



# Trends and drivers in global surface ocean pH over the past 3 decades

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**Abstract.** We report global long-term trends in surface ocean pH using a new pH data set computed by combining  $f\text{CO}_2$  observations from the Surface Ocean  $\text{CO}_2$  Atlas (SOCAT) version 2 with surface alkalinity estimates based on temperature and salinity. Trends were determined over the periods 1981–2011 and 1991–2011 for a set of 17 biomes using a weighted linear least squares method. We observe significant decreases in surface ocean pH in  $\sim 70\%$  of all biomes and a mean rate of decrease of  $0.0018 \pm 0.0004 \text{ yr}^{-1}$  for 1991–2011. We are not able to calculate a global trend for 1981–2011 because too few biomes have enough data for this. In half the biomes, the rate of change is commensurate with the trends expected based on the assumption that the surface ocean pH change is only driven by the surface ocean  $\text{CO}_2$  chemistry remaining in a transient equilibrium with the increase in atmospheric  $\text{CO}_2$ . In the remaining biomes, deviations from such equilibrium may reflect that the trend of surface ocean  $f\text{CO}_2$  is not equal to that of the atmosphere, most notably in the equatorial Pacific Ocean, or may reflect changes in the oceanic buffer (Revelle) factor. We conclude that well-planned and long-term sustained observational networks are key to reliably document the ongoing and future changes in ocean carbon chemistry due to anthropogenic forcing.

## 1 Introduction

The concentration of atmospheric carbon dioxide ( $\text{CO}_2$ ) is rapidly increasing due to the burning of fossil fuels, cement production, and land use changes (Le Quéré et al., 2014). This drives a net flux of  $\text{CO}_2$  into the ocean, causing the dissolved inorganic carbon (DIC) concentration to increase, which drives a decrease in pH and in the concentration of the carbonate ion ( $\text{CO}_3^{2-}$ , Doney et al., 2009b; Zeebe and Wolf-Gladrow, 2001). These changes in the ocean inorganic carbon chemistry, collectively referred to as ocean acidification (Gattuso and Hansson, 2011), are a source of concern due to their potential impact on organisms, ecosystems, and biogeochemical cycles (Doney et al., 2009a). Hereafter, we refer to the inorganic carbon chemistry in the ocean as  $\text{CO}_2$  chemistry. In contrast to the surface ocean fugacity of carbon dioxide ( $f\text{CO}_2$ ), for which many studies have analyzed the long-term trends, both regionally and globally (e.g., Fay and McKinley, 2013; Le Quéré, 2010; Lenton et al., 2012; Takahashi et al., 2009b), only a handful of regional studies have so far been published on long-term pH trends (Bates, 2007; Dore et al., 2009; Gonzalez-Davila et al., 2007; Olafsson et al., 2010).

The most extensive assessment to date is the one of Bates et al. (2014). They described changes in ocean  $\text{CO}_2$  chemistry variables at seven, mostly tropical/subtropical, time-series stations, all of which have been operational for at least 2 decades. Their analysis shows that while there are regional differences, these open ocean time-series show relatively similar trends in DIC,  $f\text{CO}_2$ , and pH. At the tropi-

cal and subtropical open ocean stations (Bates, 2007; Dore et al., 2009; Gonzalez-Davila et al., 2010), ocean pH is decreasing at a rate of  $0.0017 \pm 0.0002 \text{ yr}^{-1}$ . At the high-latitude stations, however, a more variable picture emerges. While the pH trend in the Iceland Sea follows the rate observed at the lower latitude stations, the trend in the Irminger Sea (Olafsson et al., 2010) is nearly twice as large, i.e.,  $-0.0026 \pm 0.0006 \text{ yr}^{-1}$ . Thus, in a global analysis, we expect a complex spatial pattern of long-term trends, yet hitherto unknown.

The absence of a global analysis of long-term trends is largely a consequence of the lack of direct surface ocean pH measurements, which is in sharp contrast to the situation for surface ocean  $f\text{CO}_2$ , for which data products contain several million observations (Bakker et al., 2014; Pfeil et al., 2013; Takahashi et al., 2009a). This limitation can be overcome by using computed pH, obtained by combining the very large data products of  $f\text{CO}_2$  with estimates of surface alkalinity. Lauvset and Gruber (2014) demonstrated that for the North Atlantic, this approach is able to produce rather accurate estimates of surface ocean pH. Takahashi et al. (2014) came to the same result globally. Even though the use of pH computed from  $f\text{CO}_2$  generates a global data set containing millions of pH observations, the resulting data are still sparse in time and space on a global scale, making the determination of global long-term trends challenging. For surface  $p\text{CO}_2$  this challenge has historically been overcome by binning the data into a very coarse grid (order of  $5^\circ$ – $10^\circ$  in latitude and longitude) by, for example, Lenton et al. (2012), Takahashi et al. (2002), and Takahashi et al. (2009b); however, more recently Fay and McKinley (2013) proposed to aggregate the data into biomes. This type of aggregation is more likely to capture the correct long-term dynamics of a region, as one expects a biome to respond in a more coherent manner to perturbations than a region defined by a latitude/longitude range.

Given the absence of a global observation-based analysis of pH trends, models have so far been the only source of information. The Norwegian Earth System Model (NorESM1-ME), as part of the Coupled Model Intercomparison Project phase 5 (CMIP5, Taylor et al., 2012), simulates a global average pH decrease of  $0.0017 \text{ yr}^{-1}$  (1981–2011), which is largely commensurate with observations reported from the time series stations. A recent study using ten different CMIP5 models, including NorESM1-ME, showed that all models give similar global average pH trends – both in the historical and future scenarios (Bopp et al., 2013).

This secular pH trend of  $-0.0017 \text{ yr}^{-1}$  and the low spread between models is expected for an ocean where (i) the surface ocean  $f\text{CO}_2$  follows that in the atmosphere due to the sufficiently rapid exchange of the excess  $\text{CO}_2$  between the atmosphere and the surface ocean, and (ii) where the change in the buffer (Revelle) factor remains spatially uniform, as the partial derivative  $\partial[\text{H}^+]/\partial f\text{CO}_2$  is directly related to this quantity (Orr, 2011; Sarmiento and Gruber, 2006). A change

in the buffer (Revelle) factor is expected since much of the  $\text{CO}_2$  newly added to the surface ocean from the atmosphere will be titrated away by  $\text{CO}_3^{2-}$ , causing a decrease in its concentration. This decreases the ability of the surface ocean to “buffer” the pH against further uptake of  $\text{CO}_2$ , thus increasing the Revelle factor (Sarmiento and Gruber, 2006). However, regional variations in how the Revelle factor changes may occur. Bates et al. (2014) show, for example, not only variations of the pH trends between the high- and low latitude time series, but also that the trends in Revelle factor are different, indicating that other factors are influencing the Revelle factor. These factors are mainly those processes that affect DIC and alkalinity, such as changes in ocean productivity and calcification, while changes in temperature and salinity are of minor importance (Sarmiento and Gruber, 2006).

Local and regional changes in the buffer (Revelle) factor are driven by the changing, and spatially varying, ratio of DIC to alkalinity. Spatial changes in this ratio have the potential to decouple the pH trends from those of the surface ocean  $f\text{CO}_2$  (Orr, 2011), potentially causing a more variable pattern in the pH trends. The complex spatial variability, identified by Bates et al. (2014) and others (e.g., Tjiputra et al., 2014) supports this hypothesis. This also shows that analyses of global pH trends, including the regional distribution of changes and the dynamics of the changing ocean  $\text{CO}_2$  system, are required for a comprehensive understanding. Global analyses are also necessary for the validation of model results, for underpinning and interpreting response studies from organism to ecosystem level, and for optimizing the planning of continued and future observational networks.

