values <- c(rep(0,K), c(mean(y), 0), rep(1,2), 3)
```

46. Laplace Regression

This linear regression specifies that \mathbf{y} is Laplace-distributed, where it is usually Gaussian or normally-distributed. It has been claimed that it should be surprising that the normal distribution became the standard, when the Laplace distribution usually fits better and has wider tails (Kotz, Kozubowski, and Podgorski 2001). Another popular alternative is to use the t-distribution (see Robust Regression in section 84), though it is more computationally expensive to estimate, because it has three parameters. The Laplace distribution has only two parameters, location and scale like the normal distribution, and is computationally easier to fit. This example could be taken one step further, and the parameter vector β could be Laplace-distributed. Laplace's Demon recommends that users experiment with replacing the normal distribution with the Laplace distribution.

46.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{L}(\mu, \sigma^2) \\
 \mu &= \mathbf{X}\beta \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \sigma &\sim \mathcal{HC}(25)
 \end{aligned}$$

46.2. Data

```
N <- 10000
J <- 5
```

```

X <- matrix(1,N,J)
for (j in 2:J) {X[,j] <- rnorm(N,runif(1,-3,3),runif(1,0.1,1))}
beta <- runif(J,-3,3)
e <- rlaplace(N,0,0.1)
y <- tcrossprod(X, t(beta)) + e
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), sigma=0))
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  sigma <- runif(1)
  return(c(beta, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma, y=y)

```

46.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  LL <- sum(dlaplace(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rlaplace(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

46.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1)
```

47. Latent Dirichlet Allocation

47.1. Form

$$\begin{aligned}
 \mathbf{Y}_{m,n} &\sim \text{CAT}(\phi[\mathbf{Z}_{m,n},]), \quad m = 1, \dots, M, \quad n = 1, \dots, N \\
 \mathbf{Z}_{m,n} &\sim \text{CAT}(\theta_{m,1:K}) \\
 \phi_{k,v} &\sim \mathcal{D}(\beta) \\
 \theta_{m,k} &\sim \mathcal{D}(\alpha) \\
 \alpha_k &= 1, \quad k = 1, \dots, K \\
 \beta_v &= 1, \quad v = 1, \dots, V
 \end{aligned}$$

47.2. Data

```

K <- 2 #Number of (latent) topics
M <- 4 #Number of documents in corpus
N <- 15 #Maximum number of (used) words per document
V <- 5 #Maximum number of occurrences of any word (Vocabulary size)
Y <- matrix(rcat(M*N,rep(1/V,V)), M, N)
rownames(Y) <- paste("doc", 1:nrow(Y), sep="")
colnames(Y) <- paste("word", 1:ncol(Y), sep="")
#Note: Y is usually represented as w, a matrix of word counts.
if(min(Y) == 0) Y <- Y + 1 #A zero cannot occur, Y must be 1,2,...,V.
V <- max(Y) #Maximum number of occurrences of any word (Vocabulary size)
alpha <- rep(1,K) # hyperparameters (constant)
beta <- rep(1,V)
mon.names <- "LP"
parm.names <- as.parm.names(list(phi=matrix(0,K,V), theta=matrix(0,M,K),
  Z=matrix(0,M,N)))
pos.phi <- grep("phi", parm.names)
pos.theta <- grep("theta", parm.names)
pos.Z <- grep("Z", parm.names)
PGF <- function(Data) {
  phi <- matrix(runif(Data$J*Data$V), Data$K, Data$V)
  phi <- phi / rowSums(phi)
  theta <- matrix(runif(Data$M*Data$K), Data$M, Data$K)
  theta <- theta / rowSums(theta)
  z <- rcat(Data$M*Data$N, rep(1/Data$K,Data$K))
  return(c(as.vector(phi), as.vector(theta), z))}
MyData <- list(K=K, M=M, N=N, PGF=PGF, V=V, Y=Y, alpha=alpha, beta=beta,
  mon.names=mon.names, parm.names=parm.names, pos.phi=pos.phi,
  pos.theta=pos.theta, pos.Z=pos.Z)

```

47.3. Model

```

Model <- function(parm, Data)
{

```

```

### Parameters
phi <- matrix(interval(parm[Data$pos.phi], 0, 1), Data$K, Data$V)
phi <- phi / rowSums(phi)
parm[Data$pos.phi] <- as.vector(phi)
theta <- matrix(interval(parm[Data$pos.theta], 0, 1), Data$M, Data$K)
theta <- theta / rowSums(theta)
parm[Data$pos.theta] <- as.vector(theta)
Z <- matrix(parm[Data$pos.Z], Data$M, Data$N)
### Log-Prior
phi.prior <- sum(ddirichlet(phi, beta, log=TRUE))
theta.prior <- sum(ddirichlet(theta, alpha, log=TRUE))
### Log-Likelihood
LL <- Z.prior <- 0
Yhat <- Data$Y
for (m in 1:Data$M) {for (n in 1:Data$N) {
  Z.prior + Z.prior + dcat(Z[m,n], theta[m,], log=TRUE)
  LL <- LL + dcat(Data$Y[m,n], as.vector(phi[Z[m,n],]), log=TRUE)
  Yhat[m,n] <- rcat(1, as.vector(phi[Z[m,n],]))}}
### Log-Posterior
LP <- LL + phi.prior + theta.prior + Z.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP, yhat=Yhat, parm=parm)
return(Modelout)
}

```

47.4. Initial Values

```
Initial.Values <- c(rep(1/V,K*V), rep(1/K,M*K), rcat(M*N,rep(1/K,K)))
```

48. Linear Regression

48.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\
 \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \sigma &\sim \mathcal{HC}(25)
 \end{aligned}$$

48.2. Data

```

N <- 10000
J <- 5
X <- matrix(1,N,J)
for (j in 2:J) {X[,j] <- rnorm(N,runif(1,-3,3),runif(1,0.1,1))}

```

```

beta <- runif(J,-3,3)
e <- rnorm(N,0,0.1)
y <- tcrossprod(X, t(beta)) + e
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), sigma=0))
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  sigma <- runif(1)
  return(c(beta, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma, y=y)

```

48.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  sigma.prior <- dgamma(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

48.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1)
```

49. Linear Regression, Frequentist

By eliminating prior probabilities, a frequentist linear regression example is presented. Although frequentism is not endorsed here, the purpose of this example is to illustrate how the **LaplacesDemon** package can be used for Bayesian or frequentist inference.

49.1. Form

$$\mathbf{y} \sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2)$$

$$\boldsymbol{\mu} = \mathbf{X}\boldsymbol{\beta}$$

49.2. Data

```

N <- 10000
J <- 5
X <- matrix(1,N,J)
for (j in 2:J) {X[,j] <- rnorm(N,runif(1,-3,3),runif(1,0.1,1))}
beta <- runif(J,-3,3)
e <- rnorm(N,0,0.1)
y <- tcrossprod(X, t(beta)) + e
mon.names <- "LL"
parm.names <- as.parm.names(list(beta=rep(0,J), sigma=0))
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  sigma <- runif(1)
  return(c(beta, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma, y=y)

```

49.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma, 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
  Modelout <- list(LP=LL, Dev=-2*LL, Monitor=LL,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

49.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1)
```

50. Linear Regression, Hierarchical Bayesian

50.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\ \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \\ \beta_j &\sim \mathcal{N}(\gamma, \delta), \quad j = 1, \dots, J \\ \gamma &\sim \mathcal{N}(0, 1000) \\ \delta &\sim \mathcal{HC}(25) \\ \sigma &\sim \mathcal{HC}(\tau) \\ \tau &\sim \mathcal{HC}(25) \end{aligned}$$

50.2. Data

```
data(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(log(demonsnacks[,c(1,4,10)]+1)))
J <- ncol(X)
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), gamma=0, delta=0, sigma=0,
  tau=0))
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.delta <- grep("delta", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.tau <- grep("tau", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  gamma <- rnorm(1)
  delta <- runif(1)
  sigma <- runif(1)
  tau <- runif(1)
  return(c(beta, gamma, delta, sigma, tau))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.gamma=pos.gamma,
```



```
pos.delta=pos.delta, pos.sigma=pos.sigma, pos.tau=pos.tau, y)
```

50.3. Model

```
Model <- function(parm, Data)
{
  ### Hyperparameters
  gamma <- parm[Data$pos.gamma]
  delta <- interval(parm[Data$pos.delta], 1e-100, Inf)
  parm[Data$pos.delta] <- delta
  parm[Data$pos.tau] <- tau <- interval(parm[Data$pos.tau], 1e-100, Inf)
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Hyperprior
  gamma.prior <- dnormv(gamma, 0, 1000, log=TRUE)
  delta.prior <- dhalfcauchy(delta, 25, log=TRUE)
  tau.prior <- dhalfcauchy(tau, 25, log=TRUE)
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, gamma, delta, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, tau, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + gamma.prior + delta.prior + sigma.prior +
    tau.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}
```

50.4. Initial Values

```
Initial.Values <- c(rep(0,J), 0, rep(1,3))
```

51. Linear Regression, Multilevel

51.1. Form

$$\mathbf{y} \sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2)$$

$$\mu_i = \mathbf{X}_i \boldsymbol{\beta}_{\mathbf{m}[i], 1:J}$$

$$\begin{aligned}\beta_{g,1:J} &\sim \mathcal{N}_J(\gamma, \Omega^{-1}), \quad g = 1, \dots, G \\ \Omega &\sim \mathcal{W}_{J+1}(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_J \\ \gamma_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\ \sigma &\sim \mathcal{HC}(25)\end{aligned}$$

where \mathbf{m} is a vector of length N , and each element indicates the multilevel group ($g = 1, \dots, G$) for the associated record.

51.2. Data

```
N <- 30
J <- 2 ### Number of predictors (including intercept)
G <- 2 ### Number of Multilevel Groups
X <- cbind(1, matrix(rnorm(N*(J-1),0,1),N,J-1))
Sigma <- matrix(runif(J*J,-1,1),J,J)
diag(Sigma) <- runif(J,1,5)
Sigma <- as.positive.definite(Sigma)
gamma <- runif(J,-1,1)
beta <- matrix(NA,G,J)
for (g in 1:G) {beta[g,] <- rmvn(1, gamma, Sigma)}
m <- rcat(N, rep(1/G,G)) ### Multilevel group indicator
y <- rowSums(beta[m,] * X) + rnorm(N,0,0.1)
S <- diag(J)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=matrix(0,G,J), gamma=rep(0,J),
  sigma=0, U=S), uppertri=c(0,0,0,1))
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  U <- rwishartc(Data$J+1, Data$S)
  gamma <- rnorm(Data$J)
  beta <- as.vector(rmvnpc(Data$G, gamma, U))
  sigma <- runif(1)
  return(c(beta, gamma, sigma, U[upper.tri(U, diag=TRUE)]))
}
MyData <- list(G=G, J=J, N=N, PGF=PGF, S=S, X=X, m=m, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.gamma=pos.gamma,
  pos.sigma=pos.sigma, y=y)
```

51.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
```

```

beta <- matrix(parm[Data$pos.beta], Data$G, Data$J)
gamma <- parm[Data$pos.gamma]
sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
parm[Data$pos.sigma] <- sigma
U <- as.parm.matrix(U, Data$J, parm, Data, chol=TRUE)
diag(U) <- exp(diag(U))
### Log-Prior
U.prior <- dwishartc(U, Data$J+1, Data$S, log=TRUE)
beta.prior <- sum(dmvnpc(beta, gamma, U, log=TRUE))
gamma.prior <- sum(dnormv(gamma, 0, 100, log=TRUE))
sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
### Log-Likelihood
mu <- rowSums(beta[Data$m,] * Data$X)
LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
### Log-Posterior
LP <- LL + U.prior + beta.prior + gamma.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

51.4. Initial.Values

```

Initial.Values <- c(rep(0,G*J), rep(0,J), 1,
  upper.triangle(S, diag=TRUE))

```

52. Linear Regression with Full Missingness

With ‘full missingness’, there are missing values for both the dependent variable (DV) and at least one independent variable (IV). The ‘full likelihood’ approach to full missingness is excellent as long as the model is identifiable. When it is not identifiable, imputation may be done in a previous stage, such as with the MISS function. In this example, matrix α is for regression effects for IVs, vector β is for regression effects for the DV, vector γ is for missing values for IVs, and δ is for missing values for the DV.

52.1. Form

$$\begin{aligned}
 \mathbf{y}^{imp} &\sim \mathcal{N}(\nu, \sigma_j^2) \\
 \mathbf{X}^{imp} &\sim \mathcal{N}(\mu, \sigma_{-j}^2) \\
 \nu &= \mathbf{X}^{imp} \beta \\
 \mu &= \mathbf{X}^{imp} \alpha \\
 \mathbf{y}^{imp} &= \begin{cases} \delta & \text{if } \mathbf{y}^{mis} \\ \mathbf{y}^{obs} & \end{cases}
 \end{aligned}$$

$$\mathbf{X}^{imp} = \begin{cases} \gamma & \text{if } \mathbf{X}^{mis} \\ \mathbf{X}^{obs} & \end{cases}$$

$$\alpha_{j,l} \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, (J-1), \quad l = 1, \dots, (J-1)$$

$$\beta_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

$$\gamma_m \sim \mathcal{N}(0, 1000), \quad m = 1, \dots, M$$

$$\delta_p \sim \mathcal{N}(0, 1000), \quad p = 1, \dots, P$$

$$\sigma_j \sim \mathcal{HC}(25), \quad j = 1, \dots, J$$

52.2. Data

```

N <- 100
J <- 5
X <- matrix(runif(N*J,-2,2),N,J); X[,1] <- 1 #Design matrix X
M <- matrix(round(runif(N*J)-0.45),N,J); M[,1] <- 0 #Missing indicators
X <- ifelse(M == 1, NA, X) #Simulated X gets missings according to M
beta.orig <- runif(J,-2,2)
y <- as.vector(tcrossprod(X, t(beta.orig)) + rnorm(N,0,0.1))
y[sample(1:N, round(N*.05))] <- NA
m <- ifelse(is.na(y), 1, 0) #Missing indicator for vector y
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=matrix(0,J-1,J-1),
  beta=rep(0,J),
  gamma=rep(0,sum(is.na(X))),
  delta=rep(0,sum(is.na(y))),
  sigma=rep(0,J)))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.delta <- grep("delta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm((Data$J-1)*(Data$J-1))
  beta <- rnorm(Data$J)
  gamma <- rnorm(sum(is.na(Data$X)))
  delta <- rnorm(sum(is.na(Data$y)), mean(Data$y, na.rm=TRUE), 1)
  sigma <- runif(Data$J)
  return(c(alpha, beta, gamma, delta, sigma))
}
MyData <- list(J=J, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.beta=pos.beta,
  pos.gamma=pos.gamma, pos.delta=pos.delta, pos.sigma=pos.sigma, y=y)

```

52.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- matrix(parm[Data$pos.alpha], Data$J-1, Data$J-1)
  beta <- parm[Data$pos.beta]
  gamma <- parm[Data$pos.gamma]
  delta <- parm[Data$pos.delta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  gamma.prior <- sum(dnormv(gamma, 0, 1000, log=TRUE))
  delta.prior <- sum(dnormv(delta, 0, 1000, log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  mu <- X.imputed <- Data$X
  X.imputed[which(is.na(X.imputed))] <- gamma
  y.imputed <- Data$y
  y.imputed[which(is.na(y.imputed))] <- delta
  for (j in 2:Data$J) {mu[,j] <- tcrossprod(X.imputed[,-j],
    t(alpha[, (j-1)]))}
  nu <- tcrossprod(X.imputed, t(beta))
  LL <- sum(dnorm(X.imputed[,-1], mu[,-1],
    matrix(sigma[1:(Data$J-1)], Data$N, Data$J-1), log=TRUE),
    dnorm(y.imputed, nu, sigma[Data$J], log=TRUE))
  ### Log-Posterior
  LP <- LL + alpha.prior + beta.prior + gamma.prior + delta.prior +
    sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(nu), nu, sigma[Data$J]), parm=parm)
  return(Modelout)
}

```

52.4. Initial Values

```

Initial.Values <- c(rep(0, (J-1)^2), rep(0,J), rep(0, sum(is.na(X))),
  rep(0, sum(is.na(y))), rep(1,J))

```

53. Linear Regression with Missing Response

This is an introductory example to missing values using data augmentation with auxiliary variables. The dependent variable, or response, has both observed values, \mathbf{y}^{obs} , and missing values, \mathbf{y}^{mis} . The α vector is for missing value imputation, and enables the use of the full-likelihood by augmenting the state with these auxiliary variables. In the model form, M is

used to denote the number of missing values, though it is used as an indicator in the data.

53.1. Form

$$\mathbf{y}^{imp} \sim \mathcal{N}(\mu, \sigma^2)$$

$$\mathbf{y}^{imp} = \begin{cases} \alpha & \text{if } \mathbf{y}^{mis} \\ \mathbf{y}^{obs} & \end{cases}$$

$$\mu = \mathbf{X}\beta$$

$$\alpha_m \sim \mathcal{N}(0, 1000), \quad m = 1, \dots, M$$

$$\beta_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

$$\sigma \sim \mathcal{HC}(25)$$

53.2. Data

```
data(demonsnacks)
N <- nrow(demonsnacks)
J <- ncol(demonsnacks)
y <- log(demonsnacks$Calories)
y[sample(1:N, round(N*0.05))] <- NA
M <- ifelse(is.na(y), 1, 0)
X <- cbind(1, as.matrix(demonsnacks[,c(1,3:10)]))
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,sum(M)), beta=rep(0,J),
  sigma=0))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(sum(Data$M), mean(y, na.rm=TRUE), 1)
  beta <- rnorm(Data$J)
  sigma <- runif(1)
  return(c(alpha, beta, sigma))
}
MyData <- list(J=J, M=M, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.beta=pos.beta,
  pos.sigma=pos.sigma, y=y)
```

53.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
```

```

alpha <- parm[Data$pos.alpha]
beta <- parm[Data$pos.beta]
sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
parm[Data$pos.sigma] <- sigma
### Log-Prior
alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
sigma.prior <- dgamma(sigma, 25, log=TRUE)
### Log-Likelihood
mu <- tcrossprod(Data$X, t(beta))
y.imputed <- Data$y
y.imputed[which(is.na(Data$y))] <- alpha
LL <- sum(dnorm(y.imputed, mu, sigma, log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

53.4. Initial Values

```
Initial.Values <- c(rep(0,sum(M)), rep(0,J), 1)
```

54. Linear Regression with Missing Response via ABB

The Approximate Bayesian Bootstrap (ABB), using the `ABB` function, is used to impute missing values in the dependent variable (DV), or response, given a propensity score. In this example, vector α is used to estimate propensity score η , while vector β is for regression effects, and vector γ has the monitored missing values. For more information on ABB, see the `ABB` function.

54.1. Form

$$\begin{aligned}
 \mathbf{y}^{imp} &\sim \mathcal{N}(\mu, \sigma^2) \\
 \mathbf{y}^{imp} &= \begin{cases} \gamma & \text{if } \mathbf{y}^{mis} \\ \mathbf{y}^{obs} & \end{cases} \\
 \mu &= \mathbf{X}\beta \\
 \gamma &\sim p(\mathbf{y}^{obs} | \mathbf{y}^{obs}, \mathbf{y}^{mis}, \eta) \\
 \eta &= \frac{1}{1 + \exp(-\nu)} \\
 \nu &= \mathbf{X}\alpha
 \end{aligned}$$

$$\alpha_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

$$\beta_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

$$\sigma \sim \mathcal{HC}(25)$$

54.2. Data

```

data(demonsnacks)
N <- nrow(demonsnacks)
J <- ncol(demonsnacks)
y <- log(demonsnacks$Calories)
y[sample(1:N, round(N*0.05))] <- NA
M <- ifelse(is.na(y), 1, 0)
X <- cbind(1, as.matrix(demonsnacks[,c(1,3:10)]))
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- c("LP",paste("gamma[",1:sum(is.na(y)),"]",sep=""))
parm.names <- as.parm.names(list(alpha=rep(0,J), beta=rep(0,J), sigma=0))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(Data$J)
  beta <- rnorm(Data$J)
  sigma <- runif(1)
  return(c(alpha, beta, sigma))
}
MyData <- list(J=J, M=M, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.beta=pos.beta,
  pos.sigma=pos.sigma, y=y)

```

54.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  sigma.prior <- dgamma(sigma, 25, log=TRUE)
  ### Log-Likelihood
  y.imputed <- Data$y
  mu <- tcrossprod(Data$X, t(beta))

```



```

nu <- as.vector(tcrossprod(Data$X, t(alpha)))
eta <- invlogit(nu)
breaks <- as.vector(quantile(eta, probs=c(0,0.2,0.4,0.6,0.8,1)))
B <- matrix(breaks[-length(breaks)], length(Data$y), 5, byrow=TRUE)
z <- rowSums(eta >= B)
for (i in 1:5) {
  if(any(is.na(Data$y[which(z == i)]))) {
    imp <- unlist(ABB(Data$y[which(z == i)]))
    y.imputed[which({z == i} & is.na(Data$y))] <- imp}}
gamma <- y.imputed[which(is.na(Data$y))]
LL <- sum(dbern(Data$M, eta, log=TRUE),
  dnorm(y.imputed, mu, sigma, log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,gamma),
  yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

54.4. Initial Values

```
Initial.Values <- c(rep(0,J), rep(0,J), 1)
```

55. Linear Regression with Power Priors

Power priors (Ibrahim and Chen 2000) are a class of informative priors when relevant historical data is available. Power priors may be used when it is desirable to take historical data into account while analyzing similar, current data. Both the current data, \mathbf{y} and \mathbf{X} , and historical data, \mathbf{y}_h and \mathbf{X}_h , are included in the power prior analysis, where h indicates historical data. Each data set receives its own likelihood function, though the likelihood of the historical data is raised to an exponential power, $\alpha \in [0, 1]$. In this example, α is a constant.

55.1. Form

$$\mathbf{y} \sim \mathcal{N}(\mu, \sigma^2)$$

$$\mathbf{y}_h \sim \mathcal{N}(\mu_h, \sigma^2)^\alpha$$

$$\mu = \mathbf{X}\beta$$

$$\mu_h = \mathbf{X}_h\beta$$

$$\alpha = 0.5$$

$$\beta_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

$$\sigma \sim \mathcal{HC}(25)$$

55.2. Data

```

N <- 100
J <- 5 #Number of predictors, including the intercept
X <- Xh <- matrix(1,N,J)
for (j in 2:J) {
  X[,j] <- rnorm(N,runif(1,-3,3),runif(1,0.1,1))
  Xh[,j] <- rnorm(N,runif(1,-3,3),runif(1,0.1,1))}
beta.orig <- runif(J,-3,3)
e <- rnorm(N,0,0.1)
yh <- as.vector(tcrossprod(beta.orig, Xh) + e)
y <- as.vector(tcrossprod(beta.orig, X) + e)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), sigma=0))
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  sigma <- runif(1)
  return(c(beta, sigma))
}
MyData <- list(alpha=0.5, J=J, PGF=PGF, X=X, Xh=Xh, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma, y=y,
  yh=yh)

```

55.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  muh <- tcrossprod(Data$Xh, t(beta))
  mu <- tcrossprod(Data$X, t(beta))
  LL <- sum(Data$alpha*dnorm(Data$yh, muh, sigma, log=TRUE) +
    dnorm(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
}

```

```
return(Modelout)
}
```

55.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1)
```

56. Linear Regression with Zellner's g-Prior

For more information on Zellner's g-prior, see the documentation for the `dzellner` function.

56.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\ \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \\ \boldsymbol{\beta} &\sim \mathcal{N}_J(0, g\sigma^2(\mathbf{X}^T\mathbf{X})^{-1}) \\ g &\sim \mathcal{HG}(\alpha), \quad \alpha = 3 \\ \sigma &\sim \mathcal{HC}(25) \end{aligned}$$

56.2. Data

```
data(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(demonsnacks[,c(1,3:10)]))
J <- ncol(X)
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), g0=0, sigma=0))
pos.beta <- grep("beta", parm.names)
pos.g <- grep("g0", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  g0 <- runif(1)
  sigma <- runif(1)
  return(c(beta, g0, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.g=pos.g, pos.sigma=pos.sigma,
  y=y)
```

56.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  parm[Data$pos.g] <- g <- interval(parm[Data$pos.g], 1e-100, Inf)
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Hyperprior
  g.prior <- dhyperg(g, alpha=3, log=TRUE)
  ### Log-Prior
  beta.prior <- dzellner(beta, g, sigma, Data$X, log=TRUE)
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + g.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

56.4. Initial Values

```
Initial.Values <- c(rep(1,J), rep(1,2))
```

57. LSTAR

This is a Logistic Smooth-Threshold Autoregression (LSTAR), and is specified with a transition function that includes γ as the shape parameter, \mathbf{y} as the transition variable, θ as the location parameter, and d as the delay parameter.

57.1. Form

$$\mathbf{y}_t \sim \mathcal{N}(\mu_t, \sigma^2), \quad t = 1, \dots, T$$

$$\mu_t = \pi_t(\alpha_1 + \phi_1 \mathbf{y}_{t-1}) + (1 - \pi_t)(\alpha_2 + \phi_2 \mathbf{y}_{t-1}), \quad t = 2, \dots, T$$

$$\pi_t = \frac{1}{1 + \exp(-(\gamma(\mathbf{y}_{t-d} - \theta)))}$$

$$\alpha_j \sim \mathcal{N}(0, 1000) \in [\mathbf{y}_{min}, \mathbf{y}_{max}], \quad j = 1, \dots, 2$$

$$\frac{\phi_j + 1}{2} \sim \mathcal{BETA}(1, 1), \quad j = 1, \dots, 2$$

$$\gamma \sim \mathcal{HC}(25)$$

$$\theta \sim \mathcal{U}(\mathbf{y}_{min}, \mathbf{y}_{max})$$

$$\pi_1 \sim \mathcal{U}(0.001, 0.999)$$

$$\sigma \sim \mathcal{HC}(25)$$

57.2. Data

