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## ASPECTS OF THE NEUTRON NOISE ANALYSIS IN VIEW OF DETERMINING THE KINETIC PARAMETERS OF THE NUCLEAR REACTORS WITH THERMAL NEUTRONS AT IFA - BUCHAREST

## I. METHOD OF THE REDUCED VARIANCE

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I. METHOD OF THE REDUCED VARIANCE

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Abstract : In the first part of this paper, a succint introduction is made to the fundamental concepts specific of the study of the stochastic processes. The non-Poissonian character is emphasized of the probability distribution of the number of neutrons recorded by a chain of counters within a given the interval, i.e. of the number representing the physically observable that is holding informations concerning the fluctuations of the neutron populations in the reactor. The difficult problem of collecting the experimental data in view of determining the re duced variance of the physically observable is solved by the authors by means of an electronic equipment which is sorting these experimental data in such a way that the relative recording frequencies of a siven sumber of neutrons is At interval is directly determined. In the second part of the paper che fundamontal diagram is described of the analysis chain after which a minute description is given of the fundamental diagram as well as of the signal diagram supplied by the stochastic multi-channel analyser achie vad by printed circuits in the form of a WIN module.

## I. INTRODUCTION

The conditions under which the neutron population within a nuclear reactor can be kept constant as a function of time (steady operating conditions of the neutron population) result from the general evolution equation of the neutron number within the active region of the reactor; this equation can be expressed as follows / 1 / :

(1) 
$$\hat{p}N - \hat{D}N + S = \frac{\partial}{\partial t}N$$

where, in symbolical writing, the rates of appearance and disappearance of the neutrons are  $R_p = \hat{P}N$ , respectively  $R_D = \hat{D}N$  and where appears moreover a term expressing the contribution of an excitation source of intensity S.

The production rate of the neutrons is proportional to the fission rate of the fuel nuclei.

The neutron population within the reactor is constant as a function of time  $\left(\frac{\partial}{\partial t}N = 0\right)$ , in the absence of an excitation (S = 0), only as long as an equilibrium is achieved between rates  $R_d$  and  $R_p$ . The state of the reactor when these conditions are fulfilled is called critical state.

If, however, for achieving the steady operating conditions of the neutron population, the presence is neces sary of an additional neutron source, the reactor is said to be in a subcritical state. The intensity of this neutron source is a function of want of balance between rates  $R_p$  and  $R_d$  and of the level of the neutron populating within the environment. For a given level of the population, source S is a measure of

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the subcriticallity. The intensity of the source is expressed by

$$S = R_{-1} - R_{-p}$$

Actually, even under operating conditions that at a macroscopic scale are called steady, the neutron population in the reactor displays certain fluctuations due to the stochastic character of the microscopic interaction processes between neutrons and nu clei within the environment. For instance, the probability for a neutron to provoke the fission of a fuel nucleous depends on several factors such as:

- the position within the reactor active area of the respective fuel nucleous;
- the composition of the active zone around the res pective spot;
- the microscopic cross-sections of the neutrons-nuclei interactions depending on the energy of the neutrons.

On the other hand, the function of the neutron spec tral distribution is strongly influenced by the form and the geometrical dimensions of the active area as well as by the relative concentrations of the materials forming it.

Owing to these factors, the fissions of the fuel nu clei do not occur at regular time intervals. The existence of a dispersion of the time intervals between the fissions offers a partial explanations of the fluctuations existing in the neutron population within the reastor. As a result of the fission of a fuel nucleus provoked by a neutron existing in the reactor, two fission fragments and a number  $v_p$  of neutrons (prompt neutrons) are simultaneously produced. The fission fragments are radioactive nuclei in an excited state, the atomic numbers of which are related by :

(3) 
$$Z_1 + Z_2 = Z$$
 (Law of conservation of the electric charge)

It is experimentally found that the nuclei of a fissile material that splits in apparently identical conditions (with neutrons of the same energy) do not emit every time the same pair of fission fragments, nor to they emit the same number  $v_p$  of prompt neutrons.

These relative frequencies represent as a matter of fact estimations of the probability that as the result of a fission the nucleus of a given species would appear, respectively of the probability that the tission would release a given number of prompt neutrons.

When these probabilities are known, it is possible to determine the statistical mean value of the prompt neutrons produced by the fission of a fuel nucleus.

$$\widetilde{v}_{p} = \widetilde{\widetilde{L}} \quad v_{p} \cdot p_{v} \\ v_{n} = 1 \qquad p$$

Quantity  $\tilde{v}_p$  is a constant of the material, very interesting as concerns the self maintaining of the chain (fission) reactions within the active region of the reactor, i.e. of the reactor working. Beside the prompt neutrons, the fission of a fuel nu cleus releases also neutrons due to the radioactive disintegration of the fission fragments (delayed neutrons). A fission fragment releases a single neutron by disintegration. It is a random phenomenon since there is a disper sion in the time interval between the disintegration moment of the nucleus and the appearance of the neutron owing to the fission . This dispersion is characteristic of the phenomenon of radioactive disintegration.