Here, we take advantage of the approach of Lauvset and Gruber (2014) to determine global ocean pH trends, and their drivers, using pH data calculated from the more than 10 million observations of surface ocean  $f\text{CO}_2$  that have been made available through the Surface Ocean  $\text{CO}_2$  Atlas (SOCAT) project (Bakker et al., 2014; Pfeil et al., 2013). Although pH is the main parameter of interest,  $f\text{CO}_2$  has been carried through all our analyses in order to determine how  $\text{CO}_2$  chemistry causes the evolution of pH to differ from that expected from  $f\text{CO}_2$  alone. Finally we use the long-term pH trends derived from a global earth system model, the NorESM1-ME, in order to illustrate how important spatial variability is for the representativeness of our trend results.

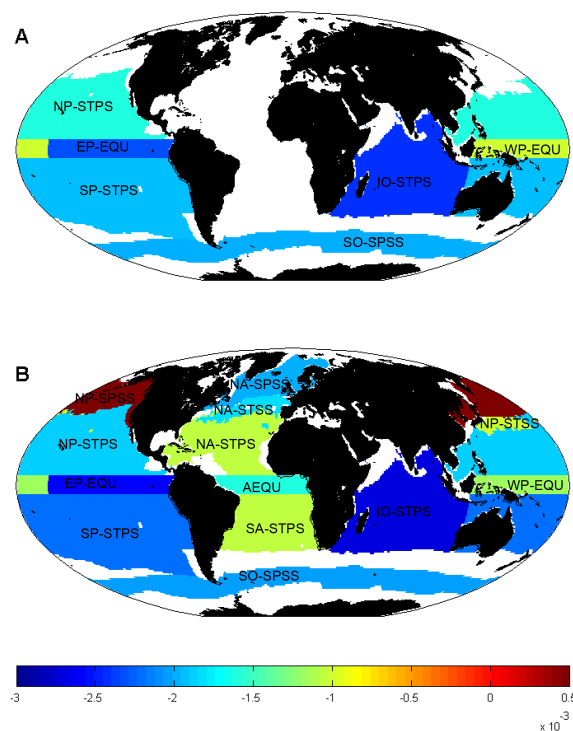
## 2 Data and methods

We calculated pH in the surface ocean by a two-step calculation using observations of  $f\text{CO}_2$ , sea surface temperature (SST), and sea surface salinity (SSS) from SOCAT version 2, (Bakker et al., 2014). In the first step, alkalinity was calculated from SSS and SST using the algorithms developed by Lee et al. (2006) and Nondal et al. (2009). The Nondal et al. (2009) algorithms were developed specifically for the high-latitude ( $> 60^\circ \text{ N}$ ) Atlantic Ocean, and were used only

there. Whenever no measured SSS was available in the SOCATv2 data set, the climatological World Ocean Atlas SSS value (Antonov et al., 2010) – which is included in the SOCATv2 data product – was used instead. The SOCAT SSS data have not been quality controlled and might therefore be biased. Lauvset and Gruber (2014) showed that this potential bias does not greatly affect the precision of the pH trends. It may affect the accuracy of the calculation, but for our purpose of determining long-term trends, the accuracy (i.e., the lack of bias in the data) is of less importance as long as the precision is good enough, and assuming that any bias remains constant over time. In the second step, pH on the total scale at in situ temperatures was calculated from the estimated alkalinity and the observed  $f\text{CO}_2$  using CO2SYS (Lewis and Wallace, 1998). We used the  $K_1$  and  $K_2$  constants from Mehrbach et al. (1973) refit by Dickson and Millero (1987), and the borate-to-salinity ratio from Uppström (1974). Since we use CO2SYS, this calculation also gives us dissolved inorganic carbon (DIC) and all other variables of the ocean carbon chemistry system.

Quite a few of the data fall outside the valid ranges for input data for the Lee et al. (2006) and Nondal et al. (2009) alkalinity algorithms and are lost in this step. There remain 7 381 013 data points of pH (and alkalinity) over the global ocean in the time period 1973–2011. The  $f\text{CO}_2$  trends have been estimated using only data points which have a calculated pH value in order to avoid spurious differences when comparing these trends to those of pH. The global calculation error (precision) for pH is  $0.0032 \pm 0.0005$ , and the calculated pH compares well to observed pH at crossover locations in the Atlantic Ocean (Lauvset and Gruber, 2014). Before analysis, the pH data were bin averaged into monthly  $1^\circ \times 1^\circ$  bins, using no extrapolation or interpolation of the data. The global data set was divided into the 17 ocean biomes, defined (using mixed layer depth, sea surface temperature, and Chlorophyll  $a$  concentrations) by Fay and McKinley (2014), as shown in Fig. 1. Here, we only evaluate trends in the open ocean. Data from coastal regions shallower than 250 m, based on the ETOPO2 bathymetry, and those with salinity  $< 20$  were removed.

In each biome, a least squares linear regression weighted with Tukey's bisquare method was used to determine the long-term pH trend. For the long-term trend determination, we required each biome to have at least three observations in each decade (1981–1990, 1991–2000, and 2001–2011). While this criterion was met in only 8 biomes for the period 1981–2011, 15 had sufficient data for the period 1991–2011. Both ordinary and weighted least squares regressions were carried out, but we chose a weighted least squares regression over an ordinary least squares regression since this is less sensitive to outliers in the data. This makes the statistics of the regression more robust, and generally this choice does not significantly affect the results presented here. All regression results are presented with the standard error of the slope, which represents its 68 % confidence interval, and the root



**Figure 1.** A map of the Fay and McKinley (2014) biomes which have (a) a statistically significant pH trend in the period 1981–2011, and (b) the biomes with a statistically significant pH trend in the period 1991–2011.

mean square error (RMSE). The RMSE is used as a measure of interannual variability.

Before the regression analysis was carried out, two corrections were applied to the data: deseasonalization and removal of spatial bias. The importance of these corrections, particularly in data sparse biomes such as those in the Southern Ocean, was recently highlighted by Fay et al. (2014). The seasonal cycle in the data was removed following Takahashi et al. (2009b), using the long-term average seasonal cycle as contained in our data for each biome. However, we find that using the climatological seasonal cycle – calculated using the Takahashi  $p\text{CO}_2$  climatology (Takahashi et al., 2009a) – does not significantly affect the results. To correct for any spatial bias in the large scale biomes, the difference between the climatological value in each  $1^\circ \times 1^\circ$  bin and the biome-mean climatological value was subtracted from the observed value in each  $1^\circ \times 1^\circ$  bin. There is no difference between this method and simply subtracting the climatological value in each  $1^\circ \times 1^\circ$  bin, but our approach retains the absolute values in each biome. It should be noted that the computed trends in some biomes are sensitive to which climatological data is used for the spatial bias correction: subtracting the climatological value vs. subtracting the long-term average in each  $1^\circ \times 1^\circ$  bin. Mostly, this is because in some  $1^\circ \times 1^\circ$  bins, the

long-term average is biased towards the last decade, which has significantly more data than earlier periods.