```

data(demonfx)
y <- as.vector((log(as.matrix(demonfx[,1]))))
T <- length(y)
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,2), phi=rep(0,2), gamma=0,
  theta=0, pi=0, sigma=0))
pos.alpha <- grep("alpha", parm.names)
pos.phi <- grep("phi", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.theta <- grep("theta", parm.names)
pos.pi <- grep("pi", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- runif(2,min(Data$y),max(Data$y))
  phi <- runif(2, -1, 1)
  gamma <- runif(1)
  theta <- runif(1,min(Data$y),max(Data$y))
  pi <- runif(1, 0.001, 0.999)
  sigma <- runif(1)
  return(c(alpha, phi, gamma, theta, pi, sigma))
}
MyData <- list(PGF=PGF, T=T, mon.names=mon.names, parm.names=parm.names,
  pos.alpha=pos.alpha, pos.phi=pos.phi, pos.gamma=pos.gamma,
  pos.theta=pos.theta, pos.pi=pos.pi, pos.sigma=pos.sigma, y=y)

```

57.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- interval(parm[Data$pos.alpha], min(Data$y), max(Data$y))
  parm[Data$pos.alpha] <- alpha
  phi <- interval(parm[Data$pos.phi], -1, 1)
  gamma <- interval(parm[Data$pos.gamma], 1e-100, Inf)
  parm[Data$pos.gamma] <- gamma
  theta <- interval(parm[Data$pos.theta], min(Data$y), max(Data$y))
  parm[Data$pos.theta] <- theta
  pi <- interval(parm[Data$pos.pi], 0.001, 0.999)
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)

```

```

parm[Data$pos.sigma] <- sigma
### Log-Prior
alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
phi.prior <- sum(dbeta((phi+1)/2, 1, 1, log=TRUE))
gamma.prior <- dhalfcauchy(gamma, 25, log=TRUE)
theta.prior <- dunif(theta, min(Data$y), max(Data$y), log=TRUE)
pi.prior <- dunif(pi, 0.001, 0.999, log=TRUE)
sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
### Log-Likelihood
pi <- c(pi, 1 / (1 + exp(-(gamma*(Data$y[-Data$T]-theta))))))
mu <- pi * c(alpha[1], alpha[1] + phi[1]*Data$y[-Data$T]) +
      (1-pi) * c(alpha[2], alpha[2] + phi[2]*Data$y[-Data$T])
LL <- sum(dnorm(Data$y[-1], mu[-1], sigma, log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + phi.prior + gamma.prior + theta.prior +
      pi.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
                yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

57.4. Initial Values

```
Initial.Values <- c(rep(mean(y),2), rep(0.5,2), 1, mean(y), 0.5, 1)
```

58. MANCOVA

Since this is a multivariate extension of ANCOVA, please see the ANCOVA example in section 2 for a univariate introduction.

58.1. Form

$$\begin{aligned}
 \mathbf{Y}_{i,1:J} &\sim \mathcal{N}_K(\mu_{i,1:J}, \Sigma), \quad i = 1, \dots, N \\
 \mu_{i,k} &= \alpha_k + \beta_{k,\mathbf{X}[i,1]} + \gamma_{k,\mathbf{X}[i,1]} + \mathbf{X}_{1:N,3:(C+J)} \delta_{k,1:C} \\
 \epsilon_{i,k} &= \mathbf{Y}_{i,k} - \mu_{i,k} \\
 \alpha_k &\sim \mathcal{N}(0, 1000), \quad k = 1, \dots, K \\
 \beta_{k,l} &\sim \mathcal{N}(0, \sigma_1^2), \quad l = 1, \dots, (L-1) \\
 \beta_{1:K,L} &= - \sum_{l=1}^{L-1} \beta_{1:K,l} \\
 \gamma_{k,m} &\sim \mathcal{N}(0, \sigma_2^2), \quad m = 1, \dots, (M-1)
 \end{aligned}$$

$$\gamma_{1:K,M} = - \sum_{m=1}^{M-1} \beta_{1:K,m}$$

$$\delta_{k,c} \sim \mathcal{N}(0, 1000)$$

$$\Omega \sim \mathcal{W}_{K+1}(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_K$$

$$\Sigma = \Omega^{-1}$$

$$\sigma_{1:J} \sim \mathcal{HC}(25)$$

58.2. Data

```

C <- 2 #Number of covariates
J <- 2 #Number of factors (treatments)
K <- 3 #Number of endogenous (dependent) variables
L <- 4 #Number of levels in factor (treatment) 1
M <- 5 #Number of levels in factor (treatment) 2
N <- 100
X <- matrix(c(rcat(N, rep(1/L,L)), rcat(N, rep(1/M,M))),
            runif(N*C,0,1)), N, J + C)
alpha <- runif(K,-1,1)
beta <- matrix(runif(K*L,-2,2), K, L)
beta[,L] <- -rowSums(beta[, -L])
gamma <- matrix(runif(K*M,-2,2), K, M)
gamma[,M] <- -rowSums(gamma[, -M])
delta <- matrix(runif(K*C), K, C)
Y <- matrix(NA,N,K)
for (k in 1:K) {
  Y[,k] <- alpha[k] + beta[k,X[,1]] + gamma[k,X[,2]] +
  tcrossprod(delta[k,], X[, -c(1,2)]) + rnorm(1,0,0.1)}
S <- diag(K)
mon.names <- c("LP", "s.o.beta", "s.o.gamma", "s.o.epsilon",
  as.parm.names(list(s.beta=rep(0,K), s.gamma=rep(0,K),
  s.epsilon=rep(0,K))))
parm.names <- as.parm.names(list(alpha=rep(0,K), beta=matrix(0,K,(L-1)),
  gamma=matrix(0,K,(M-1)), delta=matrix(0,K,C), U=diag(K),
  sigma=rep(0,2)), uppertri=c(0,0,0,0,1,0))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.delta <- grep("delta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(Data$K)
  sigma <- runif(2)
  beta <- rnorm(Data$K*(Data$L-1), 0, sigma[1])
  gamma <- rnorm(Data$K*(Data$M-1), 0, sigma[2])

```

```

delta <- rnorm(Data$K*Data$C)
U <- rwishartc(Data$K+1, Data$S)
return(c(alpha, beta, gamma, delta, U[upper.tri(U, diag=TRUE)],
        sigma))
}
MyData <- list(C=C, J=J, K=K, L=L, M=M, N=N, PGF=PGF, S=S, X=X, Y=Y,
             mon.names=mon.names, parm.names=parm.names, pos.alpha=pos.alpha,
             pos.beta=pos.beta, pos.gamma=pos.gamma, pos.delta=pos.delta,
             pos.sigma=pos.sigma)

```

58.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- matrix(c(parm[Data$pos.beta], rep(0,Data$K)), Data$K, Data$L)
  beta[,Data$L] <- -rowSums(beta[, -Data$L])
  gamma <- matrix(c(parm[Data$[pos.gamma]],
                  rep(0,Data$K)), Data$K, Data$M)
  gamma[,Data$M] <- -rowSums(gamma[, -Data$M])
  delta <- matrix(parm[Data$pos.delta], Data$K, Data$C)
  U <- as.parm.matrix(U, Data$K, parm, Data, chol=TRUE)
  diag(U) <- exp(diag(U))
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
  beta.prior <- sum(dnorm(beta, 0, sigma[1], log=TRUE))
  gamma.prior <- sum(dnorm(gamma, 0, sigma[2], log=TRUE))
  delta.prior <- sum(dnormv(delta, 0, 1000, log=TRUE))
  U.prior <- dwishartc(U, Data$K+1, Data$S, log=TRUE)
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  mu <- matrix(0,Data$N,Data$K)
  for (k in 1:Data$K) {
    mu[,k] <- alpha[k] + beta[k,Data$X[,1]] + gamma[k,Data$X[,2]] +
      tcrossprod(Data$X[, -c(1,2)], t(delta[k,]))}
  LL <- sum(dmvnpc(Data$Y, mu, U, log=TRUE))
  ### Variance Components, Omnibus
  s.o.beta <- sd(as.vector(beta))
  s.o.gamma <- sd(as.vector(gamma))
  s.o.epsilon <- sd(as.vector(Data$Y - mu))
  ### Variance Components, Univariate
  s.beta <- sqrt(.rowVars(beta))
  s.gamma <- sqrt(.rowVars(gamma))
}

```



```
s.epsilon <- sqrt(.colVars(Data$Y - mu))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior + gamma.prior + delta.prior +
  U.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP, s.o.beta, s.o.gamma,
  s.o.epsilon, s.beta, s.gamma, s.epsilon),
  yhat=rmvnp(nrow(mu), mu, U), parm=parm)
return(Modelout)
}
```

58.4. Initial Values

```
Initial.Values <- c(rep(0,K), rep(0,K*(L-1)), rep(0,K*(M-1)),
  rep(0,C*K), upper.triangle(S, diag=TRUE), rep(1,2))
```

59. MANOVA

Since this is a multivariate extension of ANOVA, please see the two-way ANOVA example in section 4 for a univariate introduction.

59.1. Form

$$\mathbf{Y}_{i,1:J} \sim \mathcal{N}_K(\mu_{i,1:J}, \Omega^{-1}), \quad i = 1, \dots, N$$

$$\mu_{i,k} = \alpha_k + \beta_{k,\mathbf{x}[i,1]} + \gamma_{k,\mathbf{x}[i,1]}$$

$$\epsilon_{i,k} = \mathbf{Y}_{i,k} - \mu_{i,k}$$

$$\alpha_k \sim \mathcal{N}(0, 1000), \quad k = 1, \dots, K$$

$$\beta_{k,l} \sim \mathcal{N}(0, \sigma_1^2), \quad l = 1, \dots, (L - 1)$$

$$\beta_{1:K,L} = - \sum_{l=1}^{L-1} \beta_{1:K,l}$$

$$\gamma_{k,m} \sim \mathcal{N}(0, \sigma_2^2), \quad m = 1, \dots, (M - 1)$$

$$\gamma_{1:K,M} = - \sum_{m=1}^{M-1} \beta_{1:K,m}$$

$$\Omega \sim \mathcal{W}_{K+1}(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_K$$

$$\sigma_{1:J} \sim \mathcal{HC}(25)$$

59.2. Data

```
J <- 2 #Number of factors (treatments)
K <- 3 #Number of endogenous (dependent) variables
```

```

L <- 4 #Number of levels in factor (treatment) 1
M <- 5 #Number of levels in factor (treatment) 2
N <- 100
X <- cbind(rcat(N, rep(1/L,L)), rcat(N, rep(1/M,M)))
alpha <- runif(K,-1,1)
beta <- matrix(runif(K*L,-2,2), K, L)
beta[,L] <- -rowSums(beta[, -L])
gamma <- matrix(runif(K*M,-2,2), K, M)
gamma[,M] <- -rowSums(gamma[, -M])
Y <- matrix(NA,N,K)
for (k in 1:K) {
  Y[,k] <- alpha[k] + beta[k,X[,1]] + gamma[k,X[,2]] + rnorm(1,0,0.1)}
S <- diag(K)
mon.names <- c("LP", "s.o.beta", "s.o.gamma", "s.o.epsilon",
  as.parm.names(list(s.beta=rep(0,K), s.gamma=rep(0,K),
  s.epsilon=rep(0,K))))
parm.names <- as.parm.names(list(alpha=rep(0,K), beta=matrix(0,K,(L-1)),
  gamma=matrix(0,K,(M-1)), U=diag(K), sigma=rep(0,2)),
  uppertri=c(0,0,0,1,0))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(Data$K)
  sigma <- runif(2)
  beta <- rnorm(Data$K*(Data$L-1), 0, sigma[1])
  gamma <- rnorm(Data$K*(Data$M-1), 0, sigma[2])
  U <- rwishartc(Data$K+1, Data$S)
  return(c(alpha, beta, gamma, U[upper.tri(U, diag=TRUE)], sigma))
}
MyData <- list(J=J, K=K, L=L, M=M, N=N, PGF=PGF, S=S, X=X, Y=Y,
  mon.names=mon.names, parm.names=parm.names, pos.alpha=pos.alpha,
  pos.beta=pos.beta, pos.gamma=pos.gamma, pos.sigma=pos.sigma)

```

59.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- matrix(c(parm[Data$pos.beta], rep(0,Data$K)),
  beta[,Data$L] <- -rowSums(beta[, -Data$L])
  gamma <- matrix(c(parm[Data$pos.gamma],
  rep(0,Data$K)), Data$K, Data$M)
  gamma[,Data$M] <- -rowSums(gamma[, -Data$M])

```

```

U <- as.parm.matrix(U, Data$K, parm, Data, chol=TRUE)
diag(U) <- exp(diag(U))
sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
parm[Data$pos.sigma] <- sigma
### Log-Prior
alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
beta.prior <- sum(dnorm(beta, 0, sigma[1], log=TRUE))
gamma.prior <- sum(dnorm(gamma, 0, sigma[2], log=TRUE))
U.prior <- dwishartc(U, Data$K+1, Data$S, log=TRUE)
sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
### Log-Likelihood
mu <- matrix(0,Data$N,Data$K)
for (k in 1:Data$K) {
  mu[,k] <- alpha[k] + beta[k,Data$X[,1]] + gamma[k,Data$X[,2]]}
LL <- sum(dmvnpc(Data$Y, mu, U, log=TRUE))
### Variance Components, Omnibus
s.o.beta <- sd(as.vector(beta))
s.o.gamma <- sd(as.vector(gamma))
s.o.epsilon <- sd(as.vector(Data$Y - mu))
### Variance Components, Univariate
s.beta <- sqrt(.rowVars(beta))
s.gamma <- sqrt(.rowVars(gamma))
s.epsilon <- sqrt(.colVars(Data$Y - mu))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior + gamma.prior + U.prior +
  sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP, s.o.beta, s.o.gamma,
  s.o.epsilon, s.beta, s.gamma, s.epsilon),
  yhat=rmvnp(nrow(mu), mu, U), parm=parm)
return(Modelout)
}

```

59.4. Initial Values

```

Initial.Values <- c(rep(0,K), rep(0,K*(L-1)), rep(0,K*(M-1)),
  upper.triangle(S, diag=TRUE), rep(1,2))

```

60. Mixed Logit

60.1. Form

$$y_i \sim \text{CAT}(\mathbf{p}_{i,1:J}), \quad i = 1, \dots, N$$

$$\mathbf{p}_{i,j} = \frac{\phi_{i,j}}{\sum_{j=1}^J \phi_{i,j}}$$

$$\begin{aligned}\phi &= \exp(\mu) \\ \mu_{i,j} &= \beta_{j,1:K,i} \mathbf{X}_{i,1:K} \in [-700, 700], \quad i = 1, \dots, N, \quad j = 1, \dots, (J-1) \\ \mu_{i,J} &= 0 \\ \beta_{j,k,i} &\sim \mathcal{N}(\zeta_{j,k}^\mu, \zeta_{j,k}^\sigma 2_{j,k}), \quad i = 1, \dots, N, \quad j = 1, \dots, (J-1), \quad k = 1, \dots, K \\ \zeta_{j,k}^\mu &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, (J-1), \quad k = 1, \dots, K \\ \zeta_{j,k}^\sigma &\sim \mathcal{HC}25), \quad j = 1, \dots, (J-1), \quad k = 1, \dots, K\end{aligned}$$

60.2. Data

```
data(demonchoice)
y <- as.numeric(demonchoice[,1])
X <- cbind(1, as.matrix(demonchoice[,2:3]))
for (j in 2:ncol(X)) X[,j] <- CenterScale(X[,j])
N <- length(y)
J <- length(unique(y)) #Number of categories in y
K <- ncol(X) #Number of predictors (including the intercept)
S <- diag(J-1)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=array(0, dim=c(J-1,K,N)),
  zeta.mu=matrix(0,J-1,K), zeta.sigma=matrix(0,J-1,K)))
pos.beta <- grep("beta", parm.names)
pos.zeta.mu <- grep("zeta.mu", parm.names)
pos.zeta.sigma <- grep("zeta.sigma", parm.names)
PGF <- function(Data) {
  zeta.mu <- matrix(rnorm((Data$J-1)*Data$K), Data$J-1, Data$K)
  zeta.sigma <- matrix(runif((Data$J-1)*Data$K), Data$J-1, Data$K)
  beta <- array(rnorm((Data$J-1)*Data$K*Data$N),
    dim=c(Data$J-1, Data$K, Data$N))
  return(c(beta, as.vector(zeta.mu), as.vector(zeta.sigma)))
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, S=S, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.zeta.mu=pos.zeta.mu,
  pos.zeta.sigma=pos.zeta.sigma, y=y)
```

60.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  beta <- array(parm[Data$pos.beta], dim=c(Data$J-1, Data$K, Data$N))
  zeta.mu <- matrix(parm[Data$pos.zeta.mu], Data$J-1, Data$K)
  zeta.sigma <- matrix(interval(parm[Data$pos.zeta.sigma], 1e-100, Inf),
    Data$J-1, Data$K)
```

```

parm[Data$pos.zeta.sigma] <- as.vector(zeta.sigma)
### Log-Hyperprior
zeta.mu.prior <- sum(dnormv(zeta.mu, 0, 1000, log=TRUE))
zeta.sigma.prior <- sum(dhalfcauchy(zeta.sigma, 25, log=TRUE))
### Log-Prior
beta.prior <- sum(dnorm(beta, zeta.mu, zeta.sigma, log=TRUE))
### Log-Likelihood
mu <- matrix(0, Data$N, Data$J)
for (j in 1:(Data$J-1)) mu[,j] <- rowSums(Data$X * t(beta[j, , ]))
mu <- interval(mu, -700, 700, reflect=FALSE)
phi <- exp(mu)
p <- phi / rowSums(phi)
LL <- sum(dcat(Data$y, p, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + zeta.mu.prior + zeta.sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP, yhat=rcat(nrow(p), p),
  parm=parm)
return(Modelout)
}

```

60.4. Initial Values

```
Initial.Values <- c(rep(0,(J-1)*K*N), rep(0,(J-1)*K), rep(1,(J-1)*K))
```

61. Mixture Model, Finite

This finite mixture model (FMM) imposes a multilevel structure on each of the J regression effects in β , so that mixture components share a common residual standard deviation, ν_m . Identifiability is gained at the expense of some shrinkage. The record-level mixture membership parameter vector, θ , is a vector of discrete parameters. Discrete parameters are not supported in all algorithms. The example below is updated with the Slice sampler.

61.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\mu, \sigma^2) \\
 \mu_i &= \mathbf{X}_{i,1:J} \beta_{\theta[i],1:J}, \quad i = 1, \dots, N \\
 \theta_i &\sim \mathcal{CAT}(\pi_{1:M}), \quad i = 1, \dots, N \\
 \beta_{m,j} &\sim \mathcal{N}(0, \nu_m^2), \quad j = 1, \dots, J, \quad m = 2, \dots, M \\
 \beta_{1,j} &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \nu_m &\sim \mathcal{HC}(25), \quad m = 1, \dots, M \\
 \sigma &\sim \mathcal{HC}(25) \\
 \pi_{1:M} &\sim \mathcal{D}(\alpha_{1:M})
 \end{aligned}$$

$$\alpha_m = 1$$

61.2. Data

```

data(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(log(demonsnacks[,c(1,4,10)]+1)))
M <- 2 #Number of mixtures
N <- length(y) #Number of records
J <- ncol(X) #Number of predictors, including the intercept
for (j in 2:J) X[,j] <- CenterScale(X[,j])
alpha <- rep(1,M) #Prior probability of mixing probabilities
mon.names <- "LP"
parm.names <- as.parm.names(list(theta=rep(0,N), beta=matrix(0,M,J),
  nu=rep(0,M), sigma=0))
pos.theta <- grep("theta", parm.names)
pos.beta <- grep("beta", parm.names)
pos.nu <- grep("nu", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  theta <- rcat(Data$N, rep(1/Data$M, Data$M))
  nu <- runif(Data$M)
  beta <- rnormv(Data$M*Data$J, 0,
    cbind(1000, matrix(nu, Data$M, Data$J-1)))
  sigma <- runif(1)
  return(c(theta, beta, nu, sigma))
}
MyData <- list(J=J, M=M, N=N, PGF=PGF, X=X, alpha=alpha,
  mon.names=mon.names, parm.names=parm.names, pos.theta=pos.theta,
  pos.beta=pos.beta, pos.nu=pos.nu, pos.sigma=pos.sigma, y=y)

```

61.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- matrix(parm[Data$pos.beta], Data$M, Data$J)
  theta <- parm[Data$pos.theta]
  parm[Data$pos.nu] <- nu <- interval(parm[Data$pos.nu], 1e-100, Inf)
  pi <- rep(0, Data$M)
  tab <- table(theta)
  pi[as.numeric(names(tab))] <- as.vector(tab)
  pi <- pi / sum(pi)
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior

```

```

beta.prior <- sum(dnormv(beta, 0,
  cbind(1000, matrix(nu, Data$M, Data$J-1)), log=TRUE))
theta.prior <- sum(dcat(theta, p=pi, log=TRUE))
pi.prior <- ddirichlet(pi, Data$alpha, log=TRUE)
nu.prior <- sum(dhalfcauchy(nu, 25, log=TRUE))
sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
### Log-Likelihood
mu <- rowSums(beta[theta,] * Data$X)
LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + theta.prior + pi.prior + nu.prior +
  sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

61.4. Initial Values

```
Initial.Values <- c(rcat(N,rep(1/M,M)), rep(0,M*J), rep(1,M), 1)
```

62. Mixture Model, Infinite

This infinite mixture model (IMM) uses a Dirichlet process via truncated stick-breaking. The record-level mixture membership parameter vector, θ , is a vector of discrete parameters. Discrete parameters are not supported in all algorithms. The example below is updated with the Slice sampler.

62.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\
 \mu_i &= \mathbf{X}_{i,1:J} \boldsymbol{\beta}_{\theta[i],1:J}, \quad i = 1, \dots, N \\
 \theta_i &\sim \mathcal{CAT}(\boldsymbol{\pi}_{1:M}), \quad i = 1, \dots, N \\
 \beta_{m,j} &\sim \mathcal{N}(0, \nu_m^2), \quad j = 1, \dots, J, \quad m = 2, \dots, M \\
 \beta_{1,j} &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \nu_m &\sim \mathcal{HC}(25), \quad m = 1, \dots, M \\
 \sigma &\sim \mathcal{HC}(25) \\
 \pi &= \text{Stick}(\delta) \\
 \delta_m &\sim \mathcal{BETA}(1, \gamma), m = 1, \dots, (M-1) \\
 \gamma &\sim \mathcal{G}(\alpha, \iota)
 \end{aligned}$$

$$\alpha \sim \mathcal{HC}(25)$$

$$\iota \sim \mathcal{HC}(25)$$

62.2. Data

```

data(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(log(demonsnacks[,c(1,4,10)]+1)))
M <- 3 #Maximum number of mixtures to explore
N <- length(y) #Number of records
J <- ncol(X) #Number of predictors, including the intercept
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- c("LP", as.parm.names(list(pi=rep(0,M))))
parm.names <- as.parm.names(list(theta=rep(0,N), delta=rep(0,M-1),
  beta=matrix(0,M,J), nu=rep(0,M), sigma=0, alpha=0, iota=0, gamma=0))
pos.theta <- grep("theta", parm.names)
pos.delta <- grep("delta", parm.names)
pos.beta <- grep("beta", parm.names)
pos.nu <- grep("nu", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.alpha <- grep("alpha", parm.names)
pos.iota <- grep("iota", parm.names)
pos.gamma <- grep("gamma", parm.names)
PGF <- function(Data) {
  nu <- runif(Data$M)
  beta <- rnormv(Data$M*Data$J, 0,
    cbind(1000, matrix(nu, Data$M, Data$J-1)))
  sigma <- runif(1)
  alpha <- runif(1)
  iota <- runif(1)
  gamma <- rgamma(1, alpha, iota)
  delta <- rev(sort(rbeta(Data$M-1, 1, gamma)))
  theta <- rcat(Data$N, Stick(delta))
  return(c(theta, delta, beta, nu, sigma, alpha, iota, gamma))
}
MyData <- list(J=J, M=M, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.theta=pos.theta, pos.delta=pos.delta,
  pos.beta=pos.beta, pos.nu=pos.nu, pos.sigma=pos.sigma,
  pos.alpha=pos.alpha, pos.iota=pos.iota, pos.gamma=pos.gamma, y=y)

```

62.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperhyperparameters

```



```

alpha <- interval(parm[Data$pos.alpha], 1e-100, Inf)
parm[Data$pos.alpha] <- alpha
iota <- interval(parm[Data$pos.iota], 1e-100, Inf)
parm[Data$pos.iota] <- iota
### Hyperparameters
delta <- interval(parm[Data$pos.delta], 1e-10, 1-1e-10)
parm[Data$pos.delta] <- delta
gamma <- interval(parm[Data$pos.gamma], 1e-100, Inf)
parm[Data$pos.gamma] <- gamma
parm[Data$pos.nu] <- nu <- interval(parm[Data$pos.nu], 1e-100, Inf)
### Parameters
beta <- matrix(parm[Data$pos.beta], Data$M, Data$J)
theta <- parm[Data$pos.theta]
pi <- Stick(delta)
sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
parm[Data$pos.sigma] <- sigma
### Log-Hyperhyperprior
alpha.prior <- dhalfcauchy(alpha, 25, log=TRUE)
iota.prior <- dhalfcauchy(iota, 25, log=TRUE)
### Log-Hyperprior
delta.prior <- dStick(delta, gamma, log=TRUE)
gamma.prior <- dgamma(gamma, alpha, iota, log=TRUE)
nu.prior <- sum(dhalfcauchy(nu, 25, log=TRUE))
### Log-Prior
beta.prior <- sum(dnormv(beta, 0,
  cbind(1000, matrix(nu, Data$M, Data$J-1)), log=TRUE))
theta.prior <- sum(dcat(theta, pi, log=TRUE))
sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
### Log-Likelihood
mu <- rowSums(beta[theta,]*Data$X)
LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + delta.prior + theta.prior + nu.prior +
  sigma.prior + alpha.prior + iota.prior + gamma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,pi),
  yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

62.4. Initial Values

```

Initial.Values <- c(rcat(N, rev(sort(rStick(M-1,1))))), rep(0.5,M-1),
  rep(0,M*J), rep(1,M), rep(1,4))

```

63. Multinomial Logit

63.1. Form

$$\mathbf{y}_i \sim \mathcal{CAT}(\mathbf{p}_{i,1:J}), \quad i = 1, \dots, N$$

$$\mathbf{p}_{i,j} = \frac{\phi_{i,j}}{\sum_{j=1}^J \phi_{i,j}}, \quad \sum_{j=1}^J \mathbf{p}_{i,j} = 1$$

$$\phi = \exp(\boldsymbol{\mu})$$

$$\mu_{i,J} = 0, \quad i = 1, \dots, N$$

$$\mu_{i,j} = \mathbf{X}_{i,1:K} \boldsymbol{\beta}_{j,1:K} \in [-700, 700], \quad j = 1, \dots, (J-1)$$

$$\boldsymbol{\beta}_{j,k} \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, (J-1), \quad k = 1, \dots, K$$