The above mentioned phenomena emphasize the stochastic character of the neutrons interactions with the surrounding nuclei and chiefly of the fuel nuclei fission.

The stochastic character of the microscopic processes peculiar to the neutron multiplication within a reactor explains at the macroscopic scale, the presence of the fluctuations of the neutron population in the reactor. This kind of fluctuations takes the name of neutron noise / 2 /.

There results that the neutron population N(t), exis ting at moment t considered with respect to an arbitrarily chosen origin, within a critical or subcritical reactor, where the statistical balance of the neutron population is nevertheless achieved presents fluctuations around a mean value  $\widetilde{N(t)}$  expressed by

(5) 
$$H(t) = \hat{H} \{ N(t) \} = m_{N}$$

where M is the operator of the statistic mean

(6) 
$$\hat{H} \{ \mathbf{H} (\mathbf{t}) \} = \int_{0}^{\infty} \mathbf{x} \, d \mathbf{F}_{\mathbf{H}} (\mathbf{x})$$

and  $\mathbf{F}_{W}(\mathbf{x})$  is the distribution function of the random variable W(t).

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The neutron population N(t) within the reactor at time t should therefore not be interpreted in a deterministic sense. Since it is a random variable, the reactor state is determined if its moments of any order are known. To characterize the evolution as a function of time of the neutron popu lation, it is necessar; to use a set of such functions depen ding on time and on an infinity of random parameters. Such a family of functions defines a stochastic process / 3 /.

Under conditions easily fulfilled in experiments, the study of a stochastic process reduces itself to the study of a single sample function of the process. With reference to a single sample function, time mean values can be defined such as :

(7) 
$$\langle N(t) \rangle = \lim_{T \to \infty} \frac{1}{2T} \int_{-T}^{T} N(t) dt$$
 (time mean value)

(8) 
$$\langle N^2(t) \rangle = \lim_{T \to \infty} \frac{1}{2T} \int_{-T}^{T} N^2(t) dt$$
 (mean square value called also total mean power)

(9) 
$$R_{N}(\tau) = \lim_{T \to \infty} \frac{1}{2T} \int_{-T}^{T} H(t) H(t+\tau) dt$$
 (auto correlation function)

A stochastic process is called ergodic when all the time means exist and coincide , independently of the used sample, with the corresponding statistical means. If the process is for instance (aussian, it is easy to show that the condition :

(10) 
$$\int_{-\infty}^{\infty} /R_{\rm H}(\tau) - R_{\rm H}/^2 d\tau < \infty$$

involves ergodicity.

The ergodicity property is very important since it reduces the evaluation of the statistical means of a stochastic process to the estimation of time means by taking into account a single sample of the process. Such an experiment is only valuable if the process is ergodic and stationary or at least stationary in a restricted sense.

A stochastic process is stationary in a restricted sense if all its statistical properties are invariant with regard to translations occuring along the time axis; in particular when distribution function  $P_{\rm N}(x)$  and statistical mean  $m_{\rm N}$  do not depend on time while the selfcorrelation function

(11) 
$$\tilde{H} \{ H(t_1) . N(t_2) \} = R_H(t_1, t_2)$$

depends only on difference  $t_2 - z_1 = \tau$ , the process is stationary of the second order.

The fluctuations of the neutron population within a reactor amound the mean value are strongly influenced by certain factors depending both on the composition of the active zone and on the geometrical size of the latter. It could be expected that the analysis of the neutron noise would permit to determine der tain global characteristics of the active zone which would prove usefull for describing the transient behaviour of the reactor as a deterministic physical system. These characteristics that dan be deduced from moments of the 1 and 11 order of the neutron propulation are usually called kinetic parameters.

The state variable in regard to which equation of evolution (1) is written in therefore the statistical mean of the neutron population.

The solution of the equation of evolution is influenced by the values the following parameters' :

- ap the decrease constant of the prompt neucron population;
  - 1 the mean life of the neutrons
- $\beta_{ef}^{-}$  the effective fraction of the delayed neutrons  $k_p^{-}$  the effective multiplication factor of the prompt neutrons, or, depending on these, the subcritical prompt reactivity of the reactor.

(12) 
$$\rho_{\rm p} = \frac{1 - k_{\rm p}}{k_{\rm p}}$$

The mathematical theory of the neutron noise in a reactor makes use of the usual methods and concepts for the study of any stochastic processes. As any mathematical theory for the study of physi cal phenomena, it is based on hypotheses liable to simplify the mathematic relations between the state variables occuring in the description of the respective system. The theoretical relations are however valuable only as long as they are verified by experiments. Therefore, a mathematic theory appears as useful as concerns the development and the diversifying of the methods for the measurements of the parameters of a physical system.