A statistical test was performed to test the necessity of these corrections: results after applying one or both corrections were compared to results after applying none using a one-way analysis of variance (ANOVA, see e.g., Vijayvargiya, 2009). A statistically significant change in the slope and its standard error was interpreted as making the correction(s) necessary. The deseasonalization removes scatter in the data and leads to more robust regressions by reducing the standard error of the slope in all biomes. This correction does not significantly ( $p$  value  $< 0.05$ ) affect the long-term trend in any biome or time period, however. The spatial bias correction has no statistically significant impact on the long-term trend in most biomes, but because it reduces the standard error and increases the  $r^2$  in six biomes, we decided to continue applying it. The long-term pH trend is also much more sensitive to this correction than the  $f\text{CO}_2$  trend, mostly because the pH trend is very small and thus more sensitive to any data correction.

The pH change expected from a certain change in  $f\text{CO}_2$  was calculated using  $\Delta\text{pH}/\Delta f\text{CO}_2 = \partial\text{pH}/\partial f\text{CO}_2$ . The partial derivative was estimated in CO2SYS using 0.01  $\mu\text{atm}$  increments in  $f\text{CO}_2$ . Since both the  $f\text{CO}_2$  and pH trends are inextricably coupled to DIC change, what we in reality calculate here is the pH change incurred by a change in DIC equivalent to the given  $f\text{CO}_2$  trend when alkalinity, SST, and SSS remain constant. We used the same equation to evaluate what global average  $f\text{CO}_2$  change the global long-term trend in pH is consistent with, but by using  $-0.001$  incremental changes in pH.

In each biome, the long-term trend in pH was decomposed into the effects of changes in SST, SSS, alkalinity, and DIC. First, the impact of each of these drivers on the  $f\text{CO}_2$  trend was determined following Takahashi et al. (1993), Eqs. (2–5), we then converted our results to the impacts on  $[\text{CO}_2]$  and on  $[\text{H}^+]$  following Eq. (1.5.87) in Zeebe and Wolf-Gladrow (2001), and finally we determined the impact on pH. The DIC data and dissociation constants required for these calculations were calculated in CO2SYS from the  $f\text{CO}_2$  and alkalinity pair in the same calculation that gave us pH.

To test the effect of the highly variable spatial and temporal coverage of the observational data on the results we have used the NorESM1-ME Earth system model, which prognostically simulates the seawater  $\text{CO}_2$  chemistry. A detailed description and evaluation of the model simulation is available in Tjiputra et al. (2013). We examined the model simulation for the 1981–2011 period based on the CMIP5 historical and future RCP8.5 scenarios, where the atmospheric  $\text{CO}_2$  concentration is used as the boundary condition. We binned the model monthly output into the same  $1^\circ \times 1^\circ$  bins and used the same method to calculate and decompose the long-term trends in each biome as we used for the observational data – including the two-step pH calculation described above. Two sets of model trends were determined. For the first, we used

the fully sampled model output, referred to here as the “fully sampled trend”. For the second set, we subsampled the model output according to the observational coverage, i.e., only data from monthly grid cells corresponding to those where real observations have been obtained were used. The “subsampled trend” was then computed from these subsampled model data. The comparison of these two informs us of how sensitive the calculated trends are to the variable data coverage.

### 3 Results and discussion

#### 3.1 Long-term trends in pH

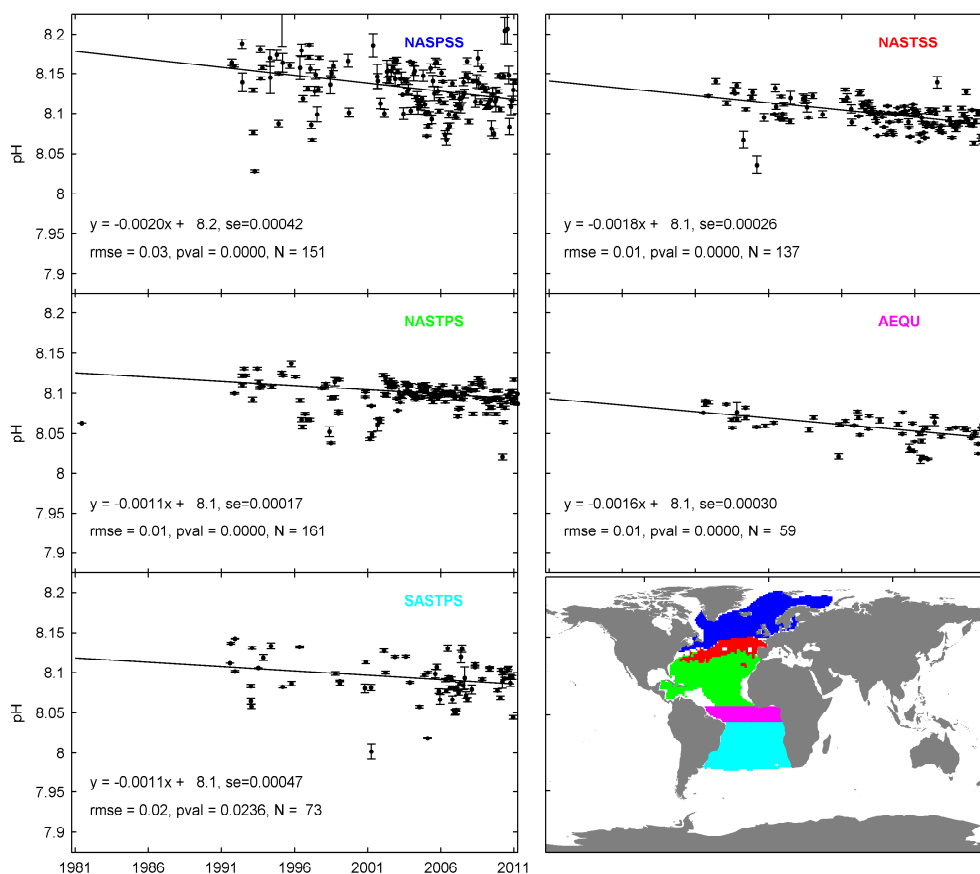
We find statistically significant trends in 6 out of the 8 biomes with sufficient data for the period 1981–2011, and for 13 out of the 15 biomes with sufficient data for the period 1991–2011 (Fig. 1 with the numerical values in Table 1). As shown in Figs. 2–4, the data coverage in each biome is generally very good after 1990, but often spotty prior to this year. These figures also reveal a substantial amount of interannual variability around the determined trends, with RMSE values of between 0.01 and 0.04 pH units, i.e., roughly similar magnitude as the cumulative trend over the 20 to 30 years of analysis. No robust analyses were possible for the North Pacific ice-covered (NP-ICE) and North Atlantic ice-covered (NA-ICE) biomes, due to the lack of data ( $< 20$  data points) hence they are not further discussed in the paper. Unfortunately, these are the Arctic biomes where the earliest impacts of ocean acidification are expected (Steinacher et al., 2009).

The regions with sufficient data, but without statistically significant trends, i.e., the North Pacific subpolar seasonally stratified (NP-SPSS) biome for the period 1981–2011, and the Southern Ocean subtropical seasonally stratified (SO-STSS) and ice-covered (SO-ICE) biomes for the period 1991–2011, are characterized by large RMSEs and a substantial amount of decadal variability, which likely masks the long-term trends. In addition to these three biomes where the trends are statistically indistinguishable from zero, the South Pacific subtropical permanently stratified (SP-STPS) biome is likely biased by its low data density, and will not be further discussed. This decision was corroborated by comparing the pH trend in the fully sampled model results with the subsampled model results (Fig. 5): the SP-STPS biome is the only one where the difference in these trends is statistically significant at the 95 % confidence level.