63.2. Data

```

data(demonchoice)
y <- as.numeric(demonchoice[,1])
X <- cbind(1, as.matrix(demonchoice[,2:3]))
for (j in 2:ncol(X)) X[,j] <- CenterScale(X[,j])
N <- length(y)
J <- length(unique(y)) #Number of categories in y
K <- ncol(X) #Number of predictors (including the intercept)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=matrix(0,J-1,K)))
PGF <- function(Data) {
  beta <- rnorm((Data$J-1)*Data$K)
  return(beta)
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, y=y)

```

63.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- matrix(parm, Data$J-1, Data$K)
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  ### Log-Likelihood
  mu <- matrix(0, Data$N, Data$J)
  mu[,-Data$J] <- tcrossprod(Data$X, beta)
  mu <- interval(mu, -700, 700, reflect=FALSE)
}

```

```

phi <- exp(mu)
p <- phi / rowSums(phi)
LL <- sum(dcat(Data$y, p, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP, yhat=rcat(nrow(p), p),
  parm=parm)
return(Modelout)
}

```

63.4. Initial Values

```
Initial.Values <- c(rep(0, (J-1)*K))
```

64. Multinomial Logit, Nested

64.1. Form

$$\mathbf{y}_i \sim \mathcal{CAT}(\mathbf{P}_{i,1:J}), \quad i = 1, \dots, N$$

$$\mathbf{P}_{1:N,1} = \frac{\mathbf{R}}{\mathbf{R} + \exp(\alpha \mathbf{I})}$$

$$\mathbf{P}_{1:N,2} = \frac{(1 - \mathbf{P}_{1:N,1})\mathbf{S}_{1:N,1}}{\mathbf{V}}$$

$$\mathbf{P}_{1:N,3} = \frac{(1 - \mathbf{P}_{1:N,1})\mathbf{S}_{1:N,2}}{\mathbf{V}}$$

$$\mathbf{R}_{1:N} = \exp(\mu_{1:N,1})$$

$$\mathbf{S}_{1:N,1:2} = \exp(\mu_{1:N,2:3})$$

$$\mathbf{I} = \log(\mathbf{V})$$

$$\mathbf{V}_i = \sum_{k=1}^K \mathbf{S}_{i,k}, \quad i = 1, \dots, N$$

$$\mu_{1:N,1} = \mathbf{X}\boldsymbol{\iota} \in [-700, 700]$$

$$\mu_{1:N,2} = \mathbf{X}\boldsymbol{\beta}_{2,1:K} \in [-700, 700]$$

$$\boldsymbol{\iota} = \alpha\boldsymbol{\beta}_{1,1:K}$$

$$\alpha \sim \mathcal{E}\mathcal{X}\mathcal{P}(1) \in [0, 2]$$

$$\beta_{j,k} \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, (J-1) \quad k = 1, \dots, K$$

where there are $J = 3$ categories of \mathbf{y} , $K = 3$ predictors, \mathbf{R} is the non-nested alternative, \mathbf{S} is the nested alternative, \mathbf{V} is the observed utility in the nest, α is effectively 1 - correlation and

has a truncated exponential distribution, and ι is a vector of regression effects for the isolated alternative after α is taken into account. The third alternative is the reference category.

64.2. Data

```

data(demonchoice)
y <- as.numeric(demonchoice[,1])
X <- cbind(1, as.matrix(demonchoice[,2:3]))
for (j in 2:ncol(X)) X[,j] <- CenterScale(X[,j])
N <- length(y)
J <- length(unique(y)) #Number of categories in y
K <- ncol(X) #Number of predictors (including the intercept)
mon.names <- c("LP", as.parm.names(list(iota=rep(0,K))))
parm.names <- as.parm.names(list(alpha=0, beta=matrix(0,J-1,K)))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
PGF <- function(Data) {
  alpha <- rtrunc(1, "exp", a=0, b=2, rate=1)
  beta <- rnorm((Data$J-1)*Data$K)
  return(c(alpha, beta))
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.beta=pos.beta, y=y)

```

64.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperparameters
  alpha.rate <- 1
  ### Parameters
  parm[Data$pos.alpha] <- alpha <- interval(parm[Data$pos.alpha],0,2)
  beta <- matrix(parm[Data$pos.beta], Data$J-1, Data$K)
  ### Log-Prior
  alpha.prior <- dtrunc(alpha, "exp", a=0, b=2, rate=alpha.rate,
    log=TRUE)
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  ### Log-Likelihood
  mu <- P <- matrix(0, Data$N, Data$J)
  iota <- alpha * beta[1,]
  mu[,1] <- tcrossprod(Data$X, t(iota))
  mu[,2] <- tcrossprod(Data$X, t(beta[2,]))
  mu <- interval(mu, -700, 700, reflect=FALSE)
  R <- exp(mu[,1])
  S <- exp(mu[,-1])
  V <- rowSums(S)
  I <- log(V)
}

```

```

P[,1] <- R / (R + exp(alpha*I))
P[,2] <- (1 - P[,1]) * S[,1] / V
P[,3] <- (1 - P[,1]) * S[,2] / V
LL <- sum(dcat(Data$y, P, log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,iota),
  yhat=rcat(nrow(P), P), parm=parm)
return(Modelout)
}

```

64.4. Initial Values

```
Initial.Values <- c(0.5, rep(0.1,(J-1)*K))
```

65. Multinomial Probit

65.1. Form

$$\mathbf{W}_{i,1:(J-1)} \sim \mathcal{N}_{J-1}(\mu_{i,1:(J-1)}, \Sigma), \quad i = 1, \dots, N$$

$$\mathbf{W}_{i,j} \in \begin{cases} [0,10] & \text{if } y_i = j \\ [-10,0] & \end{cases}$$

$$\mu_{1:N,j} = \mathbf{X}\beta_{j,1:K}$$

$$\Sigma = \mathbf{U}^T \mathbf{U}$$

$$\beta_{j,k} \sim \mathcal{N}(0, 10), \quad j = 1, \dots, (J-1), \quad k = 1, \dots, K$$

$$\mathbf{U}_{j,k} \sim \mathcal{N}(0, 1), \quad j = 1, \dots, (J-1), \quad k = 1, \dots, (J-1), \quad j \geq k, \quad j \neq k = 1$$

65.2. Data

```

data(demonchoice)
y <- as.numeric(demonchoice[,1])
X <- cbind(1, as.matrix(demonchoice[,2:3]))
for (j in 2:ncol(X)) X[,j] <- CenterScale(X[,j])
N <- length(y)
J <- length(unique(y)) #Number of categories in y
K <- ncol(X) #Number of predictors (including the intercept)
S <- diag(J-1)
U <- matrix(NA,J-1,J-1)
U[upper.tri(U, diag=TRUE)] <- 0
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=matrix(0,(J-1),K),

```

```

      U=U, W=matrix(0,N,J-1)))
parm.names <- parm.names[-which(parm.names == "U[1,1]")]
pos.beta <- grep("beta", parm.names)
pos.U <- grep("U", parm.names)
pos.W <- grep("W", parm.names)
PGF <- function(Data) {
  beta <- rnorm((Data$J-1)*Data$K)
  U <- rnorm((Data$J-2) + (factorial(Data$J-1) /
    (factorial(Data$J-1-2)*factorial(2))))
  W <- matrix(runif(Data$N*(Data$J-1),-10,0), Data$N, Data$J-1)
  Y <- as.indicator.matrix(Data$y)
  W <- ifelse(Y[,-Data$J] == 1, abs(W), W)
  return(c(beta, U, as.vector(W)))}
MyData <- list(J=J, K=K, N=N, PGF=PGF, S=S, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.U=pos.U, pos.W=pos.W,
  y=y)

```

65.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- matrix(parm[Data$pos.beta], Data$J-1, Data$K)
  u <- c(0, parm[Data$pos.U])
  U <- diag(Data$J-1)
  U[upper.tri(U, diag=TRUE)] <- u
  diag(U) <- exp(diag(U))
  Sigma <- t(U) %*% U
  Sigma[1,] <- Sigma[,1] <- U[1,]
  W <- matrix(parm[Data$pos.W], Data$N, Data$J-1)
  Y <- as.indicator.matrix(Data$y)
  temp <- which(Y[,-c(Data$J)] == 1)
  W[temp] <- interval(W[temp], 0, 10)
  temp <- which(Y[,-c(Data$J)] == 0)
  W[temp] <- interval(W[temp], -10, 0)
  parm[Data$pos.W] <- as.vector(W)
  ### Log-Prior
  beta.prior <- sum(dnorm(beta, 0, 10, log=TRUE))
  U.prior <- sum(dnorm(u[-1], 0, 1, log=TRUE))
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, beta)
  #eta <- exp(cbind(mu,0))
  #p <- eta / rowSums(eta)
  LL <- sum(dmvn(W, mu, Sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + U.prior

```

```

Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=max.col(cbind(rmvn(nrow(mu), mu, Sigma),0)), parm=parm)
return(Modelout)
}

```

65.4. Initial Values

```
Initial.Values <- GIV(Model, MyData, PGF=TRUE)
```

66. Multiple Discrete-Continuous Choice

This form of a multivariate discrete-continuous choice model was introduced in [Kim, Allenby, and Rossi \(2002\)](#) and referred to as a variety model. The original version is presented with log-normally distributed errors, but a gamma regression form is used here instead, which has always mixed better in testing. Note that the γ parameters are fixed here, as recommended for identifiability in future articles by these authors.

66.1. Form

$$\mathbf{Y} \sim \mathcal{G}(\lambda\tau, \tau)$$

$$\lambda_{i,j} = \exp(\mathbf{Z}_{i,j} \log(\psi) \mathbf{1}_{m[i],j}) + \mathbf{X} \mathbf{1}_{i,1:K} \log(\beta) + \mathbf{X} \mathbf{2}_{i,1:L} \log(\delta) (\mathbf{Y}_{i,j} + \gamma_j)^\alpha, \quad i = 1, \dots, N, j = 1, \dots, J$$

$$\alpha_j \sim \mathcal{U}(0, 1), \quad j = 1, \dots, J$$

$$\log(\beta_k) \sim \mathcal{N}(0, 1000), \quad k = 1, \dots, K$$

$$\gamma_j = 1, \quad j = 1, \dots, J$$

$$\log(\delta_{j,l}) \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, (J-1), \quad l = 1, \dots, L$$

$$\log(\psi 0_j) \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

$$\log(\psi 1_{g,j}) \sim \mathcal{N}_J(\log(\psi 0), \Omega^{-1}), \quad g = 1, \dots, G, \quad j = 1, \dots, J$$

$$\Omega \sim \mathcal{W}_{J+1}(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_J$$

$$\tau_j \sim \mathcal{HC}(25), \quad j = 1, \dots, J$$

66.2. Data

```

G <- 6 #Number of Multilevel Groups (decision-makers, households, etc.)
J <- 3 #Number of products
K <- 4 #Number of product attributes
L <- 5 #Number of decision-maker attributes
N <- 30 #Number of records
X1 <- matrix(rnorm(N*K), N, K) #Product attributes
X2 <- matrix(rnorm(N*L), N, L) #Decision-maker attributes
Sigma <- matrix(runif((J-1)*(J-1), -1, 1), J-1, J-1)

```

```

diag(Sigma) <- runif(J-1,1,5)
Sigma <- as.positive.definite(Sigma) / 100
alpha <- runif(J)
log.beta <- rnorm(K,0,0.1)
log.delta <- matrix(rnorm((J-1)*L,0,0.1), J-1, L)
log.psi0 <- rnorm(J)
log.psi1 <- rmvn(G, log.psi0, Sigma)
m <- rcat(N, rep(1/G,G)) # Multilevel group indicator
Z <- as.indicator.matrix(m)
Y <- matrix(0, N, J)
Y <- round(exp(tcrossprod(Z, t(cbind(log.psi1,0)))) +
  matrix(tcrossprod(X1, t(log.beta)), N, J) +
  tcrossprod(X2, rbind(log.delta, colSums(log.delta)*-1))) *
  (Y + 1)^matrix(alpha,N,J,byrow=TRUE) +
  matrix(rnorm(N*J,0,0.1),N,J))
S <- diag(J)
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,J), log.beta=rep(0,K),
  log.delta=matrix(0,J-1,L), log.psi0=rep(0,J),
  log.psi1=matrix(0,G,J), tau=rep(0,J), U=S),
  uppertri=c(0,0,0,0,0,0,1))
pos.alpha <- grep("alpha", parm.names)
pos.log.beta <- grep("log.beta", parm.names)
pos.log.delta <- grep("delta", parm.names)
pos.log.psi0 <- grep("log.psi0", parm.names)
pos.log.psi1 <- grep("log.psi1", parm.names)
pos.tau <- grep("tau", parm.names)
PGF <- function(Data) {
  alpha <- runif(Data$J,0.9,1)
  log.beta <- rnorm(Data$K,0,0.1)
  log.delta <- rnorm((Data$J-1)*Data$L,0,0.1)
  log.psi0 <- rnorm(Data$J)
  U <- rwishartc(Data$J+1, Data$S)
  log.psi1 <- as.vector(rmvnpc(Data$G, log.psi0, U))
  tau <- runif(Data$J)
  return(c(alpha, log.beta, log.delta, log.psi0, log.psi1, tau,
    U[upper.tri(U, diag=TRUE)]))
}
MyData <- list(G=G, J=J, K=K, L=L, N=N, PGF=PGF, S=S, X1=X1, X2=X2, Y=Y,
  Z=Z, m=m, mon.names=mon.names, parm.names=parm.names,
  pos.alpha=pos.alpha, pos.log.beta=pos.log.beta,
  pos.log.delta=pos.log.delta, pos.log.psi0=pos.log.psi0,
  pos.log.psi1=pos.log.psi1, pos.tau=pos.tau)

```


66.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  parm[Data$pos.alpha] <- alpha <- interval(parm[Data$pos.alpha], 0, 1)
  log.beta <- parm[Data$pos.log.beta]
  log.delta <- matrix(parm[Data$pos.log.delta], Data$J-1, Data$L)
  log.psi0 <- parm[Data$pos.log.psi0]
  log.psi1 <- matrix(parm[Data$pos.log.psi1], Data$G, Data$J)
  parm[Data$pos.tau] <- tau <- interval(parm[Data$pos.tau], 1e-100, Inf)
  U <- as.parm.matrix(U, Data$J, parm, Data, chol=TRUE)
  diag(U) <- exp(diag(U))
  lambda <- tcrossprod(Data$Z, t(log.psi1)) +
    matrix(tcrossprod(Data$X1, t(log.beta)), Data$N, Data$J) +
    tcrossprod(Data$X2, rbind(log.delta, colSums(log.delta)*-1))
  ### Log-Prior
  U.prior <- dwishartc(U, Data$J+1, Data$S, log=TRUE)
  alpha.prior <- sum(dunif(alpha, 0, 1, log=TRUE))
  log.beta.prior <- sum(dnormv(log.beta, 0, 1000, log=TRUE))
  log.delta.prior <- sum(dnormv(log.delta, 0, 1000, log=TRUE))
  log.psi0.prior <- sum(dnormv(log.psi0, 0, 1000, log=TRUE))
  log.psi1.prior <- sum(dmvnpc(lambda,
    matrix(log.psi0, Data$N, Data$J, byrow=TRUE), U, log=TRUE))
  tau.prior <- sum(dhalfcauchy(tau, 25, log=TRUE))
  ### Log-Likelihood
  alpha <- matrix(alpha, Data$N, Data$J, byrow=TRUE)
  lambda <- exp(lambda)*(Data$Y + 1)^alpha
  tau <- matrix(tau, Data$N, Data$J, byrow=TRUE)
  LL <- sum(dgamma(Data$Y+1, lambda*tau, tau, log=TRUE))
  ### Log-Posterior
  LP <- LL + U.prior + alpha.prior + log.beta.prior + log.delta.prior +
    log.psi0.prior + log.psi1.prior + tau.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rgamma(prod(dim(lambda)), lambda*tau, tau)-1,
    parm=parm)
  return(Modelout)
}

```

66.4. Initial Values

```

Initial.Values <- c(runif(J,0.9,1), rnorm(K,0,0.1),
  rnorm((J-1)*L,0,0.1), rnorm(J,0,0.1),
  rmvnpc(G, rnorm(J,0,0.1), rwishartc(J+1,S)), runif(J),
  upper.triangle(rwishartc(J+1,S), diag=TRUE))

```

67. Multivariate Binary Probit

67.1. Form

$$\mathbf{W}_{i,1:J} \sim \mathcal{N}_J(\mu_{i,1:J}, \Omega^{-1}), \quad i = 1, \dots, N$$

$$\mathbf{W}_{i,j} \in \begin{cases} [0,10] & \text{if } y_i = j \\ [-10,0] & \end{cases}$$

$$\mu_{1:N,j} = \mathbf{X}\beta_{j,1:K}$$

$$\Omega = \rho^{-1}$$

$$\beta_{j,k} \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, (J-1), \quad k = 1, \dots, K$$

$$\beta_{J,k} = -\sum_{j=1}^{J-1} \beta_{j,k}$$

$$\rho \sim \mathcal{U}(-1, 1)$$

67.2. Data

```

N <- 30
J <- 2 #Number of binary dependent variables
K <- 3 #Number of columns to be in design matrix X
X <- cbind(1, matrix(rnorm(N*(K-1),0,1), N, K-1))
beta <- matrix(rnorm(J*K), J, K)
mu <- tcrossprod(X, beta)
u <- runif(length(which(upper.tri(diag(J)) == TRUE)), -1, 1)
rho <- diag(J)
rho[upper.tri(rho)] <- u
rho[lower.tri(rho)] <- t(rho)[lower.tri(rho)]
rho <- as.positive.semidefinite(rho)
Omega <- as.inverse(rho)
U <- chol(Omega)
W <- interval(rmvnpc(N, mu, U) + matrix(rnorm(N*J,0,0.1), N, J),
  -10, 10)
Y <- 1 * (W >= 0)
apply(Y, 2, table)
mon.names <- "LP"
rho <- matrix(NA, J, J)
rho[upper.tri(rho)] <- 0
parm.names <- as.parm.names(list(beta=matrix(0,J,K), rho=rho,
  W=matrix(0,N,J)))
pos.beta <- grep("beta", parm.names)
pos.rho <- grep("rho", parm.names)
pos.W <- grep("W", parm.names)
PGF <- function(Data) {

```

```

beta <- rnorm(Data$J*Data$K)
rho <- rep(0, length(which(upper.tri(diag(Data$J)))))
W <- matrix(runif(Data$N*Data$J,-10,0), Data$N, Data$J)
W <- ifelse(Y == 1, abs(W), W)
return(c(beta, rho, as.vector(W)))}
MyData <- list(J=J, K=K, N=N, PGF=PGF, X=X, Y=Y,
  mon.names=mon.names, parm.names=parm.names, pos.beta=pos.beta,
  pos.rho=pos.rho, pos.W=pos.W)

```

67.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- matrix(parm[Data$pos.beta], Data$J, Data$K)
  u <- interval(parm[Data$pos.rho], -1, 1)
  rho <- diag(MyData$J)
  rho[upper.tri(rho)] <- u
  rho[lower.tri(rho)] <- t(rho)[lower.tri(rho)]
  if(is.positive.semidefinite(rho) == FALSE)
    rho <- as.positive.semidefinite(rho)
  parm[Data$pos.rho] <- upper.triangle(rho)
  Omega <- as.inverse(rho)
  U <- chol(Omega)
  W <- matrix(parm[Data$pos.W], Data$N, Data$J)
  W[Data$Y == 0] <- interval(W[Data$Y == 0], -10, 0)
  W[Data$Y == 1] <- interval(W[Data$Y == 1], 0, 10)
  parm[Data$pos.W] <- as.vector(W)
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  rho.prior <- sum(dunif(u, -1, 1, log=TRUE))
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, beta)
  LL <- sum(dmvnpc(W, mu, U, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + rho.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=1*(rmvnpc(nrow(mu), mu, U) >= 0), parm=parm)
  return(Modelout)
}

```

67.4. Initial Values

```
Initial.Values <- GIV(Model, MyData, PGF=TRUE)
```

68. Multivariate Laplace Regression

68.1. Form

$$\mathbf{Y}_{i,k} \sim \mathcal{L}_K(\mu_{i,k}, \Sigma), \quad i = 1, \dots, N; \quad k = 1, \dots, K$$

$$\mu_{i,k} = \mathbf{X}_{1:N,k} \beta_{k,1:J}$$

$$\Sigma = \Omega^{-1}$$

$$\Omega \sim \mathcal{W}_{K+1}(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_K$$

$$\beta_{k,j} \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

68.2. Data

```
data(mtcars)
Y <- as.matrix(mtcars[,c(1,7)])
X <- cbind(1, as.matrix(mtcars[,c(3,4,6)]))
N <- nrow(Y) #Number of records
J <- ncol(X) #Number of columns in design matrix
K <- ncol(Y) #Number of DVs
S <- diag(K)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=matrix(0,K,J), U=diag(K)),
  uppertri=c(0,1))
pos.beta <- grep("beta", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$K*Data$J)
  U <- rwishartc(Data$K+1, Data$S)
  return(c(beta, U[upper.tri(U, diag=TRUE)]))
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, S=S, X=X, Y=Y, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta)
```

68.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  beta <- matrix(parm[Data$pos.beta], Data$K, Data$J)
  U <- as.parm.matrix(U, Data$K, parm, Data, chol=TRUE)
  diag(U) <- exp(diag(U))
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  U.prior <- dwishart(U, Data$K+1, Data$S, log=TRUE)
  ### Log-Likelihood
```

```

mu <- tcrossprod(Data$X, beta)
LL <- sum(dmvlc(Data$Y, mu, U, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + U.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rmvlc(nrow(mu), mu, U), parm=parm)
return(Modelout)
}

```

68.4. Initial Values

```
Initial.Values <- c(rep(0,J*K), upper.triangle(S, diag=TRUE))
```

69. Multivariate Poisson Regression

69.1. Form

$$\begin{aligned}
 \mathbf{Y}_{i,k} &\sim \mathcal{P}(\lambda_{i,k}), \quad i = 1, \dots, N \quad k = 1, \dots, K \\
 \lambda_{i,k} &= \exp(\mathbf{X}_{i,k} \beta_{k,1:J} + \gamma_{i,k}), \quad i = 1, \dots, N, \quad k = 1, \dots, K \\
 \beta_{k,j} &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J, \quad k = 1, \dots, K \\
 \gamma_{i,1:K} &\sim \mathcal{N}_K(0, \Omega^{-1}), \quad i = 1, \dots, N \\
 \Omega &\sim \mathcal{W}_{K+1}(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_K
 \end{aligned}$$

69.2. Data

```

N <- 20 #Number of records
J <- 4 #Number of columns in design matrix
K <- 3 #Number of DVs
X <- matrix(runif(N*J),N,J); X[,1] <- 1
beta <- matrix(rnorm(K*J),K,J)
Omega <- matrix(runif(K*K),K,K); diag(Omega) <- runif(K,1,K)
Omega <- as.symmetric.matrix(Omega)
gamma <- rmvnp(N, 0, Omega)
Y <- round(exp(tcrossprod(X, beta) + gamma))
S <- diag(K)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=matrix(0,K,J), gamma=matrix(0,N,K),
  U=S), uppertri=c(0,0,1))
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$K*Data$J)

```

```

gamma <- rnorm(Data$N*Data$K)
U <- rwishartc(Data$K+1, Data$S)
return(c(beta, gamma, U[upper.tri(U, diag=TRUE)]))
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, S=S, X=X, Y=Y, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.gamma=pos.gamma)

```

69.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- matrix(parm[Data$pos.beta], Data$K, Data$J)
  gamma <- matrix(parm[Data$pos.gamma], Data$N, Data$K)
  U <- as.parm.matrix(U, Data$K, parm, Data, chol=TRUE)
  diag(U) <- exp(diag(U))
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  gamma.prior <- sum(dmvnpc(gamma, 0, U, log=TRUE))
  U.prior <- dwishartc(U, Data$K+1, Data$S, log=TRUE)
  ### Log-Likelihood
  lambda <- exp(tcrossprod(Data$X, beta) + gamma)
  LL <- sum(dpois(Data$Y, lambda, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + gamma.prior + U.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rpois(prod(dim(lambda)), lambda), parm=parm)
  return(Modelout)
}

```

69.4. Initial Values

```
Initial.Values <- c(rep(0,K*J), rep(0,N*K), rep(0,K*(K+1)/2))
```

70. Multivariate Regression

70.1. Form

$$\mathbf{Y}_{i,k} \sim \mathcal{N}_K(\mu_{i,k}, \Sigma), \quad i = 1, \dots, N; \quad k = 1, \dots, K$$

$$\mu_{i,k} = \mathbf{X}_{1:N,k} \beta_{k,1:J}$$

$$\Sigma \sim \mathcal{HW}_2(\gamma, 1e6)$$

$$\beta_{k,j} \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J, \quad k = 1, \dots, K$$

70.2. Data

```

data(mtcars)
Y <- as.matrix(mtcars[,c(1,7)])
X <- cbind(1, as.matrix(mtcars[,c(3,4,6)]))
N <- nrow(Y) #Number of records
J <- ncol(X) #Number of columns in design matrix
K <- ncol(Y) #Number of DVs
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=matrix(0,K,J), gamma=rep(0,K),
  U=diag(K)), uppertri=c(0,0,1))
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$K*Data$J)
  gamma <- runif(Data$K)
  U <- rhuangwandc(2, gamma, rep(1,Data$K))
  return(c(beta, gamma, U[upper.tri(U, diag=TRUE)]))
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, X=X, Y=Y, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.gamma=pos.gamma)

```

70.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- matrix(parm[Data$pos.beta], Data$K, Data$J)
  gamma <- interval(parm[Data$pos.gamma], 1e-100, Inf)
  parm[Data$pos.gamma] <- gamma
  U <- as.parm.matrix(U, Data$K, parm, Data, chol=TRUE)
  diag(U) <- exp(diag(U))
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  HW.prior <- dhuangwandc(U, 2, gamma, rep(1e6,Data$K), log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, beta)
  LL <- sum(dmvnc(Data$Y, mu, U, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + U.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rmvnc(nrow(mu), mu, U), parm=parm)
  return(Modelout)
}

```

70.4. Initial Values

```
Initial.Values <- c(rep(0,J*K), rep(1,K), rep(0,K*(K+1)/2))
```

71. Negative Binomial Regression

This example was contributed by Jim Robison-Cox.