Therefore, a mathematical theory should operate with state variables of the physical system, susceptible to be physical observables.

When studying the neutron noise, a physical observable (which is also a random variable) can be the number of neutron re-

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corded during a time interval  $\Delta t$  by a counting device provided with a neutron detector and placed within the active area. Let  $z(\Delta t)$  be this observable. The study of the fluctuations of the neutron population is thus reduced to the study of the statistical properties of random variable  $z(\Delta t) = 2$ .

Should the neutron detector be placed in the neighbourhood of a radioactive source, the probability distribution func tion of the random variable 1(t) would be of the Poisson type, i. e.

(13) Prob 
$$\left(z(\Delta t) = c(\Delta t)\right) = \frac{(\tilde{k})^{c} e^{-\tilde{k}}}{c!} = p_{c}(\Delta t)$$

Poisson's distribution function is characterised by a single parameter ( $\tilde{A}$ ), which is the statistic mean of the random variable.

Therefore  $\tilde{n}$  is the mean value of the number of neutrons  $z(\Delta t)$ , recorded by the counting device during the interval  $\Delta t$ .

The moments of higher order of variable  $z(\Delta t)$  can in turn be expressed as a function of only parameter  $\widehat{A}$  of Poisson's distribution :

(14) 
$$P_{C}(\Delta t, t) = \frac{(\tilde{v}\Delta t)^{C}}{Ct} \exp(-\tilde{v}\Delta t)$$

where  $\tilde{\mathbf{v}}$  the mean counting rate.

For instance, the variance of the random variable de -

(15) 
$$\hat{H} \left\{ \left[ z \left( \Delta t \right) - \hat{H} z \left( \Delta t \right) \right]^2 \right\} = \hat{\nabla} \left\{ z \left( \Delta t \right) \right\} = \sigma_{z}^2 (\Delta t)$$

is equal to its mean value, respectively

(16) 
$$\hat{\mathcal{H}} \left\{ \left[ \Xi(\Delta t) - \widetilde{\mathcal{H}(\Delta t)} \right]^2 \right\} = \hat{\mathcal{H}} \left\{ \Xi(\Delta t) \right\}$$

or

$$\sigma_{\Xi(\Delta t)}^{2} = E_{\Xi(\Delta t)}$$

If the neutron detector of the counting device is placed within the active some of a reactor, the probability distribution of random variable  $z(\Delta t)$  will evince a deviation from Poisson's distribution due to the chain fission reactions within the active some.

In other words, the neutron detector can only record during the interval At neutrons belonging to the same chain of fissions. The hypothesis of the independence of the events ( an "" events means here the recording of a neutron), characteristic of a Poisson's stochastic process, can no more be accepted as true. Because of the correlated events, relation

(18) 
$$\frac{\sigma_{x(\Delta t)}^{2}}{m_{x(\Delta t)}} = 1$$

is no verified any more. In such a situation, relation (18) should become / 4, 2, 5, 6, 11 /.

(19) 
$$\frac{\sigma_{z(\Delta t)}^2}{m_{z(\Delta t)}} = 1 + \psi(\Delta t)$$

Term  $\psi(\Delta t)$  represents the deviation of the probability distribution of variable  $x(\Delta t)$  from Poisson's distribution and becomes increasingly important as the weight of the correlated events increases. In this paper, we present a method of analysis of the neutron noise within a nuclear reactor that has the attention of the Laboratory of Reactor Physics at IFA, with a view to be used for the experimental studies in the lattices of teac tors with thermal neutrons.

It is the matter with the method of the reduced variance of the number of neutrons  $x(\Delta t)$  recorded by a counting device during a time interval  $\Delta t$ .

In the first part of the paper, the theoretical bares of the method are expounded. The second part is devoted to describing a chain of analysis of the fluctuations and especially of the stochastic multichannel selector, an achievement of the Laboratory of Nuclear Electronics co-operating with that of Reactor Physics.

It should be pointed out that experimentally neither the probability distribution of random variable  $z(\Delta t)$  not its statistic moments are accesible to direct measurements.

Only estimations of these statistic characteristics can be determined. As it will be shown further in this paper, the dependence of the reduced variance on the length of the recording time interval  $\Delta t$ , takes the form / 7, 8 /.

(2c) 
$$\frac{\frac{1}{n} \sum_{i=1}^{n} c_{i}^{2}(\Delta t) - (\frac{1}{n} \sum_{i=1}^{n} c_{i}(\Delta t))}{\frac{1}{n} \sum_{i=1}^{n} c_{i}(\Delta t)} = 1 + \gamma(\Delta t)$$

Relation (20) is only valid for small values of interval At, for which the contribution of the delayed neutrons can be disregarded.