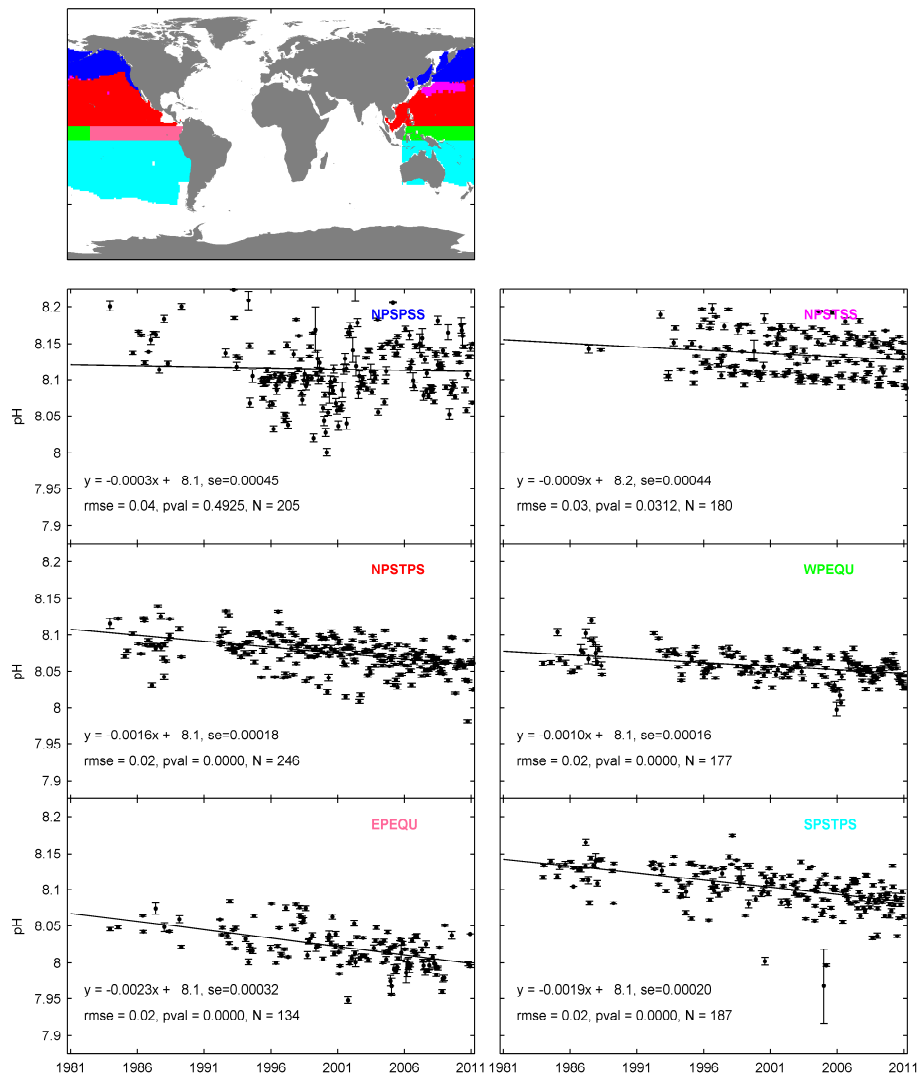
Since we are not able to calculate statistically significant trends in all 17 biomes, we are also unable to calculate a global average trend. For the period, 1991–2011 only the Arctic and parts of the Southern Ocean have no statistically significant results, and the area-weighted average pH decrease of the remaining 13 biomes (Table 1), is  $0.0018 \pm 0.0004 \text{ yr}^{-1}$ . For the period, 1981–2011 the number of biomes with trend estimates is quite small, but al-

**Table 1.** Results and statistics of the regression analysis of  $f\text{CO}_2$  ( $\mu\text{atm}$ ) and  $\text{pH}_{\text{insitu}}$  versus time. Bold text indicates biomes where the results are not statistically significant (95 % confidence). No number is given if a biome does not have enough data to calculate the trend in a given time period.

Region	1981–2011				1991–2011			
	pH		$f\text{CO}_2$		pH		$f\text{CO}_2$	
	Slope	RMSE	Slope	RMSE	Slope	RMSE	Slope	RMSE
NP-SPSS	<b>-0.0003 ± 0.0005</b>	<b>0.041</b>	1.20 ± 0.17	16.2	0.0013 ± 0.0005	0.038	0.74 ± 0.22	16.1
NP-STSS	–	–	1.30 ± 0.15	10.5	-0.0010 ± 0.0005	0.031	1.37 ± 0.13	8.9
NP-STPS	-0.0016 ± 0.0002	0.020	1.51 ± 0.09	10.3	-0.0019 ± 0.0002	0.018	1.52 ± 0.12	9.9
WP-EQU	-0.0010 ± 0.0002	0.016	1.54 ± 0.19	17.8	-0.0012 ± 0.0002	0.015	1.59 ± 0.27	17.3
EP-EQU	-0.0023 ± 0.0003	0.023	2.94 ± 0.41	28.2	-0.0026 ± 0.0002	0.023	3.51 ± 0.51	27.9
SP-STPS	-0.0019 ± 0.0002	0.020	1.34 ± 0.11	12.0	-0.0022 ± 0.0003	0.020	1.12 ± 0.18	12.3
NA-SPSS	–	–	1.18 ± 0.22	15.4	-0.0020 ± 0.0004	0.028	1.11 ± 0.22	14.2
NA-STSS	–	–	1.78 ± 0.20	12.3	-0.0018 ± 0.0003	0.015	1.79 ± 0.20	12.5
NA-STPS	–	–	1.42 ± 0.12	8.5	-0.0011 ± 0.0002	0.012	1.44 ± 0.12	8.6
A-EQU	–	–	1.86 ± 0.35	16.6	-0.0016 ± 0.0003	0.014	1.81 ± 0.32	15.7
SA-STPS	–	–	1.06 ± 0.37	16.7	-0.0011 ± 0.0005	0.024	0.99 ± 0.37	17.0
IO-STPS	-0.0024 ± 0.0004	0.023	1.49 ± 0.25	13.6	-0.0027 ± 0.0005	0.025	1.55 ± 0.26	13.5
SO-STSS	<b>-0.0006 ± 0.0004</b>	<b>0.032</b>	1.78 ± 0.11	10.8	<b>-0.0004 ± 0.0004</b>	<b>0.032</b>	1.82 ± 0.12	10.8
SO-SPSS	-0.0020 ± 0.0002	0.020	1.44 ± 0.10	9.1	-0.0021 ± 0.0002	0.020	1.46 ± 0.11	9.0
SO-ICE	–	–	0.34 ± 0.31	24.4	<b>-0.0002 ± 0.0004</b>	<b>0.029</b>	0.23 ± 0.34	24.3