71.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{NB}(\mu, \kappa) \\ p &= \frac{\kappa}{\kappa + \mu} \\ \mu &= \exp(\mathbf{X}\beta) \\ \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\ \kappa &\sim \mathcal{HC}(25) \in (0, \infty] \end{aligned}$$

71.2. Data

```
N <- 100
J <- 5 #Number of predictors, including the intercept
kappa.orig <- 2
beta.orig <- runif(J,-2,2)
X <- matrix(runif(J*N,-2, 2), N, J); X[,1] <- 1
mu <- exp(tcrossprod(X, t(beta.orig)) + rnorm(N))
p <- kappa.orig / (kappa.orig + mu)
y <- rbinom(N, size=kappa.orig, mu=mu)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), kappa=0))
pos.beta <- grep("beta", parm.names)
pos.kappa <- grep("kappa", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  kappa <- runif(1)
  return(c(beta, kappa))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.kappa=pos.kappa, y=y)
```

71.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
```



```

beta <- parm[Data$pos.beta]
parm[Data$J + 1] <- kappa <- interval(parm[Data$pos.kappa],
  .Machine$double.xmin, Inf)
### Log-Prior
beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
kappa.prior <- dhalfcauchy(kappa, 25, log=TRUE)
### Log-Likelihood
mu <- as.vector(exp(tcrossprod(Data$X, t(beta))))
#p <- kappa / (kappa + mu)
LL <- sum(dnbinom(Data$y, size=kappa, mu=mu, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + kappa.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnbinom(length(mu), size=kappa, mu=mu), parm=parm)
return(Modelout)
}

```

71.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1)
```

72. Normal, Multilevel

This is Gelman's school example (Gelman, Carlin, Stern, and Rubin 2004). Note that **LaplacesDemon** is slower to converge than WinBUGS through the **R2WinBUGS** package (Gelman 2013), an R package on CRAN. This example is very sensitive to the prior distributions. The recommended, default, half-Cauchy priors with scale 25 on scale parameters overwhelms the likelihood, so uniform priors are used.

72.1. Form

$$\begin{aligned}
 \mathbf{y}_j &\sim \mathcal{N}(\theta_j, \sigma_j^2), \quad j = 1, \dots, J \\
 \theta_j &\sim \mathcal{N}(\theta_\mu, \theta_\sigma^2) \\
 \theta_\mu &\sim \mathcal{N}(0, 1000000) \\
 \theta_{\sigma[j]} &\sim \mathcal{N}(0, 1000) \\
 \sigma &\sim \mathcal{U}(0, 1000)
 \end{aligned}$$

72.2. Data

```

J <- 8
y <- c(28.4, 7.9, -2.8, 6.8, -0.6, 0.6, 18.0, 12.2)
sd <- c(14.9, 10.2, 16.3, 11.0, 9.4, 11.4, 10.4, 17.6)

```

```

mon.names <- "LP"
parm.names <- as.parm.names(list(theta=rep(0,J), theta.mu=0,
  theta.sigma=0))
pos.theta <- 1:J
pos.theta.mu <- J+1
pos.theta.sigma <- J+2
PGF <- function(Data) {
  theta.mu <- rnorm(1)
  theta.sigma <- runif(1)
  theta <- rnorm(Data$J, theta.mu, theta.sigma)
  return(c(theta, theta.mu, theta.sigma))
}
MyData <- list(J=J, PGF=PGF, mon.names=mon.names, parm.names=parm.names,
  pos.theta=pos.theta, pos.theta.mu=pos.theta.mu,
  pos.theta.sigma=pos.theta.sigma, sd=sd, y=y)

```

72.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperparameters
  theta.mu <- parm[Data$pos.theta.mu]
  theta.sigma <- interval(parm[Data$pos.theta.sigma], 1e-100, Inf)
  parm[Data$pos.theta.sigma] <- theta.sigma
  ### Parameters
  theta <- parm[Data$pos.theta]
  ### Log-Hyperprior
  theta.mu.prior <- dnormp(theta.mu, 0, 1.0E-6, log=TRUE)
  theta.sigma.prior <- dunif(theta.sigma, 0, 1000, log=TRUE)
  ### Log-Prior
  theta.prior <- sum(dnorm(theta, theta.mu, theta.sigma, log=TRUE))
  sigma.prior <- sum(dunif(Data$sd, 0, 1000, log=TRUE))
  ### Log-Likelihood
  LL <- sum(dnorm(Data$y, theta, Data$sd, log=TRUE))
  ### Log-Posterior
  LP <- LL + theta.prior + theta.mu.prior + theta.sigma.prior +
    sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(theta), theta, Data$sd), parm=parm)
  return(Modelout)
}

```

72.4. Initial Values

```

Initial.Values <- c(rep(mean(y),J), mean(y), 1)

```

73. Ordinal Logit

73.1. Form

$$\begin{aligned}
 \mathbf{y}_i &\sim \mathcal{CAT}(P_{i,1:J}) \\
 P_{,J} &= 1 - Q_{,(J-1)} \\
 P_{,j} &= |Q_{,j} - Q_{,(j-1)}|, \quad j = 2, \dots, (J-1) \\
 P_{,1} &= Q_{,1} \\
 Q &= \frac{1}{1 + \exp(\mu)} \\
 \mu_{,j} &= \delta_j - \mathbf{X}\beta, \quad \in [-5, 5] \\
 \beta_k &\sim \mathcal{N}(0, 1000), \quad k = 1, \dots, K \\
 \delta_j &\sim \mathcal{N}(0, 1) \in [(j-1), j] \in [-5, 5], \quad j = 1, \dots, (J-1)
 \end{aligned}$$

73.2. Data

```

data(demonsnacks)
N <- nrow(demonsnacks)
J <- 3 #Number of categories in y
X <- as.matrix(demonsnacks[,c(1,3:10)])
K <- ncol(demonsnacks) #Number of columns in design matrix X
y <- log(demonsnacks$Calories)
y <- ifelse(y < 4.5669, 1, ifelse(y > 5.5268, 3, 2)) #Discretize
for (k in 1:K) X[,k] <- CenterScale(X[,k])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,K), delta=rep(0,J-1)))
pos.beta <- grep("beta", parm.names)
pos.delta <- grep("delta", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$K)
  delta <- sort(rnorm(Data$J-1))
  return(c(beta, delta))
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.delta=pos.delta, y=y)

```

73.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters

```

```

beta <- parm[Data$pos.beta]
delta <- interval(parm[Data$pos.delta], -5, 5)
delta <- sort(delta)
parm[Data$pos.delta] <- delta
### Log-Prior
beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
delta.prior <- sum(dtrunc(delta, "norm", a=-5, b=5, log=TRUE,
  mean=0, sd=1)
### Log-Likelihood
mu <- matrix(delta, Data$N, Data$J-1, byrow=TRUE) -
  matrix(tcrossprod(Data$X, t(beta)), Data$N, Data$J-1)
P <- Q <- invlogit(mu)
P[, -1] <- abs(Q[, -1] - Q[, -(Data$J-1)])
P <- cbind(P, 1 - Q[, (Data$J-1)])
LL <- sum(dcat(Data$y, P, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + delta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP, yhat=rcat(nrow(P), P)
  parm=parm)
return(Modelout)
}

```

73.4. Initial Values

```
Initial.Values <- c(rep(0,K), seq(from=-1, to=1, len=(J-1)))
```

74. Ordinal Probit

74.1. Form

$$\mathbf{y}_i \sim \mathcal{CAT}(P_{i,1:J})$$

$$P_{,J} = 1 - Q_{,(J-1)}$$

$$P_{,j} = |Q_{,j} - Q_{,(j-1)}|, \quad j = 2, \dots, (J-1)$$

$$P_{,1} = Q_{,1}$$

$$Q = \phi(\mu)$$

$$\mu_{,j} = \delta_j - \mathbf{X}\beta, \quad \in [-5, 5]$$

$$\beta_k \sim \mathcal{N}(0, 1000), \quad k = 1, \dots, K$$

$$\delta_j \sim \mathcal{N}(0, 1) \in [(j-1), j] \in [-5, 5], \quad j = 1, \dots, (J-1)$$

74.2. Data

```

data(demonsnacks)
N <- nrow(demonsnacks)
J <- 3 #Number of categories in y
X <- as.matrix(demonsnacks[,c(1,3:10)])
K <- ncol(demonsnacks) #Number of columns in design matrix X
y <- log(demonsnacks$Calories)
y <- ifelse(y < 4.5669, 1, ifelse(y > 5.5268, 3, 2)) #Discretize
for (k in 1:K) X[,k] <- CenterScale(X[,k])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,K), delta=rep(0,J-1)))
pos.beta <- grep("beta", parm.names)
pos.delta <- grep("delta", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$K)
  delta <- sort(rnorm(Data$J-1))
  return(c(beta, delta))
}
MyData <- list(J=J, K=K, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.delta=pos.delta, y=y)

```

74.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  delta <- interval(parm[Data$pos.delta], -5, 5)
  delta <- sort(delta)
  parm[Data$pos.delta] <- delta
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  delta.prior <- sum(dtrunc(delta, "norm", a=-5, b=5, log=TRUE,
    mean=0, sd=1))
  ### Log-Likelihood
  mu <- matrix(delta, Data$N, Data$J-1, byrow=TRUE) -
    matrix(tcrossprod(Data$X, t(beta)), Data$N, Data$J-1)
  P <- Q <- pnorm(mu)
  P[, -1] <- abs(Q[, -1] - Q[, -(Data$J-1)])
  P <- cbind(P, 1 - Q[, (Data$J-1)])
  LL <- sum(dcat(Data$y, P, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + delta.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP, yhat=rcat(nrow(P), P)
    parm=parm)
}

```

```

return(Modelout)
}

```

74.4. Initial Values

```
Initial.Values <- c(rep(0,K), seq(from=-1, to=1, len=(J-1)))
```

75. Panel, Autoregressive Poisson

75.1. Form

$$\begin{aligned}
\mathbf{Y} &\sim \mathcal{P}(\Lambda) \\
\Lambda_{1:N,1} &= \exp(\alpha + \beta \mathbf{x}) \\
\Lambda_{1:N,t} &= \exp(\alpha + \beta \mathbf{x} + \rho \log(\mathbf{Y}_{1:N,t-1})), \quad t = 2, \dots, T \\
\alpha_i &\sim \mathcal{N}(\alpha_\mu, \alpha_\sigma^2), \quad i = 1, \dots, N \\
\alpha_\mu &\sim \mathcal{N}(0, 1000) \\
\alpha_\sigma &\sim \mathcal{HC}(25) \\
\beta &\sim \mathcal{N}(0, 1000) \\
\rho &\sim \mathcal{N}(0, 1000)
\end{aligned}$$

75.2. Data

```

N <- 10
T <- 10
alpha <- rnorm(N,2,0.5)
rho <- 0.5
beta <- 0.5
x <- runif(N,0,1)
Y <- matrix(NA,N,T)
Y[,1] <- exp(alpha + beta*x)
for (t in 2:T) {Y[,t] <- exp(alpha + beta*x + rho*log(Y[,t-1]))}
Y <- round(Y)
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,N), alpha.mu=0,
  alpha.sigma=0, beta=0, rho=0))
pos.alpha <- 1:N
pos.alpha.mu <- grep("alpha.mu", parm.names)
pos.alpha.sigma <- grep("alpha.sigma", parm.names)
pos.beta <- grep("beta", parm.names)
pos.rho <- grep("rho", parm.names)

```

```

PGF <- function(Data) {
  alpha.mu <- rnorm(1)
  alpha.sigma <- runif(1)
  alpha <- rnorm(Data$N, alpha.mu, alpha.sigma)
  beta <- rnorm(1)
  rho <- rnorm(1)
  return(c(alpha, alpha.mu, alpha.sigma, beta, rho))
}
MyData <- list(N=N, PGF=PGF, T=T, Y=Y, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.alpha.mu=pos.alpha.mu,
  pos.alpha.sigma=pos.alpha.sigma, pos.beta=pos.beta, pos.rho=pos.rho,
  x=x)

```

75.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperparameters
  alpha.mu <- parm[Data$pos.alpha.mu]
  alpha.sigma <- interval(parm[Data$pos.alpha.sigma], 1e-100, Inf)
  parm[Data$pos.alpha.sigma] <- alpha.sigma
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- parm[Data$pos.beta]
  rho <- parm[Data$pos.rho]
  ### Log-Hyperprior
  alpha.mu.prior <- dnormv(alpha.mu, 0, 1000, log=TRUE)
  alpha.sigma.prior <- dhalfcauchy(alpha.sigma, 25, log=TRUE)
  ### Log-Prior
  alpha.prior <- sum(dnorm(alpha, alpha.mu, alpha.sigma, log=TRUE))
  beta.prior <- dnormv(beta, 0, 1000, log=TRUE)
  rho.prior <- dnormv(rho, 0, 1000, log=TRUE)
  ### Log-Likelihood
  Lambda <- Data$Y
  Lambda[,1] <- exp(alpha + beta*x)
  Lambda[,2:Data$T] <- exp(alpha + beta*Data$x +
    rho*log(Data$Y[,1:(Data$T-1)]))
  LL <- sum(dpois(Data$Y, Lambda, log=TRUE))
  ### Log-Posterior
  LP <- LL + alpha.prior + alpha.mu.prior + alpha.sigma.prior +
    beta.prior + rho.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rpois(prod(dim(Lambda)), parm=parm)
  return(Modelout)
}

```

75.4. Initial Values

```
Initial.Values <- c(rep(0,N), 0, 1, 0, 0)
```

76. Penalized Spline Regression

This example applies penalized splines to one predictor in a linear regression. The user selects the degree of the polynomial, D , and the number of knots, K .

76.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{N}(\mu, \sigma_1^2) \\ \mu &= \mathbf{X}\beta + \mathbf{S} \\ \mathbf{S} &= \mathbf{Z}\gamma \\ \mathbf{Z}_{i,k} &= \begin{cases} (\mathbf{x}_i - k)^D & \text{if } \mathbf{Z}_{i,k} > 0 \\ 0 & \end{cases} \\ \beta_d &\sim \mathcal{N}(0, 1000), \quad d = 1, \dots, (D+1) \\ \gamma_k &\sim \mathcal{N}(0, \sigma_2^2), \quad k = 1, \dots, K \\ \sigma_j &\sim \mathcal{HC}(25), \quad j = 1, \dots, 2 \end{aligned}$$

76.2. Data

```
N <- 100
x <- 1:N
y <- sin(2*pi*x/N) + runif(N,-1,1)
K <- 10 #Number of knots
D <- 2 #Degree of polynomial
x <- CenterScale(x)
k <- as.vector(quantile(x, probs=(1:K / (K+1))))
X <- cbind(1, matrix(x, N, D))
for (d in 1:D) {X[,d+1] <- X[,d+1]^d}
Z <- matrix(x, N, K) - matrix(k, N, K, byrow=TRUE)
Z <- ifelse(Z > 0, Z, 0); Z <- Z^D
mon.names <- c("LP", paste("S[", 1:nrow(X) ,"]", sep=""))
parm.names <- as.parm.names(list(beta=rep(0,1+D), gamma=rep(0,K),
  log.sigma=rep(0,2)))
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(1+Data$D)
  gamma <- rnorm(Data$K)
  sigma <- runif(2)
```



```

    return(c(beta, gamma, sigma))
  }
MyData <- list(D=D, K=K, N=N, PGF=PGF, Z=Z, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.gamma=pos.gamma,
  pos.sigma=pos.sigma, y=y)

```

76.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  gamma <- parm[Data$pos.gamma]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  gamma.prior <- sum(dnorm(gamma, 0, sigma[2], log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  S <- as.vector(tcrossprod(Data$Z, t(gamma)))
  mu <- as.vector(tcrossprod(Data$X, t(beta))) + S
  LL <- sum(dnorm(Data$y, mu, sigma[1], log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + gamma.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,S),
    yhat=rnorm(length(mu), mu, sigma[1]), parm=parm)
  return(Modelout)
}

```

76.4. Initial Values

```
Initial.Values <- c(rep(0,1+D), rep(0,K), c(1,1))
```

77. Poisson Regression

77.1. Form

$$y \sim \mathcal{P}(\lambda)$$

$$\lambda = \exp(\mathbf{X}\beta)$$

$$\beta_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J$$

77.2. Data

```

N <- 10000
J <- 5
X <- matrix(runif(N*J,-2,2),N,J); X[,1] <- 1
beta <- runif(J,-2,2)
y <- round(exp(tcrossprod(X, t(beta))))
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J)))
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  return(beta)
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, y=y)

```

77.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  ### Log-Likelihood
  lambda <- exp(tcrossprod(Data$X, t(beta)))
  LL <- sum(dpois(Data$y, lambda, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rpois(length(lambda), lambda), parm=parm)
  return(Modelout)
}

```

77.4. Initial Values

```
Initial.Values <- rep(0,J)
```

78. Polynomial Regression

In this univariate example, the degree of the polynomial is specified as D . For a more robust extension to estimating nonlinear relationships between y and x , see penalized spline regression in section ??.

78.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\mu, \sigma^2) \\
 \mu &= \mathbf{X}\beta \\
 \mathbf{X}_{i,d} &= \mathbf{x}_i^{d-1}, \quad d = 1, \dots, (D+1) \\
 \mathbf{X}_{i,1} &= 1 \\
 \beta_d &\sim \mathcal{N}(0, 1000), \quad d = 1, \dots, (D+1) \\
 \sigma &\sim \mathcal{HC}(25)
 \end{aligned}$$

78.2. Data

```

data(demonsnacks)
N <- nrow(demonsnacks)
D <- 2 #Degree of polynomial
y <- log(demonsnacks$Calories)
x <- log(demonsnacks[,10]+1)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,D+1), sigma=0))
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(1+Data$D)
  sigma <- runif(1)
  return(c(beta, sigma))
}
MyData <- list(D=D, N=N, PGF=PGF, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma, x=x,
  y=y)

```

78.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  X <- matrix(Data$x, Data$N, Data$D)
  for (d in 2:Data$D) {X[,d] <- X[,d]^d}
}

```

```

X <- cbind(1,X)
mu <- tcrossprod(X, t(beta))
LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

78.4. Initial Values

```
Initial.Values <- c(rep(0,D+1), 1)
```

79. Proportional Hazards Regression, Weibull

Although the dependent variable is usually denoted as t in survival analysis, it is denoted here as y so Laplace's Demon recognizes it as a dependent variable for posterior predictive checks. This example does not support censoring, but it will be included soon.

79.1. Form

$$\begin{aligned}
 y_i &\sim \mathcal{WETB}(\gamma, \mu_i), \quad i = 1, \dots, N \\
 \mu &= \exp(\mathbf{X}\beta) \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \gamma &\sim \mathcal{G}(1, 0.001)
 \end{aligned}$$

79.2. Data

```

N <- 50
J <- 5
X <- matrix(runif(N*J,-2,2),N,J); X[,1] <- 1
beta <- c(1,runif(J-1,-1,1))
y <- round(exp(tcrossprod(X, t(beta)))) + 1 # Undefined at zero
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), gamma=0))
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  gamma <- rgamma(1,1e-3)
  return(c(beta, gamma))
}

```

```
MyData <- list(J=J, N=N, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.gamma=pos.gamma, y=y)
```

79.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  gamma <- interval(parm[Data$pos.gamma], 1e-100, Inf)
  parm[Data$pos.gamma] <- gamma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  gamma.prior <- dgamma(gamma, 1, 1.0E-3, log=TRUE)
  ### Log-Likelihood
  mu <- exp(tcrossprod(Data$X, t(beta)))
  LL <- sum(dweibull(Data$y, gamma, mu, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + gamma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rweibull(length(mu), gamma, mu), parm=parm)
  return(Modelout)
}
```

79.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1)
```

80. PVAR(p)

This is a Poisson vector autoregression, with autoregressive order p , for multivariate time-series of counts. It allows for dynamic processes and accounts for overdispersion.

80.1. Form

$$\mathbf{Y}_{t,j} \sim \mathcal{P}(\lambda_{t,j}), \quad t = 1, \dots, T \quad j = 1, \dots, J$$

$$\lambda_{t,j} = \sum_{p=1}^P \Phi_{1:J,j,p} \mathbf{Y}_{t-p,j} + \exp(\alpha_j + \gamma_{t,j})$$

$$\alpha_j \sim \mathcal{N}(0, 1000)$$

$$\Phi_{i,k,p} \sim \mathcal{N}(\Phi_{i,k,p}^\mu, \Sigma_{i,k,p}), \quad i = 1, \dots, J, \quad k = 1, \dots, J, \quad p = 1, \dots, P$$

$$\gamma_{t,1:J} \sim \mathcal{N}_J(0, \Omega^{-1}), \quad t = 1, \dots, T$$

$$\Omega \sim \mathcal{W}_{J+1}(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_J$$

where Φ^μ and Σ are set according to the Minnesota prior.

80.2. Data

```

data(demonsessions)
Y.orig <- as.matrix(demonsessions)
Y <- Y.orig[1:24,1:5]
T <- nrow(Y)
J <- ncol(Y)
L <- c(1,2,3) #Autoregressive lags
P <- length(L) #Autoregressive order
Phi.mu <- array(0, dim=c(J,J,P))
Phi.mu[, , 1] <- diag(J)
S <- diag(J)
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,J),
  Phi=array(0, dim=c(J,J,P)), gamma=matrix(0,T-L[P],J), U=S),
  uppertri=c(0,0,0,1))
pos.alpha <- grep("alpha", parm.names)
pos.Phi <- grep("Phi", parm.names)
pos.gamma <- grep("gamma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(Data$J)
  Phi <- runif(Data$J*Data$J*Data$P, -1e-10, 1e-10)
  gamma <- rnorm((Data$T-Data$L[Data$P])*Data$J)
  U <- rwishartc(Data$J+1, diag(Data$J))
  return(c(alpha, Phi, gamma, U[upper.tri(U, diag=TRUE)]))
}
MyData <- list(J=J, L=L, P=P, PGF=PGF, Phi.mu=Phi.mu, S=S, T=T, Y=Y,
  mon.names=mon.names, parm.names=parm.names, pos.alpha=pos.alpha,
  pos.Phi=pos.Phi, pos.gamma=pos.gamma)

```

80.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  Phi <- array(parm[Data$pos.Phi], dim=c(Data$J, Data$J, Data$P))
  gamma <- matrix(parm[Data$pos.gamma], Data$T-Data$L[Data$P], Data$J)
  U <- as.parm.matrix(U, Data$J, parm, Data, chol=TRUE)
  diag(U) <- exp(diag(U))
  Omega <- t(U) %*% U
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
}

```

```

Sigma <- MinnesotaPrior(Data$J, lags=Data$L, lambda=1, theta=0.5,
  diag(as.inverse(Omega)))
Phi.prior <- sum(dnormv(Phi, Data$Phi.mu, Sigma, log=TRUE))
gamma.prior <- sum(dmvnp(gamma, 0, Omega, log=TRUE))
U.prior <- dwishart(Omega, Data$J+1, Data$S, log=TRUE)
### Log-Likelihood
lambda <- exp(matrix(alpha, Data$T, Data$J, byrow=TRUE) +
  rbind(matrix(0, Data$L[Data$P], Data$J), gamma))
for (p in 1:Data$P)
  lambda[(1+Data$L[p]):Data$T,] <- lambda[(1+Data$L[p]):Data$T,] +
    Data$Y[1:(Data$T-Data$L[p]),] LL <- sum(dpois(Data$Y[(1+Data$L[Data$P]):Data$T],
  lambda[(1+Data$L[Data$P]):Data$T,], log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + Phi.prior + gamma.prior + U.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rpois(prod(dim(lambda)), lambda), parm=parm)
return(Modelout)
}

```

80.4. Initial Values

```

Initial.Values <- c(rep(0,J), rep(0,J*J*P), rep(0,(T-L[P])*J),
  rep(0,J*(J+1)/2))

```

81. Quantile Regression

81.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\phi, \sigma^2) \\
 \phi &= \frac{(1-2P)}{P(1-P)}\zeta + \mu \\
 \mu &= \mathbf{X}\beta \\
 \sigma &= \frac{P(1-P)\tau}{2\zeta} \\
 \beta &\sim \mathcal{N}(0, 1000) \\
 \tau &\sim \mathcal{HC}(25) \\
 \zeta &\sim \mathcal{E}\mathcal{X}\mathcal{P}(\tau)
 \end{aligned}$$

where P is the user-specified quantile in $(0, 1)$.

81.2. Data

```

data(demonsnacks)
y <- log(demonsnacks$Calories)

```

```

X <- cbind(1, as.matrix(log(demonsnacks[,c(1,4,10)]+1)))
N <- nrow(X)
J <- ncol(X)
P <- 0.5 #Quantile in (0,1)
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), tau=0, zeta=rep(0,N)))
pos.beta <- grep("beta", parm.names)
pos.tau <- grep("tau", parm.names)
pos.zeta <- grep("zeta", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  tau <- runif(1)
  zeta <- rexp(Data$N, tau)
  return(c(beta, tau, zeta))
}
MyData <- list(J=J, N=N, P=P, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.tau=pos.tau,
  pos.zeta=pos.zeta, y=y)

```

81.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  parm[Data$pos.tau] <- tau <- interval(parm[Data$pos.tau], 1e-100, Inf)
  zeta <- interval(parm[Data$pos.zeta], 1e-100, Inf)
  parm[Data$pos.zeta] <- zeta
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  tau.prior <- dhalfcauchy(tau, 25, log=TRUE)
  zeta.prior <- sum(dexp(zeta, tau, log=TRUE))
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  phi <- (1 - 2*Data$P) / (Data$P*(1 - Data$P))*zeta + mu
  sigma <- (Data$P*(1 - Data$P)*tau) / (2*zeta)
  LL <- sum(dnorm(Data$y, phi, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + tau.prior + zeta.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(phi), phi, sigma), parm=parm)
  return(Modelout)
}

```


81.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1, rep(1,N))
```

82. Revision, Normal

This example provides both an analytic solution and numerical approximation of the revision of a normal distribution. Given a normal prior distribution (α) and data distribution (β), the posterior (γ) is the revised normal distribution. This is an introductory example of Bayesian inference, and allows the user to experiment with numerical approximation, such as with MCMC in `LaplacesDemon`. Note that, regardless of the data sample size N in this example, Laplace Approximation is inappropriate due to asymptotics since the data (β) is perceived by the algorithm as a single datum rather than a collection of data. MCMC, on the other hand, is biased only by the effective number of samples taken of the posterior.