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In this assumption function  $\phi(\Delta t)$  previously introduced is evpressed by :

(21) 
$$\psi(\Delta t) = \varepsilon D_{v} \left(\frac{k_{p}}{1-k_{p}}\right)^{2} \left(1 - \frac{1-\exp(-\alpha_{p} \Delta t)}{\alpha_{p} \Delta t}\right)$$

Except the recording interval  $\Delta t$ , function  $\psi(\Delta t)$  depends on several constants including the reactor kinetic parameters  $a_p$  and  $\rho_p$  which can be determined by suitable fitting procedures. For determining the kinetic parameters, the experimential measurement is required of the reduced variance of the number of recorded neutrons for various values of length  $\Delta t$  of the recording interval.

This method is only apparently simple. The relative error of measurement of function  $\psi(\Delta t)$  and the total number n of intervals  $\Delta t$  are related by / ? /

(22) 
$$\frac{\Delta (1 + \psi(\Delta t))}{1 + \psi(\Delta t)} = \sqrt{\frac{2}{n}}$$

If for  $\psi(\Delta t) = 0.01$  we accept a measuring error of  $10^{9}$  dure results from (22) that the necessary number of recording intervals should satisfy relation

$$n \ge 10^6$$

The achievement of the recording device to which, as suggested by (20), this experiment is leading, raises particularly intri-... cate tebhnical problems / 2, 10 /. This is the reason that di rected as towards the achieving of an electronic equipment (fig.1) capable of permitting the direct determination of the probabilility distribution of the random variable  $x(\Delta t)$  along the variation range of the recording interval  $\Delta t$ .

This equipment achieved in the IPA Laboratory of Muclear Electronics is the stochastic multichannel analyser. It represents the key-element of rn analysis chain of the fluctuations of the neutron population (fig. 2) made in fact with the help of elements available on the market.

We shall describe succintly the working of the equip rent that permits a rational enough use of relation (20) when determining  $\psi(\Delta t)$ .

Let t = 0 be the time when the START knob is pushed that releases the equipment for which the possibility is provided of stopping it either manually or controlled by a time T which can be selected by switch PRESELECTION fixed on the stochastic selector panel. The time interval (0,T) of a recording can be considered as divided into a number of an adjacent subintervals of equal duration At.

$$(23) \qquad \qquad \Delta t = \frac{1}{n}$$

Every At subinterval is separately analysed by the equipment. The number of neutrons that can be detected within such a subinterval has a higher limit imposed by the very method of processing the experimental data. According to the informa tions supplied by the literature, this method becomes inapplicable when the number of the neutrons detected within a At interval grows higher than  $2^7$ .

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If, for instance, in the first time subinterval &, neutrons are recorded ( $c \leq 2^7$ ), the content of the adress of the memory device EM95B grows by a unity. At the end of the experiment (after time T) the memory addresses bear the following distribution:

(24) 
$$\{q_c(\Delta t), c = 0, 1, ..., 127\}$$

where  $q_c(\Delta t)$  is the number of intervals - from the total n- when the detector recorded c neutrons.

When distribution  $\{q_c(\Delta t)\}$  is known, the relative frequencies can be determined of the intervals where c neutrons were recorded. These frequencies, supplied by :

(25) 
$$f_c(\Delta t) = \frac{q_c(\Delta t)}{n}$$

are estimations of the probability that the number of neutrons detected in a  $\Delta t$  time interval would be c.

Therefore, the working of the recording device is only correct as long as guantities  $q_{\mu}(\Delta t)$  verify relation:

(26) 
$$\begin{array}{c} 127\\ \Sigma \\ c=0 \end{array} \qquad \mathbf{f}_{c}(\Delta \mathbf{t}) = \mathbf{h} \\ \end{array}$$

By using the recorded distribution  $\{q_{c}(\Delta t)\}$ , we shall be able to compute the empiric moments of the I and II orders of random variable  $\pi(\Delta t)$ ; which are unbiased estimations of the statistical mean value and of the mean-square value, respectively,

(27) 
$$\mu_{z} = \frac{127}{c} c f_{c}(Lc)$$

(28) 
$$S_z^2 = \frac{127}{\Sigma} c^2 f_c(tt)$$

These quantities permit to determine function (At) by means of

(29) 
$$\psi(\Delta t) = \frac{s_z^2 - \mu_z^2}{\mu_z} - 1$$

So as to allow for additional tests concerning the correctness of working of the equipment, it is moreover necessary to record separately the total number  $N_D$  of the recorded neu-trons during the whole time T of an experiment.