**Figure 2.** Long-term pH trend (1981–2011) in the five Atlantic Ocean biomes.

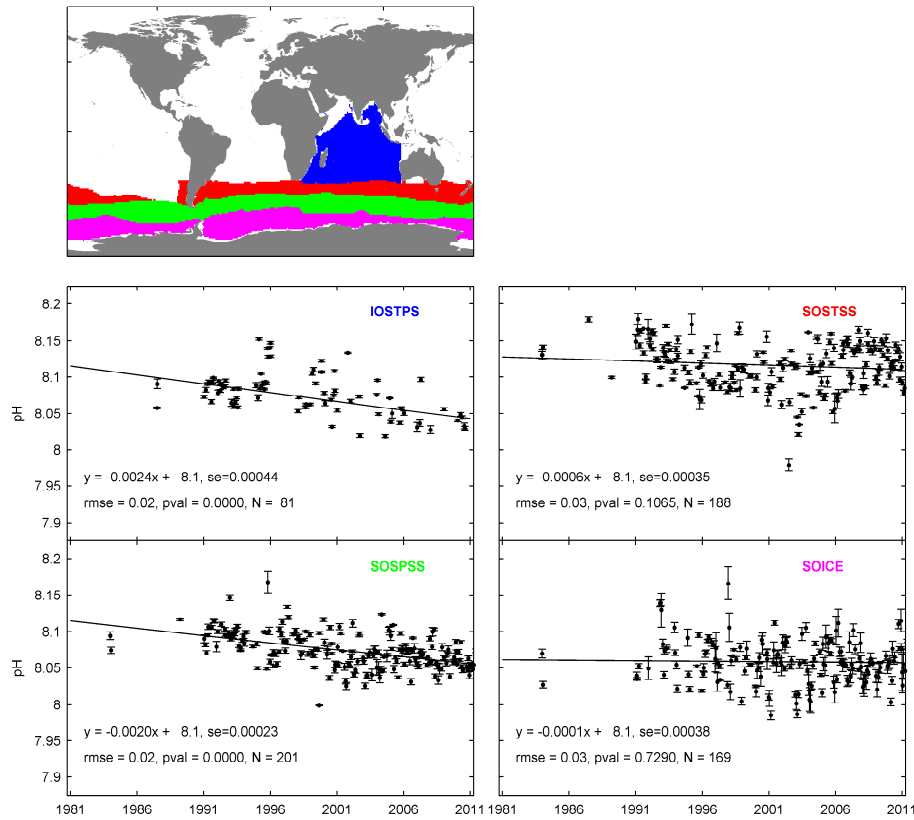


**Figure 3.** Long-term pH trend (1981–2011) in the five Pacific Ocean biomes.

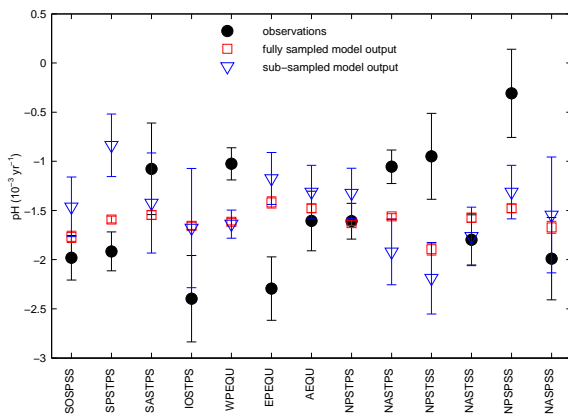
most all the Pacific Ocean biomes have results and the area-weighted pH decrease is  $0.0019 \pm 0.0001 \text{ yr}^{-1}$  between 1981 and 2011. Within the uncertainty limits, the global 1991–2011 trend is comparable to the global trend in the fully sampled NorESM1-ME model results ( $-0.0017 \text{ yr}^{-1}$ ) and to the average trend of  $-0.0018 \pm 0.0003 \text{ yr}^{-1}$  over the seven time series evaluated by Bates et al. (2014). Assuming that alkalinity, SST, and SSS remain constant, and that the change in DIC and Revelle factor remain spatially uniform, this global average pH trend corresponds to a rate of increase in surface ocean  $f\text{CO}_2$  of  $1.75 \pm 0.4 \mu\text{atm yr}^{-1}$ , which is roughly the rate of increase in atmospheric  $f\text{CO}_2$ . Regionally, however, the response of the ocean  $\text{CO}_2$  system to the atmospheric forcing is more variable (Fig. 1).

In the North Atlantic subpolar seasonally stratified (NA-SPSS) biome, the observed pH trend is  $-0.0020 \pm 0.0004 \text{ yr}^{-1}$ . This is right in between

the trend observed at the Irminger Sea time series ( $-0.0026 \pm 0.0006 \text{ yr}^{-1}$ ) and that observed at the Iceland Sea time series ( $-0.0014 \pm 0.0005 \text{ yr}^{-1}$ ) (Bates et al., 2014). Within the 68 % confidence intervals, the NA-SPSS pH trend is consistent with both of these local trends. In the North Atlantic subtropical seasonally stratified (NA-STSS) biome, there are no time series data to compare with, but its trend of  $-0.0018 \pm 0.0003 \text{ yr}^{-1}$  is consistent with a trend of  $\sim -0.0020 \text{ yr}^{-1}$  observed in the Rockall Trough by McGrath et al. (2012). In the North (NA-STPS) and South (SA-STPS) Atlantic subtropical permanently stratified biomes, the pH trend is the same, but the RMSE values indicate larger interannual variability in the southern biome (Table 1). This is likely caused by the inclusion of the Benguela upwelling region, but the full effect of this has not been quantified for the SA-STPS or any other biome. The trend identified here for the NA-STPS ( $-0.0011 \pm 0.0002 \text{ yr}^{-1}$ ) is



**Figure 4.** Long-term pH trend (1981–2011) in the Indian Ocean biome and the three Southern Ocean biomes.



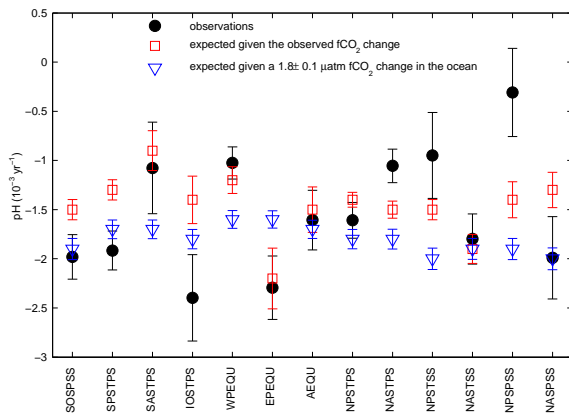
**Figure 5.** Summary of the pH trends in all biomes. The error bars show the  $1\sigma$  confidence interval.

significantly lower than the trend observed at the Bermuda Atlantic Time-series Study (BATS, Bates et al., 2014), of  $-0.0017 \pm 0.0001 \text{ yr}^{-1}$ . Unfortunately, we have no time series data for comparison in the SA-STPS biome. In the Atlantic Ocean equatorial region (A-EQU), the pH trend ( $-0.0016 \pm 0.0003 \text{ yr}^{-1}$ ) is lower than that observed at the Carbon Retention in A Colored Ocean (CARIACO) time-series station of  $-0.0025 \pm 0.0004 \text{ yr}^{-1}$  (Bates et al.,

2014), but this station is located at the very edge of the biome in a more coastal setting and not ideal for comparison.

In the Pacific Ocean, the RMSE around the fitted pH trends is generally larger than in the Atlantic Ocean (Table 1), possibly reflecting the higher interannual variability of the surface  $\text{CO}_2$  system there (see e.g., Landschützer et al. 2014 for  $p\text{CO}_2$  variability). In the North Pacific subtropical permanently stratified (NP-STPS) biome, the pH trend of  $-0.0016 \pm 0.0002 \text{ yr}^{-1}$  is the same as that observed at the Hawaii Ocean Time-series (HOT, Bates et al., 2014). The trends in the two equatorial Pacific Ocean biomes differ substantially. While the western biome (WP-EQU) has a relatively weak trend ( $-0.0010 \pm 0.0002 \text{ yr}^{-1}$ ), the eastern (EP-EQU) biome has a much stronger pH trend than any other biome except the IO-STPS. This could be related to the recent trend toward stronger and more prevalent La Niña conditions in the eastern tropical Pacific leading to stronger upwelling, and higher surface  $f\text{CO}_2$  and lower pH in this region (Rödenbeck et al., 2014).