```
### Analytic Solution
```

```
prior.mu <- 0
prior.sigma <- 10
N <- 10
data.mu <- 1
data.sigma <- 2
posterior.mu <- (prior.sigma^-2 * prior.mu + N * data.sigma^-2 * data.mu) /
  (prior.sigma^-2 + N * data.sigma^-2)
posterior.sigma <- sqrt(1/(prior.sigma^-2 + data.sigma^-2))
posterior.mu
posterior.sigma
```

82.1. Form

$$\alpha \sim \mathcal{N}(0, 10)$$

$$\beta \sim \mathcal{N}(1, 2)$$

$$\gamma = \frac{\alpha_{\sigma}^{-2}\alpha + N\beta_{\sigma}^{-2}\beta}{\alpha_{\sigma}^{-2} + N\beta_{\sigma}^{-2}}$$

82.2. Data

```
N <- 10
mon.names <- c("LP", "gamma")
parm.names <- c("alpha", "beta")
PGF <- function(Data) {
  alpha <- rnorm(1, 0, 10)
  beta <- rnorm(1, 1, 2)
  return(c(alpha, beta))
}
```

```
MyData <- list(N=N, PGF=PGF, mon.names=mon.names, parm.names=parm.names)
```

82.3. Model

```
Model <- function(parm, Data)
{
  ### Hyperparameters
  alpha.mu <- 0
  alpha.sigma <- 10
  beta.mu <- 1
  beta.sigma <- 2
  ### Parameters
  alpha <- parm[1]
  beta <- parm[2]
  ### Log-Prior
  alpha.prior <- dnorm(alpha, alpha.mu, alpha.sigma, log=TRUE)
  ### Log-Likelihood
  LL <- dnorm(beta, beta.mu, beta.sigma, log=TRUE)
  ### Posterior
  gamma <- (alpha.sigma^-2 * alpha + N * beta.sigma^-2 * beta) /
    (alpha.sigma^-2 + N * beta.sigma^-2)
  ### Log-Posterior
  LP <- LL + alpha.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,gamma),
    yhat=rnorm(1, beta.mu, beta.sigma), parm=parm)
  return(Modelout)
}
```

82.4. Initial Values

```
Initial.Values <- c(0,0)
```

83. Ridge Regression

83.1. Form

$$\mathbf{y} \sim \mathcal{N}(\mu, \sigma_1^2)$$

$$\mu = \mathbf{X}\beta$$

$$\beta_1 \sim \mathcal{N}(0, 1000)$$

$$\beta_j \sim \mathcal{N}(0, \sigma_2^2), \quad j = 2, \dots, J$$

$$\sigma_k \sim \mathcal{HC}(25), \quad k = 1, \dots, 2$$

83.2. Data

```

data(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(log(demonsnacks[,-2]+1)))
J <- ncol(X)
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), sigma=rep(0,2)))
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  sigma <- runif(2)
  return(c(beta, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma, y=y)

```

83.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnorm(beta, 0, c(1000, rep(sigma[2], Data$J-1)),
    log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  LL <- sum(dnorm(Data$y, mu, sigma[1], log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma[1]), parm=parm)
  return(Modelout)
}

```

83.4. Initial Values

```
Initial.Values <- c(rep(1,J), rep(1,2))
```

84. Robust Regression

By replacing the normal distribution with the Student t distribution, linear regression is often called robust regression. As an alternative approach to robust regression, consider Laplace regression (see section 46).

84.1. Form

$$\begin{aligned} \mathbf{y} &\sim t(\mu, \sigma^2, \nu) \\ \mu &= \mathbf{X}\beta \\ \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\ \sigma &\sim \mathcal{HC}(25) \\ \nu &\sim \mathcal{HC}(25) \end{aligned}$$

84.2. Data

```
N <- 100
J <- 5
X <- matrix(1,N,J)
for (j in 2:J) {X[,j] <- rnorm(N,runif(1,-3,3),runif(1,0.1,1))}
beta <- runif(J,-3,3)
e <- rst(N,0,1,5)
y <- tcrossprod(X, t(beta)) + e
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), sigma=0, nu=0))
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.nu <- grep("nu", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  sigma <- runif(1)
  nu <- runif(1)
  return(c(beta, sigma, nu))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma,
  pos.nu=pos.nu, y=y)
```

84.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
```

```

beta <- parm[1:Data$J]
sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
parm[Data$pos.sigma] <- sigma
parm[Data$pos.nu] <- nu <- interval(parm[Data$pos.nu], 1e-100, Inf)
### Log-Prior
beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
nu.prior <- dhalfcauchy(nu, 25, log=TRUE)
### Log-Likelihood
mu <- tcrossprod(Data$X, t(beta))
LL <- sum(dst(Data$y, mu, sigma, nu, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + sigma.prior + nu.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rst(length(mu), mu, sigma, nu), parm=parm)
return(Modelout)
}

```

84.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1, 5)
```

85. Seemingly Unrelated Regression (SUR)

The following data was used by Zellner (1962) when introducing the Seemingly Unrelated Regression methodology. This model uses the Yang-Berger prior distribution for the precision matrix of a multivariate normal distribution.

85.1. Form

$$\begin{aligned}
 \mathbf{Y}_{t,k} &\sim \mathcal{N}_K(\mu_{t,k}, \Omega^{-1}), \quad t = 1, \dots, T, \quad k = 1, \dots, K \\
 \mu_{1,t} &= \alpha_1 + \alpha_2 \mathbf{X}_{t-1,1} + \alpha_3 \mathbf{X}_{t-1,2}, \quad t = 2, \dots, T \\
 \mu_{2,t} &= \beta_1 + \beta_2 \mathbf{X}_{t-1,3} + \beta_3 \mathbf{X}_{t-1,4}, \quad t = 2, \dots, T \\
 \alpha_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J
 \end{aligned}$$

where $J = 3$, $K = 2$, and $T = 20$.

85.2. Data

```

T <- 20 #Time-periods
year <- c(1935,1936,1937,1938,1939,1940,1941,1942,1943,1944,1945,1946,
  1947,1948,1949,1950,1951,1952,1953,1954)
IG <- c(33.1,45.0,77.2,44.6,48.1,74.4,113.0,91.9,61.3,56.8,93.6,159.9,

```

```

147.2,146.3,98.3,93.5,135.2,157.3,179.5,189.6)
VG <- c(1170.6,2015.8,2803.3,2039.7,2256.2,2132.2,1834.1,1588.0,1749.4,
1687.2,2007.7,2208.3,1656.7,1604.4,1431.8,1610.5,1819.4,2079.7,
2371.6,2759.9)
CG <- c(97.8,104.4,118.0,156.2,172.6,186.6,220.9,287.8,319.9,321.3,319.6,
346.0,456.4,543.4,618.3,647.4,671.3,726.1,800.3,888.9)
IW <- c(12.93,25.90,35.05,22.89,18.84,28.57,48.51,43.34,37.02,37.81,
39.27,53.46,55.56,49.56,32.04,32.24,54.38,71.78,90.08,68.60)
VW <- c(191.5,516.0,729.0,560.4,519.9,628.5,537.1,561.2,617.2,626.7,
737.2,760.5,581.4,662.3,583.8,635.2,723.8,864.1,1193.5,1188.9)
CW <- c(1.8,0.8,7.4,18.1,23.5,26.5,36.2,60.8,84.4,91.2,92.4,86.0,111.1,
130.6,141.8,136.7,129.7,145.5,174.8,213.5)
J <- 2 #Number of dependent variables
Y <- matrix(c(IG,IW), T, J)
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,3), beta=rep(0,3),
U=diag(J)), uppertri=c(0,0,1))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(3)
  beta <- rnorm(3)
  U <- runif(Data$J*(Data$J+1)/2)
  return(c(alpha, beta, U))
}
MyData <- list(J=J, PGF=PGF, T=T, Y=Y, CG=CG, CW=CW, IG=IG, IW=IW,
VG=VG, VW=VW, mon.names=mon.names, parm.names=parm.names,
pos.alpha=pos.alpha, pos.beta=pos.beta)

```

85.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- parm[Data$pos.beta]
  U <- as.parm.matrix(U, Data$J, parm, Data, chol=TRUE)
  diag(U) <- exp(diag(U))
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  U.prior <- dyangbergerc(U, log=TRUE)
  ### Log-Likelihood
  mu <- Data$Y
  mu[-1,1] <- alpha[1] + alpha[2]*Data$CG[-Data$T] +
  alpha[3]*Data$VG[-Data$T]

```

```

mu[-1,2] <- beta[1] + beta[2]*Data$CW[-Data$T] +
  beta[3]*Data$VW[-Data$T]
LL <- sum(dmvnpc(Data$Y[-1,], mu[-1,], U, log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior + U.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rmvnpc(nrow(mu), mu, U), parm=parm)
return(Modelout)
}

```

85.4. Initial Values

```
Initial.Values <- c(rep(0,3), rep(0,3), rep(0,J*(J+1)/2))
```

86. Simultaneous Equations

This example of simultaneous equations uses Klein's Model I (Kleine 1950) regarding economic fluctuations in the United States in 1920-1941 ($N=22$). Usually, this example is modeled with 3-stage least squares (3SLS), excluding the uncertainty from multiple stages. By constraining each element in the instrumental variables matrix $\nu \in [-10, 10]$, this example estimates the model without resorting to stages. The dependent variable is matrix \mathbf{Y} , in which $\mathbf{Y}_{1,1:N}$ is \mathbf{C} or Consumption, $\mathbf{Y}_{2,1:N}$ is \mathbf{I} or Investment, and $\mathbf{Y}_{3,1:N}$ is \mathbf{Wp} or Private Wages. Here is a data dictionary:

```

A = Time Trend measured as years from 1931
C = Consumption
G = Government Nonwage Spending
I = Investment
K = Capital Stock
P = Private (Corporate) Profits
T = Indirect Business Taxes Plus Neg Exports
Wg = Government Wage Bill
Wp = Private Wages
X = Equilibrium Demand (GNP)

```

See Kleine (1950) for more information.

86.1. Form

$$\mathbf{Y} \sim \mathcal{N}_3(\boldsymbol{\mu}, \boldsymbol{\Omega}^{-1})$$

$$\begin{aligned} \mu_{1,1} &= \alpha_1 + \alpha_2\nu_{1,1} + \alpha_4\nu_{2,1} \\ \mu_{1,i} &= \alpha_1 + \alpha_2\nu_{1,i} + \alpha_3\mathbf{P}_{i-1} + \alpha_4\nu_{2,i}, \quad i = 2, \dots, N \\ \mu_{2,1} &= \beta_1 + \beta_2\nu_{1,1} + \beta_4\mathbf{K}_1 \\ \mu_{2,i} &= \beta_1 + \beta_2\nu_{1,i} + \beta_3\mathbf{P}_{i-1} + \beta_4\mathbf{K}_i, \quad i = 2, \dots, N \\ \mu_{3,1} &= \gamma_1 + \gamma_2\nu_{3,1} + \gamma_4\mathbf{A}_1 \end{aligned}$$

$$\begin{aligned} \mu_{3,i} &= \gamma_1 + \gamma_2 \nu_{3,i} + \gamma_3 \mathbf{X}_{i-1} + \gamma_4 \mathbf{A}_i, \quad i = 2, \dots, N \\ \mathbf{Z}_{j,i} &\sim \mathcal{N}(\nu_{j,i}, \sigma_j^2), \quad j = 1, \dots, 3 \\ \nu_{j,1} &= \pi_{j,1} + \pi_{j,3} \mathbf{K}_1 + \pi_{j,5} \mathbf{A}_1 + \pi_{j,6} \mathbf{T}_1 + \pi_{j,7} \mathbf{G}_1, \quad j = 1, \dots, 3 \\ \nu_{j,i} &= \pi_{j,1} + \pi_{j,2} \mathbf{P}_{i-1} + \pi_{j,3} \mathbf{K}_i + \pi_{j,4} \mathbf{X}_{i-1} + \pi_{j,5} \mathbf{A}_i + \pi_{j,6} \mathbf{T}_i + \pi_{j,7} \mathbf{G}_i, \quad i = 1, \dots, N, \quad j = 1, \dots, 3 \\ \alpha_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, 4 \\ \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, 4 \\ \gamma_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, 4 \\ \pi_{j,i} &\sim \mathcal{N}(0, 1000) \in [-10, 10], \quad j = 1, \dots, 3, \quad i = 1, \dots, N \\ \sigma_j &\sim \mathcal{HC}(25), \quad j = 1, \dots, 3 \\ \Omega &\sim \mathcal{W}_4(\mathbf{S}), \quad \mathbf{S} = \mathbf{I}_3 \end{aligned}$$

86.2. Data

```

N <- 22
A <- c(-11,-10,-9,-8,-7,-6,-5,-4,-3,-2,-1,0,1,2,3,4,5,6,7,8,9,10)
C <- c(39.8,41.9,45,49.2,50.6,52.6,55.1,56.2,57.3,57.8,55,50.9,45.6,46.5,
      48.7,51.3,57.7,58.7,57.5,61.6,65,69.7)
G <- c(2.4,3.9,3.2,2.8,3.5,3.3,3.3,4,4.2,4.1,5.2,5.9,4.9,3.7,4,4.4,2.9,4.3,
      5.3,6.6,7.4,13.8)
I <- c(2.7,-0.2,1.9,5.2,3,5.1,5.6,4.2,3,5.1,1,-3.4,-6.2,-5.1,-3,-1.3,2.1,2,
      -1.9,1.3,3.3,4.9)
K <- c(180.1,182.8,182.6,184.5,189.7,192.7,197.8,203.4,207.6,210.6,215.7,
      216.7,213.3,207.1,202,199,197.7,199.8,201.8,199.9,201.2,204.5)
P <- c(12.7,12.4,16.9,18.4,19.4,20.1,19.6,19.8,21.1,21.7,15.6,11.4,7,11.2,
      12.3,14,17.6,17.3,15.3,19,21.1,23.5)
T <- c(3.4,7.7,3.9,4.7,3.8,5.5,7,6.7,4.2,4,7.7,7.5,8.3,5.4,6.8,7.2,8.3,6.7,
      7.4,8.9,9.6,11.6)
Wg <- c(2.2,2.7,2.9,2.9,3.1,3.2,3.3,3.6,3.7,4,4.2,4.8,5.3,5.6,6,6.1,7.4,
      6.7,7.7,7.8,8,8.5)
Wp <- c(28.8,25.5,29.3,34.1,33.9,35.4,37.4,37.9,39.2,41.3,37.9,34.5,29,28.5,
      30.6,33.2,36.8,41,38.2,41.6,45,53.3)
X <- c(44.9,45.6,50.1,57.2,57.1,61,64,64.4,64.5,67,61.2,53.4,44.3,45.1,
      49.7,54.4,62.7,65,60.9,69.5,75.7,88.4)
year <- c(1920,1921,1922,1923,1924,1925,1926,1927,1928,1929,1930,1931,1932,
      1933,1934,1935,1936,1937,1938,1939,1940,1941)
Y <- matrix(c(C,I,Wp),3,N, byrow=TRUE)
Z <- matrix(c(P, Wp+Wg, X), 3, N, byrow=TRUE)
S <- diag(nrow(Y))
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,4), beta=rep(0,4),
      gamma=rep(0,4), pi=matrix(0,3,7), sigma=rep(0,3),
      U=diag(3)), uppertri=c(0,0,0,0,0,1))

```



```

pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.pi <- grep("pi", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(4)
  beta <- rnorm(4)
  gamma <- rnorm(4)
  pi <- rnorm(3*7)
  sigma <- runif(3)
  U <- rwishartc(ncol(Data$Y)+1, Data$S)
  return(c(alpha, beta, gamma, pi, sigma, U[upper.tri(U, diag=TRUE)]))
}
MyData <- list(A=A, C=C, G=G, I=I, K=K, N=N, P=P, PGF=PGF, S=S, T=T, Wg=Wg,
  Wp=Wp, X=X, Y=Y, Z=Z, mon.names=mon.names, parm.names=parm.names,
  pos.alpha=pos.alpha, pos.beta=pos.beta, pos.gamma=pos.gamma,
  pos.pi=pos.pi, pos.sigma=pos.sigma)

```

86.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  beta <- parm[Data$pos.beta]
  gamma <- parm[Data$pos.gamma]
  parm[Data$pos.pi] <- pi <- interval(parm[Data$pos.pi], -10, 10)
  pi <- matrix(pi, 3, 7)      sigma <- interval(parm[Data$pos.sigma], 1e-100,
Inf)
  parm[Data$pos.sigma] <- sigma      U <- as.parm.matrix(U, nrow(Data$S), parm,
Data, chol=TRUE)
  parm[grep("Omega", Data$parm.names)] <- upper.triangle(Omega,
  diag=TRUE)
  diag(U) <- exp(diag(U))
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  gamma.prior <- sum(dnormv(gamma, 0, 1000, log=TRUE))
  pi.prior <- sum(dnormv(pi, 0, 1000, log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  U.prior <- dwishartc(U, nrow(Data$S)+1, Data$S, log=TRUE)
  ### Log-Likelihood
  mu <- nu <- matrix(0,3,Data$N)
  for (i in 1:3) {
    nu[i,1] <- pi[i,1] + pi[i,3]*Data$K[1] + pi[i,5]*Data$A[1] +

```

```

    pi[i,6]*Data$T[1] + pi[i,7]*Data$G[1]
  nu[i,-1] <- pi[i,1] + pi[i,2]*Data$P[-Data$N] +
    pi[i,3]*Data$K[-1] + pi[i,4]*Data$X[-Data$N] +
    pi[i,5]*Data$A[-1] + pi[i,6]*Data$T[-1] +
    pi[i,7]*Data$G[-1]}
LL <- sum(dnorm(Data$Z, nu, matrix(sigma, 3, Data$N), log=TRUE))
mu[1,1] <- alpha[1] + alpha[2]*nu[1,1] + alpha[4]*nu[2,1]
mu[1,-1] <- alpha[1] + alpha[2]*nu[1,-1] +
  alpha[3]*Data$P[-Data$N] + alpha[4]*nu[2,-1]
mu[2,1] <- beta[1] + beta[2]*nu[1,1] + beta[4]*Data$K[1]
mu[2,-1] <- beta[1] + beta[2]*nu[1,-1] +
  beta[3]*Data$P[-Data$N] + beta[4]*Data$K[-1]
mu[3,1] <- gamma[1] + gamma[2]*nu[3,1] + gamma[4]*Data$A[1]
mu[3,-1] <- gamma[1] + gamma[2]*nu[3,-1] +
  gamma[3]*Data$X[-Data$N] + gamma[4]*Data$A[-1]
LL <- LL + sum(dmvnpc(t(Data$Y), t(mu), U, log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior + gamma.prior + pi.prior +
  sigma.prior + U.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=t(rmvnp(ncol(mu), t(mu), U)), parm=parm)
return(Modelout)
}

```

86.4. Initial Values

```

Initial.Values <- c(rep(0,4), rep(0,4), rep(0,4), rep(0,3*7), rep(1,3),
  upper.triangle(S, diag=TRUE))

```

87. Space-Time, Dynamic

This approach to space-time or spatiotemporal modeling applies kriging to a stationary spatial component for points in space $s = 1, \dots, S$ first at time $t = 1$, where space is continuous and time is discrete. Vector ζ contains these spatial effects. Next, SSM (State Space Model) or DLM (Dynamic Linear Model) components are applied to the spatial parameters (ϕ , κ , and λ) and regression effects (β). These parameters are allowed to vary dynamically with time $t = 2, \dots, T$, and the resulting spatial process is estimated for each of these time-periods. When time is discrete, a dynamic space-time process can be applied. The matrix Θ contains the dynamically varying stationary spatial effects, or space-time effects. Spatial coordinates are given in longitude and latitude for $s = 1, \dots, S$ points in space and measurements are taken across discrete time-periods $t = 1, \dots, T$ for $\mathbf{Y}_{s,t}$. The dependent variable is also a function of design matrix \mathbf{X} (which may also be dynamic, but is static in this example) and dynamic regression effects matrix $\beta_{1:K,1:T}$. For more information on kriging, see section 44. For more information on SSMs or DLMS, see section 92. To extend this to a large spatial data set, consider incorporating the predictive process kriging example in section 45.

87.1. Form

$$\begin{aligned}
 \mathbf{Y}_{s,t} &\sim \mathcal{N}(\mu_{s,t}, \sigma_1^2), \quad s = 1, \dots, S, \quad t = 1, \dots, T \\
 \mu_{s,t} &= \mathbf{X}_{s,1:K} \beta_{1:K,t} + \Theta_{s,t} \\
 \Theta_{s,t} &= \frac{\Sigma_{s,s,t}}{\sum_{r=1}^S \Sigma_{r,s,t}} \Theta_{s,t-1}, \quad s = 1, \dots, S, \quad t = 2, \dots, T \\
 \Theta_{s,1} &= \zeta_s \\
 \zeta &\sim \mathcal{N}_S(0, \Sigma_{1:S,1:S,1}) \\
 \Sigma_{1:S,1:S,t} &= \lambda_t^2 \exp(-\phi_t \mathbf{D})^{\kappa[t]} \\
 \sigma_j &\sim \mathcal{HC}(25), \quad j = 1, \dots, 4 \\
 \beta_{k,1} &\sim \mathcal{N}(0, 1000), \quad k = 1, \dots, K \\
 \beta_{k,t} &\sim \mathcal{N}(\beta_{k,t-1}, \tau_k^2), \quad k = 1, \dots, K, \quad t = 2, \dots, T \\
 \phi_1 &\sim \mathcal{HN}(1000) \\
 \phi_t &\sim \mathcal{N}(\phi_{t-1}, \sigma_2^2) \in [0, \infty], \quad t = 2, \dots, T \\
 \kappa_1 &\sim \mathcal{HN}(1000) \\
 \kappa_t &\sim \mathcal{N}(\kappa_{t-1}, \sigma_3^2) \in [0, \infty], \quad t = 2, \dots, T \\
 \lambda_1 &\sim \mathcal{HN}(1000) \\
 \lambda_t &\sim \mathcal{N}(\lambda_{t-1}, \sigma_4^2) \in [0, \infty], \quad t = 2, \dots, T
 \end{aligned}$$

87.2. Data

```

data(demontexas)
Y <- as.matrix(demontexas[1:20,c(18:30)])
X <- cbind(1,as.matrix(demontexas[1:20,c(1,4)])) #Static predictors
latitude <- demontexas[1:20,2]
longitude <- demontexas[1:20,3]
D <- as.matrix(dist(cbind(longitude,latitude), diag=TRUE, upper=TRUE))
S <- nrow(Y) #Number of sites, or points in space
T <- ncol(Y) #Number of time-periods
K <- ncol(X) #Number of columns in design matrix X including the intercept
mon.names <- "LP"
parm.names <- as.parm.names(list(zeta=rep(0,S), beta=matrix(0,K,T),
  phi=rep(0,T), kappa=rep(0,T), lambda=rep(0,T), sigma=rep(0,4),
  tau=rep(0,K)))
pos.zeta <- grep("zeta", parm.names)
pos.beta <- grep("beta", parm.names)
pos.phi <- grep("phi", parm.names)
pos.kappa <- grep("kappa", parm.names)
pos.lambda <- grep("lambda", parm.names)
pos.sigma <- grep("sigma", parm.names)

```

```

pos.tau <- grep("tau", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$K*Data$T, rbind(mean(Data$Y),
    matrix(0, Data$K-1, Data$T)), 1)
  phi <- rhalfnorm(Data$T, 1)
  kappa <- rhalfnorm(Data$T, 1)
  lambda <- rhalfnorm(Data$T, 1)
  Sigma <- lambda[1]*lambda[1]*exp(-phi[1]*Data$D)^kappa[1]
  zeta <- as.vector(rmvn(1, rep(0,Data$S), Sigma))
  sigma <- runif(4)
  tau <- runif(Data$K)
  return(c(zeta, beta, phi, kappa, lambda, sigma, tau))
}
MyData <- list(D=D, K=K, PGF=PGF, S=S, T=T, X=X, Y=Y, latitude=latitude,
  longitude=longitude, mon.names=mon.names, parm.names=parm.names,
  pos.zeta=pos.zeta, pos.beta=pos.beta, pos.phi=pos.phi,
  pos.kappa=pos.kappa, pos.lambda=pos.lambda, pos.sigma=pos.sigma,
  pos.tau=pos.tau)

```

87.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- matrix(parm[Data$pos.beta], Data$K, Data$T)
  zeta <- parm[Data$pos.zeta]
  parm[Data$pos.phi] <- phi <- interval(parm[Data$pos.phi], 1e-100, Inf)
  kappa <- interval(parm[Data$pos.kappa], 1e-100, Inf)
  parm[Data$pos.kappa] <- kappa
  lambda <- interval(parm[Data$pos.lambda], 1e-100, Inf)
  parm[Data$pos.lambda] <- lambda
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  parm[Data$pos.tau] <- tau <- interval(parm[Data$pos.tau], 1e-100, Inf)
  Sigma <- array(0, dim=c(Data$S, Data$S, Data$T))
  for (t in 1:Data$T) {
    Sigma[ , ,t] <- lambda[t]^2 * exp(-phi[t] * Data$D)^kappa[t]}
  ### Log-Prior
  beta.prior <- sum(dnormv(beta[,1], 0, 1000, log=TRUE),
    dnorm(beta[,-1], beta[,-Data$T], matrix(tau, Data$K,
    Data$T-1), log=TRUE))
  zeta.prior <- dmvn(zeta, rep(0,Data$S), Sigma[ , , 1], log=TRUE)
  phi.prior <- sum(dhalfnorm(phi[1], sqrt(1000), log=TRUE),
    dtrunc(phi[-1], "norm", a=0, b=Inf, mean=phi[-Data$T],
    sd=sigma[2], log=TRUE))
  kappa.prior <- sum(dhalfnorm(kappa[1], sqrt(1000), log=TRUE),

```

```

    dtrunc(kappa[-1], "norm", a=0, b=Inf, mean=kappa[-Data$T],
    sd=sigma[3], log=TRUE))
lambda.prior <- sum(dhalfnorm(lambda[1], sqrt(1000), log=TRUE),
    dtrunc(lambda[-1], "norm", a=0, b=Inf, mean=lambda[-Data$T],
    sd=sigma[4], log=TRUE))
sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
tau.prior <- sum(dhalfcauchy(tau, 25, log=TRUE))
### Log-Likelihood
mu <- tcrossprod(Data$X, t(beta))
Theta <- matrix(zeta, Data$S, Data$T)
for (t in 2:Data$T) {
  for (s in 1:Data$S) {
    Theta[s,t] <- sum(Sigma[,s,t] / sum(Sigma[,s,t]) * Theta[,t-1])}
mu <- mu + Theta
LL <- sum(dnorm(Data$Y, mu, sigma[1], log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + zeta.prior + sum(phi.prior) +
  sum(kappa.prior) + sum(lambda.prior) + sigma.prior + tau.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(prod(dim(mu)), mu, sigma[1]), parm=parm)
return(Modelout)
}

```

87.4. Initial Values