The average number of neutrons corresponding to a single time interval can therefore be equally deduced from ratio  $n D \over n$ , that should coincide with the empirical mean value

$$\frac{n_D}{n} = \mu_g$$

Lastly, the recorded data permit moreover to verify relation

(31) 
$$n_{D} = \sum_{c=1}^{127} c q_{c}(\Delta t)$$

The experimentally, obtained data permit therefore to determine function  $\psi(\Delta t)$  corresponding to a given value  $\Delta t$ of the argument.

However, the determination of the kinetic parameters of the reactor active some can only be made if the dependence of function  $\phi(\Delta t)$  on duration  $\Delta t$  18  $k \Pi^{O_H/I}$ . The experiment must therefore be repeated for all the argument values  $\Delta t_1$ ,  $\Delta t_2$ , ...,  $\Delta t_j$  to which, through relation (29), correspond the quantities:  $\psi(\Delta t_1)$ ,  $\psi(\Delta t_2)$ , ...,  $\psi(\Delta t_4)$ .

# II. METHOD OF THE VARIANCE OF DETECTED NEUTRONS NUMBERS

Let dt, and dt, be two disjoint time intervals.

We propose to deduce an expression of the average number of neutrons detected within these intervals by a neutron detector while taking into account the neutron multiplication due to the chain fission reactions. To do this, we shall introduce the following probabilities:

 $-p(m, dt_1) = the probability for m neutrons to be recorded in <math>dt_1$  $-p(m, dt_1; n, dt_2) = the joint probability for m neutrons to be$  $recorded in <math>dt_1$ , and for n neutrons to be recorded in  $dt_2$ 

 $-p(n, dt_2/m, dt_1)$  = the conditional probability for n neutrons to be recorded in  $dt_2$ , if m neutrons were recorded in  $dt_1$ .

Let r(t) be the instantaneous count rate at time t (random variable). The number of recordings in dt is then r(t) dt while the statistic mean value of the product of the number of recordings in dt<sub>1</sub> by the number of recordings in dt<sub>2</sub> will be

(32) 
$$\langle R(t_1)dt_1, h(t_2)dt_2 \rangle = \Sigma = p(n, dt_1) n, dt_2$$
  
m,n

When the time intervals are small, this mean value is approximated by the probability of a single recording, whence

(33) 
$$\langle A(t_1) A(t_2) dt_1 dt_2 \rangle \simeq p(1, dt_1; 1, dt_2)$$

We shall express this probability as the sum of probabilities of two disjoint events. The first term represents the probability of a recording in  $dt_2$  whether a neutron was recorded or not in  $dt_1$ . The second term expresses the contribution to the mean value of the events defined by a "ecording in  $dt_2$  after having been recorded in  $dt_1$ . The pairs of recordings in  $dt_1$  and  $dt_2$  can therefore be either accidental or correlated.

The uncorrelated (accidental) pairs of recordings are produced by neutrons with no common ancestor , meaning that they do not belong to the same fission chain.

On the other hand, the correlated pairs originate in recordings of neutrons yielded by the same fission reaction. This manner of interpreting the pairs of recordings will be better explained later on.

We can write :

$$(34) \quad p(1,dt_1; 1,dt_2) = p_1(1,dt_2; 1,dt_2) + p_2(1,dt_2; 1,dt_2)$$

where

(35) 
$$p_{\alpha}(1,dt_1; 1,dt_2) = p(1,dt_1) \cdot p(1,dt_2/1,dt_1)$$

represents the probability of the recording of a correlated pair of newtrons within the intervals  $dt_1$  and  $dt_2$ .

## Uncorrelated pairs

The uncorrelated pairs of recordings are due to detected neutrons that do not belong to the same fission chain.

So as to be able to deduce the probability expression  $p_a(1, dt_1; 1, dt_2)$  we shall define the following quantities:

- F = the mean fission rate within the active zone of the reactor
- $\varepsilon$  = the efficiency of the neutron detector by relations / 1 /

(36) 
$$F = \int \int dE dV \Sigma_f(\vec{R}, E) \cdot v(E) n(\vec{R}, E)$$
  
(E)  $(V_R)$ 

where

- dE.n( $\vec{n}$ ,E) is the mean value of the number of neutrons per unit volume around the point the posi tion vector of which is  $\vec{A}$  the neutron energy being of E to E + dE,
- $\Sigma_{f}(\vec{n}, E)$  is the probability with regard to the unit path that a neutron of E energy at space point  $\vec{n}$  would produce the fission of a fuel nucleus.