The Indian Ocean subtropical permanently stratified (IO-STPS) biome had a very strong pH trend the past 30 years, only rivaled by that in the EP-EQU biome, as mentioned above. There are no time series stations in the Indian Ocean to compare with, but  $f\text{CO}_2$  trends for the Indian Ocean computed by Metzl (2009) are considerably larger than what we

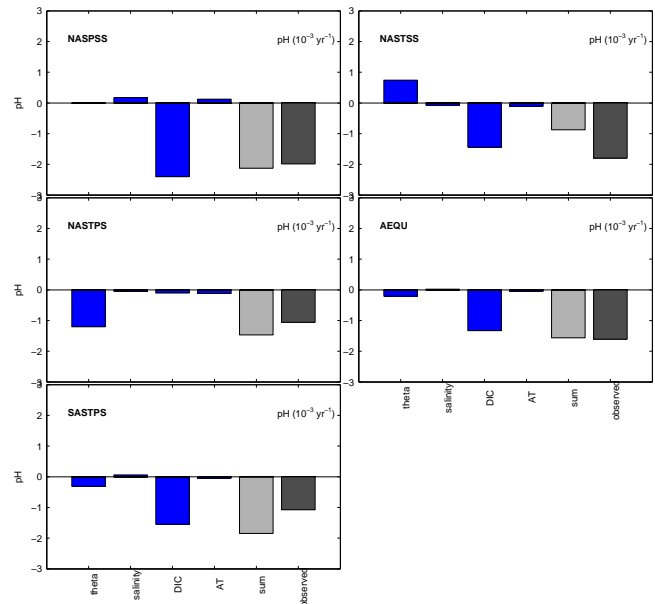


**Figure 6.** Comparison between the observed pH trend in each biome (either 1981–2011 or 1991–2011) in black, and the pH trends expected if the surface ocean  $f\text{CO}_2$  change were equal to that the atmosphere in blue and expected for the observed ocean  $f\text{CO}_2$  trends in red.

find:  $2.11 \mu\text{atm yr}^{-1}$  vs  $1.44 \pm 0.24 \mu\text{atm yr}^{-1}$ . Hence there is no reason to believe that our approach overestimates the pH trends here. It should be noted though that the trend identified by Metzl (2009) is based on data in a considerably smaller region than the IO-STPS which could account for some of the difference. In the Southern Ocean, only the subpolar seasonally stratified (SO-SPSS) biome has a statistically significant pH trend, which at  $-0.0020 \pm 0.0002 \text{ yr}^{-1}$  is comparable to that in the NA-SPSS biome. Furthermore, this trend is very similar to that calculated for this region by Takahashi et al. (2014), although they used a different method.

### 3.2 Effects of changes in carbonate chemistry

To first order, the pH trends are expected to represent the direct response to increasing oceanic DIC, as is the case for the long-term trends in surface ocean  $f\text{CO}_2$ . In order to assess how our results compare with this expectation, we have calculated two expected pH rates of change: first, the 1981–2011 change in pH resulting from a surface ocean  $f\text{CO}_2$  rate of change equal to that in the atmosphere ( $1.8 \pm 0.1 \mu\text{atm yr}^{-1}$ ) while keeping all other variables constant at their 1981 values; and second, the change in pH that would be expected if the pH change mirrored the observed  $f\text{CO}_2$  change in each biome provided that all other variables were kept at their 1981 values. The first expected pH change reflects how pH should change if the change in atmospheric  $\text{CO}_2$  was the sole driver for the change in ocean pH. The second expected pH change reflects how pH should change if the oceanic  $f\text{CO}_2$  changes were allowed to depart from the atmospheric ones but  $f\text{CO}_2$  change remaining the only driver of pH change. Figure 6 shows both expected pH changes along with the observed pH change in each biome. Only the 13 biomes that have statistically significant pH trends for either 1981–2011 or 1991–2011 (Fig. 1)

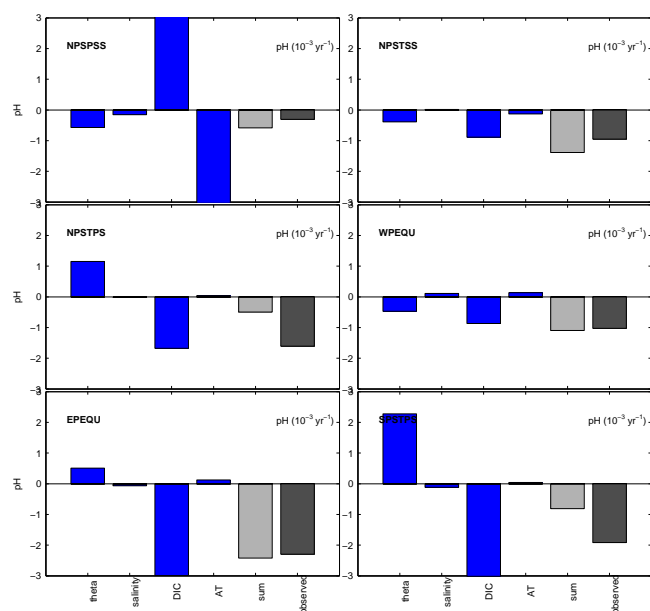


**Figure 7.** The long-term trends in pH from Fig. 2 decomposed into the contributions from SST, SSS, alkalinity, and DIC. Also shown is the sum of the four contributions and the actual observed trend. Note that the trend has been multiplied by 1000 for easier visualization.

are discussed further. When the atmospheric  $\text{CO}_2$  increase is assumed to be the only driver for the pH changes, we find that in 7 of the 13 biomes, the observed pH trends significantly differ from the expected pH change. This is due to the uncertainty in the observed trends, to associated changes in the  $\text{CO}_2$  chemistry, or to the surface ocean  $f\text{CO}_2$  trends being significantly different from that in the atmosphere. However, the observed pH trends also significantly differ from the expected pH change calculated using the observed  $f\text{CO}_2$  trend in 6 of the 13 biomes (Fig. 6). Only 3 of the biomes are the same in both cases. Thus, the surface ocean  $f\text{CO}_2$  trend not exactly mirroring the atmospheric cannot explain the discrepancy between expected and observed pH trends in most biomes. It may be an explanation in the equatorial Pacific biomes (EP-EQU and WP-EQU) where there is no discrepancy between observed and expected pH trends when the observed  $f\text{CO}_2$  trend is used to calculate the expected pH change (Fig. 6), but a significant difference exists when an atmospheric rate of change is assumed.