```

Initial.Values <- c(rep(0,S), rep(c(mean(Y),rep(0,K-1)),T), rep(1,T),
  rep(1,T), rep(1,T), rep(1,4), rep(1,K))

```

88. Space-Time, Nonseparable

This approach to space-time or spatiotemporal modeling applies kriging both to the stationary spatial and temporal components, where space is continuous and time is discrete. Matrix Ξ contains the space-time effects. Spatial coordinates are given in longitude and latitude for $s = 1, \dots, S$ points in space and measurements are taken across time-periods $t = 1, \dots, T$ for $\mathbf{Y}_{s,t}$. The dependent variable is also a function of design matrix \mathbf{X} and regression effects vector β . For more information on kriging, see section 44. This example uses a nonseparable, stationary covariance function in which space and time are separable only when $\psi = 0$. To extend this to a large space-time data set, consider incorporating the predictive process kriging example in section 45.

88.1. Form

$$\mathbf{Y}_{s,t} \sim \mathcal{N}(\mu_{s,t}, \sigma_1^2), \quad s = 1, \dots, S, \quad t = 1, \dots, T$$

$$\mu = \mathbf{X}\beta + \Xi$$

$$\begin{aligned} \Xi &\sim \mathcal{N}_{ST}(\Xi_\mu, \Sigma) \\ \Sigma &= \sigma_2^2 \exp \left(-\frac{\mathbf{D}_S^\kappa}{\phi_1} - \frac{\mathbf{D}_T^\lambda}{\phi_2} - \psi \frac{\mathbf{D}_S^\kappa \mathbf{D}_T^\lambda}{\phi_1 \phi_2} \right) \\ \beta_k &\sim \mathcal{N}(0, 1000), \quad k = 1, \dots, K \\ \phi_j &\sim \mathcal{U}(1, 5), \quad j = 1, \dots, 2 \\ \sigma_j &\sim \mathcal{HC}(25), \quad j = 1, \dots, 2 \\ \psi &\sim \mathcal{HC}(25) \\ \Xi_\mu &= 0 \\ \kappa &= 1, \quad \lambda = 1 \end{aligned}$$

88.2. Data

```

data(demontexas)
Y <- as.matrix(demontexas[1:10,c(18:30)])
X <- cbind(1,as.matrix(demontexas[1:10,c(1,4)])) #Static predictors
latitude <- demontexas[1:10,2]
longitude <- demontexas[1:10,3]
S <- nrow(Y) #Number of sites, or points in space
T <- ncol(Y) #Number of time-periods
K <- ncol(X) #Number of columns in design matrix X including the intercept
D.S <- as.matrix(dist(cbind(rep(longitude,T),rep(latitude,T)), diag=TRUE,
  upper=TRUE))
D.T <- as.matrix(dist(cbind(rep(1:T,each=S),rep(1:T,each=S)), diag=TRUE,
  upper=TRUE))
mon.names <- "LP"
parm.names <- as.parm.names(list(Xi=matrix(0,S,T), beta=rep(0,K),
  phi=rep(0,2), sigma=rep(0,2), psi=0))
pos.Xi <- grep("Xi", parm.names)
pos.beta <- grep("beta", parm.names)
pos.phi <- grep("phi", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.psi <- grep("psi", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$K, c(mean(Data$Y),rep(0,Data$K-1)), 1)
  phi <- runif(2,1,5)
  sigma <- runif(2)
  psi <- runif(1)
  kappa <- 1
  lambda <- 1
  Sigma <- sigma[2]*sigma[2] * exp(-(Data$D.S / phi[1])^kappa -
    (Data$D.T / phi[2])^lambda -
    psi*(Data$D.S / phi[1])^kappa * (Data$D.T / phi[2])^lambda)
  Xi <- as.vector(rmvn(1, rep(0,Data$S*Data$T), Sigma))
  return(c(Xi, beta, phi, sigma, psi))
}

```

```

}
MyData <- list(D.S=D.S, D.T=D.T, K=K, PGF=PGF, S=S, T=T, X=X, Y=Y,
  latitude=latitude, longitude=longitude, mon.names=mon.names,
  parm.names=parm.names, pos.Xi=pos.Xi, pos.beta=pos.beta,
  pos.phi=pos.phi, pos.sigma=pos.sigma, pos.psi=pos.psi)

```

88.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperparameters
  Xi.mu <- rep(0,Data$S*Data$T)
  ### Parameters
  beta <- parm[Data$pos.beta]
  Xi <- parm[Data$pos.Xi]
  kappa <- 1; lambda <- 1
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  parm[Data$pos.phi] <- phi <- interval(parm[Data$pos.phi], 1, 5)
  parm[Data$pos.psi] <- psi <- interval(parm[Data$pos.psi], 1e-100, Inf)
  Sigma <- sigma[2]*sigma[2] * exp(-(Data$D.S / phi[1])^kappa -
    (Data$D.T / phi[2])^lambda -
    psi*(Data$D.S / phi[1])^kappa * (Data$D.T / phi[2])^lambda)
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  Xi.prior <- dmvn(Xi, Xi.mu, Sigma, log=TRUE)
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  phi.prior <- sum(dunif(phi, 1, 5, log=TRUE))
  psi.prior <- dhalfcauchy(psi, 25, log=TRUE)
  ### Log-Likelihood
  Xi <- matrix(Xi, Data$S, Data$T)
  mu <- as.vector(tcrossprod(Data$X, t(beta))) + Xi
  LL <- sum(dnorm(Data$Y, mu, sigma[1], log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + Xi.prior + sigma.prior + phi.prior + psi.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(prod(dim(mu)), mu, sigma[1]), parm=parm)
  return(Modelout)
}

```

88.4. Initial Values

```

Initial.Values <- c(rep(0,S*T), c(mean(Y),rep(0,K-1)), rep(1,2), rep(1,2),
  1)

```

89. Space-Time, Separable

This introductory approach to space-time or spatiotemporal modeling applies kriging both to the stationary spatial and temporal components, where space is continuous and time is discrete. Vector ζ contains the spatial effects and vector θ contains the temporal effects. Spatial coordinates are given in longitude and latitude for $s = 1, \dots, S$ points in space and measurements are taken across time-periods $t = 1, \dots, T$ for $\mathbf{Y}_{s,t}$. The dependent variable is also a function of design matrix \mathbf{X} and regression effects vector β . For more information on kriging, see section 44. This example uses separable space-time covariances, which is more convenient but usually less appropriate than a nonseparable covariance function. To extend this to a large space-time data set, consider incorporating the predictive process kriging example in section 45.

89.1. Form

$$\begin{aligned} \mathbf{Y}_{s,t} &\sim \mathcal{N}(\mu_{s,t}, \sigma_1^2), \quad s = 1, \dots, S, \quad t = 1, \dots, T \\ \mu_{s,t} &= \mathbf{X}_{s,1:T} \beta + \zeta_s + \Theta_{s,t} \\ \Theta_{s,1:T} &= \theta \\ \theta &\sim \mathcal{N}_N(\theta_\mu, \Sigma_T) \\ \Sigma_T &= \sigma_3^2 \exp(-\phi_2 \mathbf{D}_T)^\lambda \\ \zeta &\sim \mathcal{N}_N(\zeta_\mu, \Sigma_S) \\ \Sigma_S &= \sigma_2^2 \exp(-\phi_1 \mathbf{D}_S)^\kappa \\ \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, 2 \\ \sigma_k &\sim \mathcal{HC}(25), \quad k = 1, \dots, 3 \\ \phi_k &\sim \mathcal{U}(1, 5), \quad k = 1, \dots, 2 \\ \zeta_\mu &= 0 \\ \theta_\mu &= 0 \\ \kappa &= 1, \quad \lambda = 1 \end{aligned}$$

89.2. Data

```
data(demontexas)
Y <- as.matrix(demontexas[1:20,c(18:30)])
X <- cbind(1,as.matrix(demontexas[1:20,c(1,4)])) #Static predictors
latitude <- demontexas[1:20,2]
longitude <- demontexas[1:20,3]
S <- nrow(Y) #Number of sites, or points in space
T <- ncol(Y) #Number of time-periods
K <- ncol(X) #Number of columns in design matrix X including the intercept
D.S <- as.matrix(dist(cbind(longitude,latitude), diag=TRUE, upper=TRUE))
D.T <- as.matrix(dist(cbind(c(1:T),c(1:T)), diag=TRUE, upper=TRUE))
```



```

mon.names <- "LP"
parm.names <- as.parm.names(list(zeta=rep(0,S), theta=rep(0,T),
  beta=rep(0,K), phi=rep(0,2), sigma=rep(0,3)))
pos.zeta <- grep("zeta", parm.names)
pos.theta <- grep("theta", parm.names)
pos.beta <- grep("beta", parm.names)
pos.phi <- grep("phi", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$K, c(mean(Data$Y),rep(0,Data$K-1)), 1)
  phi <- runif(2,1,5)
  sigma <- runif(3)
  kappa <- 1
  lambda <- 1
  Sigma.S <- sigma[2]^2 * exp(-phi[1] * Data$D.S)^kappa
  Sigma.T <- sigma[3]^2 * exp(-phi[2] * Data$D.T)^lambda
  zeta <- as.vector(rmvn(1, rep(0,Data$S), Sigma.S))
  theta <- as.vector(rmvn(1, rep(0,Data$T), Sigma.T))
  return(c(zeta, theta, beta, phi, sigma))
}
MyData <- list(D.S=D.S, D.T=D.T, K=K, PGF=PGF, S=S, T=T, X=X, Y=Y,
  latitude=latitude, longitude=longitude, mon.names=mon.names,
  parm.names=parm.names, pos.zeta=pos.zeta, pos.theta=pos.theta,
  pos.beta=pos.beta, pos.phi=pos.phi, pos.sigma=pos.sigma)

```

89.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperparameters
  zeta.mu <- rep(0,Data$S)
  theta.mu <- rep(0,Data$T)
  ### Parameters
  beta <- parm[Data$pos.beta]
  zeta <- parm[Data$pos.zeta]
  theta <- parm[Data$pos.theta]
  kappa <- 1; lambda <- 1
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  parm[Data$pos.phi] <- phi <- interval(parm[Data$pos.phi], 1, 5)
  Sigma.S <- sigma[2]^2 * exp(-phi[1] * Data$D.S)^kappa
  Sigma.T <- sigma[3]^2 * exp(-phi[2] * Data$D.T)^lambda
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  zeta.prior <- dmvn(zeta, zeta.mu, Sigma.S, log=TRUE)
  theta.prior <- dmvn(theta, theta.mu, Sigma.T, log=TRUE)
}

```

```

sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
phi.prior <- sum(dunif(phi, 1, 5, log=TRUE))
### Log-Likelihood
Theta <- matrix(theta, Data$S, Data$T, byrow=TRUE)
mu <- as.vector(tcrossprod(Data$X, t(beta))) + zeta + Theta
LL <- sum(dnorm(Data$Y, mu, sigma[1], log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + zeta.prior + theta.prior + sigma.prior +
      phi.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
                yhat=rnorm(prod(dim(mu)), mu, sigma[1]), parm=parm)
return(Modelout)
}

```

89.4. Initial Values

```
Initial.Values <- c(rep(0,S), rep(0,T), rep(0,2), rep(1,2), rep(1,3))
```

90. Spatial Autoregression (SAR)

The spatial autoregressive (SAR) model in this example uses areal data that consists of first-order neighbors that were specified and converted from point-based data with longitude and latitude coordinates.

90.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\
 \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\phi}\mathbf{z} \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \phi &\sim \mathcal{U}(-1, 1) \\
 \sigma &\sim \mathcal{HC}(25)
 \end{aligned}$$

90.2. Data

```

N <- 100
latitude <- runif(N,0,100); longitude <- runif(N,0,100)
J <- 3 #Number of predictors, including the intercept
X <- matrix(runif(N*J,0,3), N, J); X[,1] <- 1
beta.orig <- runif(J,0,3); phi <- runif(1,0,1)
D <- as.matrix(dist(cbind(longitude, latitude), diag=TRUE, upper=TRUE))
W <- exp(-D) #Inverse distance as weights
W <- ifelse(D == 0, 0, W)
epsilon <- rnorm(N,0,1)

```

```

y <- tcrossprod(X, t(beta.orig)) + sqrt(latitude) + sqrt(longitude) +
  epsilon
Z <- W / matrix(rowSums(W), N, N) * matrix(y, N, N, byrow=TRUE)
z <- rowSums(Z)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), phi=0, sigma=0))
pos.beta <- grep("beta", parm.names)
pos.phi <- grep("phi", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  phi <- runif(1,-1,1)
  sigma <- runif(1)
  return(c(beta, phi, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, latitude=latitude, longitude=longitude,
  mon.names=mon.names, parm.names=parm.names, pos.beta=pos.beta,
  pos.phi=pos.phi, pos.sigma=pos.sigma, y=y, z=z)

```

90.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  parm[Data$pos.phi] <- phi <- interval(parm[Data$pos.phi], -1, 1)
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  phi.prior <- dunif(phi, -1, 1, log=TRUE)
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta)) + phi*Data$z
  LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + phi.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

90.4. Initial Values

```
Initial.Values <- c(rep(0,J), 0.5, 1)
```

91. STARMA(p,q)

The data in this example of a space-time autoregressive moving average (STARMA) are coordinate-based, and the adjacency matrix \mathbf{A} is created from K nearest neighbors. Otherwise, an adjacency matrix may be specified as usual for areal data. Spatial coordinates are given in longitude and latitude for $s = 1, \dots, S$ points in space and measurements are taken across time-periods $t = 1, \dots, T$ for $\mathbf{Y}_{s,t}$.

91.1. Form

$$\begin{aligned} \mathbf{Y} &\sim \mathcal{N}(\mu, \sigma^2) \\ \mu_{s,t} &= \sum_{j=1}^J \mathbf{X}_{s,t,j} \beta_j + \sum_{l=1}^L \sum_{p=1}^P \phi_{l,p} \mathbf{W}_{s,t-p,l}^1 + \sum_{m=1}^M \sum_{q=1}^Q \theta_{m,q} \mathbf{W}_{s,t-q,m}^2, \quad j = 1, \dots, J, \quad s = 1, \dots, S, \quad t = p, \dots, T \\ \mathbf{W}_{1:S,1:T,l}^1 &= \mathbf{V}_{1:S,1:S,l} \mathbf{Y}, \quad l = 1, \dots, L \\ \mathbf{W}_{1:S,1:T,m}^2 &= \mathbf{V}_{1:S,1:S,m} \epsilon, \quad m = 1, \dots, M \\ \epsilon &= \mathbf{Y} - \mu \\ \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\ \phi_{l,p} &\sim \mathcal{U}(-1, 1), \quad l = 1, \dots, L, \quad p = 1, \dots, P \\ \sigma &\sim \mathcal{HC}(25) \\ \theta_{m,q} &\sim \mathcal{N}(0, 1000), \quad m = 1, \dots, M, \quad q = 1, \dots, Q \end{aligned}$$

where \mathbf{V} is an adjacency array that is scaled so that each row sums to one, β is a vector of regression effects, ϕ is a matrix of autoregressive space-time parameters, σ is the residual variance, and θ is a matrix of moving average space-time parameters.

91.2. Data

```
data(demontexas)
Y <- t(diff(t(as.matrix(demontexas[,c(18:30)])))) #Note this is not stationary
X <- array(1, dim=c(369,13-1,3))
X[, , 2] <- CenterScale(demontexas[,1])
X[, , 3] <- demontexas[,4]
latitude <- demontexas[,2]
longitude <- demontexas[,3]
S <- nrow(Y) #Number of sites, or points in space
T <- ncol(Y) #Number of time-periods
J <- dim(X)[3] #Number of columns in design matrix X including the intercept
K <- 5 #Number of nearest neighbors
L <- 2 #Spatial autoregressive order
M <- 2 #Spatial moving average order
P <- 2 #Autoregressive order
Q <- 2 #Moving average order
D <- as.matrix(dist(cbind(longitude, latitude), diag=TRUE, upper=TRUE))
```

```

A <- V <- array(0, dim=c(nrow(D),ncol(D),P))
W1 <- array(0, dim=c(S,T,max(L,M)))
for (l in 1:max(L,M)) {
  A[, , l] <- exp(-D)
  A[, , l] <- apply(A[, , l], 1, rank)
  A[, , l] <- ifelse(A[, , l] > (l-1)*K & A[, , l] <= l*K, 1, 0)
  V[, , l] <- A[, , l] / rowSums(A[, , l])
  V[, , l] <- ifelse(is.nan(V[, , l]), 1/ncol(V[, , l]), V[, , l])
  W1[, , l] <- tcrossprod(V[, , l], t(Y))}
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), phi=matrix(0,L,P), sigma=0,
  theta=matrix(0,M,Q)))
pos.beta <- grep("beta", parm.names)
pos.phi <- grep("phi", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.theta <- grep("theta", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  phi <- runif(Data$L*Data$P,-1,1)
  sigma <- runif(1)
  theta <- rnorm(Data$M*Data$Q)
  return(c(beta, phi, sigma, theta))
}
MyData <- list(J=J, K=K, L=L, M=M, P=P, Q=Q, PGF=PGF, S=S, T=T, V=V, W1=W1,
  X=X, Y=Y, latitude=latitude, longitude=longitude, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.phi=pos.phi,
  pos.sigma=pos.sigma, pos.theta=pos.theta)

```

91.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  phi <- matrix(interval(parm[Data$pos.phi], -1, 1), Data$L, Data$P)
  parm[Data$pos.phi] <- as.vector(phi)
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  theta <- matrix(parm[Data$pos.theta], Data$M, Data$Q)
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  phi.prior <- sum(dunif(phi, -1, 1, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  theta.prior <- sum(dnormv(theta, 0, 1000, log=TRUE))
  ### Log-Likelihood
  mu <- beta[1]*Data$X[, , 1]

```

```

for (j in 2:Data$J) mu <- mu + beta[j]*Data$X[, , j]
for (l in 1:Data$L) {for (p in 1:Data$P) {
  mu[,-c(1:p)] <- mu[,-c(1:p)] +
    phi[l,p]*Data$W1[, 1:(Data$T - p), l]}}
epsilon <- Data$Y - mu
for (m in 1:Data$M) {
  W2 <- tcrossprod(Data$V[, , m], t(epsilon))
  for (q in 1:Data$Q) {
    mu[,-c(1:q)] <- mu[,-c(1:q)] +
      theta[m,q]*W2[,1:(Data$T - q)]}}
LL <- sum(dnorm(Data$Y[,-c(1:max(Data$P,Data$Q))],
  mu[,-c(1:max(Data$P,Data$Q))], sigma, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + phi.prior + sigma.prior + theta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(prod(dim(mu)), mu, sigma), parm=parm)
return(Modelout)
}

```

91.4. Initial Values

```
Initial.Values <- c(rep(0,J), rep(0,L*P), 1, rep(0,M*Q))
```

92. State Space Model (SSM), Linear Regression

92.1. Form

$$\begin{aligned}
 \mathbf{y}_t &\sim \mathcal{N}(\mu_t, \sigma_{J+1}^2), \quad t = 1, \dots, T \\
 \mu &= \mathbf{X}\beta \\
 \beta_{t,j} &\sim \mathcal{N}(\mu_j + \phi_j(\beta_{t-1,j} - \mu_j), \sigma_j^2), \quad t = 2, \dots, T, \quad j = 1, \dots, J \\
 \beta_{1,j} &\sim \mathcal{N}(\mu_j + \phi_j(b_j^0 - \mu_j), \sigma_j^2), \quad j = 1, \dots, J \\
 b_j^0 &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \mu_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \phi_j &\sim \mathcal{BETA}(20, 1.5) \quad j = 1, \dots, J \\
 \sigma_j &\sim \mathcal{HC}(25), \quad j = 1, \dots, (J+1)
 \end{aligned}$$

92.2. Data

```

data(demonfx)
y <- demonfx[1:50,1]

```

```

X <- cbind(1, as.matrix(demonfx[1:50,2:3]))
T <- nrow(X)
J <- ncol(X)
mon.names <- "LP"
parm.names <- as.parm.names(list(b0=rep(0,J), beta=matrix(0,T,J),
  mu=rep(0,J), phi=rep(0,J), sigma=rep(0,J+1)))
pos.b0 <- grep("b0", parm.names)
pos.beta <- grep("beta", parm.names)
pos.mu <- grep("mu", parm.names)
pos.phi <- grep("phi", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  b0 <- rnorm(Data$J)
  beta <- c(rnorm(Data$T,mean(Data$y),1), rnorm(Data$T*(Data$J-1)))
  mu <- rnorm(Data$J)
  phi <- runif(Data$J, -1, 1)
  sigma <- runif(Data$J+1)
  return(c(beta, mu, phi, sigma))
}
MyData <- list(J=J, PGF=PGF, T=T, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.b0=pos.b0, pos.beta=pos.beta,
  pos.mu=pos.mu, pos.phi=pos.phi, pos.sigma=pos.sigma, y=y)

```

92.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  b0 <- parm[Data$pos.b0]
  beta <- matrix(parm[Data$pos.beta], Data$T, Data$J)
  mu <- parm[Data$pos.mu]
  parm[Data$pos.phi] <- phi <- interval(parm[Data$pos.phi], -1, 1)
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  b0.prior <- sum(dnormv(b0, 0, 1000, log=TRUE))
  beta.prior <- sum(dnorm(beta, matrix(mu, Data$T, Data$J, byrow=TRUE) +
    matrix(phi, Data$T, Data$J, byrow=TRUE) *
    (rbind(b0, beta[-Data$T,]) -
    matrix(mu, Data$T, Data$J, byrow=TRUE)),
    matrix(sigma[1:Data$J], Data$T, Data$J, byrow=TRUE), log=TRUE))
  mu.prior <- sum(dnormv(mu, 0, 1000, log=TRUE))
  phi.prior <- sum(dbeta((phi+1)/2, 20, 1.5, log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  mu <- rowSums(beta*Data$X)

```

```

LL <- sum(dnorm(Data$y, mu, sigma[Data$J+1], log=TRUE))
yhat <- rnorm(length(mu), mu, sigma[Data$J+1]) #Fitted
#yhat <- rnorm(length(mu), rowSums(matrix(rnorm(Data$T*Data$J,
      # rbind(b0, beta[-Data$T,]), matrix(sigma[-Data$J], Data$T, Data$J,
      # byrow=TRUE))), Data$T, Data$J) * Data$X), sigma[Data$J+1]) #One-step
ahead
### Log-Posterior
LP <- LL + b0.prior + beta.prior + mu.prior + phi.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP, yhat=yhat, parm=parm)
return(Modelout)
}

```

92.4. Initial Values

```

Initial.Values <- c(rep(0,J), rep(mean(y),T), rep(0,T*(J-1)), rep(0,J),
  rep(0,J), rep(1,J+1))

```

93. State Space Model (SSM), Local Level

The local level model is the simplest, non-trivial example of a state space model (SSM). As such, this version of a local level SSM has static variance parameters.

93.1. Form

$$\begin{aligned}
 \mathbf{y}_t &\sim \mathcal{N}(\mu_t, \sigma_1^2), & t = 1, \dots, T \\
 \mu_t &\sim \mathcal{N}(\mu_{t-1}, \sigma_2^2), & t = 2, \dots, T \\
 \mu_1 &\sim \mathcal{N}(0, 1000) \\
 \sigma_j &\sim \mathcal{HC}(25), & j = 1, \dots, 2
 \end{aligned}$$

93.2. Data

```

T <- 20
T.m <- 14
mu.orig <- rep(0,T)
for (t in 2:T) {mu.orig[t] <- mu.orig[t-1] + rnorm(1,0,1)}
y <- mu.orig + rnorm(T,0,0.1)
y[(T.m+2):T] <- NA
mon.names <- rep(NA, (T-T.m))
for (i in 1:(T-T.m)) mon.names[i] <- paste("yhat[",(T.m+i),"]", sep="")
parm.names <- as.parm.names(list(mu=rep(0,T), sigma=rep(0,2)))
pos.mu <- grep("mu", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {

```



```

mu <- rnorm(Data$T)
sigma <- runif(2)
return(c(mu, sigma))
}
MyData <- list(PGF=PGF, T=T, T.m=T.m, mon.names=mon.names,
  parm.names=parm.names, pos.mu=pos.mu, pos.sigma=pos.sigma, y=y)
Dyn <- matrix(paste("mu[" ,1:T, "]" ,sep=""), T, 1)

```

93.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  mu <- parm[Data$pos.mu]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  mu.prior <- sum(dnorm(mu[1], 0, 1000, log=TRUE),
    dnorm(mu[-1], mu[-Data$T], sigma[2], log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  LL <- sum(dnorm(Data$y[1:Data$T.m], mu[1:Data$T.m], sigma[1],
    log=TRUE))
  yhat <- rnorm(length(mu), c(mu[1], rnorm(Data$T-1, mu[-Data$T],
    sigma[2])), sigma[1]) #One-step ahead    ### Log-Posterior
  LP <- LL + mu.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=mu[(Data$T.m+1):Data$T],
    yhat=yhat, parm=parm)
  return(Modelout)
}

```

93.4. Initial Values

```
Initial.Values <- c(rep(0,T), rep(1,2))
```

94. State Space Model (SSM), Local Linear Trend

The local linear trend model is a state space model (SSM) that extends the local level model to include a dynamic slope parameter. For more information on the local level model, see section 93. This example has static variance parameters.