(37) 
$$\varepsilon = \frac{\int dE \int \Sigma_{D}(\vec{n}, E) v(E) n(\vec{n}, E) dV}{\int dE \int \Sigma_{f}(\vec{n}, E) v(E) n(\vec{n}, E) dV}$$

The detector efficiency is therefore equal to the ratio between the mean rate of the detector recording and the mean rate of the reactor fissions. The probability of a recording in dt is

(38) 
$$p_d(1, dt) = \epsilon P dt$$

so that the probability of a uncorrelated pair recording in  $dt_1$  and  $dt_2$  is supplied by

(39) 
$$P_a(1,dt_1; 1,dt_2) = \varepsilon^2 P^2 dt_1 dt_2$$

## Correlated pairs

Let  $\Pi(t_k, t)$  be the probability with regard to the unit time that a neutron produced within the reactor at time t as the result of a fission would have a descendent (itself being one of them) within the system at time  $t_k$ .

The probability of one count in  $dt_1$  from a progeny of a fission which emits  $v_p$  prompt neutrons at time t is :

(40) 
$$v_p p dt \Pi(t_1,t) = \frac{dt_1 c}{v_p}$$

The probability of a recording in  $dt_1$  to be followed by another one in  $dt_2$  from a progeny of the same fission at time t is :

(41) 
$$(v_{p}-1) \pi(t_{2},t) - \frac{cdt_{2}}{v_{2}}$$

The probability of the correlated pair: is therefore :

(42) 
$$\frac{v_{p}(v_{p}-1)}{(v_{p})^{2}} Pe^{2} dt dt_{1} dt_{2} I(t_{2},t) I(t_{1},t)$$

The fission reaction of a fuel nucleus at time t has released  $v_p$  neutrons but with probability  $p_{v_p}$ , so that to the above expression we must apply the operator of the "statistic mean" whence

(43) 
$$\frac{\overline{v_p(v_p-1)}}{(\overline{v_p})^2} \mathbf{F} c^2 dt = t_1 dt_2 \mathbf{I}(t_2,t) \mathbf{I}(t_1,t)$$

The probability of a recording in  $dt_1$  and of another one in  $dt_2$ due to neutrons released by earlied fissions of nuclei is supplied by integration

(44) 
$$p_{c}(1,dt_{1}; 1,dt_{2}) = P \epsilon^{2} D_{v} dt_{1}, dt_{2} \int_{-\infty}^{t_{1}} \pi(t_{1},t) \pi(t_{2},t) dt$$

The statistic mean value of the number of both uncorrelated and correlated pairs recorded in  $dt_1$  and  $dt_2$  is therefore

(45) 
$$\langle \pi(t_1) \pi(t_2) dt_1 dt_2 \rangle = \epsilon^2 p^2 dt_1 dt_2 + p \epsilon^2 D_1 dt_1 dt_2 \int_{-\infty}^{t_1} \pi(t_1, t) \pi(t_2, t) dt_1 dt_2 \int_{-\infty}^{t_1} \pi(t_1, t) \pi(t_2, t) dt_1 dt_2 \int_{-\infty}^{t_2} \pi(t_1, t) \pi(t_2, t) dt_1 dt_2 \int_{-\infty}^{t_1} \pi(t_1, t) \pi(t_2, t) dt_1 dt_2 \int_{-\infty}^{t_2} \pi(t_1, t) \pi(t_2, t) dt_1 dt_2 \int_{-\infty}^{t_1} \pi(t_1, t) \pi(t_2, t) dt_1 dt_2 \int_{-\infty}^{t_2} \pi(t_1, t) \pi(t_2, t) dt_1 dt_2 \int_{-\infty}^{t_2} \pi(t_1, t) \pi(t_2, t) dt_1 dt_2 \int_{-\infty}^{t_2} \pi(t_1, t) \pi(t_2, t) dt_2 \int_{-\infty}^{t_2} \pi(t) dt_$$

Probabilities  $R(t_1,t)$  and  $R(t_2,t)$  are solutions of the following kinetic equations / 5, 8 / :

$$\frac{d\pi}{dt} = \frac{p-\beta}{\Lambda} \pi + z\lambda_1 C_1 + \delta(t)$$
(1)

(47) 
$$\frac{dc_1}{dt} = -\lambda_1 c_1 + \frac{s_1}{h} \mathbf{x}$$

In the case of s small t, the "source" of delayed seutrons can be assumed as approximately constant so that the solution of the above system of equations is

(48) 
$$\Pi(t,0) = e^{-\alpha_0 t}$$
, where  $\alpha_0 = \frac{\rho - \beta}{\Lambda}$ 

The general solution of the equations takes the form

(49) 
$$\Pi(t,0) = \Lambda \Sigma \Lambda_{\chi} \exp(-n_{k}t)$$
(k)

where  $\lambda_k$  and  $\alpha_k$  are constants resulting from the cransfer function of the reactor