The observed pH trend is more often smaller than that expected for the ocean mirroring the atmospheric  $f\text{CO}_2$  change than vice versa. Only the EP-EQU and IO-STPS biomes have observed pH changes larger than those expected (Fig. 6). Our hypothesis is that the differences between the observed and expected pH trends are caused by changes in the spatial variations in the ratio of DIC to alkalinity, which leads to spatial changes in the buffer (Revelle) factor. In the biomes where the observed trend differs from the expected trend, there are indications which point to such changes. When the differ-



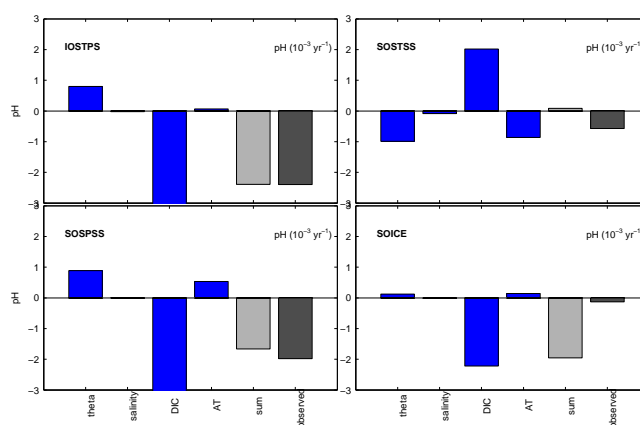


**Figure 8.** The long-term trends in pH from Fig. 3 decomposed into the contributions from SST, SSS, alkalinity, and DIC. Also shown is the sum of the four contributions and the actual observed trend. Note that the trend has been multiplied by 1000 for easier visualization.

ence is negative (i.e., the observed trend is smaller than the expected), the decrease in Revelle factor, for example, is stronger the larger the difference. However, given the combined calculation errors, generally high level of noise in our data, and relatively few data points, only some of these indications are statistically significant. Further analysis of these spatial patterns needs to be undertaken using independent pH data, preferably direct measurements, in order to quantify any possible biases in the results due to our pH being a calculated variable. A combination of SOCAT data with repeat hydrography and time-series data would be ideal but this is outside the scope of this study.

### 3.3 Major driving forces behind the observed pH and trends

The decomposition of the  $f\text{CO}_2$  and pH trends (Figs. 7–9) confirms that in all biomes, the long-term increase in DIC is by far the dominant driver for the long-term pH changes. Knowledge about the changes in ocean DIC therefore is the most important in understanding – and predicting – changes in ocean pH (Table 2). This is not unexpected since the open ocean is in – or very close to – chemical equilibrium with the atmosphere (Lauvset and Gruber, 2014). Thus the surface ocean is taking up  $\text{CO}_2$  from the atmosphere in order to re-establish a chemical equilibrium, leading to a corresponding increase in  $f\text{CO}_2$  and DIC. It must be noted that since we do not have measurements of alkalinity, this parameter is calculated from SST and SSS, and the relatively large uncer-



**Figure 9.** The long-term trends in pH from Fig. 4 decomposed into the contributions from SST, SSS, alkalinity, and DIC. Also shown is the sum of the four contributions and the actual observed trend. Note that the trend has been multiplied by 1000 for easier visualization.

tainties in these calculations may add a degree of uncertainty to the decomposition. Due to a lack of independent data this is not further evaluated in this study.

In the Atlantic Ocean biomes, the second most important driver is SST (Fig. 7), which mostly has a positive change and therefore has limited the DIC increase required to maintain an  $f\text{CO}_2$  growth rate similar to that in the atmosphere. SST is the second most important driver in the Pacific Ocean biomes also (except in the NP-SPSS, Fig. 8), but here SST decreased in many biomes leading to an enhanced increase in DIC through  $\text{CO}_2$  uptake from the atmosphere. In the Southern Ocean biomes, alkalinity changes have a significant impact on the trends (Fig. 9), which also modulates the DIC changes. Decreasing alkalinity over time increases  $f\text{CO}_2$  so that the DIC change required to maintain a sea surface  $f\text{CO}_2$  growth rate similar to the atmospheric is reduced.

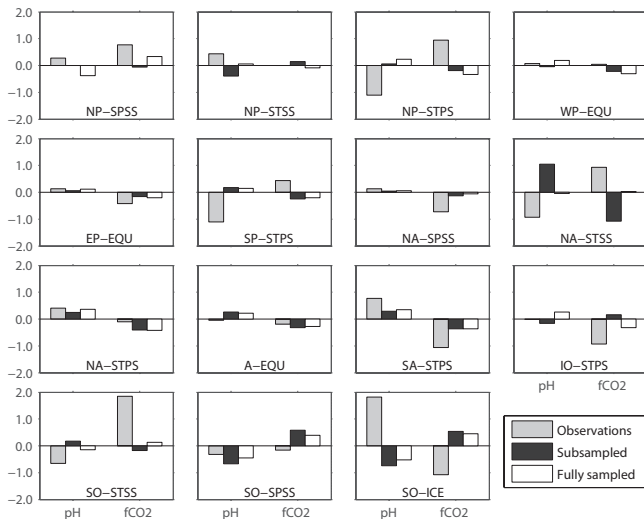
In most biomes, there is a residual between the sum of the four components and the observed trend (Fig. 10). Lenton et al. (2012) performed a similar analysis and attributed such residuals to the use of a spatial mean Revelle factor, the approximations underlying the Takahashi et al. (1993) equations, and the assumption of linear trends in all variables. We tested whether variable data coverage is also an important contributor to this residual by subsampling the NorESM1-ME simulated pH data and comparing the resulting 1981–2011 decomposition with the decomposition determined using the full model output. Figure 10 illustrates that in most biomes, there are similar residuals between the sum of the four components and the actual trends in the sub-sampled and fully sampled model fields as well. We can, therefore, find no evidence to show that poor data coverage is of major importance in determining what drives the change in surface ocean pH.

**Table 2.** Decomposition of the  $f\text{CO}_2$  and  $\text{pH}_{\text{insitu}}$  trends into their major drivers. The units are  $\mu\text{atm yr}^{-1}$  and  $10^{-3}$  pH units  $\text{yr}^{-1}$  respectively.

Region	pH					$f\text{CO}_2$				
	Theta	Salinity	DIC	Alkalinity	Sum	Theta	Salinity	DIC	Alkalinity	Sum
NP-SPSS	-0.57	-0.15	3.18	-3.04	-0.58	0.52	0.14	-2.89	2.76	0.53
NP-STSS	-0.39	0.02	-0.89	-0.13	-1.39	0.38	-0.02	0.87	0.13	1.37
NP-STPS	1.15	-0.02	-1.68	0.04	-0.50	-1.19	0.02	1.73	-0.05	0.51
WP-EQU	-0.47	0.11	-0.87	0.14	-1.10	0.53	-0.12	0.97	-0.15	1.23
EP-EQU	0.51	-0.07	-2.99	0.13	-2.42	-0.63	0.08	3.68	-0.15	2.99
SP-STPS	2.28	-0.11	-3.02	0.04	-0.81	-2.47	0.12	3.28	-0.05	0.88
NA-SPSS	-0.02	0.17	-2.41	0.12	-2.13	0.01	-0.16	2.17	-0.11	1.91
NA-STSS	0.74	-0.07	-1.43	-0.11	-0.87	-0.72	0.07	1.40	0.10	0.85
NA-STPS	-1.20	-0.05	-0.10	-0.12	-1.47	1.29	0.05	0.11	0.13	1.57
A-EQU	-0.21	0.02	-1.33	-0.05	-1.56	0.24	-0.03	1.53	0.06	1.80
SA-STPS	-0.31	0.06	-1.55	-0.05	-1.85	0.34	-0.07	1.69	0.05	2.02
IO-STPS	0.80	-0.02	-3.23	0.06	-2.39	-0.79	0.02	3.22	-0.06	2.38
SO-STSS	-0.99	-0.08	2.02	-0.86	0.09	0.88	0.08	-1.81	0.77	-0.08
SO-SPSS	0.89	0.01	-3.09	0.53	-1.66	-0.83	-0.01	2.89	-0.50	1.56
SO-ICE	0.13	-0.01	-2.22	0.15	-1.95	-0.12	0.01	2.02	-0.13	1.78

**Table 3.** Decomposition of the 2001–2011  $f\text{CO}_2$  and  $\text{pH}_{\text{insitu}}$  trends in the Southern Ocean into their major drivers. The units are  $\mu\text{atm yr}^{-1}$  and  $10^{-3}$  pH units  $\text{yr}^{-1}$  respectively. Observed pH trend is in pH units per year.