94.1. Form

$$\mathbf{y}_t \sim \mathcal{N}(\mu_t, \sigma_1^2), \quad t = 1, \dots, T$$

$$\begin{aligned}\mu_t &\sim \mathcal{N}(\mu_{t-1} + \delta_{t-1}, \sigma_2^2), \quad t = 2, \dots, T \\ \mu_1 &\sim \mathcal{N}(0, 1000) \\ \delta_t &\sim \mathcal{N}(\delta_{t-1}, \sigma_3^2), \quad t = 2, \dots, T \\ \delta_1 &\sim \mathcal{N}(0, 1000) \\ \sigma_j &\sim \mathcal{HC}(25), \quad j = 1, \dots, 3\end{aligned}$$

94.2. Data

```
T <- 20
T.m <- 14
mu.orig <- delta.orig <- rep(0,T)
for (t in 2:T) {
  delta.orig[t] <- delta.orig[t-1] + rnorm(1,0,0.1)
  mu.orig[t] <- mu.orig[t-1] + delta.orig[t-1] + rnorm(1,0,1)}
y <- mu.orig + rnorm(T,0,0.1)
y[(T.m+2):T] <- NA
mon.names <- rep(NA, (T-T.m))
for (i in 1:(T-T.m)) mon.names[i] <- paste("yhat[",(T.m+i),"]", sep="")
parm.names <- as.parm.names(list(mu=rep(0,T), delta=rep(0,T),
  sigma=rep(0,3)))
pos.mu <- grep("mu", parm.names)
pos.delta <- grep("delta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  mu <- rnorm(Data$T)
  delta <- rnorm(Data$T)
  sigma <- runif(3)
  return(c(mu, delta, sigma))
}
MyData <- list(PGF=PGF, T=T, T.m=T.m, mon.names=mon.names,
  parm.names=parm.names, pos.mu=pos.mu, pos.delta=pos.delta,
  pos.sigma=pos.sigma, y=y)
```

94.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  mu <- parm[Data$pos.mu]
  delta <- parm[Data$pos.delta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  mu.prior <- sum(dnormv(mu[1], 0, 1000, log=TRUE),
```

```

    dnorm(mu[-1], mu[-Data$T]+delta[-Data$T], sigma[2],
    log=TRUE))
delta.prior <- sum(dnormv(delta[1], 0, 1000, log=TRUE),
    dnorm(delta[-1], delta[-Data$T], sigma[3], log=TRUE))
sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
### Log-Likelihood
LL <- sum(dnorm(Data$y[1:Data$T.m], mu[1:Data$T.m], sigma[1],
    log=TRUE))
yhat <- rnorm(length(mu), c(mu[1], rnorm(Data$T-1, mu[-Data$T],
    sigma[2])), sigma[1]) #One-step ahead
### Log-Posterior
LP <- LL + mu.prior + delta.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=mu[(Data$T.m+1):Data$T],
    yhat=yhat, parm=parm)
return(Modelout)
}

```

94.4. Initial Values

```
Initial.Values <- c(rep(0,T), rep(0,T), rep(1,3))
```

95. State Space Model (SSM), Stochastic Volatility (SV)

95.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(0, \sigma^2) \\
 \sigma^2 &= \frac{1}{\exp(\theta)} \\
 \beta &= \exp(\mu/2) \\
 \theta_1 &\sim \mathcal{N}(\mu + \phi(\alpha - \mu), \tau) \\
 \theta_t &\sim \mathcal{N}(\mu + \phi(\theta_{t-1} - \mu), \tau), \quad t = 2, \dots, T \\
 \alpha &\sim \mathcal{N}(\mu, \tau) \\
 \phi &\sim \mathcal{U}(-1, 1) \\
 \mu &\sim \mathcal{N}(0, 10) \\
 \tau &\sim \mathcal{HC}(25)
 \end{aligned}$$

95.2. Data

```
T <- 20
y <- rep(10,T); epsilon <- rnorm(T,0,1)
```

```

for (t in 2:T) {y[t] <- 0.8*y[t-1] + epsilon[t-1]}
mon.names <- c("LP",paste("sigma2[",1:T,"]",sep=""))
parm.names <- as.parm.names(list(theta=rep(0,T), alpha=0, phi=0, mu=0,
    tau=0))
pos.theta <- grep("theta", parm.names)
pos.alpha <- grep("alpha", parm.names)
pos.phi <- grep("phi", parm.names)
pos.mu <- grep("mu", parm.names)
pos.tau <- grep("tau", parm.names)
PGF <- function(Data) {
  phi <- runif(1,-1,1)
  mu <- rnorm(1)
  tau <- runif(1)
  alpha <- rnorm(1, mu, tau)
  theta <- rnorm(Data$T, mu + phi*(alpha - mu), tau)
  return(c(theta, alpha, phi, mu, tau))
}
MyData <- list(PGF=PGF, T=T, mon.names=mon.names, parm.names=parm.names,
pos.alpha=pos.alpha, pos.phi=pos.phi,
pos.mu=pos.mu, pos.tau=pos.tau y=y)
Dyn <- matrix(paste("theta[",1:T,"]",sep=""), T, 1)

```

pos.theta=pos

95.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  theta <- parm[Data$pos.theta]
  alpha <- parm[Data$pos.alpha]
  parm[Data$pos.phi] <- phi <- interval(parm[Data$pos.phi], -1, 1)
  mu <- parm[Data$pos.mu]
  parm[Data$pos.tau] <- tau <- interval(parm[Data$pos.tau], 1e-100, Inf)
  ### Log-Prior
  alpha.prior <- dnormv(alpha, mu, tau, log=TRUE)
  theta.prior <- sum(dnormv(theta[1], mu + phi*(alpha-mu), tau,
    log=TRUE), dnormv(theta[-1], mu + phi*(theta[-Data$T]-mu), tau,
    log=TRUE))
  phi.prior <- dunif(phi, -1, 1, log=TRUE)
  mu.prior <- dnormv(mu, 0, 10, log=TRUE)
  tau.prior <- dhalfcauchy(tau, 25, log=TRUE)
  ### Log-Likelihood
  beta <- exp(mu / 2)
  sigma2 <- 1 / exp(theta)
  LL <- sum(dnormv(Data$y, 0, sigma2, log=TRUE))
  ### Log-Posterior
  LP <- LL + alpha.prior + theta.prior + phi.prior + mu.prior +

```

```

    tau.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP, sigma2),
    yhat=rnormv(length(Data$y), 0, sigma2), parm=parm)
  return(Modelout)
}

```

95.4. Initial Values

```
Initial.Values <- c(rep(0,T), 0, 0, 0, 1)
```

96. Threshold Autoregression (TAR)

96.1. Form

$$\begin{aligned}
 \mathbf{y}_t &\sim \mathcal{N}(\nu_t, \sigma^2), \quad t = 1, \dots, T \\
 \nu_t &= \begin{cases} \alpha_1 + \phi_1 \mathbf{y}_{t-1}, & t = 1, \dots, T \quad \text{if } t \geq \theta \\ \alpha_2 + \phi_2 \mathbf{y}_{t-1}, & t = 1, \dots, T \quad \text{if } t < \theta \end{cases} \\
 \alpha_j &\sim \mathcal{N}(0, 1000) \in [-1, 1], \quad j = 1, \dots, 2 \\
 \phi_j &\sim \mathcal{N}(0, 1000), \in [-1, 1], \quad j = 1, \dots, 2 \\
 \theta &\sim \mathcal{U}(2, T - 1) \\
 \sigma &\sim \mathcal{HC}(25)
 \end{aligned}$$

96.2. Data

```

data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
T <- length(y)
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,2), phi=rep(0,2), theta=0,
  sigma=0))
pos.alpha <- grep("alpha", parm.names)
pos.phi <- grep("phi", parm.names)
pos.theta <- grep("theta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rtrunc(2, "norm", a=-1, b=1, mean=0, sd=1)
  phi <- rtrunc(2, "norm", a=-1, b=1, mean=0, sd=1)
  theta <- runif(1,2,Data$T-1)
  sigma <- runif(1)
  return(c(alpha, phi, theta, sigma))
}

```

```
MyData <- list(PGF=PGF, T=T, mon.names=mon.names, parm.names=parm.names,
  pos.alpha=pos.alpha, pos.phi=pos.phi, pos.theta=pos.theta,
  pos.sigma=pos.sigma, y=y)
```

96.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  parm[Data$pos.alpha] <- alpha <- interval(parm[Data$pos.alpha], -1, 1)
  parm[Data$pos.phi] <- phi <- interval(parm[Data$pos.phi], -1, 1)
  theta <- interval(parm[Data$pos.theta], 2, Data$T-1)
  parm[Data$pos.theta] <- theta
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- sum(dtrunc(alpha, "norm", a=-1, b=1, mean=0,
    sd=sqrt(1000), log=TRUE))
  phi.prior <- sum(dtrunc(phi, "norm", a=-1, b=1, mean=0,
    sd=sqrt(1000), log=TRUE))
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
  phi.prior <- sum(dnormv(phi, 0, 1000, log=TRUE))
  theta.prior <- dunif(theta, 2, Data$T-1, log=TRUE)
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- matrix(0, Data$T, 2)
  mu[,1] <- c(alpha[1], alpha[1] + phi[1]*Data$y[-Data$T])
  mu[,2] <- c(alpha[2], alpha[2] + phi[2]*Data$y[-Data$T])
  nu <- mu[,2]; temp <- which(1:Data$T < theta)
  nu[temp] <- mu[temp,1]
  LL <- sum(dnorm(Data$y[-1], nu[-1], sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + alpha.prior + phi.prior + theta.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(nu), nu, sigma), parm=parm)
  return(Modelout)
}
```

96.4. Initial Values

```
Initial.Values <- c(rep(0,4), T/2, 1)
```

97. Time Varying AR(1) with Chebyshev Series

This example consists of a first-order autoregressive model, AR(1), with a time-varying pa-

parameter (TVP) ϕ , that is a Chebyshev series constructed from a linear combination of orthonormal Chebyshev time polynomials (CTPs) and parameter vector β . The user creates basis matrix \mathbf{P} , specifying polynomial degree D and time T . Each column is a CTP of a different degree, and the first column is restricted to 1, the linear basis. CTPs are very flexible for TVPs, and estimate quickly because each is orthogonal, unlike simple polynomials and splines.

97.1. Form

$$\begin{aligned} \mathbf{y}_t &\sim \mathcal{N}(\mu_t, \sigma^2), \quad t = 1, \dots, T \\ \mu_t &= \alpha + \phi_{t-1} \mathbf{y}_{t-1} \\ \phi_t &= \mathbf{P}\beta \\ \alpha &\sim \mathcal{N}(0, 1000) \\ \beta_d &\sim \mathcal{N}(0, 1000), \quad d = 1, \dots, (D+1) \\ \sigma &\sim \mathcal{HC}(25) \end{aligned}$$

97.2. Data

```
data(demonfx)
y <- as.vector(diff(log(as.matrix(demonfx[1:261,1]))))
D <- 6 #Maximum degree of Chebyshev time polynomials
T <- length(y)
P <- matrix(1, T, D+1)
for (d in 1:D) {P[,d+1] <- sqrt(2)*cos(d*pi*(c(1:T)-0.5)/T)}
mon.names <- c("LP", "ynew", as.parm.names(list(phi=rep(0,T-1))))
parm.names <- as.parm.names(list(alpha=0, beta=rep(0,D+1), sigma=0))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(1)
  beta <- rnorm(Data$D+1)
  sigma <- runif(1)
  return(c(alpha, beta, sigma))
}
MyData <- list(D=D, P=P, PGF=PGF, T=T, mon.names=mon.names,
  parm.names=parm.names, pos.alpha=pos.alpha, pos.beta=pos.beta,
  pos.sigma=pos.sigma, y=y)
```

97.3. Model

```
Model <- function(parm, Data)
{
```

```

### Parameters
alpha <- parm[Data$pos.alpha]
beta <- parm[Data$pos.beta]
sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
parm[Data$pos.sigma] <- sigma
### Log-Prior
alpha.prior <- dnormv(alpha, 0, 1000, log=TRUE)
beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
### Log-Likelihood
phi <- tcrossprod(Data$P[-Data$T,], t(beta))
mu <- c(alpha, alpha + phi*Data$y[-Data$T])
ynew <- rnorm(1, alpha + tcrossprod(Data$P[Data$T,], t(beta))*
  Data$y[Data$T], sigma)
LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP,ynew,phi),
  yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

97.4. Initial Values

```
Initial.Values <- c(rep(0,D+2), 1)
```

98. Variable Selection, BAL

This approach to variable selection is one of several forms of the Bayesian Adaptive Lasso (BAL). The lasso applies shrinkage to exchangeable scale parameters, γ , for the regression effects, β .

98.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\
 \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \\
 \beta_1 &\sim \mathcal{L}(0, 1000) \\
 \beta_j &\sim \mathcal{L}(0, \gamma_j), \quad j = 2, \dots, J \\
 \gamma_j &\sim \mathcal{G}^{-1}(\delta, \tau), \quad \in [0, \infty] \\
 \delta &\sim \mathcal{HC}(25) \\
 \tau &\sim \mathcal{HC}(25) \\
 \sigma &\sim \mathcal{HC}(25)
 \end{aligned}$$

98.2. Data

```

data(demonsnacks)
J <- ncol(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(demonsnacks[,c(1,3:10)]))
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), gamma=rep(0,J-1), delta=0,
  tau=0, sigma=0))
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.delta <- grep("delta", parm.names)
pos.tau <- grep("tau", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  delta <- runif(1)
  tau <- runif(1)
  sigma <- runif(1)
  gamma <- rinvgamma(Data$J-1, delta, tau)
  beta <- rlaplace(Data$J, 0, c(1,gamma))
  return(c(beta, gamma, delta, tau, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.gamma=pos.gamma,
  pos.delta=pos.delta, pos.tau=pos.tau, pos.sigma=pos.sigma, y=y)

```

98.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperhyperparameters
  delta <- interval(parm[Data$pos.delta], 1e-100, Inf)
  parm[Data$pos.delta] <- delta
  parm[Data$pos.tau] <- tau <- interval(parm[Data$pos.tau], 1e-100, Inf)
  ### Hyperparameters
  gamma <- interval(parm[Data$pos.gamma], 1e-100, Inf)
  parm[Data$pos.gamma] <- gamma
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Hyperhyperprior
  delta.prior <- dhalfcauchy(delta, 25, log=TRUE)
  tau.prior <- dhalfcauchy(tau, 25, log=TRUE)
  ### Log-Hyperprior

```

```

gamma.prior <- sum(dinvgamma(gamma, delta, tau, log=TRUE))
### Log-Prior
beta.prior <- sum(dlaplace(beta, 0, c(1000, gamma), log=TRUE))
sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
### Log-Likelihood
mu <- tcrossprod(Data$X, t(beta))
LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior + gamma.prior + delta.prior + tau.prior +
      sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
                 yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

98.4. Initial Values

```
Initial.Values <- c(rep(0,J), rep(0,J-1), rep(1,3))
```

99. Variable Selection, Horseshoe

99.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\
 \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \\
 \beta_1 &\sim \mathcal{N}(0, 1000) \\
 \beta_j &\sim \mathcal{HS}(\lambda_j, \tau), \quad j = 2, \dots, J \\
 \lambda_j &\sim \mathcal{HC}(1), \quad j = 2, \dots, J \\
 \tau &\sim \mathcal{HC}(1) \\
 \sigma &\sim \mathcal{HC}(25)
 \end{aligned}$$

99.2. Data

```

data(demonsnacks)
J <- ncol(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(demonsnacks[,c(1,3:10)]))
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), lambda=rep(0,J-1),
                                tau=0, sigma=0))

```

```

pos.beta <- grep("beta", parm.names)
pos.lambda <- grep("lambda", parm.names)
pos.tau <- grep("tau", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  lambda <- runif(Data$J-1)
  tau <- runif(1)
  sigma <- runif(1)
  return(c(beta, lambda, tau, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.lambda=pos.lambda,
  pos.tau=pos.tau, pos.sigma=pos.sigma, y=y)

```

99.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  lambda <- interval(parm[Data$pos.lambda], 1e-100, Inf)
  parm[Data$pos.lambda] <- lambda
  parm[Data$pos.tau] <- tau <- interval(parm[Data$pos.tau], 1e-100, Inf)
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta[1], 0, 1000, log=TRUE),
    dhs(beta[-1], lambda, tau, log=TRUE))
  lambda.prior <- sum(dhalfcauchy(lambda, 1, log=TRUE))
  tau.prior <- dhalfcauchy(tau, 1, log=TRUE)
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

99.4. Initial Values

```

Initial.Values <- c(rep(0,J), rep(1,J-1), rep(1,2))

```

100. Variable Selection, LASSO

100.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\ \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \\ \beta_1 &\sim \mathcal{N}(0, 1000) \\ \beta_j &\sim \mathcal{LASSO}(0, \sigma, \tau, \lambda_j), \quad j = 2, \dots, J \end{aligned}$$

100.2. Data

```
data(demonsnacks)
J <- ncol(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(demonsnacks[,c(1,3:10)]))
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), sigma=0, tau=rep(0,J-1),
  lambda=0))
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
pos.tau <- grep("tau", parm.names)
pos.lambda <- grep("lambda", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  sigma <- runif(1)
  tau <- runif(Data$J-1)
  lambda <- runif(1)
  return(c(beta, sigma, tau, lambda))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma,
  pos.tau=pos.tau, pos.lambda=pos.lambda, y=y)
```

100.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  parm[Data$pos.tau] <- tau <- interval(parm[Data$pos.tau], 1e-100, Inf)
```

```

lambda <- interval(parm[Data$pos.lambda], 1e-100, Inf)
parm[Data$pos.lambda] <- lambda
### Log-Prior
beta.prior <- sum(dnorm(beta[1], 0, 1000, log=TRUE),
  dlasso(beta[-1], sigma, tau, lambda, log=TRUE))
### Log-Likelihood
mu <- tcrossprod(Data$X, t(beta))
LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
### Log-Posterior
LP <- LL + beta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(length(mu), mu, sigma), parm=parm)
return(Modelout)
}

```

100.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1, rep(1,J-1), 1)
```

101. Variable Selection, RJ

This example uses the RJ (Reversible-Jump) algorithm of the `LaplacesDemon` function for variable selection and Bayesian Model Averaging (BMA). Other MCMC algorithms will not perform variable selection with this example, as presented. This is an example of variable selection in a linear regression. The only difference between the following example, and the example of linear regression (48), is that RJ specifications are also included for the RJ algorithm, and that the RJ algorithm must be used.

101.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\
 \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \sigma &\sim \mathcal{HC}(25)
 \end{aligned}$$

101.2. Data

```

N <- 1000
J <- 100 #Number of predictors, including the intercept
X <- matrix(1,N,J)
for (j in 2:J) {X[,j] <- rnorm(N,runif(1,-3,3),runif(1,0.1,1))}
beta.orig <- runif(J,-3,3)
zero <- sample(2:J, round(J*0.9)) #Assign most parameters to be zero

```

```

beta.orig[zero] <- 0
e <- rnorm(N,0,0.1)
y <- as.vector(tcrossprod(beta.orig, X) + e)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), sigma=0))
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  sigma <- runif(1)
  return(c(beta, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma, y=y)
### Reversible-Jump Specifications bin.n <- J-1 #Maximum allowable model size
bin.p <- 0.4 #Most probable size: bin.p x bin.n is binomial mean and median
parm.p <- rep(1/J,J+1)
selectable=c(0, rep(1,J-1), 0)

```

101.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  LL <- sum(dnorm(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

101.4. Initial Values

```

Initial.Values <- GIV(Model, MyData, PGF=TRUE)

```

102. Variable Selection, SSVS

This example uses a modified form of the random-effects (or global adaptation) Stochastic Search Variable Selection (SSVS) algorithm presented in [O’Hara and Sillanpaa \(2009\)](#), which selects variables according to practical significance rather than statistical significance. Here, SSVS is applied to linear regression, though this method is widely applicable. For J variables, each regression effect β_j is conditional on γ_j , a binary inclusion variable. Each β_j is a discrete mixture distribution with respect to $\gamma_j = 0$ or $\gamma_j = 1$, with precision 100 or $\beta_\sigma = 0.1$, respectively. As with other representations of SSVS, these precisions may require tuning.

The binary inclusion variables are discrete parameters, and discrete parameters are not supported in all algorithms. The example below is updated with the Slice sampler.

When the goal is to select the best model, each $\mathbf{X}_{1:N,j}$ is retained for a future run when the posterior mean of $\gamma_j \geq 0.5$. When the goal is model-averaging, the results of this model may be used directly, which would please L. J. Savage, who said that “models should be as big as an elephant” ([Draper 1995](#)).

102.1. Form

$$\begin{aligned} \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\ \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \\ \beta_1 &\sim \mathcal{N}(0, 1000) \\ (\beta_j | \gamma_j) &\sim (1 - \gamma_j)\mathcal{N}(0, 0.01) + \gamma_j\mathcal{N}(0, \beta_\sigma^2) \quad j = 2, \dots, J \\ \beta_\sigma &\sim \mathcal{HC}(25) \\ \gamma_j &\sim \mathcal{BERN}(1/(J - 1)), \quad j = 1, \dots, (J - 1) \\ \sigma &\sim \mathcal{HC}(25) \end{aligned}$$

102.2. Data

```
data(demonsnacks)
J <- ncol(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(demonsnacks[,c(1,3:10)]))
for (j in 2:J) X[,j] <- CenterScale(X[,j])
mon.names <- c("LP", "min.beta.sigma")
parm.names <- as.parm.names(list(beta=rep(0,J), gamma=rep(0,J-1),
  b.sd=0, sigma=0))
pos.beta <- grep("beta", parm.names)
pos.gamma <- grep("gamma", parm.names)
pos.b.sd <- grep("b.sd", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
  gamma <- rep(1,Data$J-1)
```

```

b.sd <- rnorm(1)
sigma <- runif(1)
return(c(beta, gamma, b.sd, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.gamma=pos.gamma,
  pos.b.sd=pos.b.sd, pos.sigma=pos.sigma, y=y)

```

102.3. Model

```

Model <- function(parm, Data)
{
  ### Hyperparameters
  beta.sigma <- interval(parm[Data$pos.b.sd], 1e-100, Inf)
  parm[Data$pos.b.sd] <- beta.sigma
  ### Parameters
  beta <- parm[Data$pos.beta]
  gamma <- parm[Data$pos.gamma]
  beta.sigma <- rep(beta.sigma, Data$J-1)
  beta.sigma[gamma == 0] <- 0.1
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  ### Log-Hyperprior
  beta.sigma.prior <- sum(dhalfcauchy(beta.sigma, 25, log=TRUE))
  ### Log-Prior
  beta.prior <- sum(dnorm(beta, 0, c(sqrt(1000), beta.sigma, log=TRUE)))
  gamma.prior <- sum(dbern(gamma, 1/(Data$J-1), log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta*c(1,gamma)))
  LL <- sum(dnorm(y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + beta.sigma.prior + gamma.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=c(LP, min(beta.sigma)),
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

102.4. Initial Values

```
Initial.Values <- c(rep(0,J), rep(1,J-1), rep(1,2))
```

103. VARMA(p,q) - SSVS

Stochastic search variable selection (SSVS) is applied to VARMA parameters. Note that the constants for the mixture variances are typically multiplied by the posterior standard

deviations from an unrestricted VARMA that was updated previously, and these are not included in this example. Since an unrestricted VARMA model may be difficult to identify, this should be performed only on the AR parameters.

103.1. Form

$$\mathbf{Y}_{t,j} \sim \mathcal{N}(\mu_{t,j}, \sigma_j^2), \quad t = 1, \dots, T, \quad j = 1, \dots, J$$

$$\mu_{t,j} = \alpha_j + \sum_{p=1}^P \Gamma_{1:J,j,p}^{\Phi} \Phi_{1:J,j,p} \mathbf{Y}_{t-p,j} + \sum_{q=1}^Q \Gamma_{1:J,j,q}^{\Theta} \Theta_{1:J,j,q} \epsilon_{t-q,j}$$

$$\alpha_j \sim \mathcal{N}(0, 1000)$$

$$\Gamma_{i,k,p}^{\Phi} \sim \mathcal{BERN}(0.5), \quad i = 1, \dots, J, \quad k = 1, \dots, J, \quad p = 1, \dots, P$$

$$(\Phi_{i,k,p} | \Gamma_{i,k,p}^{\Phi}) \sim (1 - \Gamma_{i,k,p}^{\Phi}) \mathcal{N}(0, 0.01) + \Gamma_{i,k,p}^{\Phi} \mathcal{N}(0, 10), \quad i = 1, \dots, J, \quad k = 1, \dots, J, \quad p = 1, \dots, P$$

$$\Gamma_{i,k,q}^{\Theta} \sim \mathcal{BERN}(0.5), \quad i = 1, \dots, J, \quad k = 1, \dots, J, \quad q = 1, \dots, Q$$

$$(\Theta_{i,k,q} | \Gamma_{i,k,q}^{\Theta}) \sim (1 - \Gamma_{i,k,q}^{\Theta}) \mathcal{N}(0, 0.01) + \Gamma_{i,k,q}^{\Theta} \mathcal{N}(0, 10), \quad i = 1, \dots, J, \quad k = 1, \dots, J, \quad q = 1, \dots, Q$$

$$\sigma_j \sim \mathcal{HC}(25)$$

103.2. Data

```
data(demonfx)
Y.orig <- as.matrix(demonfx[,1:3])
Y <- diff(log(Y.orig[1:100,]))
Y.scales <- sqrt(.colVars(Y))
Y <- Y / matrix(Y.scales, nrow(Y), ncol(Y), byrow=TRUE)
T <- nrow(Y)
J <- ncol(Y)
L.P <- c(1,5,20) #Autoregressive lags
L.Q <- c(1,2) #Moving average lags
P <- length(L.P) #Autoregressive order
Q <- length(L.Q) #Moving average order
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,J),
  Gamma.phi=array(0, dim=c(J,J,P)), Phi=array(0, dim=c(J,J,P)),
  Gamma.theta=array(0, dim=c(J,J,Q)), Theta=array(0, dim=c(J,J,Q)),
  sigma=rep(0,J)))
pos.alpha <- grep("alpha", parm.names)
pos.Gamma.phi <- grep("Gamma.phi", parm.names)
pos.Phi <- grep("Phi", parm.names)
pos.Gamma.theta <- grep("Gamma.theta", parm.names)
pos.Theta <- grep("Theta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
```

```

alpha <- rnorm(Data$J)
Gamma.phi <- rep(1, Data$J*Data$J*Data$P)
Phi <- runif(Data$J*Data$J*Data$P, -1, 1)
Gamma.theta <- rep(1, Data$J*Data$J*Data$Q)
Theta <- rnorm(Data$J*Data$J*Data$Q)
sigma <- runif(Data$J)
return(c(alpha, Gamma.phi, Phi, Gamma.theta, Theta, sigma))      }
MyData <- list(J=J, L.P=L.P, L.Q=L.Q, P=P, Q=Q, PGF=PGF, T=T, Y=Y,
  mon.names=mon.names, parm.names=parm.names, pos.alpha=pos.alpha,
  pos.Gamma.phi=pos.Gamma.phi, pos.Phi=pos.Phi,
  pos.Gamma.theta=pos.Gamma.theta, pos.Theta=pos.Theta,
  pos.sigma=pos.sigma)

```

103.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  Gamma.phi <- array(parm[Data$pos.Gamma.phi],
    dim=c(Data$J, Data$J, Data$P))
  Phi.Sigma <- Gamma.phi * 10
  Phi.Sigma[Gamma.phi == 0] <- 0.1
  Phi <- array(parm[Data$pos.Phi], dim=c(Data$J, Data$J, Data$P))
  Gamma.theta <- array(parm[Data$pos.Gamma.theta],
    dim=c(Data$J, Data$J, Data$Q))
  Theta.Sigma <- Gamma.theta * 10
  Theta.Sigma[Gamma.theta == 0] <- 0.1
  Theta <- array(parm[Data$pos.Theta], dim=c(Data$J, Data$J, Data$Q))
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
  Gamma.phi.prior <- sum(dbern(Gamma.phi, 0.5, log=TRUE))
  Phi.prior <- sum(dnorm(Phi, 0, Phi.Sigma, log=TRUE))
  Gamma.theta.prior <- sum(dbern(Gamma.theta, 0.5, log=TRUE))
  Theta.prior <- sum(dnorm(Theta, 0, Theta.Sigma, log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  mu <- matrix(alpha, Data$T, Data$J, byrow=TRUE)
  for (p in 1:Data$P)
    mu[(1+Data$L.P[p]):Data$T,] <- mu[(1+Data$L.P[p]):Data$T,] +
      Data$Y[1:(Data$T-Data$L.P[p]),] %*%
      (Gamma.phi[, , p] * Phi[, , p])
  epsilon <- Data$Y - mu
  for (q in 1:Data$Q)

```

```

mu[(1+Data$L.Q[q]):Data$T,] <- mu[(1+Data$L.Q[q]):Data$T,] +
  epsilon[1:(Data$T-Data$L.Q[q]),] %*%
  (Gamma.theta[, , q] * Theta[, , q])
Sigma <- matrix(sigma, Data$T, Data$J, byrow=TRUE)
LL <- sum(dnorm(Data$Y[(1+Data$L.P[Data$P]):Data$T,],
  mu[(1+Data$L.P[Data$P]):Data$T,],
  Sigma[(1+Data$L.P[Data$P]):Data$T,], log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + Gamma.phi.prior + Phi.prior +
Gamma.theta.prior + Theta.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(prod(dim(mu)), mu, Sigma), parm=parm)
return(Modelout)
}

```

103.4. Initial Values

```

Initial.Values <- c(colMeans(Y), rep(1,J*J*P), runif(J*J*P,-1,1),
  rep(1,J*J*Q), rep(0,J*J*Q), rep(1,J))

```

104. VAR(p)-GARCH(1,1)-M

The Minnesota prior is applied to the VAR parameters, and the multivariate GARCH component is estimated with asymmetric BEKK. Compared to VAR(p) or VARMA(p,q), this is computationally intensive. However, it also tends to result in a substantial improvement when time for computation is feasible. This model also performs well when SSVS is applied to all parameters except \mathbf{C} , though it is even more computationally intensive, and is not shown here.