(50) 
$$T(s) = \frac{1 - \sum_{\substack{(i) \\ (i) \\ s + \lambda_i}} \frac{s \beta_i}{s + \lambda_i}}{s (t + \sum_{\substack{(i) \\ s + \lambda_i}} \frac{\beta_i}{s + \lambda_i}) - \rho} \frac{\lambda_k}{s + \alpha_k}$$

With the former notations, a<sub>o</sub> represents the decrease constant of the population of prompt neutrons

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When t is small the general solution coincides with the particular one if  $\lambda_0 = 1$ . We have then

(52) 
$$\Pi(t_1,t) = \overset{\sharp}{L} \beta_k \exp\left[-\alpha_k (t_1-t)\right]$$
  
k=0

(53) 
$$I(t_2,t) = \sum_{k=0}^{2} s_k \exp \left[-a_k(t_2-t)\right]$$

Relation (45) becomes

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$$(54) < \mathfrak{r}(\mathfrak{t}_{1}) \mathfrak{r}(\mathfrak{t}_{2}) d\mathfrak{t}_{1} d\mathfrak{t}_{2} > = \varepsilon^{2} \mathfrak{r}^{2} d\mathfrak{t}_{1} d\mathfrak{t}_{2} + \mathfrak{r} \varepsilon^{2} \mathfrak{D}_{v} d\mathfrak{t}_{1} d\mathfrak{t}_{2} \mathfrak{r} \frac{\mathfrak{B}_{1} \mathfrak{B}_{1}}{\mathfrak{a}_{1} \mathfrak{a}_{j}} \exp \left[-\mathfrak{a}_{1} (\mathfrak{t}_{2} - \mathfrak{t}_{1})\right]$$

Introducing notation

$$Q_{i} = \frac{\Gamma}{j=1} \frac{B_{j}}{\alpha_{i} + \alpha_{j}}$$

we shall obtain

$$(55) \langle \mathcal{X}(t_1) \pi(t_2) dt_1 dt_2 \rangle = \mathbb{P}^2 \varepsilon^2 dt_1 dt_2 + \mathbb{P} \varepsilon^2 D_v dt_1 dt_2 = \frac{\mathbb{E} A_1 \Omega_1 \exp \left[-\alpha_1 (t_2 - t_1)\right]}{(1)}$$

,

Let us denote by  $\Delta t$  a time interval including interval  $(t_1, t_2)$  and let  $c(\Delta t)$  be the number of recordings within this interval. The number of recording pairs will be

$$C_{c(\Delta t)}^{(2)} = \frac{c(\Delta t) \cdot [c(\Delta t) - 1]}{2}$$

and the mean number of pairs will be :

(56) 
$$\langle \frac{c(\Delta t) \cdot [c(\Delta t) - 1]}{2} \rangle = \int_{t_2=0}^{\Delta t} \int_{t_1=0}^{\Delta t_2} dt_1 \langle \pi(t_1) \cdot \pi(t_2) \rangle$$

(57) 
$$\frac{1}{2} \langle c(\Delta t), [c(\Delta t)-1] \rangle = \frac{e^2 p^2 \Delta t}{2} + p e^2 D_v \Delta t \qquad \frac{h_+ 1}{k} \frac{h_k Q_k}{k} \left(1 - \frac{1 - \exp(-\alpha_k \Delta t)}{\alpha_k \Delta t}\right)$$

But  $\varepsilon P \Delta t$  represents the mean number of recordings within interval  $\Delta t$ , i.e.

(58) 
$$\langle c(\Delta t) \rangle = \epsilon P \Delta t$$

and therefore

(59) 
$$\frac{\langle c^{2}(\Delta t) \rangle - \langle c(\Delta t) \rangle^{2}}{\langle c(\Delta t) \rangle} = 1 + 2\varepsilon D_{v} \sum_{k=1}^{n+1} \frac{A_{k}Q_{k}}{\alpha_{k}} (1 - \frac{1 - \exp(\alpha_{k}\Delta t)}{\alpha_{k}\Delta t})$$

or

(60) 
$$\frac{\langle c^2(\Delta t) \rangle - \langle c(\Delta t) \rangle^2}{\langle c(\Delta t) \rangle} = 1 + \sum_{k=1}^{n+1} Y_k (1 - \frac{1 - \exp(-a_k \Delta t)}{a_k \Delta t})$$

which was deduced by E.N.Courant, R.L.Wallace, D.R Frisch, D.J. Liffler and R.N.Albrach in the form

(61) 
$$\frac{\langle c^2(\Delta t) \rangle - \langle c(\Delta t) \rangle^2}{\langle c(\Delta t) \rangle} = 1 + \sum_{\substack{k \neq k}} \psi_k(\Delta t)$$

## Structure and performances of the system

For data accumulation, the memory register BM 96-B (4096 addresses) is used which in an accumulation regime has a capacity of  $10^6$ /adress.