Region	pH					Observed	$f\text{CO}_2$					Observed
	Theta	Salinity	DIC	Alkalinity	Sum		Theta	Salinity	DIC	Alkalinity	Sum	
SO-STSS	-3.7	-0.79	8.46	-2.48	1.49	$0.0032 \pm 0.0010$	3.4	0.73	-7.78	2.28	-1.37	$1.56 \pm 0.39$
SO-SPSS	1.11	-0.07	-1.28	-0.05	-0.29	$-0.0011 \pm 0.0006$	-1.06	0.07	1.22	0.05	0.28	$0.89 \pm 0.22$
SO-ICE	-0.62	-0.05	1	-0.39	-0.06	$0.0006 \pm 0.0009$	0.59	0.05	-0.95	0.37	0.06	$0.21 \pm 0.74$

**Figure 10.** The residual between the actual pH trends and the sum of the four decomposition parts (SSS, SST, DIC, ALK). In gray is the residual for the observations, in black the residual for the subsampled model output, and in white the residual for the fully sampled model output.

### 3.4 Recent changes in the Southern Ocean biomes

In contrast to the majority of the global ocean biomes, trends within the SO-STSS and SO-ICE biomes do not appear statistically significant over the past 2 decades (Table 1). This can be linked to strong interannual and decadal variations (Fig. 3). This is consistent with the changing  $f\text{CO}_2$  trends revealed in a recent study by Fay et al. (2014) as well as previous findings of a change in the  $\text{CO}_2$  sink in this region (e.g., Fay and McKinley, 2013; Landschützer et al., 2014). In order to investigate these recent changes in the trend in the Southern Ocean, we also decompose the 2001–2011 trends in the Southern Ocean biomes (Table 3).

In the SO-STSS biome, there is no significant change in pH over the 30-year period, but from Fig. 3 it is seen that there is a decrease until  $\sim 2000$  and then an increase over the last decade. Over the last decade (Table 3), we find that the contributions of the individual parameters to the overall trend in pH are amplified. Temperature and DIC changes remain the strongest drivers, and of these the forcing from DIC has increased strongest over the last decade. We hence conclude that the increase in pH over the past decade in the SO-STSS biome is due to the decreasing DIC concentrations dominating over the thermally induced reduction in pH.

In the SO-SPSS biome, the pH trend appears to become less steep over the last decade (Fig. 3), which is consistent with the well-documented trend changes in  $f\text{CO}_2$ . In this biome, we find a less-negative DIC driven pH trend in the period 2001–2011 compared to the period 1981–2011, indicating a reduced positive trend in DIC over this decade. This supports the conclusion drawn by Fay and McKinley (2013) that a reduction in vertical DIC supply causes a weakening of both  $f\text{CO}_2$  and pH trends in this region. In the SO-ICE biome, the sign of the non-thermal drivers appears to change within the last decade, potentially driven by the recent Antarctic ice melt and ice-sheet-melting-driven iron fertilization (Death et al., 2014).

### 3.5 Spatial variability

In both the observations and the sub-sampled model results, we see significant regional differences in the pH trends (Fig. 5). Note that the actual simulated pH trends in each biome are not directly comparable with the observed trends since the model is a coupled climate model, which simulates its own internal climate variability. We therefore compare the fully sampled and the sub-sampled model results, and the fully sampled model results show much more uniform pH trends (Fig. 5). While these differences are mostly statistically indistinguishable within the uncertainties, it highlights the need for a careful consideration of representativeness when comparing model-derived future changes and trends based on data. Figure 5 shows that in the IO-STPS and WP-EQU biomes, the sub-sampled trend is within  $\pm 0.0001$  of the fully sampled pH trends, and an ANOVA analysis shows that only in the SP-STPS biome are the two model trends significantly different. Thus the trends based on the existing observational coverage are overall representative of the respective biomes, and it is unlikely that there are major biases in our results due to low data density. However, the uncertainties in the long-term pH trend estimates remain large, both in observations and the model (Fig. 5) and this prohibits a mechanistic understanding the observed changes in most biomes. Improved sampling strategies are necessary to reduce these uncertainties and thereby improve our understanding of surface ocean  $\text{CO}_2$  chemistry changes today and in the future.

This highlights the importance of both maintaining the observational networks already in place – like the voluntary observing ship (VOS) network in the North Atlantic (Watson et al., 2009) – and instigating new ones in less well-covered ocean regions. Of particular importance is improved data coverage in the southern Pacific Ocean (SP-STPS) where the data density as of today is too low for a robust analysis of long-term pH trends.

## 4 Conclusions

Global surface ocean pH changes over the past 30 years cannot be calculated as there are too few data in many biomes. For the past twenty years on the other hand, we find that the surface ocean pH has decreased by on average  $0.0018 \pm 0.0004 \text{ yr}^{-1}$ , excluding the Arctic and high-latitude Southern Ocean. There are however large regional variations with trends ranging from  $-0.0024 \text{ yr}^{-1}$  in the Indian Ocean (IO-STPS) biome to no significant change in the polar Southern Ocean (SO-ICE) biome. Our estimated global trend is comparable to the trends found at time-series stations and to the global average trend in the NorESM1-ME model. In all biomes, the pH trend is predominantly driven by changes in DIC, implying that the surface ocean pH decline is a direct response to the increasing uptake of atmospheric  $\text{CO}_2$ . Despite this, the  $f\text{CO}_2$  and pH trends do not exactly mirror each other, which is potentially linked to trends in the surface ocean buffer (Revelle) factor over the past few decades. In some biomes, this leads to smaller pH changes than expected from the  $f\text{CO}_2$  change, while in others regions, the pH changes are larger than expected. Thus, knowledge of both the changing ocean DIC and the changing ocean buffer (Revelle) factor is important for understanding and accurately determining the changing ocean pH.

There are regional differences in the pH trends. It is likely that these are caused by spatial heterogeneity in the concurrent changes in buffer (Revelle) factor, while spatial heterogeneity in the surface ocean  $f\text{CO}_2$  trends seems to have only a minor effect. Our comparison between fully sampled model and sub-sampled output from the NorESM1-ME model indicates that variable data coverage only presents a major problem in the South Pacific. This nicely highlights the overall success of the scientific community in creating observational networks that reduce data coverage issues. The many scientific studies arising from this effort – among many others the recent publications by Nakaoka et al. (2013), Landschützer et al. (2013, 2014), and Schuster et al. (2013) – show that we have come a long way in understanding how ocean  $\text{CO}_2$  chemistry is evolving in a world perturbed by fossil fuel emissions. The uncertainties in the trends presented here are, however, substantial and this largely prevents a more thorough understanding of current changes. Filling the remaining gaps in our surface ocean data is, therefore, still of great importance.

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