104.1. Form

$$\mathbf{Y}_{t,1:J} \sim \mathcal{N}_J(\mu_{t,1:J}, \mathbf{H}_{1:J,1:J,t})$$

$$\mu_{t,j} = \alpha_j + \sum_{p=1}^P \Phi_{1:J,j,p} \mathbf{Y}_{t-p,j} + \sum \mathbf{H}_{1:J,j,t-1} \delta_{1:J,j}$$

$$\mathbf{H}_{:,t} = \Omega + \mathbf{A}^T \epsilon_{t-1} \epsilon_{t-1}^T \mathbf{A} + \mathbf{B}^T \mathbf{H}_{:,t-1} \mathbf{B} + \mathbf{D}^T \zeta_{t-1} \zeta_{t-1}^T \mathbf{D}, \quad t = 2, \dots, T$$

$$\Omega = \mathbf{C} \mathbf{C}^T$$

$$\alpha_j \sim \mathcal{N}(0, 1000)$$

$$\delta_{i,k} \sim \mathcal{N}(0, 1000), \quad i = 1, \dots, J, \quad k = 1, \dots, J$$

$$\Phi_{i,k,p} \sim \mathcal{N}(\Phi_{i,k,p}^\mu, \Sigma_{i,k,p}), \quad i = 1, \dots, J, \quad k = 1, \dots, J, \quad p = 1, \dots, P$$

$$\mathbf{C}_{i,j} \sim \mathcal{N}(0, 100)$$

$$\mathbf{A}_{i,j} \sim \mathcal{N}(0, 100)$$

$$\mathbf{B}_{i,j} \sim \mathcal{N}(0, 100)$$

$$\mathbf{D}_{i,j} \sim \mathcal{N}(0, 100)$$

where Φ has a Minnesota prior, \mathbf{C} is lower-triangular with positive-only diagonal elements, and $\mathbf{A}_{1,1}$, $\mathbf{B}_{1,1}$, and $\mathbf{D}_{1,1}$ must be positive.

104.2. Data

```

data(demonfx)
Y.orig <- as.matrix(demonfx[,1:3])
Y <- diff(log(Y.orig[1:100,]))
Y.scales <- sqrt(.colVars(Y))
Y <- Y / matrix(Y.scales, nrow(Y), ncol(Y), byrow=TRUE)
T <- nrow(Y)
J <- ncol(Y)
L <- c(1,5,20) #Autoregressive lags
P <- length(L) #Autoregressive order
Phi.mu <- array(0, dim=c(J,J,P))
Phi.mu[, , 1] <- diag(J)
C <- matrix(NA, J, J)
C[lower.tri(C, diag=TRUE)] <- 0
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,J), delta=matrix(0,J,J),
  Phi=array(0, dim=c(J,J,P)), C=C, A=matrix(0,J,J), B=matrix(0,J,J),
  D=matrix(0,J,J)))
pos.alpha <- grep("alpha", parm.names)
pos.delta <- grep("delta", parm.names)
pos.Phi <- grep("Phi", parm.names)
pos.C <- grep("C", parm.names)
pos.A <- grep("A", parm.names)
pos.B <- grep("B", parm.names)
pos.D <- grep("D", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(Data$J)
  delta <- rnorm(Data$J*Data$J)
  Phi <- runif(Data$J*Data$J*Data$P, -1, 1)
  C <- runif(Data$J*(Data$J+1)/2)
  A <- as.vector(diag(Data$J)) + runif(Data$J*Data$J, -0.1, 0.1)
  B <- as.vector(diag(Data$J)) + runif(Data$J*Data$J, -0.1, 0.1)
  D <- as.vector(diag(Data$J)) + runif(Data$J*Data$J, -0.1, 0.1)
  return(c(alpha, delta, Phi, C, A, B, D))
}
MyData <- list(J=J, L=L, P=P, PGF=PGF, Phi.mu=Phi.mu, T=T, Y=Y,
  mon.names=mon.names, parm.names=parm.names, pos.alpha=pos.alpha,
  pos.delta=pos.delta, pos.Phi=pos.Phi, pos.C=pos.C, pos.A=pos.A,

```

```
pos.B=pos.B, pos.D=pos.D)
```

104.3. Model

```
Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  delta <- matrix(parm[Data$pos.delta], Data$J, Data$J)
  Phi <- array(parm[Data$pos.Phi], dim=c(Data$J, Data$J, Data$P))
  C <- matrix(0, Data$J, Data$J)
  C[lower.tri(C, diag=TRUE)] <- parm[Data$pos.C]
  diag(C) <- abs(diag(C))
  parm[Data$pos.C] <- C[lower.tri(C, diag=TRUE)]
  Omega <- C %*% t(C)
  A <- matrix(parm[Data$pos.A], Data$J, Data$J)
  A[1,1] <- abs(A[1,1])
  parm[Data$pos.A] <- as.vector(A)
  B <- matrix(parm[Data$pos.B], Data$J, Data$J)
  B[1,1] <- abs(B[1,1])
  parm[Data$pos.B] <- as.vector(B)
  D <- matrix(parm[Data$pos.D], Data$J, Data$J)
  D[1,1] <- abs(D[1,1])
  parm[Data$pos.D] <- as.vector(D)
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
  delta.prior <- sum(dnormv(delta, 0, 1000, log=TRUE))
  Sigma <- MinnesotaPrior(Data$J, lags=Data$L, lambda=1,
    theta=0.5, sqrt(diag(Omega)))
  Phi.prior <- sum(dnormv(Phi, Data$Phi.mu, Sigma, log=TRUE))
  C.prior <- sum(dnormv(C[lower.tri(C, diag=TRUE)], 0, 100, log=TRUE))
  A.prior <- sum(dnormv(A, 0, 100, log=TRUE))
  B.prior <- sum(dnormv(B, 0, 100, log=TRUE))
  D.prior <- sum(dnormv(D, 0, 100, log=TRUE))
  ### Log-Likelihood
  mu <- matrix(alpha, Data$T, Data$J, byrow=TRUE)
  for (p in 1:Data$P)
    mu[(1+Data$L[p]):Data$T,] <- mu[(1+Data$L[p]):Data$T,] +
      Data$Y[1:(Data$T-Data$L[p]),] %*% Phi[, , p]
  LL <- 0
  Yhat <- Data$Y
  H <- array(Omega, dim=c(Data$J, Data$J, Data$T))
  for (t in 2:Data$T) {
    eps <- Data$Y - mu
    zeta <- matrix(interval(eps, -Inf, 0, reflect=FALSE), Data$T,
      Data$J)
```

```

part1 <- t(A) %*% eps[t-1,] %*% t(eps[t-1,]) %*% A
part2 <- t(B) %*% H[, , t-1] %*% B
part3 <- t(D) %*% zeta[t-1,] %*% t(zeta[t-1,]) %*% D
H0 <- Omega + part1 + part2 + part3
H0[upper.tri(H0, diag=TRUE)] <- t(H0)[upper.tri(H0, diag=TRUE)]
H[, , t] <- H0
mu[t-1,] <- mu[t-1,] + colMeans(H[, , t-1]*delta)
Sigma <- MinnesotaPrior(Data$J, lags=Data$L, lambda=1,
  theta=0.5, sqrt(diag(H[, , t])))
Phi.prior <- Phi.prior + sum(dnormv(Phi, Data$Phi.mu, Sigma,
  log=TRUE))
LL <- LL + dmvn(Y[t,], mu[t,], H[, , t], log=TRUE)
Yhat[t,] <- rmvn(1, mu[t,], H[, , t])
}

Phi.prior <- Phi.prior / Data$T
### Log-Posterior
LP <- LL + alpha.prior + delta.prior + Phi.prior + C.prior +
  A.prior + B.prior + D.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP, yhat=Yhat, parm=parm)
return(Modelout)
}

```

104.4. Initial Values

```

Initial.Values <- c(colMeans(Y), rnorm(J*J), runif(J*J*P,-1,1),
  runif(J*(J+1)/2), as.vector(diag(J)), as.vector(diag(J)),
  as.vector(diag(J)))

```

105. VAR(p) - Minnesota Prior

This is an example of a vector autoregression or VAR with P lags that uses the Minnesota prior to estimate Σ .

105.1. Form

$$\mathbf{Y}_{t,j} \sim \mathcal{N}(\mu_{t,j}, \sigma_j^2), \quad t = 1, \dots, T, \quad j = 1, \dots, J$$

$$\mu_{t,j} = \alpha_j + \sum_{p=1}^P \Phi_{1:J,j,p} \mathbf{Y}_{t-p,j}$$

$$\mathbf{y}_j^{new} = \alpha_j + \Phi_{1:J,j} \mathbf{Y}_{T,j}$$

$$\alpha_j \sim \mathcal{N}(0, 1000)$$

$$\Phi_{i,k,p} \sim \mathcal{N}(\Phi_{i,k,p}^\mu, \Sigma_{i,k,p}), \quad i = 1, \dots, J, \quad k = 1, \dots, J, \quad p = 1, \dots, P$$

$$\sigma_j \sim \mathcal{HC}(25)$$

where Φ^μ and Σ are set according to the Minnesota prior.

105.2. Data

```

data(demonfx)
Y.orig <- as.matrix(demonfx[,1:3])
Y <- diff(log(Y.orig[1:100,]))
Y.scales <- sqrt(.colVars(Y))
Y <- Y / matrix(Y.scales, nrow(Y), ncol(Y), byrow=TRUE)
T <- nrow(Y)
J <- ncol(Y)
L <- c(1,5,20) #Autoregressive lags
P <- length(L) #Autoregressive order
Phi.mu <- array(0, dim=c(J,J,P))
Phi.mu[, , 1] <- diag(J)
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,J),
  Phi=array(0, dim=c(J,J,P)), sigma=rep(0,J)))
pos.alpha <- grep("alpha", parm.names)
pos.Phi <- grep("Phi", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(Data$J)
  Phi <- runif(Data$J*Data$J*Data$P, -1, 1)
  sigma <- runif(Data$J)
  return(c(alpha, Phi, sigma))
}
MyData <- list(J=J, L=L, P=P, PGF=PGF, Phi.mu=Phi.mu, T=T, Y=Y,
  mon.names=mon.names, parm.names=parm.names, pos.alpha=pos.alpha,
  pos.Phi=pos.Phi, pos.sigma=pos.sigma)

```

105.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  Phi <- array(parm[Data$pos.Phi], dim=c(Data$J, Data$J, Data$P))
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
  Sigma <- MinnesotaPrior(Data$J, lags=Data$L, lambda=1, theta=0.5,
    sigma)
  Phi.prior <- sum(dnormv(Phi, Data$Phi.mu, Sigma, log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))

```

```

### Log-Likelihood
mu <- matrix(alpha, Data$T, Data$J, byrow=TRUE)
for (p in 1:Data$P) {
  mu[(1+Data$L[p]):Data$T,] <- mu[(1+Data$L[p]):Data$T,] +
  Data$Y[1:(Data$T-Data$L[p]),] %% Phi[ , , p]}
Sigma <- matrix(sigma, Data$T, Data$J, byrow=TRUE)
LL <- sum(dnorm(Data$Y[(1+Data$L[Data$P]):Data$T,],
  mu[(1+Data$L[Data$P]):Data$T,],
  Sigma[(1+Data$L[Data$P]):Data$T,], log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + Phi.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
  yhat=rnorm(prod(dim(mu)), mu, Sigma), parm=parm)
return(Modelout)
}

```

105.4. Initial Values

```
Initial.Values <- c(as.vector(colMeans(Y)), rep(0,J*J*P), rep(1,J))
```

106. VAR(p) - SSVS

Stochastic search variable selection (SSVS) is applied to VAR autoregressive parameters. Note that the constants for the mixture variances are typically multiplied by the posterior standard deviations from an unrestricted VAR that was updated previously, and these are not included in this example.

106.1. Form

$$\mathbf{Y}_{t,j} \sim \mathcal{N}(\mu_{t,j}, \sigma_j^2), \quad t = 1, \dots, T, \quad j = 1, \dots, J$$

$$\mu_{t,j} = \alpha_j + \sum_{p=1}^P \Gamma_{1:J,j,p} \Phi_{1:J,j,p} \mathbf{Y}_{t-p,j}$$

$$\alpha_j \sim \mathcal{N}(0, 1000)$$

$$\Gamma_{i,k,p} \sim \mathcal{BERN}(0.5), \quad i = 1, \dots, J, \quad k = 1, \dots, J, \quad p = 1, \dots, P$$

$$(\Phi_{i,k,p} | \Gamma_{i,k,p}) \sim (1 - \Gamma_{i,k,p}) \mathcal{N}(0, 0.01) + \Gamma_{i,k,p} \mathcal{N}(0, 10), \quad i = 1, \dots, J, \quad k = 1, \dots, J, \quad p = 1, \dots, P$$

$$\sigma_j \sim \mathcal{HC}(25)$$

106.2. Data

```

data(demonfx)
Y.orig <- as.matrix(demonfx[,1:3])
Y <- diff(log(Y.orig[1:100,]))

```



```

Y.scales <- sqrt(.colVars(Y))
Y <- Y / matrix(Y.scales, nrow(Y), ncol(Y), byrow=TRUE)
T <- nrow(Y)
J <- ncol(Y)
L <- c(1,5,20) #Autoregressive lags
P <- length(L) #Autoregressive order
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,J),
Gamma=array(0, dim=c(J,J,P)), Phi=array(0, dim=c(J,J,P)),
sigma=rep(0,J)))
pos.alpha <- grep("alpha", parm.names)
pos.Gamma <- grep("Gamma", parm.names)
pos.Phi <- grep("Phi", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(Data$J)
  Gamma <- rep(1, Data$J*Data$J*Data$P)
  Phi <- runif(Data$J*Data$J*Data$P, -1, 1)
  sigma <- runif(Data$J)
  return(c(alpha, Gamma, Phi, sigma))
} MyData <- list(J=J, L=L, P=P, PGF=PGF, T=T, Y=Y, mon.names=mon.names,
parm.names=parm.names, pos.alpha=pos.alpha, pos.Gamma=pos.Gamma,
pos.Phi=pos.Phi, pos.sigma=pos.sigma)

```

106.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  alpha <- parm[Data$pos.alpha]
  Gamma <- array(parm[Data$pos.Gamma], dim=c(Data$J, Data$J, Data$P))
  Phi.Sigma <- Gamma * 10
  Phi.Sigma[Gamma == 0] <- 0.1
  Phi <- array(parm[Data$pos.Phi], dim=c(Data$J, Data$J, Data$P))
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 1000, log=TRUE))
  Gamma.prior <- sum(dbern(Gamma, 0.5, log=TRUE))
  Phi.prior <- sum(dnorm(Phi, 0, Phi.Sigma, log=TRUE))
  sigma.prior <- sum(dhalfcauchy(sigma, 25, log=TRUE))
  ### Log-Likelihood
  mu <- matrix(alpha, Data$T, Data$J, byrow=TRUE)
  for (p in 1:Data$P)
    mu[(1+Data$L[p]):Data$T,] <- mu[(1+Data$L[p]):Data$T,] +
      Data$Y[1:(Data$T-Data$L[p]),] %*% (Gamma[, , p]*Phi[, , p])

```

```

Sigma <- matrix(sigma, Data$T, Data$J, byrow=TRUE)
LL <- sum(dnorm(Data$Y[(1+Data$L[Data$P]):Data$T,],
             mu[(1+Data$L[Data$P]):Data$T,],
             Sigma[(1+Data$L[Data$P]):Data$T,], log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + Gamma.prior + Phi.prior + sigma.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
                yhat=rnorm(prod(dim(mu)), mu, Sigma), parm=parm)
return(Modelout)
}

```

106.4. Initial Values

```
Initial.Values <- c(colMeans(Y), rep(1,J*J*P), runif(J*J*P,-1,1), rep(1,J))
```

107. Weighted Regression

It is easy enough to apply record-level weights to the likelihood. Here, weights are applied to the linear regression example in section 48.

107.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2) \\
 \boldsymbol{\mu} &= \mathbf{X}\boldsymbol{\beta} \\
 \beta_j &\sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J \\
 \sigma &\sim \mathcal{HC}(25)
 \end{aligned}$$

107.2. Data

```

data(demonsnacks)
N <- nrow(demonsnacks)
J <- ncol(demonsnacks)
y <- log(demonsnacks$Calories)
X <- cbind(1, as.matrix(demonsnacks[,c(1,3:10)]))
for (j in 2:J) X[,j] <- CenterScale(X[,j])
w <- c(rep(1,5), 0.2, 1, 0.01, rep(1,31))
w <- w * (sum(w) / N)
mon.names <- "LP"
parm.names <- as.parm.names(list(beta=rep(0,J), sigma=0))
pos.beta <- grep("beta", parm.names)
pos.sigma <- grep("sigma", parm.names)
PGF <- function(Data) {
  beta <- rnorm(Data$J)
}

```

```

sigma <- runif(1)
return(c(beta, sigma))
}
MyData <- list(J=J, PGF=PGF, X=X, mon.names=mon.names,
  parm.names=parm.names, pos.beta=pos.beta, pos.sigma=pos.sigma, w=w,
  y=y)

```

107.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  beta <- parm[Data$pos.beta]
  sigma <- interval(parm[Data$pos.sigma], 1e-100, Inf)
  parm[Data$pos.sigma] <- sigma
  ### Log-Prior
  beta.prior <- sum(dnormv(beta, 0, 1000, log=TRUE))
  sigma.prior <- dhalfcauchy(sigma, 25, log=TRUE)
  ### Log-Likelihood
  mu <- tcrossprod(Data$X, t(beta))
  LL <- sum(w * dnorm(Data$y, mu, sigma, log=TRUE))
  ### Log-Posterior
  LP <- LL + beta.prior + sigma.prior
  Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
    yhat=rnorm(length(mu), mu, sigma), parm=parm)
  return(Modelout)
}

```

107.4. Initial Values

```
Initial.Values <- c(rep(0,J), 1)
```

108. Zero-Inflated Poisson (ZIP)

108.1. Form

$$\begin{aligned}
 \mathbf{y} &\sim \mathcal{P}(\Lambda_{1:N,2}) \\
 \mathbf{z} &\sim \mathcal{BERN}(\Lambda_{1:N,1}) \\
 z_i &= \begin{cases} 1 & \text{if } y_i = 0 \\ 0 & \text{otherwise} \end{cases} \\
 \Lambda_{i,2} &= \begin{cases} 0 & \text{if } \Lambda_{i,1} \geq 0.5 \\ \Lambda_{i,2} & \text{otherwise} \end{cases}
 \end{aligned}$$

$$\Lambda_{1:N,1} = \frac{1}{1 + \exp(-\mathbf{X}_1\alpha)}$$

$$\Lambda_{1:N,2} = \exp(\mathbf{X}_2\beta)$$

$$\alpha_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J_1$$

$$\beta_j \sim \mathcal{N}(0, 1000), \quad j = 1, \dots, J_2$$

108.2. Data

```

N <- 1000
J1 <- 4
J2 <- 3
X1 <- matrix(runif(N*J1,-2,2),N,J1); X1[,1] <- 1
X2 <- matrix(runif(N*J2,-2,2),N,J2); X2[,1] <- 1
alpha <- runif(J1,-1,1)
beta <- runif(J2,-1,1)
p <- invlogit(tcrossprod(X1, t(alpha)) + rnorm(N,0,0.1))
mu <- round(exp(tcrossprod(X2, t(beta)) + rnorm(N,0,0.1)))
y <- ifelse(p > 0.5, 0, mu)
z <- ifelse(y == 0, 1, 0)
mon.names <- "LP"
parm.names <- as.parm.names(list(alpha=rep(0,J1), beta=rep(0,J2)))
pos.alpha <- grep("alpha", parm.names)
pos.beta <- grep("beta", parm.names)
PGF <- function(Data) {
  alpha <- rnorm(Data$J1)
  beta <- rnorm(Data$J2)
  return(c(alpha, beta))
}
MyData <- list(J1=J1, J2=J2, N=N, PGF=PGF, X1=X1, X2=X2,
  mon.names=mon.names, parm.names=parm.names, pos.alpha=pos.alpha,
  pos.beta=pos.beta, y=y, z=z)

```

108.3. Model

```

Model <- function(parm, Data)
{
  ### Parameters
  parm[Data$pos.alpha] <- alpha <- interval(parm[Data$pos.alpha], -5, 5)
  parm[Data$pos.beta] <- beta <- interval(parm[Data$pos.beta], -5, 5)
  ### Log-Prior
  alpha.prior <- sum(dnormv(alpha, 0, 5, log=TRUE))
  beta.prior <- sum(dnormv(beta, 0, 5, log=TRUE))
  ### Log-Likelihood
  Lambda <- matrix(NA, Data$N, 2)
  Lambda[,1] <- invlogit(tcrossprod(Data$X1, t(alpha)))

```

```

Lambda[,2] <- exp(tcrossprod(Data$X2, t(beta))) + 1e-100
Lambda[which(Lambda[,1] >= 0.5),2] <- 0
LL <- sum(dbern(Data$z, Lambda[,1], log=TRUE),
         dpois(Data$y, Lambda[,2], log=TRUE))
### Log-Posterior
LP <- LL + alpha.prior + beta.prior
Modelout <- list(LP=LP, Dev=-2*LL, Monitor=LP,
               yhat=rpois(nrow(Lambda), Lambda[,2]), parm=parm)
return(Modelout)
}

```

108.4. Initial Values

```
Initial.Values <- GIV(Model, MyData, n=10000)
```

References

- Congdon P (2003). *Applied Bayesian Modelling*. John Wiley & Sons, West Sussex, England.
- Draper D (1995). “Assessment and Propagation of Model Uncertainty.” *Journal of the Royal Statistical Society*, **B 57**(1), 45–97.
- Gelman A (2013). **R2WinBUGS**: *Running WinBUGS and OpenBUGS from R / S-PLUS*. R package version 2.1-19, URL <http://cran.r-project.org/web/packages/R2WinBUGS/index.html>.
- Gelman A, Carlin J, Stern H, Rubin D (2004). *Bayesian Data Analysis*. 2nd edition. Chapman & Hall, Boca Raton, FL.
- Ibrahim J, Chen M (2000). “Power Prior Distributions for Regression Models.” *Statistical Science*, **15**, 46–60.
- Kim J, Allenby G, Rossi P (2002). “Modeling Consumer Demand for Variety.” *Marketing Science*, **21**(3), 229–250.
- Kleine L (1950). *Economic Fluctuations in the United States 1921-1940*. John Wiley & Sons, New York, New York.
- Kotz S, Kozubowski T, Podgorski K (2001). *The Laplace Distribution and Generalizations: A Revisit with Applications to Communications, Economics, Engineering, and Finance*. Birkhauser, Boston.
- O’Hara R, Sillanpaa M (2009). “A Review of Bayesian Variable Selection Methods: What, How and Which.” *Journal of Bayesian Analysis*, **4**(1), 85–118.
- Spiegelhalter D, Thomas A, Best N, Lunn D (2003). *WinBUGS User Manual, Version 1.4*. MRC Biostatistics Unit, Institute of Public Health and Department of Epidemiology and Public Health, Imperial College School of Medicine, UK.

Statisticat LLC (2015). **LaplacesDemon**: *Complete Environment for Bayesian Inference*. R package version 15.03.19, URL <http://www.bayesian-inference.com/software>.

Zellner A (1962). “An Efficient Method of Estimating Seemingly Unrelated Regression Equations and Tests for Aggregation Bias.” *Journal of the American Statistical Association*, **57**, 348–368.

Affiliation:

Statisticat, LLC

Farmington, CT

E-mail: defunct

URL: <https://web.archive.org/web/20141224051720/http://www.bayesian-inference.com/index>