The complete structure of the system (fig. 1) is formed by :

- the analyser itself, conceived as module %IF (fig. 2)

- the memory device BM96-B

- the neutron detector D with a 0,5 µs time resolution

- the time-mark generator G, with a period  $\Delta t_{G} = 50 \ \mu s = 100 \ ms$ - the  $n_{D}$  counter with a  $10^{9}$  capacity for recording the total number of analysed pulses (detector counts).

The analyser preselects the information accumulation received from the detector in recording intervals T of  $10^3$ ,  $10^4$ ,  $10^5$  or  $10^6$  or as case may be any time interval.

The analysis chain can accept up to 127 pulses from the detector during an interval  $\Delta t_G$  (for a number higher than 127 pulses, the 127-th address will increase by 1 its content for each recording interval.

#### Operation

During the feeding of the installation, the START-STOP (SP) switch is in position 1 (STOP) so that the binary counter with seven bits (NB) and flip-flops PF1, FF2 and FF3 are held at the zero state (fig. 2). The equipment is reset (decimal counter n) with the help of knob SERO. At the same time, counter  $n_D$  is separately reset too. For data accumulation, swith SP is put on position 2 (START).

When the first pulse comes from time-mark generator G (SP being on 2), FP1 passes on the back wave front, on "1" by opening gate  $P_2$  when the first pulse from G achieve its and (fig. 3). It is necessary that the first pulse should not pass so as to ensure the equality of all the access periods of the pulses from detector D. The second pulse from G pass through  $P_2$  where it undergoes a delay  $t_1 = 200$  ns; its wave front releases mono-flip-flop MS (the duration of the pulse produced by MS is  $\delta_{MS}=1$  µs). The back wave of the MS shapes with the help of delay  $t_2$  a needle pulse of loo ns, the back wave of which passes FF2 and FF3 in "1". From this moment forward, gate  $P_1$  is opered and the pulses from D are recorded on  $n_p$  and NB.

When the third pulse comes from G, the sequence of events produced by the second pulse is repeated. On the other hand, FT2 is reset and blocks gate  $P_1$  and since FT3 is in state "1", all impulses coming from G, beginning by the third one are recorded on NE. At the same time, MS releases on the first front the memory cycle (IM). During this cycle (of 16  $\pm$  2 µs) the NB contents is transferred to the address register RA of memory device BM 96-B, the RA address content passes into information register RI; this content increased by 1 is returned to the specified address in RA.

The negative impulse from reset MB, while its back wave brings again in "1" PP2 which opens again gate  $P_1$ .

Delay  $t_1$  is necessary for the transient processes in WB to be at an end before the namery cycle is released.

The flip-flop FT3 avoids the releasing of the memory cycle when the second pulse arrives from G which would have resulted in the recording of an event at address zero. On the other hand, the second pulse is not recorded on W2 either.

The duration of the recording period is At = At<sub>G</sub> =  $\Lambda_{MR}$ .

If in a  $At_G$  interval, detector D supplied to the output sure than 127 pulses, at the 127-th pulse gate P<sub>3</sub> opens and blocks the passage of the pulses through P<sub>1</sub> so that the event will be re-

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gistered on address 127. The frequent inaccurate recording of events on address 127 results in the lighting of laps OVERLOAD. In this case, the frequency of time-mark generator G should be increased.

When the contents of (n) reach  $10^3$ ,  $10^4$ ,  $10^5$  or  $10^6$ (as a function of the position of swith PRESELECTION), PS becomes "1", the aignalling lamp of the self-stopping (SA) and PP1 and PF2 are therefore reseted zero; the last memory cycle is released, after the recording the NB content is reset and FF3 is reset to: (which avoids the possibility of an accidental release of the memory cycle as well as of the altering of the accumulated data). The experiment, is thus "automatically" stopped.

For a new experiment, SP is again set on 1, the equipment is reset after wich SP is passed in position 2.

The information accumulation can also be stopped " manually" at any time by putting SPon prosition 1. The interrupted experiment can however be continued by simply putiing SP again on position 2.

## Characteristics of the analyser input/output pulses

- output from time-mark generator and detector

- polarity positive
- amplitude 2 + 30 V
- deration (at 1.5 Wolts) >0,2 µs at the generator 3and 0.22 µs at the detector

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- output towards the counter

- polarity positive
- amplitude 3,5 V

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- duration - the same as that of the pulses from the

## detector

- towards memory device BM96-B

- polarity positive
- amplitude 7,4 V + 10%
- rise time 0.3 µs
- fall time 40 ns
- duration 1 µs.

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Pig. 1

 $1 - \frac{1}{2} \tan \left(\frac{1}{2} + \frac{1}{2} + \frac{1}{2}$ 



Fundamental diagram of the stochastic analyser





The signal diagram in the stochastic analyser.